**Meta-analysis**

**Dietary glycaemic index and glycaemic load in relation to the risk of type 2 diabetes: a meta-analysis of prospective cohort studies**

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**Abstract**

Epidemiological studies of dietary glycaemic index (GI) and glycaemic load (GL) in relation to diabetes risk have yielded inconsistent results. We aimed to examine the associations between dietary GI and GL and the risk of type 2 diabetes by conducting a meta-analysis of prospective cohort studies. Relevant studies were identified by a PubMed database search up to February 2011. Reference lists from retrieved articles were also reviewed. We included prospective cohort studies that reported risk estimates with 95% CI for the associations between dietary GI and GL and the risk of type 2 diabetes. Either a fixed- or random-effects model was used to compute the summary relative risk (RR). We identified thirteen prospective cohort studies of dietary GI or GL related to diabetes risk. The summary RR of type 2 diabetes for the highest category of the GI compared with the lowest was 1·16 (95% CI 1·06, 1·26; n = 12), with moderate evidence of heterogeneity (P = 0·02, I² = 50·8 %). For the GL, the summary RR was 1·20 (95% CI 1·11, 1·30; n = 12), with little evidence of heterogeneity (P = 0·10, I² = 34·8 %). No evidence of publication bias was observed. In addition, the associations persisted and remained statistically significant in the sensitivity analyses. In conclusion, the present meta-analysis provides further evidence in support of significantly positive associations between dietary GI and GL and the risk of type 2 diabetes. Reducing the intake of high-GI foods may bring benefits in diabetes prevention.

**Key words:** Diabetes mellitus; Glycaemic index; Glycaemic load; Meta-analysis

Type 2 diabetes has been a growing public health problem worldwide. A large body of evidence has demonstrated that the presence of diabetes doubles the risk of a wide range of vascular diseases. In addition, diabetes is moderately associated with death from cancers of the liver, pancreas, ovary, colorectum, lung, bladder and breast. Preventive strategies are urgently needed to reduce the huge burden of diabetes.

Diet is widely believed to play an important role in the development of type 2 diabetes. Dietary glycaemic index (GI) and glycaemic load (GL) have received considerable attention for their potential contribution to the diabetes epidemic. The GI, introduced by Jenkins et al., ranks the carbohydrate content of individual foods according to their postprandial glycaemic effects. Consumption of high-GI diets is associated with high blood glucose and insulin concentrations, thereby eventually resulting in glucose intolerance and high diabetes risk. The GL, which is the product of the GI of a food item and the available carbohydrate content, quantifies the overall glycaemic effect and insulin demand.

Prospective cohort studies assessing the effects of dietary GI and GL on the risk of type 2 diabetes have yielded inconsistent results. A previous meta-analysis of seven studies found that diets with a high GI or GL significantly increased the risk of type 2 diabetes. During the past few years, the number of original studies linking dietary GI and GL to diabetes has doubled. However, not all confirmed the previous findings, and the controversy on this topic continues. With accumulating evidence, we therefore aimed to examine the associations between dietary GI and GL and the risk of type 2 diabetes by conducting an updated meta-analysis of prospective cohort studies.

**Materials and methods**

**Literature search**

We conducted a PubMed database search up to February 2011 to identify published studies of dietary GI and GL and the
risk of type 2 diabetes, using the search terms 'glycemic index' or 'glycemic load' in combination with 'diabetes'. No restrictions were imposed. In addition, we reviewed the reference lists of retrieved articles. Efforts were made to obtain additional data by contacting original authors.

**Study selection**

Studies were eligible for the present meta-analysis if they met the following criteria: a prospective cohort design, which is less prone to bias than a retrospective one; the exposure of interest was dietary GI or GL; the outcome of interest was incidence of type 2 diabetes; risk estimates and associated 95% CI (or data to calculate them) were reported.

**Data extraction**

The following information was independently extracted by two authors: the first author's last name, publication year; study population, location, length of follow-up; number of cases and participants; assessment of exposure and outcome; most fully adjusted risk estimates with corresponding 95% CI from the multivariable model for each category of exposure or for exposure as a continuous variable; statistical adjustment for the potential confounding factors.

**Statistical analysis**

The relative risk (RR) was used as the common measure of association across studies. For four studies (12,13,15,21) that analysed exposure as a continuous variable, RR for categories of exposure were obtained from original authors or estimated based on the exposure range (i.e. the highest and lowest categories of exposure) reported in the primary studies. A sensitivity analysis was performed by excluding these four studies to test the robustness of the results.

We calculated the $Q$ and $I^2$ statistics to examine statistical heterogeneity across studies. $I^2$ is the proportion of total variation explained by between-study variation (23). Either a fixed-effects model, the weight of each study is equal to 1, whereas in the random-effects model, an extra term is added to the variance according to the method proposed by DerSimonian & Laird (23). In addition, we examined the influence of a single study on the combined risk estimates by omitting one study and analysing the remainders in each turn. Potential publication bias was assessed by visual inspection of Begg’s funnel plots and by the Begg rank correlation and Egger linear regression tests (24,25). All analyses were performed with the use of STATA version 11.0 (StataCorp, College Station, TX, USA). All statistical tests were two-sided and $P<0.05$ was considered statistically significant, except where otherwise specified.

**Results**

**Study characteristics**

We identified sixteen potentially relevant prospective cohort studies of GI and GL and type 2 diabetes. Of these sixteen studies, three were excluded because the outcome was gestational diabetes mellitus (27) or because of overlapping publications from the same study population (28,29), thus leaving thirteen studies (23–26) (twelve for GI and twelve for GL) for the final analysis. Characteristics of the selected studies are presented in Table 1. The association between GI and GL and diabetes risk was the primary outcome in all primary studies. The thirteen prospective cohort studies were published between 1997 and 2010. Among them, eight studies were conducted in the USA, two in Europe, two in Australia and one in China. The number of cases diagnosed in the original studies ranged from 138 to 8587, and the size of the cohort ranged from 1833 to 91,249. Of these thirteen studies, seven enrolled both sexes, one included men only and five included women only. The median length of follow-up ranged from 4 to 14 years. All studies used FFQ in dietary assessment. Diabetes ascertainment was based on self-report of physician diagnosis in most studies, but the majority of cases were confirmed in validation studies. The major confounding factors of interest included age, BMI, physical activity, smoking, and intakes of total energy, alcohol and dietary fibre.

**Glycaemic index, glycaemic load and the risk of type 2 diabetes**

The multivariable-adjusted RR for each study and the combined RR comparing the highest with the lowest categories of the GI and GL are presented in Figs. 1 and 2. Among the included studies, six studies found a statistically significant association between GI and an increased risk of type 2 diabetes. The GI was significantly associated with an increased diabetes risk in four studies. The summary RR of type 2 diabetes for the highest category of GI compared with the lowest was 1.16 (95% CI 1.06, 1.26; $n$ 12), with moderate evidence of heterogeneity ($P=0.05, I^2=50.8\%$). For the GL, the summary RR was 1.20 (95% CI 1.11, 1.30; $n$ 12), with little evidence of heterogeneity ($P=0.10, I^2=34.8\%$). Visual inspection of Begg’s funnel plots did not show important asymmetry for either exposure. The Begg rank correlation and Egger linear regression tests did not indicate evidence of publication bias (all $P>0.30$).

**Sensitivity analysis**

We also performed sensitivity analyses to test the robustness of the findings. Restricting analysis to studies that adjusted for total fibre or cereal fibre yielded a summary RR of 1.14 (95% CI 1.03, 1.26; $n$ 10) for the GI and 1.16 (95% CI 1.05, 1.29; $n$ 9) for the GL. Restricting analysis to studies that were conducted in the USA yielded a summary RR of 1.17 (95% CI 1.02, 1.34; $n$ 7) for the GI and 1.19 (95% CI 1.08, 1.31; $n$ 8) for the GL. Restricting analysis to studies that analysed exposure as categorical variables yielded a summary...
<table>
<thead>
<tr>
<th>Study</th>
<th>Population (cases)</th>
<th>Duration (years)</th>
<th>Dietary assessment</th>
<th>Case ascertainment</th>
<th>Exposure</th>
<th>Adjustments</th>
</tr>
</thead>
<tbody>
<tr>
<td>Salmeron et al. (1997)(9)</td>
<td>42 759 men aged 40–75 years, USA (523)</td>
<td>6</td>
<td>FFQ</td>
<td>Confirmed self-report</td>
<td>GI and GL</td>
<td>Age, BMI, smoking, physical activity, family history of diabetes, and intakes of alcohol, total energy and cereal fibre</td>
</tr>
<tr>
<td>Salmeron et al. (1997)(10)</td>
<td>65 173 women aged 40–65 years, USA (951)</td>
<td>6</td>
<td>FFQ</td>
<td>Confirmed self-report</td>
<td>GI and GL</td>
<td>Age, BMI, smoking, physical activity, family history of diabetes, and intakes of alcohol, total energy and cereal fibre</td>
</tr>
<tr>
<td>Meyer et al. (2000)(11)</td>
<td>35 988 women aged 55–69 years, USA (1141)</td>
<td>6</td>
<td>FFQ</td>
<td>Confirmed self-report</td>
<td>GI and GL</td>
<td>Age, BMI, WHR, education, smoking, physical activity, and intakes of alcohol, total energy and total fibre</td>
</tr>
<tr>
<td>Stevens et al. (2002)(12)</td>
<td>12 251 adults aged 45–64 years, USA (1447)</td>
<td>9</td>
<td>FFQ</td>
<td>Confirmed self-report</td>
<td>GI and GL</td>
<td>Age, BMI, sex, field centre, education, smoking, physical activity and cereal fibre intake</td>
</tr>
<tr>
<td>Hodge et al. (2004)(13)</td>
<td>34 641 adults aged 40–69 years, Australia (365)</td>
<td>4</td>
<td>FFQ</td>
<td>Confirmed self-report</td>
<td>GI and GL</td>
<td>Age, BMI, WHR, weight change, education, sex, country of birth, physical activity, family history of diabetes, and intakes of alcohol and total energy</td>
</tr>
<tr>
<td>Schulze et al. (2004)(14)</td>
<td>91 249 women aged 26–46 years, USA (741)</td>
<td>8</td>
<td>FFQ</td>
<td>Confirmed self-report</td>
<td>GI and GL</td>
<td>Age, BMI, physical activity, history of high blood pressure, oral contraceptive use, hormone therapy, family history of diabetes, and intakes of alcohol, fat, total energy, cereal fibre, caffeine and Mg</td>
</tr>
<tr>
<td>Barclay et al. (2007)(15)</td>
<td>1833 adults aged &gt;49 years, Australia (138)</td>
<td>10</td>
<td>FFQ</td>
<td>Glucose levels, use of diabetic medication and self-report</td>
<td>GI</td>
<td>Age, BMI, sex, family history of diabetes, smoking, physical activity, TAG, HDL and vegetable fibre intake</td>
</tr>
<tr>
<td>Krishnan et al. (2007)(16)</td>
<td>59 000 women aged 21–69 years, USA (1938)</td>
<td>8</td>
<td>FFQ</td>
<td>Confirmed self-report</td>
<td>GI and GL</td>
<td>Age, BMI, family history of diabetes, smoking, physical activity, and intakes of cereal fibre, protein, total fat and total energy</td>
</tr>
<tr>
<td>Mosdol et al. (2007)(17)</td>
<td>7321 adults aged 39–63 years, UK (329)</td>
<td>13</td>
<td>FFQ</td>
<td>Glucose levels, use of diabetic medication and self-report</td>
<td>GI and GL</td>
<td>Age, BMI, WHR, sex, employment, physical activity, smoking, and intakes of alcohol, fibre, carbohydrates and total energy</td>
</tr>
<tr>
<td>Villegas et al. (2007)(18)</td>
<td>64 227 women aged 40–70 years, China (1808)</td>
<td>4-6</td>
<td>FFQ</td>
<td>Confirmed self-report</td>
<td>GI and GL</td>
<td>Age, WHR, smoking, physical activity, income, education, occupation, hypertension, and intakes of alcohol and total energy</td>
</tr>
<tr>
<td>Sahyoun et al. (2008)(19)</td>
<td>3075 adults aged 70–79 years, USA (99)</td>
<td>4</td>
<td>FFQ</td>
<td>Glucose levels, use of diabetic medication and self-report</td>
<td>GI and GL</td>
<td>Age, BMI, sex, race, clinical site, education, physical activity, smoking, baseline fasting glucose, and intakes of alcohol and cereal fibre</td>
</tr>
<tr>
<td>Hopping et al. (2010)(20)</td>
<td>75 512 adults aged 45–75 years, USA (8587)</td>
<td>14</td>
<td>FFQ</td>
<td>Glucose levels, use of diabetic medication and self-report</td>
<td>GL</td>
<td>BMI, physical activity, education, race and total energy intake</td>
</tr>
<tr>
<td>Slujs et al. (2010)(21)</td>
<td>37 846 adults aged 21–70 years, The Netherlands (915)</td>
<td>10</td>
<td>FFQ</td>
<td>Confirmed self-report</td>
<td>GI and GL</td>
<td>Age, BMI, waist circumference, sex, physical activity, smoking, mean systolic blood pressure, education, family history of diabetes, and intakes of total energy, vitamin C, vitamin E, protein, fat, alcohol and total fibre</td>
</tr>
</tbody>
</table>

WHR, waist:hip ratio.
RR of 1.19 (95% CI 1.04, 1.36; n 8) for the GI and 1.22 (95% CI 1.10, 1.36; n 9) for the GL. Further analyses investigating the influence of a single study on the overall risk estimate by omitting one study in each turn yielded a narrow range of RR from 1.13 (95% CI 1.04, 1.22) to 1.18 (95% CI 1.09, 1.29) for the GI and from 1.18 (95% CI 1.09, 1.28) to 1.23 (95% CI 1.15, 1.32) for the GL.

Discussion

The present meta-analysis of prospective cohort studies provides further evidence that higher dietary GI and GL increases the risk of type 2 diabetes. The highest GI and GL exposure compared with the lowest was associated with a 16 and 20% increase in diabetes risk, respectively. In addition, the associations persisted and remained statistically significant in the sensitivity analyses.

A major strength of the meta-analysis is that the present findings are based on prospective cohort studies. This minimises the possibility of recall and selection biases, which are always of concern in retrospective studies. Compared with the previous meta-analysis(22), the risk estimates reported in the present study were a bit smaller. However, with accumulating evidence and enlarged sample size, we have enhanced statistical power to provide more precise and reliable risk estimates relating dietary GI and GL to diabetes risk.

Limitations of the present meta-analysis should also be acknowledged while interpreting the results. First, residual confounding is always of concern in observational studies. Although several important potential confounding factors, including age, BMI, physical activity and smoking, have been widely controlled in original studies, dietary factors were not sufficiently considered. For instance, higher GI was found to be correlated with lower intake of Mg(9,10,14), and mounting evidence has suggested that lower Mg intake increases the risk of type 2 diabetes(30). Yet, few studies(14) have adjusted for Mg in the multivariable analysis. In addition, it is possible that a low-GI diet is rich in plant foods containing many phytochemicals that may protect against diabetes, which could result in attenuated associations. Therefore, we could not exclude the likelihood that unmeasured or inaccurately measured factors may be responsible for the findings.

Second, misclassification error and bias may have weakened the strength of the associations. Misclassification of dietary assessment is not avoidable and misclassification of diabetes cases was likely to occur given that the majority of diabetes ascertainment were based on self-reports. In addition, most cohort studies measured dietary intakes at baseline only, and the lack of repeated dietary assessment during the follow-up period could also result in misclassification. Third, we observed heterogeneity among studies of GI and diabetes risk. This may be partly explained by variations in characteristics of study populations, measurements of dietary intakes and ranges of dietary GI exposure. We attempted to explore possible sources of variations in the results using sensitivity analyses, but significant heterogeneity remained unsatisfactorily explained. Nevertheless, differences in risk estimates among individual studies were mainly in the magnitude rather than the direction of the association. Fourth, because GI were highly correlated with carbohydrate intake,
Glycaemic index and risk of type 2 diabetes

<table>
<thead>
<tr>
<th>Study</th>
<th>RR</th>
<th>95% CI</th>
<th>Weight (%)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Salmeron et al. (1997, Female)</td>
<td>1·25</td>
<td>0·90, 1·73</td>
<td>4·71</td>
</tr>
<tr>
<td>Salmeron et al. (1997, Male)</td>
<td>1·47</td>
<td>1·16, 1·86</td>
<td>7·59</td>
</tr>
<tr>
<td>Meyer et al. (2000)</td>
<td>0·95</td>
<td>0·78, 1·16</td>
<td>9·43</td>
</tr>
<tr>
<td>Stevens et al. (2002, White)</td>
<td>1·10</td>
<td>0·89, 1·36</td>
<td>8·53</td>
</tr>
<tr>
<td>Stevens et al. (2002, Black)</td>
<td>0·97</td>
<td>0·71, 1·32</td>
<td>5·12</td>
</tr>
<tr>
<td>Hodge et al. (2004)</td>
<td>0·91</td>
<td>0·52, 1·59</td>
<td>1·87</td>
</tr>
<tr>
<td>Schulze et al. (2004)</td>
<td>1·33</td>
<td>0·92, 1·92</td>
<td>3·89</td>
</tr>
<tr>
<td>Krishnan et al. (2007)</td>
<td>1·22</td>
<td>0·98, 1·51</td>
<td>8·50</td>
</tr>
<tr>
<td>Mosdol et al. (2007)</td>
<td>0·80</td>
<td>0·51, 1·26</td>
<td>2·73</td>
</tr>
<tr>
<td>Villegas et al. (2007)</td>
<td>1·34</td>
<td>1·13, 1·58</td>
<td>11·34</td>
</tr>
<tr>
<td>Sahyoun et al. (2008)</td>
<td>1·30</td>
<td>0·61, 2·76</td>
<td>1·07</td>
</tr>
<tr>
<td>Hopping et al. (2010, Male)</td>
<td>1·16</td>
<td>0·99, 1·36</td>
<td>11·95</td>
</tr>
<tr>
<td>Hopping et al. (2010, Female)</td>
<td>1·41</td>
<td>1·15, 1·73</td>
<td>9·12</td>
</tr>
<tr>
<td>Slujs et al. (2010)</td>
<td>1·27</td>
<td>1·12, 1·45</td>
<td>14·16</td>
</tr>
<tr>
<td>Overall (I² = 34·8 %, P = 0·104)</td>
<td>1·20</td>
<td>1·11, 1·30</td>
<td>100·00</td>
</tr>
</tbody>
</table>

Fig. 2. Relative risks (RR) for the association between dietary glycaemic load and type 2 diabetes in prospective cohort studies.

it was difficult to determine to what extent the effect of GL on diabetes risk was explained by carbohydrate intake. Finally, publication bias could affect results of any meta-analyses. Yet formal statistical tests suggested little evidence of this bias in the present meta-analysis.

The exact mechanisms by which high-GI foods may increase the risk of type 2 diabetes are uncertain. A high-GI diet produces high concentrations of blood glucose and increased insulin demand. This could result in glucose intolerance and insulin resistance and eventually lead to a higher risk of type 2 diabetes. Furthermore, a high-GI diet could increase postprandial NEFA release, thereby directly increasing insulin resistance. In addition, there is evidence that higher dietary GI is positively associated with subsequent gain in waist circumference, which is a strong predictor of insulin resistance and diabetes risk.

On the other hand, metabolic studies suggest that a low-GI diet may improve glycaemic control. A recent meta-analysis of seven randomised controlled trials conducted in diabetes has shown that a low-GI diet, compared with a control diet, significantly decreased glycated HbA1c, a long-term measure of blood glucose levels. Although the reported decrease in glycated HbA1c was only 0·4 % (95 % CI 0·2, 0·7) in that meta-analysis, such a reduction has potential clinical importance as lowering glycated HbA1c by 0·6 % was demonstrated to contribute to risk reductions of 32 % for diabetes-related clinical endpoints, 42 % for diabetes-related deaths and 36 % for all-cause mortality.

The adverse effects of a high-GI diet may increase with an individual's underlying degree of insulin resistance. Thus, several factors related to insulin resistance as well as diabetes risk, such as cereal fibre intake, physical activity and BMI, may have impacts on the associations between GI and GL and diabetes risk. Epidemiological studies have provided compelling evidence that lower cereal fibre intake is associated with an increased diabetes risk. Previous studies have consistently suggested that the positive associations between GI and GL and diabetes risk were more pronounced among those with a lower intake of cereal fibre, with significant interaction. Physical activity has strong effects on glucose tolerance and insulin sensitivity and can significantly reduce the risk of developing type 2 diabetes. As expected, the associations of GI and GL with diabetes risk appeared to be more evident in participants with lower BMI. Of note, the numbers of cases were relatively small in these stratified analyses, and therefore chance may, at least in part, account for these findings. Nevertheless, as potential effect modifiers, these variables need to be closely considered in subsequent studies for a better understanding of the relationships of dietary GI and GL to diabetes risk.

In conclusion, the present meta-analysis of prospective cohort studies provides further evidence in support of significantly positive associations between dietary GI and GL and the risk of type 2 diabetes. Given the wide exposure and the huge burden of diabetes worldwide, reducing the intake of high-GI foods, particularly refined carbohydrates, may bring potential benefits in diabetes prevention among the general population.
Acknowledgements

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References


