A dynamic model for the outbreaks of hand, foot, and mouth disease in Taiwan

C.-C. LAI1,2, D.-S. JIANG3, H.-M. WU2 AND H.-H. CHEN2*

1 Emergency Department, Taipei City Hospital, Ren-Ai Branch, Taiwan
2 Graduate Institute of Epidemiology and Preventive Medicine, College of Public Health, National Taiwan University, Division Biostatistics, Taipei Taiwan
3 Field Epidemiology Training Program, Centres for Disease Control, Taiwan

Received 21 December 2014; Final revision 5 October 2015; Accepted 7 October 2015; first published online 16 November 2015

SUMMARY

The first large outbreak of hand, foot, and mouth disease (HFMD) with severe complications primarily caused by enterovirus 71 was reported in Taiwan in 1998. Surveillance of HFMD to evaluate the spread of HFMD with and without infection control policy is needed. We developed a new dynamic epidemic Susceptible-Infected-Recovered (SIR) model to fit the surveillance data on containing valuable information on the severity of HFMD in order to accurately estimate the basic reproductive number ($R_0$) of HFMD. After fitting the empirical data, in conjunction with other relevant parameters extracted from the literature, the estimated transmission coefficients were close to $5 \times 10^{-7}$ (per day) and the proportion of severe HFMD cases ranged between 0 and 0.0036 (per day). Taking into account the distribution of all parameters considered in our dynamic epidemic model, the $R_0$ computed was 1.37 (95% confidence interval 0.24–5.84), suggesting a higher likelihood of the spread of HFMD if no infection control policy is provided. The isolation strategy against the spread of HFMD not only delayed the epidemic peak with the delayed time ranging from 4 weeks for only 20% isolation to 47 weeks for 100% isolation but also reduced total number of HFMD cases with the percentage of reduction ranging from 1.3% for only 20% isolation to 13.3% for 100% isolation. The proposed model can also be flexible for evaluating the effectiveness of two other possible policies for containing HFMD, quarantine and vaccination (if the vaccine can be developed).

Key words: Compartmental model, enterovirus 71, hand, foot, and mouth disease (HFMD), reproductive number, SIR model.

INTRODUCTION

Hand, foot, and mouth disease (HFMD) is, to a large extent, caused by coxsackievirus A16 (CVA16) and, to a lesser extent, associated with other serotypes such as A4, A5, A9, A10, B2, and B5 viruses as well as echovirus and enterovirus 71 (EV71) [1–3]. Humans are the only natural host for these enteroviruses that are transmitted through two possible modes, personal oral–faecal route and oral–oral route [4]. The majority of diseases afflicting children are typically benign and self-limiting but HFMD caused by EV71 may be more likely to result in severe complications such as acute flaccid paralysis, encephalitis, meningitis, myocarditis and pulmonary oedema.

* Author for correspondence: H.-H. Chen, Ph.D., Graduate Institute of Epidemiology and Preventive Medicine, College of Public Health, National Taiwan University, Taiwan, 5 Fl. No. 17, Hsu Chow Road, Taipei, Taiwan, 10055. (Email: chenlin@ntu.edu.tw)

This is an Open Access article, distributed under the terms of the Creative Commons Attribution licence (http://creativecommons.org/licenses/by/4.0/), which permits unrestricted re-use, distribution, and reproduction in any medium, provided the original work is properly cited.
Large outbreaks associated with EV71, first described in 1974 [5], have been heralded in Japan since the early 1970s [6]. After around two decades, EV71 epidemic activity has been spread throughout the Asia-Pacific region producing large outbreaks such as the 1998 Taiwanese HFMD outbreak caused by two main types of virus, CVA16 and EV71. The age distribution and clinical features of infections (such as the proportion of HFMD and herpangina) between EV71 and CVA16 were similar. It is therefore difficult to clinically distinguish between HFMD caused by EV71 and that caused by CVA16, not to mention the asymptomatic but still infectious HFMD cases. In a Taiwanese seroepidemiological study of 46% people of all ages who were infected with EV71, about 71% showed no clinical symptoms [7]. The contrast between both types is that illnesses resulting from EV71 were more likely to lead to a significantly greater frequency of serious complications and fatalities than those caused by CVA16 [8, 9]. It is therefore important to make use of HFMD with severe complications to back-estimate symptomatic and asymptomatic but infectious cases resulting from infectives in the absence of intervention in order to calculate the basic reproductive number ($R_0$, the expected number of secondary cases produced by an index case in a completely susceptible population).

$R_0$ is a useful indicator for assessing the spread of HFMD because it is a function of the number of contacts per unit time, the probability of transmission per contact, and the duration of the infectious period [10]. If $R_0 > 1$, an outbreak occurs. However, if $R_0 < 1$, an outbreak is unlikely. Thus, if a sustained disease control reduces the transmission intensity by a factor that exceeds $R_0$, HFMD will eventually be eliminated. Although $R_0$ is widely estimated via the Susceptible-Infected-Recovered (SIR) model for other infectious diseases [10], its application in delineating the natural course of the infection–disease process and evaluating the preventive strategy containing HFMD is still limited because it is difficult to distinguish the pathogenic HFMD cases (including asymptomatic but still infectious cases and symptomatic cases) attributed to different causes of serotypes as indicated above from the empirical surveillance data. Therefore, it is intractable to estimate the relevant parameters for deriving $R_0$ using this traditional compartmental model. To cope with this problem, we developed a new modified SIR model by adding information on clinical symptomatic and severe cases to fit the empirical surveillance data in order to accurately estimate the parameters governing the evolution from susceptibility to infection, recovery, and finally the development of severe HFMD. These parameters, while providing information on baseline infection and disease process related to the spread of HFMD in the absence of intervention, enables assessment of the effectiveness of intervention in containing the outbreak of HFMD. The aims of this article consist of two parts. In the development and estimation process, we first build up the new proposed dynamic model. We then used empirical data from a wave of outbreaks on HFMD to train the parameters from susceptible, infection, asymptomatic disease, symptomatic disease until the final sequelae, recovery, HFMD, herpangina, and severe HFMD. We were able to project the value of the basic $R_0$ without effective infection control strategy based on parameters estimated above. In the application part, the impact of HFMD (e.g. the delayed time for the spread of HFMD and a decrease in the number of HFMD cases) can be evaluated by the different policies of the isolation strategy.

**MATERIALS AND METHODS**

**Data sources**

In Taiwan, the government has established a physician-based sentinel surveillance system, representing 8-7% of primary physicians for HFMD since 1998; all severe cases of enteroviruses must be reported to the Centers of Disease Control (CDC) through a hospital-based reporting system [9, 11]. A severe case was defined as follows: patients were hospitalized if, in addition to HFMD or herpangina, they showed signs or symptoms indicating more serious illness including high fever (body temperature $\geq 38 ^\circ C$), vomiting, tachypnoea, and indications of neurological complications or cardiopulmonary complications [11]. Active household investigations of secondary clinical cases were performed after each severe case was reported. There was no change over time regarding the surveillance system.

The data of HFMD obtained from the Taiwan CDC were used to estimate the parameters pertaining to the compartmental SIR model (see Supplementary material). Figure 1 shows the number of patients with HFMD or herpangina reported by the sentinel surveillance system between 1999 and 2008. We modelled
the results of previous outbreaks of HFMD to estimate the basic reproductive number ($R_0$) in 2000, 2001, 2005 and 2008 due to the large outbreaks of severe cases with HFMD in these calendar years.

Model specification

We modified the SIR model to evaluate the behaviour of the virus causing HFMD and the severe cases obtained from the surveillance system of Taiwan CDC. The compartmental model based on the characteristics of HFMD is shown in Figure 2 with the detailed relationships and notations. In brief, the population was divided into nine compartments namely, individuals susceptible to enterovirus (S), infectious cases before developing the disease (I), asymptomatic disease with infectiousness (AS-I), immunity after infected AS-I (R), infectious cases after developing symptoms (EV-I), symptomatic cases with immunity (EV-R) after recovery, typical cases with HFMD, severe cases and other mild symptoms or herpangina. Figure 2 shows the rates of transitions between these nine states. The model begins with the state of susceptible individuals (S) and then enters the infectious state (I) when infected. The subsequent transitions include I to AS-I or EV-I. Note that AS-I may end up with R. The most novel of this compartmental model is to add the EV-R compartment that represents a transient state for the recovery from EV-I to EV-R. If there is no recovery, the final outcome would be ascertained as herpangina, HFMD, or severe cases. Note that the exact dwelling time remaining in the infectious period in the EV-I state would be difficult to estimate if the EV-R compartment is lacking. From the viewpoint of population, the number of susceptible individuals is also affected by the death rate of the population. In addition, HFMD is a common infectious disease of infants and children, which could be largely affected by the birth rate.

The nonlinear ordinary differential equations of the HFMD model were used to describe the dynamic changes of each state. The expressions are detailed in the Supplementary material. The model assumes that the durations of these compartmental periods are exponentially distributed, giving the average duration by taking the reciprocal of the transition rates of leaving these states. We assume all enteroviruses causing HFMD lead to the similar distributions of clinical symptoms except for severe cases.

Derivation of basic reproductive number ($R_0$)

We used a next-generation method \cite{10} to derive $R_0$ in models with several disjointed compartments. The detailed process of deriving $R_0$ is elaborated upon in the second part of the Supplementary material [see equations (10)–(12)]. The $R_0$ derived from our model is a function of the transmission rate ($\beta$), the death rate ($\mu$), the transition rate ($\alpha$) from infectious cases before developing symptoms (I) to infectious cases after developing symptoms (EV-I) or asymptomatic
cases with infectiousness (AS-I), the proportion of symptomatic cases ($\rho$), the rate of recovery from AS-I ($\tau_a$), and the rate of recovery from EV-I ($\tau_s$).

The formula of deriving $R_0$ is summarized as follows:

$$R_0 = \frac{\beta}{\mu + \alpha} + \frac{\beta \alpha (1 - \rho)}{(\mu + \alpha)(\mu + \tau_a)} + \frac{\beta \alpha \rho}{(\mu + \alpha)(\mu + \tau_s)}.$$  

The numerators of the formula consist of three parts, the transmission rate before developing disease ($\beta$), the product of the transmission rate with the rate of developing asymptomatic disease [$\beta \times \alpha \times (1 - \rho)$], and the product of the transmission rate with the rate of developing symptomatic disease ($\beta \times \alpha \times \rho$). The denominator of the first part is the inverse of the death rate plus the transition rate from infectious cases before developing symptoms ($\mu + \alpha$). In addition to $\mu + \alpha$, the denominators of the second and third parts are incorporated with the inverse of the death rate plus the rates of recovery for asymptomatic cases ($\mu + \tau_a$) and the inverse of the death rate plus the rates of recovery for symptomatic cases ($\mu + \tau_s$), respectively. We estimated the parameters of this dynamic model proposed in Figure 2 using Matlab version 7.10·0·499 (MathWorks, USA).

### Input of model parameters

Table 1 shows the input of parameters used in the modified dynamic SIR model shown in Figure 2. The population size (N) was identical to the Taiwanese population. The birth and death rates were obtained from Taiwan public health vital statistics [12]. The proportion of symptomatic cases was ~94% in children and 47% in adults [13]. EV71 seroprevalence rates ranged from 57% to 67% in adults and children aged >6 years, from 0% to 36% in children aged <3 years and from 26% to 51% in children aged 3–6 years [7]. Therefore, the proportion of susceptible (S) individuals in the total population was estimated to be ~45%.

As we assumed that most of the other parameters except the transmission coefficient ($\beta$) and the proportion of severe HFMD cases ($\delta$) were stable across studies, they were extracted from the literature. The parameters on the transition rate from I to EV-I or AS-I ($\alpha$), the recovery rate from AS-I ($\tau_a$), and the recovery rate from EV-I ($\tau_s$), were elicited using the following information. The incubation period of HFMD is about 1–7 days [14]. The infectious period typically occurs during the acute stage of HFMD, and it persists for several weeks [15]. If the infectious period of HFMD is the same as that of polio (14–42 days) [15], we estimated the
The parameters for calculating coef of τ case estimate was therefore assigned as 0.35. The range EV-I or AS-I) was at least 0.167 (per day), and the best-ferred that the range of  cases, the children were aged <6 years, and fi severe HMFD cases) to examine whether they were

transition from EV-R to herpangina was similar to that of HFMD but different from that of EV-R to severe illnesses resulting from EV71 were more likely to be severe than those resulting from CVA16, but the distributions of HFMD, herpangina or other mild symptoms were similar [8]. This presupposes that the rate of transition from EV-R to herpangina was similar to that of HFMD but different from that of EV-R to severe cases (see Fig. 2).

Although it was assumed that these parameters cited in the literature were stable across studies, they would be conjoined with the two parameters (such as transmission coefficients and the proportion of severe HFMD cases) to examine whether they were fitted well with the empirical data. In almost all of the cases, the children were aged <6 years, and ~75% of the children were aged <3 years [8, 13]. CVA16 or EV71 resulted in 45–49% HFMD, 10% herpangina, and 6.3% (CVA16) in ~21% (EV71) of severe cases [8, 13, 16]. Such information gives the range of the proportion of severe HFMD cases (δ).

**Statistical analysis**

Given the number of the total population and the birth and death rates each year in Taiwan, we applied the proposed dynamic epidemic model to fit surveillance to data of HFMD to train the parameters of transmission coefficients and the proportion of severe HFMD cases (β, δ) in conjunction with the other parameters indicated above.

We then validated the results of applying the parameters trained from the surveillance data to assess whether the number of observed severe HFMD cases was close to that of the predicted severe cases with the Pearson’s χ² test. If the observed value was <5, we combined the adjacent values until it was >5.

As the parameters used in the model-fitting process were assigned in a deterministic manner, we attempted to take into account the uncertainty of these parameters while basic reproductive numbers (R₀) were computed. We therefore considered using the Bayesian Markov Chain Monte Carlo (MCMC) method to assign the major parameters (α, τᵣ, τₛ, ρ,
The results in sensitivity analysis after 15,000 times sampling simulation. (per day) $\delta$

Supplementary material). We assigned data from the outbreak in 2008 (Table 1; see also

Predicted severe proportion of $R$ posterior distribution of $\beta$ and $\tau_a$ to gamma distributions and $\rho$ to beta distributions. All of the computation procedures were implemented using WinBugs version 1.4.3 (www.mrc-bsu.cam.ac.uk/software/bugs/).

<table>
<thead>
<tr>
<th></th>
<th></th>
<th></th>
<th></th>
<th></th>
<th></th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td>$\beta$ (per day)</td>
<td>—</td>
<td>$5.7 \times 10^{-7}$</td>
<td>$5.6 \times 10^{-7}$</td>
<td>$4.5 \times 10^{-7}$ (&lt;34 weeks)</td>
<td>$1.2 \times 10^{-6}$ (&gt;34 weeks)</td>
<td>$5.3 \times 10^{-7}$ (&lt;25 weeks)</td>
</tr>
<tr>
<td>$\delta$ (per day)</td>
<td>0–0.21</td>
<td>0.0018</td>
<td>0.0023</td>
<td>0 (&lt;24 weeks)</td>
<td>6.75 $\times 10^{-4}$ (24–32 weeks)</td>
<td>1.35 $\times 10^{-4}$ (&gt;32 weeks)</td>
</tr>
<tr>
<td>Proportion of cases reported</td>
<td>—</td>
<td>25%</td>
<td>22%</td>
<td>22.5%</td>
<td></td>
<td>22%</td>
</tr>
<tr>
<td>Predicted severe cases</td>
<td>—</td>
<td>365·38</td>
<td>403·7</td>
<td>66·63</td>
<td>373·52</td>
<td></td>
</tr>
<tr>
<td>(reported cases)</td>
<td>(367)</td>
<td>(395)</td>
<td>(57)</td>
<td>(373)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>$R_0$</td>
<td>1·22</td>
<td>1·21</td>
<td>1·59</td>
<td>1·18</td>
<td>1·37</td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>(95% CI 0·23–5·71)</td>
<td></td>
</tr>
</tbody>
</table>

$\delta$, Proportion of severe case; $\beta$, transmission coefficient; CI, confidence interval.

* The results in sensitivity analysis after 15,000 times sampling simulation.

$\beta$) with the proper statistical distributions to derive the posterior distribution of $R_0$ based on the empirical data from the outbreak in 2008 (Table 1; see also Supplementary material). We assigned $\beta$ to an inverse gamma distribution, $\alpha$ to the gamma distribution, $\tau_a$ and $\tau_s$ to gamma distributions and $\rho$ to beta distributions. All of the computation procedures were implemented using WinBugs version 1.4.3 (www.mrc-bsu.cam.ac.uk/software/bugs/).

Impact of isolation strategy against HFMD

HFMD cannot be prevented by immunization because vaccine against these viruses has not been developed yet. Therefore, information about how to properly wash hands, early diagnosis and treatment of HFMD, and suspension of school classes during the endemic period was used to stop the spread of HFMD in Taiwan.

It was easy to evaluate the effects of the intervention, such as isolation (suspension of class), on HFMD disease prevalence by adding the isolated factor denoted by $\Omega$, indicating the isolation rate of symptomatic HFMD in the model. Therefore, $\lambda(t)$ in the dynamic model was changed to the following equation:

$\lambda(t) = \beta \times \left[ I(t) + AS - I(t) + (1 - \Omega) EV - I(t) \right],$

where $I(t)$ denotes the number of infectious cases before developing the disease at time $t$, which denotes the numbers of asymptomatic disease with infectiousness (AS-I) at time $t$; $EV(t)$ denotes the number of infectious cases after developing symptoms at time $t$, and $\Omega$ denotes the proportion of isolated EV-I. We will demonstrate how the five isolation strategies with respective proportions, 0–2, 0·5, 0·7, 0·9 or 1, delayed the outbreak of HFMD and also reduced the disease burden (i.e. total number of HFMD cases) in the epidemic curves of coming years. All these simulations were conducted using Matlab version 7.10.0.499 (MathWorks).

RESULTS

The parameters estimated by fitting the proposed HFMD dynamic model to the empirical data are listed in Table 2. The transmission coefficient ($\beta$) (per day) was $5.7 \times 10^{-7}$ in the 2000 outbreak, $5.6 \times 10^{-7}$ in the 2001 outbreak, $4.5 \times 10^{-7}$ before 34 weeks, $1.2 \times 10^{-6}$ after 34 weeks in the 2005 outbreak, and $5.3 \times 10^{-7}$ before 25 weeks, $2 \times 10^{-7}$ between 26 and 33 weeks and $4 \times 10^{-7}$ after 34 weeks in the 2008 outbreak. The estimated proportion of severe HFMD cases ranged between 0 and 0·0036 (per day). We also assumed the asymptomatic cases with infectiousness (AS-I) were equal to those of the symptomatic phase with infectiousness (EV-I). Therefore, the total time spent in I and EV-I was ~15–36 days ($1/\alpha + 1/\tau_s$). The simulated and reported HFMD cases are shown in Figure 3, which indicated a good fit with visual inspection. Figure 4 shows the results of applying the parameters to validate whether the predicted severe HFMD cases were close to those observed. The goodness-of-fit test resulted in a
\(\chi^2 = 5.84\) with 8 degrees of freedom \((P = 0.67)\) for severe cases in 2005, and a \(\chi^2 = 31.83\) with 25 degrees of freedom \((P = 0.16)\) for severe cases in 2008 (Supplementary material). These tested results indicated no evidence of absence of goodness of fit. Note that the simulated values were not in good agreement with the observed good matches in 2000 and 2001 \((P < 0.0001)\) because the initial sentinel surveillance system for HFMD was unstable or different aetiological viruses existed.

The results of the simulations on the delayed time and total number of HFMD cases with different proportions of isolation strategy \((\Omega)\), assuming that 0.2, 0.5, 0.7, 0.9 or 1 were applied, are shown in Figure 5. The delayed time for generating an epidemic curve varying with the proportion of isolation ranged from 4 weeks for 20% isolation, 13 weeks for 50%, 22 weeks for 70%, 37 weeks for 90%, and 47 weeks for 100% isolation. After 4 years of continuing the same strategy, the total number of HFMD cases fell by 1.3% for 20% isolation, 4.7% for 50%, 6.6% for 70%, 13.0% for 90%, and 13.3% for 100% isolation compared to no isolation strategy.

**Computation of \(R_0\) with the Bayesian MCMC method**

We computed the basic reproductive number \((R_0)\) by deriving the posterior distribution of \(R_0\), taking into account the variability of \(R_0\) that results from the uncertainty of the parameters incorporated into the
dynamic modified SIR model as mentioned above. The proper statistical distributions of these parameters ($\beta$, $\alpha$, $\tau_a$, $\tau_s$, $\rho$), according to the findings from these outbreaks in 2000, 2001, 2005 and 2008 (Supplementary material), were assigned as stated in the Methods section. After 5000 burn-in samples, we ran another 10 000 samples based on five parameters ($\beta$, $\alpha$, $\tau_a$, $\tau_s$, $\rho$) given a fixed death rate ($\mu$). The $R_0$ of HFMD was estimated as $1.37$ (95% confidence interval (CI) $0.24$–$5.84$) based on the data from 2008. The convergence was plotted as shown in Supplementary Figure S1. Note that the convergence was achieved after 10 000 samples (including 5000 burn-in samples). To guarantee the convergence, we ran an extra 5000 iterations, and the results were very similar.

**DISCUSSION**

We built a compartmental SIR model by incorporating information on the surveillance data to estimate the $R_0$ ($1.37$, 95% CI $0.24$–$5.84$) of HFMD using data from a series of outbreaks in different calendar years. Previous studies on the estimation of $R_0$ were limited. The isolation strategy against HFMD was not only able to delay the onset of the outbreak but also to decrease the disease burden of HFMD, i.e. the total number of HFMD cases.

---

**Fig. 3 (c, d).** The observed and predicted HFMD cases in Taiwan in (c) 2005; (d) 2008.
Ma et al. reported that the median $R_0$ of EV71 and CVA16 were 5·48 and 2·5, respectively, in Hong Kong [17]. The assumption of only one index case at the initial phase and the neglect of household transmission resulted in the overestimation of the reproductive numbers in that study. In addition, our estimated $R_0$ concurrently denoted the combined effects of HFMD and these viruses. We simulated the occurrence of HFMD throughout the entire year, and we also allowed the initial number of cases to change, not only one index case. Hence, the $R_0$ of HFMD estimated in our study was therefore lower than the value in the Ma et al. study.

One of the unique characteristics of our study was the prediction of HFMD outbreak by different calendar years. The values of beta used were $4·5 \times 10^{-7}$ (<34 weeks) and $1·2 \times 10^{-6}$ (≥34 weeks) for the two periods, but the other parameters were not changed. The two peaks of the epidemic curve could be simulated in the 2005 outbreak. According to the laboratory surveillance data, EV71 was the major cause of these outbreaks. In 2005, the major cause of the first wave was EV71, and the second wave was due to CVA16. Therefore, the different ratios of severe cases to all symptomatic cases ($\delta$) were noted. Similarly, the different transmission rates ($\beta$) in
different periods were also found. Hence, we obtained an asymmetric epidemic curve such as that in Figure 3d for the 2008 outbreak. However, the major cause was still only EV71. This is supported by the findings that the higher ratio of severe cases (δ) is related to the higher activity of EV71.

We observed a poor fit of the predicted HFMD cases in both 2000 and 2001 (Supplementary Table S1). There are two explanations for this. In the early stage, the sentinel surveillance system for HFMD may be unstable because clinical physicians were unfamiliar with the reporting system or the characteristics of severe cases of HFMD. Therefore, the proportion of severe HFMD cases fluctuated greatly in the early outbreak. The second reason is that HFMD has different aetiological viruses in different seasons and years, resulting in a wide range of prevalence figures. Both aspects account for the poor fit of the predicted severe HFMD cases in both 2000 and 2001 (Supplementary Table S1).

The proportion of severe cases (δ) was different in these outbreaks because of the different proportion of the aetiologies. The major aetiologies were EV71 and CVA16 in 2000, 2001 and 2008. The first wave of the outbreak resulted from viruses including CVA16 and EV71 in 2005. The second wave of the outbreak resulted from viruses including CVB3 and EV71 in 2005. From the epidemic curve of severe cases with enterovirus, we believe that EV71 was more active in the first outbreak wave because the

![Graph showing observed and predicted severe HFMD cases in Taiwan in 2005 and 2008.](image)
cases with EV71 had a high mortality rate, which was captured by a higher value of $\delta$.

**Implications for the control of HFMD**

Our proposed model is based not only on the dynamic infectious status but also on dynamic changes in the disease with complications. Incorporating the information about the clinical symptoms or disease complications that were commonly seen in routine surveillance data into the proposed model would make the model as accurate as possible in order to emulate the real scenario on infection as well as disease process. This is useful for forestalling the spread of HFMD according to $R_0$ and predicting the expected sequelae of subsequent outbreaks on the basis of information available from early epidemic curves with the same pathogen of virus variants. Providing such information may be of great benefit to evaluate the effectiveness of certain preventive strategies such as a health isolation strategy for policymakers as shown in the current study.

By the application of these parameters and $R_0$, we demonstrated that the isolation strategy against HFMD delayed the epidemic peak and also reduced the burden of HFMD. The proposed model is also very flexible in evaluating further aggressive preventive action such as quarantine before the development of vaccine. Moreover, our proposed dynamic model provides a good foundation for the future effectiveness and cost-effectiveness analysis of a vaccine against HFMD (EV71, CVA16) is the best policy for preventing severe enterovirus cases.

There were unavoidable biases in estimating the $R_0$ of HFMD because it was limited to the assumptions of the model, such as a homogeneous random mixing population, the fixed ratio of the number of reported cases in sentinel surveillance systems to the number of all cases, and the estimated parameters derived only from the characteristics of CVA16 and EV71. To minimize these biases, we estimated the parameters based on the comprehensive empirical data from infection until the occurrence of clinical complications. More importantly, the credibility of our model is attributed to the fact that there were reliable severe cases with enterovirus reported in the system, which enabled us to estimate the ratio of reported cases in the surveillance system to all cases in Taiwan and to further stabilize the estimation of relevant parameters contained in the model. Furthermore, HFMD has different aetiological causes, and the prevalence of the different viruses is different in different seasons, but we assumed that all of their characteristics were the same as those of EV71 or non-EV71 (CVA16). Another limitation is that the parameters of the model were only estimated from the characteristics of CVA16 and EV71 because we were more concerned with the impact of the severe cases of enterovirus, with their major aetiologies being composed of CVA16 and EV71. In addition, HFMD can be transmitted by asymptomatic persons. However, no studies reported on the difference of transmission probability between symptomatic and asymptomatic persons. For the

![Fig. 5. The results of an isolation strategy for HFMD at different isolation rates.](https://www.cambridge.org/core/core_image)
ease of computation, we assumed that there were no differences between them.

Partly because we are interested in large outbreaks of severe cases of HFMD, and partly because the characteristics of severe cases of HFMD outbreaks depended on the enteroviruses (EV71 or CVA16) in different years, we applied the dynamic epidemic model to fit the surveillance data in each year and not all years. However, bias may be present in the analysis of isolation strategy (Ω) because we did not take seasonal variation into account.

In conclusion, we show how to estimate $R_0$ in the light of the parameters of the proposed dynamic model with accuracy by incorporating the information of Taiwanese sentinel surveillance data and clinical symptoms or disease complications. The estimated $R_0$ was 1:37, suggesting a higher likelihood of the spread of HFMD if no infection control policy is envisaged. Based on our simulation results, we also demonstrated that the isolation strategy against HFMD not only delayed the spread of HFMD with a range from 4 weeks for 20% isolation to 47 weeks for 100% isolation but also reduced total number of HFMD cases with a range from 1:3% reduction for 20% isolation to 13:3% reduction for 100% isolation. The developed model can also be easily used to evaluate the effectiveness of other control strategies including quarantine and vaccination (if the vaccine can be developed) in order to reduce the chance of the spread of HFMD.

SUPPLEMENTARY MATERIAL

For supplementary material accompanying this paper visit http://dx.doi.org/10.1017/S0950268815002630.

DECLARATION OF INTEREST

None.

REFERENCES


