Special Issue Article

Childhood inhibition predicts adolescent social anxiety: Findings from a longitudinal twin study

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Abstract

An enduring issue in the study of mental health is identifying developmental processes that explain how childhood characteristics progress to maladaptive forms. We examine the role that behavioral inhibition (BI) has on social anxiety (SA) during adolescence in 868 families of twins assessed at ages 8, 13, and 15 years. Multimodal assessments of BI and SA were completed at each phase, with additional measures (e.g., parenting stress) for parents and twins. Analyses were conducted in several steps: first, we used a cross-lagged panel model to demonstrate bidirectional paths between BI and SA; second, a biometric Cholesky decomposition showed that both genetic and environmental influences on childhood BI also affect adolescent SA; next, multilevel phenotypic models tested moderation effects between BI and SA. We tested seven potential moderators of the BI to SA prediction in individual models and included only those that emerged as significant in a final conditional model examining predictors of SA. Though several main effects emerged as significant, only parenting stress had a significant interaction with BI to predict SA, highlighting the importance of environmental moderators in models examining temperamental effects on later psychological symptoms. This comprehensive assessment continues to build the prototype for such developmental psychopathology models.

Keywords: behavior inhibition; genetics; parenting; social anxiety; twins

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Introduction

An enduring issue in the mental health field is identifying mechanisms and developmental processes that explain how childhood characteristics progress to maladaptive adolescent and adult forms. Because of the research attention devoted to it, behavioral inhibition (BI) is currently our best example of a temperamental trait that biases reactions to later stressors in a way that can result in maladaptive behavioral patterns. In general, temperament reflects individual differences underlying affective reactivity and regulation (Goldsmith et al., 1987). More specifically, research on BI, defined as reticence or withdrawal in the face of novelty or threat, illustrates the complexity that challenges developmental psychopathology research. Summarizing a large literature, we observe that BI is conceptualized as a category or type (e.g., Kagan et al., 1984) or as a dimension (Trull & Durrett, 2005), with more evidence favoring dimensions (Haslam et al., 2012). With the dimensional approach, BI to social vs. nonsocial stimuli is usefully distinguished (Dyson et al., 2011), as is BI to threatening vs. everyday stimuli (Buss, 2011). Individual differences in BI are moderately but not highly stable across major developmental transitions (Pfeifer et al., 2002). Variation in BI predicts later social anxiety (SA; Pérez-Edgar & Guyer, 2014), and crucially, chronically high BI predicts susceptibility to SA (Hirshfeld et al., 1992; Poole et al., 2020). Indeed, anxiety disorders as a broader group were the sixth leading cause of disability worldwide in 2010. The burden is disproportionately borne by 15–34 year-olds (Baxter et al., 2014), where 8.6% of adolescents and 13% of adults meet diagnostic criteria for Social Anxiety Disorder (SA; Kessler et al., 2012). However, investigation of links between BI and SA across adolescence, attending to factors which contribute to genetic and experiential factors, is limited. Social factors such as parenting style and peer relationships—which we analyze in this paper—can moderate BI’s association with later anxiety (Frenkel et al., 2015). Thus, the investigation of how BI sometimes predisposes for the emergence of SA has assumed a prototypic role. BI provides the prototype for future bio-behavioral investigations that involve other temperament dimensions, complex genetic and environmental interplay that continues over development, and psychiatric outcomes, a point also highlighted by Pérez-Edgar and Fox (Pérez-Edgar & Fox, 2018). Here, we study the full range of SA, from very low to clinically significant levels, during a period when it can crucially impact adolescent functioning.

Overview of the phenotypic questions

Here, we conceptualize our focal predictor, BI, as a dimension rather than as a category, and our analyses all embrace a dimensional perspective. We are among many theorists who have treated
the measurement of BI empirically (Goldsmith & Lemery, 2000; Pfeifer et al., 2002). We will not reiterate the complexities of the types vs. dimensions debate, which also pertains to personality traits more generally and to dimensional views of psychopathology, such as the Research Domain Criteria (RDoC; Cuthbert & Insel, 2013) and the Hierarchical Taxonomy of Psychopathology (HiTOP; Kotov et al., 2017). This paper considers phenotypic and genotypic issues relating to BI and its developmental association with SA in the understudied developmental period from middle childhood into adolescence.

**In what senses does BI predict SA?**

The extant literature affirmatively answers the basic question of whether BI predicts SA. One meta-analysis shows a greater than seven-fold increase in SA specifically for BI versus non-BI individuals (Clauss & Blackford, 2012). In another study with a different sample but similar measures, we showed that a multisource measure of shyness at age 3 predicted a broader anxiety measure at 8 years (Volbrecht & Goldsmith, 2010). As we consider this issue, embedded questions become apparent. A key issue is whether single instance assessment of BI is a theoretically justified predictor in analyses of BI-to-SA associations. Although associations of single instance BI assessment might be associated with later anxiety, most theorizing about BI treats it as a trait-like dimension, so that mean levels of BI, when high, imply that chronic BI is likely. Although a single assessment of BI might sample a trait-like dimension, state effects might also contaminate or even dominate the assessment. Thus, incorporating evidence of both mean levels across time and multimodal assessment directly into the BI predictor seems a preferable strategy.

**Continuity and directionality of BI-to-SA predictions**

Predicting SA from earlier BI requires analysis of the continuity of both BI and SA over the period of prediction as well as an understanding of the directionality between these associations. Thus, an underlying model of longitudinal continuity and change must be adopted, at least implicitly. The most common and simplest analysis of continuity for our context is one that uses a cross-lagged panel approach, which provides information on continuity of both BI and SA, along with longitudinal and concurrent prediction of BI-to-SA and SA-to-BI: Does BI predict SA, or might SA predict future BI? Here, the prediction of BI for SA could be viewed as consistent with the vulnerability model, whereas the possible prediction of SA from BI would fit the complication or scar model (Tackett, 2006), which has mainly been studied for adult depression and substance use and not in the childhood internalizing domain.

**Temperament-psychopathology associations of BI-to-SA**

In addition to models of continuity, another deeper issue is the models for temperament-psychopathology associations, which largely derive from the personality-psychopathology models. Three primary hypotheses depict relations between temperament and psychopathology (Clark et al., 1994; Tackett, 2006). Vulnerability models posit that certain types of temperament predispose individuals to psychopathology. Complication or scar models raise the possibility of the opposite direction-of-effect, such that emerging psychopathology changes temperament (Bianconcini & Bollen, 2018). Finally, spectrum models suggest that corresponding aspects of temperament and psychopathology lie along the same continuum and share underlying causal processes. The spectrum model has garnered the most support (Markon, 2019) although the models are not mutually exclusive, and the other models may explain relations between specific traits and disorders (Gartstein et al., 2012).

**Specificity of behavioral inhibition and SA associations**

Another question addresses the specificity of the BI-to-SA prediction: does BI predict SA specifically, or is it a more general risk factor for anxiety or for other symptom dimensions such as depression or externalizing behaviors? The specificity hypothesis is risky, as temperamental traits often predict in a transdiagnostic manner (Hankin et al., 2017; Kostyrka-Allchorne et al., 2020; Morales et al., 2021). Furthermore, SA is associated with increased risk for other psychiatric diagnoses, particularly mood, other anxiety, and substance use disorders. SA predicts increased severity of comorbid depression (Beesdo et al., 2007; Fehm et al., 2008); 87.8% of adults with SA had additional psychiatric diagnoses during a 12-month period (Fehm et al., 2008). When this striking comorbidity is considered alongside the significant impairment that accompanies subdiagnostic threshold SA, the overall impact of SA is tremendous (National Collaborating Centre for Mental Health (UK), 2013). However, a recent, questionnaire-based twin study supported the specificity hypothesis by showing that retrospectively reported childhood BI specifically predicted adolescent SA (and not other internalizing disorders); BI shared both genetic (20%) and environmental (16%) variance with adolescent anxiety (Bourdon et al., 2019), a finding that foreshadows one of our results.

**Moderation of the effects of BI on SA**

We can conceptualize any “third variable” (e.g., sex of child) that moderates the prediction of SA from BI as a source of heterogeneity that alters, by either weakening or strengthening, BI-to-SA links. Rather than deriving potential moderators from a specific theoretical framework, we conducted a literature review from which seven possible sources of heterogeneity were identified: sex of child, over-protective parenting, peer victimization, socioeconomic status of the family, family stress, pubertal timing, and parent internalizing psychopathology. These seven variables vary in nature. Over-protective parenting and peer victimization are specific, proximal measures of experience within and outside the family environment, respectively. Family stress and SES are more general, more distal, and likely correlated variables that are often examined in the literature. Child’s pubertal timing and mothers’ and fathers’ internalizing psychopathology likely reflect both genetic and experiential effects. However, none of these potential moderators are unambiguously biological/genetic or environmental/experiential in nature. What findings in the literature empirically justify inclusion of these potential moderators in our analyses?

**Sex of child**

Sex differences in the development of SA are established, with women having both greater lifetime prevalence and severity of SA than men (Asher et al., 2017). These differences in prevalence are apparent both during adolescence (Merikangas et al., 2010) and adulthood (Xu et al., 2012), and in both community samples (Beesdo et al., 2007) and clinical samples (Yonkers et al., 2003). Similarly, BI research, including our own, indicates higher BI and anxiety for females in late childhood and adolescence.
Parenting and family dynamics play an important role in child development. Specifically, overprotective parenting can have adverse effects on emerging patterns of psychopathology, particularly relating to SA (McLeod et al., 2007; Spokas & Heimberg, 2009). Overprotective parenting includes aspects of parental behaviors that are both warm and controlling, yet excessive in the context of the child’s developmental stage; parents do not allow their child the space to engage in appropriate emotional or regulatory behaviors on their own (Degnan et al., 2008, 2010; Kiel & Maack, 2012).

The literature provides strong evidence that parenting stress is particularly salient in children’s emerging psychopathology. Worried or anxious parents may be more likely to exhibit overprotective parenting, which in turn may increase risk for child anxiety symptoms and disorders; however, findings about this association are mixed and focus predominantly on maternal parenting (Jones et al., 2021; Moore et al., 2004). Anxious mothers may be hypervigilant to perceived threats in their child’s environment and may seek to reduce their own anxious distress by exerting control over their child's behaviors or environment (Eysenck, 2013; Lindhout et al., 2006; Turner et al., 2003).

As a consequence of the issues just discussed, children of overprotective parents may not develop the skills essential for healthy coping or resilience (Van Petegem et al., 2021). In early childhood, maternal overprotective or controlling parenting is associated with increased child fear at 2 years of age (Kiel & Maack, 2012), child inhibition at 4 years of age (Rubin et al., 2002), and higher inhibition in middle school (Degnan et al., 2008). Further, mothers’ overprotective parenting moderates relations between shyness and internalizing and poor social interactions in kindergarteners (Coplan et al., 2008). Thus, we anticipate that mothers’ overprotective parenting will enhance associations between child early BI and emerging SA.

Parent internalizing psychopathology & wellbeing

Elevations in parental negative affect and lower levels of positive affect associated with internalizing diagnoses may shape risk for BI during childhood. Although diagnostic assessments of internalizing psychopathology tend to capture high negative affect effectively, they often do not accurately reflect low positive affect (Watson et al., 1988). Thus, assessing how both increased negative affect (e.g., presence of anxiety disorder) and decreased positive affect (e.g., lower wellbeing score) among parents may influence trajectories of child behavioral inhibition is essential. Specifically, in a sample of inhibited preschool aged children, elevations in parental negative affect were associated with subsequent emergence of clinically significant child anxiety symptoms (Bayer et al., 2021). Although parental depression and panic disorder relate individually to increased risk for childhood BI, the combination of parental depression and panic disorder is associated with an even larger increase in risk for BI (Rosenbaum et al., 2000). The presence of multiple parental anxiety disorders appears to increase risk similarly. For instance, parents of BI children who also qualify for an anxiety disorder diagnosis are significantly more likely to have at least two anxiety disorders of their own (Rosenbaum et al., 1992). Few studies emphasize the possible differential effects of father and mother internalizing problems, as we do in this paper.

Socioeconomic status of the family

Despite being a more distal factor than specific parenting behaviors, the impact of socioeconomic resources (SES) may increase risk for developing anxiety throughout childhood. In a study of 9700 Norwegian children aged 10–13, SA and other mental health problems were higher in a low SES group compared with a higher SES group, particularly when children were asked to perform in front of others (Karlsen et al., 2014). Here, we operationalize SES using two indicators of parents’ education and occupation. Education, rather than income, is an appropriate proxy variable for SES in studies that address family issues relating to child and adolescent mental health (McLaughlin et al., 2012). Using SES in the analysis also helps us connect our findings with other studies.

Peer victimization

Peer victimization is an experience that may complicate associations between BI and SA. Peer victimization refers to the repeated experience of being a target of intentionally aggressive behavior by peers (Oliveus, 2001). Behaviorally inhibited youth experience heightened rates of peer victimization and ostracism (Affrunti et al., 2014; Deater-Deckard, 2001; Hanish & Guerra, 2004; Kochenderfer-Ladd, 2003; Newcomb et al., 1993; Sugimura & Rudolph, 2012), and problems with peers during childhood and adolescence (Booth-Laforce et al., 2012; Hasenfratz et al., 2015). Peer victimization experience in youth is associated with higher rates of SA (Arseneault et al., 2008), as well as other internalizing and externalizing symptoms (Reijnjes et al., 2011). These adverse outcomes are not confined to childhood and adolescence; multiple studies demonstrate that the effects of peer victimization on psychopathology persist into adulthood. Given increased rates of peer victimization amongst higher BI youth, and the associations between peer victimization and SA in childhood and adolescence, including peer victimization as a potential moderator examining associations between BI and SA is strongly justified.

Pubertal timing

Limited evidence in adolescence indicates that pubertal development may have a mechanistic role in the BI–SA association. Early pubertal timing (mature pubertal status relative to age-matched youth) is associated with increased SA (Blumenthal et al., 2009) as well as other mood and anxiety disorders in adolescence (Black & Rofey, 2018). A study of 8–14-year-old girls found that mature pubertal status was associated with altered habituation to threat and safety cues during fear conditioning—demonstrating a dampened response to threat (Jackson et al., 2017). Although the underlying pubertal mechanism seems unclear, developmental maturation appears to play a role in fear extinction; thus, more investigation on puberty’s role in the BI-to-SA association is warranted.
Overview of the genetic considerations

Are BI and SA heritable phenotypes and do they share genetic underpinnings?

The literature based on differential similarity of different classes of genetic kin (e.g., members of MZ vs. DZ twin pairs) confirms a positive answer to this question. However, BI and SA are not highly heritable phenotypes, which means that their study should include frequently occurring experiences in conjunction with genetics. Background for the behavior-genetic analysis is provided by twin studies that have documented moderate genetic effects on variation in early BI (Cherny et al., 1994; Matheny, 1989) as well as studies that have documented moderate genetic effects on variation in early BI (Cherny et al., 1994; Matheny, 1989) as well as on continuity of BI through early development (Smith et al., 2012). We previously showed, in a sample that overlaps the one in the current analyses, that parent-rated and observational assessments of social and non-social fear in school-age children were genetically distinct from anger and sadness (Clifford et al., 2015). Both continuity and change in the development of BI from adolescence to adulthood were heritable in a questionnaire-based twin study (Takahashi et al., 2007). The stability of both BI and anxiety is largely driven by genetic factors, though non-shared environmental influences play a role as well (Ellis & Rothbart, 2001; Hannigan et al., 2017). A recent twin study showed that BI specifically predicted adolescent SA (and not other internalizing disorders); BI shared both genetic (20%) and environmental (16%) variance with adolescent anxiety (Bourdon et al., 2019). A two-wave twin study of late adolescents/young adults using the BIS/BAS scales demonstrated that the genetic variance in BIS scores accounted for genetic effects on both depression and generalized anxiety (measured later), whereas the genetic variance in the BAS scores predicted only depression (Takahashi et al., 2021).

Thus, biometric studies implicate genetic and non-shared environmental factors in the BI-SA association, but a study such as ours with a large sample size, multimodal assessment, and more than two ages studied is greatly needed to solidify our understanding.

Turning to SA, three meta-analyses (Hettema et al., 2001; Moreno et al., 2016; Scaini et al., 2014) examine the heritability of SA measured in a variety of ways across a variety of clinical and non-clinical samples. Low to moderate heritability estimates are typical in this literature, which extends beyond twin studies to other kinship designs. Yet evidence for genetic effects on SA are not limited to quantitative analyses of kinship similarity data. Genome-wide association studies on anxiety more broadly (Levey et al., 2020) and SA specifically (Stein et al., 2017) are being reported. In the first GWAS study of SA, Stein et al. (2017) estimated a SNP-based heritability of 12%. (Such SNP-based heritability estimates cannot be compared directly with the kinship model-based estimates reviewed above.) Meier et al. (2019) examined anxiety and stress-related phenotypes more broadly in an GWAS analysis; they estimated a SNP-based heritability of 28%. Levey et al. (2020), studying generalized anxiety as a dimensional trait, identified several genes to be associated with anxiety in a sample of 200,000 participants, and we expect future large-scale GWAS studies of SA more specifically from a consortium working on the issue.

In summary, multi-method evidence shows that unexplained non-genetic variance in SA remains to be accounted for by factors that we did not study here, but which may contribute to developing SA. Here, we advance genetic analyses by starting with bivariate biometric analyses that ask whether childhood BI’s effect on adolescent SA has a partially genetic basis and whether residual genetic variance remains for SA after the genetic contribution from BI is considered.

Approach to research questions

Although a link between BI and later SA is established (Eley et al., 2015; Kessler et al., 2005; Klein & Mumper, 2018; Paulus et al., 2015; Sylvester & Pine, 2018; White et al., 2017), understanding the complexities and contingencies in this link remains a challenge. We still lack a comprehensive understanding of the etiology and developmental course(s) of BI. Ambiguity about heterogeneity of BI effects on later development precludes clear answers to both etiological and developmental questions.

In brief, we examine whether the anticipated prediction from BI to later dimensional SA is (1) specific to SA (relative to other internalizing symptoms and externalizing symptoms) and whether associations between BI and SA were unidirectional (BI-to-SA only). To test the latter, we use a random intercept cross-lagged panel model (RI-CLPM), which also offers autoregressive paths representing continuity of BI and SA, along with their cross-prediction from middle childhood, through early adolescence, to our final assessment later in adolescence. This CLPM allows us to preliminarily examine the BI-SA and SA-BI paths that are also probed in the subsequent bivariate biometric models and multilevel models. Notably, because of recent criticisms of classic CLPM models, we followed recommendations to estimate a RI-CLPM (Berry & Willoughby, 2017). With the pattern of BI-to-SA association characterized, we interrogate genetic and environmental underpinnings of association using a bivariate biometric model. Then, we screen potential moderators for their interactive effects in the BI-to-SA prediction. The moderators that pass this screen are incorporated into a final model of the predictive nexus.

Method

Overview

The data were collected over two decades in a longitudinal project. Although BI and SA figured prominently into the design, its scope included other aspects of temperament and emotionality, as well as risk and resilience factors for child and adolescent psychopathology more generally (Schmidt, Brooker, et al., 2019; Schmidt, Lemery-Chalfant, et al., 2019). The project was funded by different mechanisms, each with their differing priorities and differing resources for follow-up. Although participants in later stages of the study were only recruited from those with earlier data, not all twin pairs could be followed up. The reduction in sample size over ages mainly reflects intentional reductions due to available funds and sub-project eligibility rather than participants declining to continue. More detail is provided in published project overviews (Schmidt et al., 2013).

Participants

Participants were recruited from statewide birth records. Average ages for the assessments were 8, 13, and 15 years, with an average sample size of 1400 individuals (700 twin pairs), plus parents. The final sample consisted of n = 1735 individual children (from 868 families) with available data in middle childhood, early, and later adolescence. All participants have data from at least one of the three time points. Figure 1 shows the sample size at different ages and available constructs. The variation in age at each assessment occasion was due to many factors (e.g., staff availability for assessment visits around the state, funding, delay for metal dental braces to be removed to allow MRI scans).

Parents (94.3% mothers) reported on their ethnic and racial identities during each visit. 91.9% were non-Hispanic/Latinx,
4.7% were Hispanic/Latinx, with 3.4% missing data. 88.0% of twins were White, 4.5% were more than one race, 3.9% were Black, 1.1% were other race, 0.5% were Native American, 0.1% were Hmong, 0.1% were other Island/Pacific Islander, and 1.7% were missing data on race. Parents were asked their own level of education indexed by number of years of school completed, ranging from 6 (grade school) to 20+ (graduate degree). At the initial Age 8 phase, mothers had completed an average of 14.36 (SD = 2.81) years of schooling, and fathers had completed an average of 14.82 (SD = 2.64) years of schooling. Median family income was in the range from $60,001 to $70,000.

Zygosity was initially diagnosed with the Zygosity Questionnaire for Young Twins (Goldsmith, 1991). When needed, we also used hospital placentation records, and observer ratings of details of physical similarity from videotapes. Pairs with ambiguous zygosity were genotyped. The sample includes 612 MZ individuals, 512 members of a same-sex dizygotic-twin pair, and 488 members of an opposite-sex dizygotic-twin pair. We analyzed all individuals with available data, regardless of whether the cotwin were available for a particular variable.

Measures

Behavior inhibition and social anxiety

BI and SA were assessed at each phase. Assessments of BI, as the focal predictor, were extensive, using questionnaires at each age and structured observation at ages 8 and 13 years. Assessments of SA were via questionnaire and structured clinical interviews with both twins and their parents.

Storytelling paradigm

We used one task from the Middle Childhood version of the Laboratory Temperament Assessment Battery (Lab-TAB; Goldsmith et al., 1999) to assess social inhibition in the presence of strangers. In Storytelling, the child stands in front of multiple seated child testers and is asked to talk about what they did the prior day, with at least one prompt given by the child tester. Items scored from videotape included child’s social reticence, verbal hesitance, and avoidance. All variables had Kappa agreement of at least 0.70, with master coders (i.e., staff members with extensive training) coding at least every 10th episode to minimize coder drift. Items were z-scored and combined to create an overall Storytelling Inhibition score for each child, with internal consistency reliability of $\alpha = 0.77$.

Observer ratings of approach and shyness

During the age 8 and age 13 home visits, two experimenters rated each twin’s general behavioral and affective tendencies. Three items reflect BI: initial approach/avoidance at the beginning of the home visit (physical movement away from the experimenter), initial shyness (facial and vocal hesitancy with the experimenter), and child shyness across the course of the visit. The first two items were scored on a Likert-type scale ranging from 0 (“Clearly moves toward the examiner; No sign of vocal, facial or postural wariness at all, smiling”) to 4 (“Moves away, hides behind mom/something; Clear-cut shyness; fearful facial expression”). The third Shyness item was based on the Behavior Rating Scales (BRS) from the Bayley Scales of Infant Development (Bayley, 1969) and was similarly scored on a 1–5 Likert scale. We averaged across observers (cross-observer $r_s = 0.62, 0.54, 0.64$, respectively), and then averaged the three items to reflect an overall Observer Rating of Behavior Inhibition.

At Age 8, mothers and fathers completed the Children’s Behavior Questionnaire (CBQ; Rothbart et al., 2001). The CBQ assesses multiple temperament dimensions, including shyness. The CBQ response scale ranges from 1 (“extremely untrue”) to 7 (“extremely true”). The Shyness scale includes items such as “Acts shy around new people” and “prefers to watch rather than join other children playing.” Internal consistency of the CBQ Shyness scale for mothers and fathers was Cronbach’s $\alpha = 0.90$ and 0.88, respectively. At ages 13 and 15, mothers and fathers completed the similar but age-appropriate Early Adolescence Temperament Questionnaire-revised (EATQ-R) (Capaldi & Rothbart, 1992; Ellis & Rothbart, 2001), which included a corresponding Shyness scale (Cronbach’s $\alpha = 0.85$, same for mothers and fathers, and 0.86, respectively, at ages 13 and 15).
Parent- and self-reported BI and SA

We also used the relevant Inhibition and Social Anxiety scales from the MacArthur Health and Behavior Questionnaire (HBQ; Armstrong et al., 2003; Essex et al., 2002) at all ages. At ages 8 and 13, both mothers and fathers rated their children’s behavior over the past six months using a three-point scale (0 = rarely, 2 = certainly applies). At ages 13 and 15, twins self-reported their own inhibited behavior. Sample items for Inhibition include the following: “When I’m around kids I don’t know, I get quiet,” and “When [child] meets new kids, he/she is shy.” Sample items for Social Anxiety include “It makes me nervous or uncomfortable when I have to do things in front of other people,” and “When [child] is around kids he/she doesn’t know, he/she feels nervous or uncomfortable.” Internal consistency reliability (α) for HBQ Inhibition scales ranged from 0.69 to 0.92, and for the Social Anxiety scales, they ranged from 0.78 to 0.95 for all reporters across ages.

Twins also completed the Multidimensional Anxiety Scale for Children (MASC; March et al., 1997) at age 13, with a limited number (n = 10) completing it at age 15 as well. The MASC includes a 9-item Social Anxiety scale to assess performance related anxiety. Sample items include “I worry about other people laughing at me,” and “I worry about getting called on in class.” Items are scored on a four-point Likert scale ranging from (0 = never true about me, 3 = often true about me) with a reliability at age 13 of α = 0.83.

Other Parent- and Self-Reported Disorders. The HBQ was also used to assess other behaviors that might help elucidate specific prediction of SA, namely the Depression, Overanxious, Conduct Disorder, Oppositional Defiant, Inattention, and Impulsivity scales. Like the SA scale, mothers, fathers, and twins reported on these behaviors across time. Sample items and reliabilities are described elsewhere (Moore et al., in press).

Diagnostic interview schedule for children – IV

We administered the Diagnostic Interview Schedule for Children, Version IV (DISC-IV) to mothers at each age and to twins during early and later adolescence. The DISC-IV is a computer-based structured diagnostic instrument with moderate to good diagnostic reliability (Shaffer et al., 2000). Here, we use symptom counts on the Social Phobia scale at each age to reflect SA. We also used symptom counts for Major Depression, Generalized Anxiety Disorder, Attention-Deficit/Hyperactivity Disorder (ADHD), Conduct Disorder, and Oppositional-Defiant Disorder in analyses examining the specificity of SA.

Derivation of composites for analyses

The variables in Figure 1 illustrate the raw materials for composites used in the analyses. Our general approach was to account for child linear age effects within assessment periods at the outset; thus, child age in months was first regressed out of each variable and standardized residuals were saved for use in calculating composites for child behaviors. To form the composites, we used unit weighting rather than differentially weighting the different measures, in the interest of replicability. Other than sex, we used only dimensional measures, including symptom counts.

To create overall composites at each age, we regressed age out of each variable and created composites for BI and SA within reporter, but separately for ages 8, 13, and 15. For example, we combined mother reported CBQ Shyness and HBQ Inhibition at Age 8, and father reported CBQ Shyness and HBQ Inhibition at Age 8, then averaged mother report, father report, observer reported BI, and coded behaviors from the Storytelling paradigm to create an overall BI composite. This process was repeated for each age period with the relevant BI and SA variables. The convergent validity of components and internal consistency of each BI and SA composite can be assessed by examining within age correlations (see Table 1). In general, there were consistent moderate to strong correlations within each composite, across reporter and measurement. A notable exception is Age 8 BI, with the Storytelling Paradigm showing little or no relation to other items within the composite. Nonetheless, observational tasks that assess similar behaviors to those that are parent or self-reported can show minimal association but still reflect similar processes across varying contexts (Gagne et al., 2011); thus, we include inhibition scored during the Storytelling task in our assessment of Age 8 BI.

For analyses of the specificity issue, this process was repeated for the following dimensional symptom domains using HBQ scales and DISC symptom counts in the same manner described for SA: depression, generalized anxiety, ADHD, conduct disorder, and oppositional defiant disorder.

Potential moderators

Overprotective parenting

Mothers’ parenting was assessed using the 35-question short form of the Child-Rearing Practices Report (CRPR; Block, 1965). Mothers rated each statement on the questionnaire from 1 (i.e., strongly disagree) to 6 (i.e., strongly agree). Overprotective parenting specifically was assessed using the inverse of the 7-item Encouraging Independence subscale. Items included “I try to let my twins make many decisions for themselves” and “I intend to give my twins a good many duties and family responsibilities.” Internal consistencies ranged from 0.49 to 0.61.

Parent internalizing psychopathology

Parents completed the Composite International Diagnostic Interview (CIDI; Robins et al., 1988) and the Multidimensional Personality Questionnaire (MPQ; Patrick et al., 2002; Tellegen, 1985). The CIDI is a structured interview that assesses symptoms of psychopathology across multiple domains according to criteria from the ICD-10 and DSM-III-R. Parents completed the interview via telephone. The CIDI has high test-retest reliability (Kessler & Üstün, 2004). Both mothers and fathers completed the modules depression, generalized anxiety, dysthymia, and social phobia. The MPQ is a self-report questionnaire with response options following a four-point Likert-type scale, ranging from “Definitely True” to “Definitely False.” We used the Well-being subscale, which is reverse scored to indicate depressive tendencies, as well as the Stress Reaction scale which is indicative of anxious behaviors in parents.

We calculated a sum score for CIDI diagnoses to reflect additive effects of parents’ internalizing psychopathology at child ages 8 and 13. The four CIDI internalizing items and the MPQ Well-being and Stress Reaction scores were then z-scored and averaged to create an overall mother and father internalizing score for each of the two variables for each parent at each age.

Parenting stress

A brief version (37-item, 5-point Likert scale) of the Parenting Stress Index (PSI; Abidin, 1990) was administered at ages 8, 13,
Hollingshead includes sex and marital status, but we limited our analysis to the Four-Factor Index (Hollingshead, 1975). Traditionally, the Hollingshead includes sex and marital status, but we limited our calculation to Hollingshead Education scores (ranging from 1 “less than seventh grade” to 7 “graduate professional training”), and Hollingshead Occupation codes (ranging from 1 “farm laborers/menial service workers” to 9 “higher executives and major professionals”). Each Occupation and Education score was weighted by 5 and 3, respectively, before computing a final raw Hollingshead score per family. Though the original Hollingshead Index was developed in 1970, we used information from the 2000 United States Census Bureau to assign codes for parents’ occupation. Descriptive statistics for family SES in the current study were as follows: age 8: mean = 45.48 (SD = 11.27), range = 8 to 66; age 13: mean = 45.93 (SD = 10.55), range = 14 to 66; age 15: mean = 46.00 (SD = 11.17), range = 17 to 66.

### Table 1. Intercorrelations within behavior inhibition and social anxiety constructed composites

<table>
<thead>
<tr>
<th>Behavior Inhibition – Age 8</th>
<th>1</th>
<th>2</th>
<th>3</th>
<th>4</th>
<th>5</th>
<th>6</th>
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</thead>
<tbody>
<tr>
<td>1. Storytelling Paradigm (Scored)</td>
<td>1</td>
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<tr>
<td>2. Observer Rated Initial Approach</td>
<td>-0.02</td>
<td>1</td>
<td></td>
<td></td>
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<tr>
<td>3. Observer Rated Initial Shyness</td>
<td>0.07***</td>
<td>0.69***</td>
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<tr>
<td>4. Observed Shyness (BRS)</td>
<td>0.14***</td>
<td>0.40***</td>
<td>0.57***</td>
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<tr>
<td>5. Father Report Shyness (CBQ)</td>
<td>0.03</td>
<td>0.19***</td>
<td>0.28***</td>
<td>0.33***</td>
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</tr>
<tr>
<td>6. Mother Report Shyness (CBQ)</td>
<td>0.05</td>
<td>0.22***</td>
<td>0.30***</td>
<td>0.35***</td>
<td>0.64***</td>
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<tr>
<td>Behavior Inhibition – Age 13</td>
<td>1</td>
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<tr>
<td>1. Observer Rated Initial Approach</td>
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<td>2. Observer Rated Initial Shyness</td>
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<td>3. Self-Report Shyness (EATQ)</td>
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<td>0.23***</td>
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</tr>
<tr>
<td>4. Father Report Shyness (EATQ)</td>
<td>0.27***</td>
<td>0.35***</td>
<td>0.33***</td>
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<tr>
<td>5. Mother Report Shyness (EATQ)</td>
<td>0.24***</td>
<td>0.33***</td>
<td>0.42***</td>
<td>0.54***</td>
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<tr>
<td>6. Father Report Inhibition (HBQ)</td>
<td>0.15***</td>
<td>0.21***</td>
<td>0.31***</td>
<td>0.64***</td>
<td>0.38***</td>
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<tr>
<td>7. Mother Report Inhibition (HBQ)</td>
<td>0.19***</td>
<td>0.25***</td>
<td>0.36***</td>
<td>0.44***</td>
<td>0.67***</td>
<td>0.38***</td>
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<tr>
<td>Behavior Inhibition – Age 15</td>
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<tr>
<td>1. Self-Report Shyness (EATQ)</td>
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<td>2. Self-Report Social Inhibition (HBQ)</td>
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<td>3. Mother Report Inhibition (HBQ)</td>
<td>0.52***</td>
<td>0.34***</td>
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</table>

| Social Anxiety – Age 8 | 1 | | | | | |
|------------------------|---|---|---|---|---|
| 1. Mother Reported Social Phobia (DISC) | n/a | | | | | |

<table>
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<tr>
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<td>1. Self-Reported Social Anxiety (HBQ)</td>
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<td>2. Self-Reported Social Anxiety (MASC)</td>
<td>0.61***</td>
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<td>3. Self-Reported Social Phobia (DISC)</td>
<td>0.56***</td>
<td>0.48***</td>
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<td>4. Father Report Social Anxiety (HBQ)</td>
<td>0.26***</td>
<td>0.26***</td>
<td>0.24***</td>
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<tr>
<td>5. Mother Report Social Anxiety (HBQ)</td>
<td>0.25***</td>
<td>0.23***</td>
<td>0.20***</td>
<td>0.37***</td>
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<tr>
<td>6. Mother Report Social Anxiety (DISC)</td>
<td>0.24***</td>
<td>0.20***</td>
<td>0.27***</td>
<td>0.56***</td>
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<td>Social Anxiety – Age 15</td>
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<tr>
<td>1. Self-Reported Social Anxiety (HBQ)</td>
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<td>3. Self-Reported Social Phobia (DISC)</td>
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<tr>
<td>4. Mother Report Social Anxiety (HBQ)</td>
<td>0.33***</td>
<td>0.58</td>
<td>0.28***</td>
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<tr>
<td>5. Mother Report Social Anxiety (DISC)</td>
<td>0.32***</td>
<td>0.62</td>
<td>0.348***</td>
<td>0.74***</td>
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</table>

Note: *p < 0.05, **p < 0.01, ***p < 0.001. The MASC at age 15 was completed by twins who had not completed the MASC at age 13. The sample size overlap is small (n = 10) with other age 15 items. Thus, the correlations, though large, are not significant.

and 15, reported by mothers (α = 0.83, 0.80, 0.80, respectively). The scales range from “Strongly Agree” to “Strongly Disagree.” We identified 5 subscales using item-level factor analysis: attachment, restrictions of role, relationships with spouse, sense of competence, and social isolation. Subscale scores were totaled at each age, with higher scores indicating high parenting stress. We z-scored these scales and combined into a mean composite to use for further analyses.

**Socioeconomic status (SES)**

To estimate SES, we used two items from the Hollingshead Four-Factor Index (Hollingshead, 1975). Traditionally, the Hollingshead includes sex and marital status, but we limited our...
Peer victimization

At age 13 and 15, twins reported on their own and their cotwin’s experiences using seven items on the HBQ that reflect peer victimization and social experiences (Carroll et al., 2019). Seven questions asked respondents to choose between two opposing options that were stated equivalently (e.g., Kids say mean things to me, or Kids don’t say mean things to me; Some kids at school verbally or physically threaten me, or Kids at school don’t verbally or physically threaten me). After selecting which option best described their experience, respondents rated it on a scale from 1 (sort of like me) to 3 (really like me). Responses were then converted to a 6-point Likert scale, with higher scores indicating more severe peer victimization. The 7 HBQ peer victimization items were averaged to calculate a total peer victimization scale score for each individual and their co-twin. Self ($\alpha = 0.88$) and co-twin ($\alpha = 0.90$) reports of peer victimization were correlated, $r = 0.58$, $p < 0.001$, so we computed a mean composite representing overall peer victimization for each adolescent.

Pubertal timing

We used a multi-method approach to measure pubertal development at ages 13 and 15. Tanner stages (Tanner, 1962), the parent-and self-reported Picture-Based Interview about Puberty (PBIP; Morris & Udry, 1980), and the Pubertal Development Scale (PDS; Petersen et al., 1988) were administered. The PBIP shows good reliability and validity for measuring pubertal development (Coleman & Coleman, 2002). Using scoring developed by Shirtcliff et al. (2009) and used in prior research with this sample (Phan et al., 2021), we calculated pubertal stage scores that mapped onto the Tanner stage scale metric. Tanner stages and the pubertal development stage were standardized within measure and then averaged to form a puberty score. A pubertal timing variable was calculated at ages 13 and 15 separately by regressing youths’ age at the time of puberty measures collected on the puberty score and then using the residuals (Mendle et al., 2019). A higher pubertal timing score represents youth who are more advanced in pubertal stage relative to same-age peers. Standardized residual scores were then used for primary analyses.

Child sex

Sex was reported at age 8 by parents. Gender identity beyond the binary was not systematically assessed.

Overview of analytic approaches

Rather than attempting to capture all elements of our investigation into a single highly elaborated model with many inter-dependent findings, we focus on a series of specific analyses to address the issues that motivated our study. We do not focus on measurement issues due to earlier publications on measurement with this sample and related samples (e.g., Gagne et al., 2011; Planalp et al., 2017).

Phenotypic analyses

Directionality of effects using cross-lagged panel analysis

We tested the reciprocal relationship between BI and SA across childhood and adolescence using a partially random intercept cross-lagged panel model (RI-CLPM; Berry & Willoughby, 2017; Hamaker et al., 2015) with Mplus v. 7.3 (Hamaker et al., 2015). In a RI-CLPM, each variable is regressed on its own lagged score from the previous measurement occasion, as well as on the lagged score of the other variable. However, to account for variability within individuals as well as between individuals, as in the standard CLPM, we estimated a random intercept for BI. Notably, because there was very little variance in Age 8 mother report of SA, the random intercept for SA was not appropriate in these data, indicated by an ill-fitting and error prone model. Therefore, we fit a partial RI-CLPM to BI but not SA (illustrated in Figure 2).

Specificity of prediction of SA from BI

Pearson correlation analyses that account for within family clustering of data using the Mplus clustering option were used to examine the association of BI at each age with concurrent and later
Figure 3. (a) describes the bivariate model of genetic and environmental influences linking childhood BI and adolescent SA. A1 and A2 = genetic influences on BI and SA, respectively; C1 and C2 = shared environmental influences on BI and SA, respectively; E1 and E2 = nonshared environmental influences on BI and SA, respectively; a11 = effect of A1 on BI; c11 = effect of C1 on BI; e11 = effect of E1 on BI; a21 = effect of A1 on SA; c21 = effect of C1 on SA; e21 = effect of E1 on SA; a22 = effect of A2 on SA; c22 = effect of C2 on SA; e22 = effect of E2 on SA. (b) illustrates only significant pathways in our model.

symptoms of general anxiety, depression, ADHD, conduct, and oppositional-defiant disorder. After r-to-z transformations, we compared the relative magnitude of the correlations between BI and SA with other psychopathology dimensions (Eid et al., 2011; Lenhard & Lenhard, 2014). No results warranted follow-up with multivariate approaches.

Multiple moderator approach for BI-to-SA predictions

In the moderation analyses, multi-level regression using SAS proc mixed accounted for familial clustering of twins. When examining family wide processes, data from members within a family are non-independent and do not fit standard regression assumptions (Gonzalez & Griffin, 2012; Planalp et al., 2019). This model accounts for the within-family similarities by imposing a compound symmetric covariance matrix to twin data. First, we examined age group (mean ages 8, 13, and 15) and sex as moderators. Then, we tested each of the other potential moderators, one at a time, along with age group and sex. Each regression at this stage included all main effects and all two-way and three-way interactions. From these five regression models, we carried the subset of moderators that showed main effects or two-way interactions with BI forward to the final model. In the final model, we probed significant interactions using simple slopes analysis (Aiken et al., 1991), as recommended by Preacher et al. (2006) for multi-level modeling.

Biometric analyses

Bivariate twin analyses. Twin ICCs (Table 3) provide an easily understood descriptive measure of twin similarity. Doubling the difference between MZ and DZ intraclass (similarity) correlations estimates heritability (Falconer & Mackay, 1996). We also used bivariate biometric model fitting with Cholesky decomposition, using Mplus (version 7.3; Muthen & Muthen, 1998) to examine the genetic and environmental bases of phenotypic covariances between BI and SA. We used a bivariate Cholesky decomposition that hypothesized a specific ordering of variables as we were interested in genetic and environmental bases of prediction of adolescent SA from childhood BI. Childhood BI was measured at age 8, and Adolescent BI was considered an average of the age 13 and 15 SA composites. Figure 3 summarizes the bivariate biometric model with the Cholesky approach. We also explain the interpretation more when the genetic results are presented in the Results section.

We also examined a bivariate model with scalar sex limitation, which allows genetic and environmental paths to vary for males and females. This model did not fit better than the model where the male and female paths were constrained to be equal, and thus we do not treat it further.

Results

The means, SDs, and intercorrelations of the BI and SA variables at each developmental stage and potential moderators are shown in Table 2. BI and SA variables were all intercorrelated, with correlations within the same age group substantial and higher than most of the stability correlations. We also note that child sex was a moderate but consistent correlate of both BI and SA, with females higher on both. Child sex also correlated with other variables, suggesting that it should figure prominently in our analyses. Table 3 shows intraclass correlations indexing cotwin similarity for BI and SA for MZ and same and opposite sex DZ twin groups. BI was highly heritable and SA, less so.

The cross-twin, cross-trait correlations, which foreshadow the bivariate genetic analyses, were modest in size. The association of BI in one twin with SA in the cotwin was always stronger in MZ than in DZ pairs, which implies genetic effects on the BI association with SA. Further interpretation of the bivariate association is reserved for the model-fitting approaches below.

Is the prediction of SA from BI unidirectional and specific?

Examination of unidirectionality using cross-lagged panel analysis

As Figure 2 shows, both BI and SA are moderately stable over time (0.35 and 0.36 for BI, and 0.30 and 0.49 for SA, from one age to the next). Moreover, the within-age, cross-trait correlations were 0.36, 0.63, and 0.14 at the three ages. When we move to the cross-lagged coefficients of most interest here, we see that the predictions from earlier SA to later BI were as strong as the reverse predictions (the predicted BI-to-SA), lending no support to a unidirectionality argument.

Specificity of BI-to-SA links

Controlling for twin interdependence, we computed Pearson correlations between the BI variables at each age with contemporaneous scores on depression and generalized (non-social) anxiety. The center column in Table 4 shows, for instance, that at age 13, BI correlated 0.21 with depression and 0.24 with generalized anxiety, whereas the correlation with SA (r = 0.53) was almost twice those magnitudes. Some level of correlation of BI with depression and
Table 2. Descriptive statistics and intercorrelations of BI, SA, and their potential moderators

<table>
<thead>
<tr>
<th>Scale</th>
<th>N</th>
<th>Mean (SD)</th>
<th>1</th>
<th>2</th>
<th>3</th>
<th>4</th>
<th>5</th>
<th>6</th>
<th>7</th>
<th>8</th>
<th>9</th>
<th>10</th>
</tr>
</thead>
<tbody>
<tr>
<td>1. BI (8)</td>
<td>1612</td>
<td>−0.006 (0.693)</td>
<td>1</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>2. BI (13)</td>
<td>959</td>
<td>0.002 (0.690)</td>
<td>0.449**</td>
<td>1</td>
<td></td>
<td></td>
<td></td>
<td></td>
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<td></td>
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<tr>
<td>3. BI (15)</td>
<td>1077</td>
<td>−0.002 (0.954)</td>
<td>0.297**</td>
<td>0.522**</td>
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<tr>
<td>4. SA (8)</td>
<td>1458</td>
<td>0.000 (0.100)</td>
<td>0.346**</td>
<td>0.251**</td>
<td>0.274**</td>
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<tr>
<td>5. SA (13)</td>
<td>947</td>
<td>−0.002 (0.683)</td>
<td>0.228**</td>
<td>0.630**</td>
<td>0.407**</td>
<td>0.323**</td>
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<tr>
<td>6. SA (15)</td>
<td>1093</td>
<td>−0.009 (0.931)</td>
<td>0.113**</td>
<td>0.395**</td>
<td>0.362**</td>
<td>0.167**</td>
<td>0.545**</td>
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<tr>
<td>7. PeerVic (13)</td>
<td>751</td>
<td>−0.001 (0.851)</td>
<td>−0.107**</td>
<td>0.023</td>
<td>−0.109**</td>
<td>−0.019</td>
<td>0.216**</td>
<td>0.184**</td>
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<tr>
<td>8. PeerVic (15)</td>
<td>302</td>
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<td>−0.232**</td>
<td>−0.236</td>
<td>0.011</td>
<td>−0.212**</td>
<td>0.039</td>
<td>0.148*</td>
<td>0.619**</td>
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<td>9. Pub. Tim. (13)</td>
<td>936</td>
<td>−0.002 (0.997)</td>
<td>−0.030</td>
<td>0.057</td>
<td>0.008</td>
<td>0.034</td>
<td>0.094**</td>
<td>0.113**</td>
<td>−0.054</td>
<td>0.050</td>
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<tr>
<td>10. Pub. Tim. (15)</td>
<td>304</td>
<td>0.000 (0.995)</td>
<td>0.019</td>
<td>−0.086</td>
<td>0.119*</td>
<td>0.050</td>
<td>0.121</td>
<td>0.080</td>
<td>−0.142</td>
<td>−0.085</td>
<td>0.552**</td>
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<td>11. SES (8)</td>
<td>1236</td>
<td>45.48 (11.27)</td>
<td>−0.050</td>
<td>−0.095*</td>
<td>0.001</td>
<td>−0.104**</td>
<td>−0.062</td>
<td>−0.041</td>
<td>−0.038</td>
<td>−0.141</td>
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<td>496</td>
<td>46.00 (11.17)</td>
<td>−0.121**</td>
<td>−0.100</td>
<td>−0.005</td>
<td>−0.011</td>
<td>−0.054</td>
<td>−0.065</td>
<td>−0.210*</td>
<td>−0.255**</td>
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<td>0.082</td>
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<td>1522</td>
<td>2.08 (0.44)</td>
<td>0.048</td>
<td>0.121**</td>
<td>0.127**</td>
<td>0.152**</td>
<td>0.176**</td>
<td>0.020</td>
<td>0.114**</td>
<td>0.154*</td>
<td>−0.033</td>
<td>−0.087</td>
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<td>14. PSI (8)</td>
<td>359</td>
<td>1.93 (0.46)</td>
<td>−0.043</td>
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<td>0.073</td>
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<td>0.156*</td>
<td>−0.040</td>
<td>0.048</td>
<td>0.233**</td>
<td>0.046</td>
<td>−0.164</td>
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<td>15. PSI (15)</td>
<td>1150</td>
<td>2.26 (0.56)</td>
<td>0.042</td>
<td>0.090**</td>
<td>0.062</td>
<td>0.078**</td>
<td>0.095**</td>
<td>0.034</td>
<td>0.025</td>
<td>0.014</td>
<td>0.090**</td>
<td>0.072</td>
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<td>16. OVP (8)</td>
<td>970</td>
<td>2.04 (0.51)</td>
<td>0.012</td>
<td>0.098**</td>
<td>0.075</td>
<td>0.056</td>
<td>0.110**</td>
<td>−0.014</td>
<td>0.060</td>
<td>0.126</td>
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<td>17. OVP (15)</td>
<td>315</td>
<td>2.17 (0.50)</td>
<td>0.076</td>
<td>0.079</td>
<td>−0.009</td>
<td>−0.041</td>
<td>0.113</td>
<td>0.033</td>
<td>0.140</td>
<td>0.032</td>
<td>0.507**</td>
<td>−0.079</td>
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<td>18. F Int. (8)</td>
<td>1326</td>
<td>−0.001 (0.61)</td>
<td>0.117**</td>
<td>0.156**</td>
<td>0.092**</td>
<td>0.137**</td>
<td>0.110**</td>
<td>0.115**</td>
<td>−0.031</td>
<td>−0.009</td>
<td>0.066</td>
<td>−0.026</td>
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<tr>
<td>19. M Int. (8)</td>
<td>1616</td>
<td>0.003 (0.55)</td>
<td>0.058*</td>
<td>0.053</td>
<td>0.095**</td>
<td>0.129**</td>
<td>0.144**</td>
<td>0.061</td>
<td>0.033</td>
<td>0.104</td>
<td>0.013</td>
<td>−0.016</td>
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<td>20. F Int. (13)</td>
<td>656</td>
<td>0.013 (0.64)</td>
<td>0.102**</td>
<td>0.113**</td>
<td>0.094**</td>
<td>0.110**</td>
<td>0.134**</td>
<td>0.094*</td>
<td>0.063</td>
<td>−0.388</td>
<td>0.013</td>
<td>−0.094</td>
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<td>21. M Int. (13)</td>
<td>834</td>
<td>0.005 (0.60)</td>
<td>0.024</td>
<td>0.099**</td>
<td>0.133**</td>
<td>0.090*</td>
<td>0.177**</td>
<td>0.077*</td>
<td>−0.010</td>
<td>0.328</td>
<td>−0.006</td>
<td>−0.269</td>
</tr>
</tbody>
</table>


12. SES (13) 0.742** 1
13. SES (15) 0.771** 0.654** 1
14. PSI (8) −0.179** −0.049 −0.138** 1
15. PSI (15) −0.156* −0.064 −0.178* 0.675** 1
16. OVP (8) −0.113** −0.088** −0.064 0.092** 0.168** 1
17. OVP (13) −0.112** −0.057 −0.022 0.072* 0.001 0.439** 1
18. OVP (15) −0.061 0.190 −0.040 0.078 0.108 0.432** 0.263 1
19. F Int. (8) −0.082** −0.051 0.052 0.169** 0.215** 0.050 0.060 0.014 1
20. M Int. (8) −0.085** −0.045 −0.094* 0.441** 0.419** 0.034 0.067* −0.012 0.124** 1
21. F Int. (13) −0.164** −0.155** −0.057 0.117** 0.047 0.081* 0.125** −0.352 0.506** 0.075 1
22. M Int. (13) −0.112** −0.106** −0.067 0.406** 0.440** 0.069 0.144** −0.069 0.134** 0.629** 0.205** 1

Note: *p < 0.05, **p < 0.01. Number (8, 13, 15) indicates age at which measured. BI = behavioral inhibition; SA = social anxiety; PeerVic = peer victimization; Pub. Tim. = pubertal timing; SES = socioeconomic status; PSI = parenting stress; OVP = overprotective parenting; F and M Int. = father or mother internalizing symptoms.
generalized anxiety symptoms should be expected, as these latter two variables show correlations in the 0.50-0.60 range with SA. We also examined specificity of BI’s prediction of SA in relation to three externalizing measures: ADHD, oppositional defiant disorder, and conduct disorder. Here, Table 4 shows mostly inverse or near-zero correlations, even when the measures were from the same age group. Small but significant negative correlations, consistent with a small protective effect of high BI for externalizing problems, emerged in the age 8 and 13 data.

Longitudinal correlations between age 8 BI and later psychopathology variables (column 1 of Table 4) followed the same patterns, with the strongest relations between BI and SA compared with other outcomes. To confirm that BI specifically predicts SA, we used r-to-z transformations, accounting for the dependency of calculating correlations within the same sample. The differences between the resulting z-transforms were compared using t-tests calculating correlations within the same sample. The differences

Is the prediction of SA from BI genetically mediated and does SA show residual genetic variance that is independent of genetic effects on BI?

Given the phenotypic associations between BI and SA that the cross-lagged panel analysis demonstrated, we next sought to examine the genetic and environmental underpinnings of the association. To simplify the phenotypic aspects of the model, we needed a single predictor and a single outcome. We also sought to separate the predictor and outcome temporally to clarify interpretation of results. As the predictor, we used childhood BI at age 8 and as the outcome, we estimated adolescent SA as the mean of SA at ages 13 and 15. The ordinary phenotypic correlation of these BI and SA composites was 0.16 (p < 0.001).

Bivariate model with Cholesky decomposition

The diagram presenting the parameter estimates (Figure 3b, with model characteristics provided in the figure note) has three interesting features. The left side shows additive (A) genetic, common environmental (C) and nonshared, or unique, environmental (E) parameter estimates that compose the phenotypic variation in BI. As anticipated by the ICCs provided in Table 3, the C estimate was empirically driven to zero and was thus set to zero. The path estimates are unsquared, but squared estimates provide a numeric value for overall heritability (genetic influence) of the trait; thus, the 0.84 path from A to childhood BI must be squared to estimate heritability (h²) of BI to be 0.71. If we also square the non-shared environmental path 0.54² = 0.29, we see that 0.71 + 0.29 = 1.00; that is, 100% of the phenotypic variance. The cross-variable paths (BI-to-SA) estimate the variance in SA accounted for by genetic and environmental factors that also impact BI. The path estimate of 0.25 for A (more than twice its standard error) demonstrates a significant genetic path to SA from earlier BI. Finally, the diagram also demonstrates substantial genetic input to variation in adolescent SA that is independent of the genetic basis of BI (0.70² = 0.49, so 49% of the unique variance of SA is attributable to genetic effects). More of the genetic variance in SA is unique than is contributed by BI.

Do specific variables from the literature moderate the prediction of SA from BI?

As noted in the Method section, we first constructed a “base model” of child sex and development period (ages 8, 13, and 15) moderating the BI-to-SA prediction, using a multilevel model that accounted for the paired nature of twin data. In this model, the age indicates the slope of SA across ages 8, 13, and 15. Results of this model are provided in Table 5. Results show the expected strong main effect of BI on SA (Est = 0.536) as well as a strong relation between BI and changes in SA (Est = −0.026); the significant effect of BI on the slope, or changes in SA, indicates that BI at the later ages was most strongly associated with SA. However, sex did not
interact with BI in the model predicting SA. Sex did, however, interact with age, such that SA increased with age for girls and decreased for boys.

With the model in Table 5 as our baseline, we tested each of the other potential moderators, sequentially, to determine if they interacted with BI to predict SA. Of the seven potential moderators, pubertal timing did not interact with BI, nor did it show a main effect; mothers’ internalizing did show a main effect, but it did not moderate the BI-to-SA prediction. Over-protective parenting, family stress, adolescent peer victimization, and SES all significantly moderated the effect of BI on SA. Fathers’ internalizing showed only a marginal (p = 0.07) moderating effect, but that effect was qualified by a significant three-way interaction involving BI and child sex. We carried over-protective parenting, family stress, adolescent peer victimization, SES, and paternal internalizing forward into a final model to examine moderation more comprehensively.

Table 6 presents the final “conditional model” which includes predictors that were significant in individual moderator analyses. Main effects that did not involve BI are noteworthy. Four significant main effects on levels in SA emerged: a strong effect of BI (Est. = 0.572), a medium effect of parenting stress (Est. = 0.183), and a modest effects of father internalizing (Est. = 0.110) and over-protective parenting (Est. = 0.103); SA was generally higher when each of these variables was also higher. Peer victimization had a significant main effect on changes in SA across time (Est. = 0.037), such that adolescents experiencing lower peer victimization at age 13 had slightly lower SA than those experiencing higher rates of peer victimization; the same relation was not found by age 15.

Our key question about moderation of the effects of BI yielded one strong interaction in the final model with parenting stress (Est. = 0.398). In general, when children were higher in BI, they also experienced more SA, particularly when parenting stress was higher at age 8. BI also interacted with Parenting Stress to predict the slope of SA, such that children with higher BI developed SA through adolescence at a faster rate when parenting stress was low. We also note that the significance level of the 2-way interaction of BI with SES was p = 0.038; this effect is small (Est. = −0.009) but notable given its independence from the parenting stress interaction.

**Discussion**

**Summary of results**

Our results are grounded in a substantial empirical literature about BI and anxiety, and our large-study findings strengthen some trends in the literature, weaken others, and provide novel findings. We begin by summarizing our answers to the research questions about BI and anxiety.
questions. We do not find evidence that the direction-of-effects in the BI-SA association is unidirectional—that is, only or predominately from BI to SA. We do conclude that BI robustly predicts SA. From a correlational perspective, this prediction is relatively specific, in the context of other internalizing problems and clearly in relation to externalizing problems. An ancillary finding is that BI does not appear to protect strongly against externalizing problems.

BI and SA are both moderately heritable, BI more than SA. The genetic influences on BI also impact differences in SA, as do the nonshared environmental effects, to a much weaker degree. Still, SA variability is influenced by other genetic effects that are independent of BI.

Our examination of predictors of SA began with the raw correlations in Table 2, which showed that all the potential moderators except pubertal timing significantly correlated with SA, as the literature would have predicted. Only one of these variables—peer victimization—was reasonably known beforehand for our sample (Carroll, 2021).

We then tested variables, including measured exposures and experiences, that could—based on the literature—moderate the rather strong prediction of SA from BI. We first tested individual moderators, separately, in a series of regression models with BI, age group, and sex. In addition to age group, five variables (over-protective parenting, family stress, adolescent peer victimization, and SES) all significantly moderated the effect of BI on SA. The second set of results was presented in Table 6. Although we refer to Table 6 as presenting the final model, every stage of the analysis yields informative results. Clearly, the effects of any variable in multiple regression depends on the other predictors available for study, and we did not exhaust all possibilities.

Why are these interactions developmentally important? As in our previous work in another sample (Essex et al., 2010), we subscribe to the argument that interactions among variables in longitudinal data may signal branching paths toward outcomes. So, for instance, we might consider high BI children to be on a general path leading probabilistically to SA, but many of those high BI children who experience peer victimization shift to a higher risk path. The high BI children in families that experience chronic stress shift to a different higher risk path and so on for other interactions over time with BI. Such a causative nexus would help explain heterogeneity in origins of SA.

### Table 5. “Base model” regression analysis of the role of age group and sex in the prediction of social anxiety from behavioral inhibition

<table>
<thead>
<tr>
<th>Predictor of Social Anxiety</th>
<th>β</th>
<th>SE</th>
<th>p</th>
<th>95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>Intercept of SA</td>
<td>-0.001</td>
<td>0.021</td>
<td>0.962</td>
<td>-0.043, 0.041</td>
</tr>
<tr>
<td>Slope of SA</td>
<td>-0.001</td>
<td>0.004</td>
<td>0.733</td>
<td>-0.010, 0.007</td>
</tr>
<tr>
<td>Effect of BI on Intercept of SA</td>
<td>0.536</td>
<td>0.031</td>
<td>&lt;0.001</td>
<td>0.476, 0.597</td>
</tr>
<tr>
<td>Effect of BI on Slope of SA</td>
<td>-0.026</td>
<td>0.006</td>
<td>&lt;0.001</td>
<td>-0.037, -0.015</td>
</tr>
<tr>
<td>Effect of Sex on Intercept of SA</td>
<td>-0.014</td>
<td>0.043</td>
<td>0.736</td>
<td>-0.009, 0.070</td>
</tr>
<tr>
<td>Effect of Sex on Slope of SA</td>
<td>0.028</td>
<td>0.008</td>
<td>0.001</td>
<td>0.011, 0.044</td>
</tr>
<tr>
<td>Effect of Sex*BI on Intercept of SA</td>
<td>0.090</td>
<td>0.145</td>
<td>0.533</td>
<td>-0.194, 0.375</td>
</tr>
<tr>
<td>Effect of Sex*BI on Slope of SA</td>
<td>-0.010</td>
<td>0.011</td>
<td>0.403</td>
<td>-0.0318, 0.013</td>
</tr>
</tbody>
</table>

Note: N = 1711 individual children with available data; SA = Social Anxiety as the outcome of interest.

### Table 6. “Conditional model” regression analysis of the role of moderators on the prediction of adolescent social anxiety from behavioral inhibition

<table>
<thead>
<tr>
<th>Predictor of Social Anxiety</th>
<th>β</th>
<th>SE</th>
<th>p</th>
<th>95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>Intercept of SA</td>
<td>-0.023</td>
<td>0.032</td>
<td>0.476</td>
<td>-0.087, 0.040</td>
</tr>
<tr>
<td>Effect of Predictor on Intercept of SA:</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>BI</td>
<td>0.572</td>
<td>0.051</td>
<td>&lt;0.0001</td>
<td>0.472, 0.673</td>
</tr>
<tr>
<td>Overprotective Parenting</td>
<td>0.103</td>
<td>0.048</td>
<td>0.033</td>
<td>0.008, 0.197</td>
</tr>
<tr>
<td>Mother Internalizing</td>
<td>0.028</td>
<td>0.067</td>
<td>0.679</td>
<td>-0.104, 0.159</td>
</tr>
<tr>
<td>Father Internalizing</td>
<td>0.110</td>
<td>0.054</td>
<td>0.043</td>
<td>0.003, 0.216</td>
</tr>
<tr>
<td>Parenting Stress</td>
<td>0.183</td>
<td>0.083</td>
<td>0.028</td>
<td>0.020, 0.346</td>
</tr>
<tr>
<td>Socioeconomic Status</td>
<td>-0.001</td>
<td>0.002</td>
<td>0.602</td>
<td>-0.006, 0.003</td>
</tr>
<tr>
<td>Peer Victimization</td>
<td>-0.037</td>
<td>0.038</td>
<td>0.323</td>
<td>-0.111, 0.037</td>
</tr>
<tr>
<td>Sex</td>
<td>0.088</td>
<td>0.067</td>
<td>0.189</td>
<td>-0.043, 0.218</td>
</tr>
<tr>
<td>BI*Parenting Stress</td>
<td>0.398</td>
<td>0.122</td>
<td>0.001</td>
<td>0.157, 0.638</td>
</tr>
<tr>
<td>BI*Socioeconomic Status</td>
<td>-0.009</td>
<td>0.004</td>
<td>0.038</td>
<td>-0.017, -0.001</td>
</tr>
<tr>
<td>BI*Sex</td>
<td>-0.036</td>
<td>0.080</td>
<td>0.650</td>
<td>-0.193, 0.121</td>
</tr>
<tr>
<td>BI*Peer Victimization</td>
<td>-0.065</td>
<td>0.049</td>
<td>0.187</td>
<td>-0.162, 0.032</td>
</tr>
<tr>
<td>BI*Father Internalizing</td>
<td>-0.008</td>
<td>0.069</td>
<td>0.905</td>
<td>-0.144, 0.127</td>
</tr>
<tr>
<td>Father Internalizing*Sex</td>
<td>0.222</td>
<td>0.107</td>
<td>0.040</td>
<td>0.010, 0.433</td>
</tr>
<tr>
<td>BI<em>Father Internalizing</em>Sex</td>
<td>-0.143</td>
<td>0.135</td>
<td>0.293</td>
<td>-0.409, 0.124</td>
</tr>
<tr>
<td>Slope of SA</td>
<td>0.006</td>
<td>0.009</td>
<td>0.471</td>
<td>-0.011, 0.023</td>
</tr>
</tbody>
</table>

| Effect of Predictor on Slope of SA: | | | |
| BI                           | 0.003 | 0.014 | 0.817  | -0.025, 0.031   |
| Father Internalizing         | -0.007 | 0.015 | 0.665  | -0.037, 0.023   |
| Mother Internalizing         | -0.006 | 0.018 | 0.727  | -0.042, 0.030   |
| Parenting Stress             | -0.031 | 0.022 | 0.157  | -0.075, 0.012   |
| Peer Victimization           | 0.037 | 0.010 | 0.000  | 0.017, 0.057    |
| Sex                          | 0.011 | 0.018 | 0.537  | -0.024, 0.045   |
| Father Internalizing*Sex     | 0.222 | 0.107 | 0.040  | 0.010, 0.433    |
| BI*Parenting Stress          | -0.068 | 0.030 | 0.026  | -0.127, -0.008  |
| BI*Socioeconomic Status      | 0.002 | 0.001 | 0.102  | -0.001, 0.005   |
| BI*Father Internalizing*Sex  | -0.055 | 0.030 | 0.066  | -0.114, 0.004   |

Note: SA = Social Anxiety; BI = behavioral inhibition; Sex is coded 1 = female, –1 = male.

High BI need not be a necessary condition for SA. Table 6 suggests that, independently of BI, over-protective parenting and parenting stress predict SA, and some interactions are also predictive of SA independently of BI. Thus, high BI is clearly not a sufficient condition for developing SA. Children with high BI who do not develop SA might not have experienced bullying or over-protective parenting, and their parents might not have experienced stress in the parenting role.

**Cautions against strong inference are warranted in some instances**

The generalizability of research on twins is sometimes questioned; however, twins differ minimally from singletons in personality (Johnson et al., 2002), cognitive ability (Christensen et al., 2006), and clinical conditions (Kendler & Prescott, 2006). Ours
is a community sample; confining attention to clinically diagnosed SA might yield different results. Our large sample size leads to very modest effect sizes being statistically significant. The SES range is wider than in some samples, and rural representation is notable. The sample racially and ethnically matches Wisconsin’s 2018 census estimates, which means a lack of strong diversity in race and ethnicity. This lack of diversity severely limits generalizability to nonwhite groups.

Our sample size was likely under-powered for the bivariate sex limitation model (Neale & Cardon, 2013), and the resulting lack of significant difference in fit from the model where male and female paths were constrained to be equal should not close the door to examining these sex differences in genetic paths in a larger twin sample or with different methods. Indeed, the rejected sex limitation model suggested slightly stronger genetic mediation of the BI-to-SA path for males than for females, a finding that deserves more investigation.

**Measurement considerations**

As Figure 2 shows, we used core instruments to assess BI and SA at each age group. These instruments were modified (by the original authors) to make them age-appropriate (for example, the EATQ shyness items differ somewhat from the CBQ shyness items). Observations and reactions to a structured situation were possible at the younger age groups, and an additional questionnaire was used to improve the coverage for SA at the oldest age group. We strove for best-feasible-estimates for each age group, as opposed to factorial invariance. Thus, some instability across age groups could be due to measurement differences, which we judge to be favorably counter balanced by enhanced validity at each age group.

Our primary a priori concern was the measure of over-protective parenting, which was measured as the inverse of parental questionnaire items about encouraging independence. Although over-protective parenting nevertheless emerged as influential in our analyses, its role might have been underestimated relative to what we might have found with a more direct, multi-source, measurement.

**Design considerations**

We only studied a subset of the behavioral moderators that have been suggested for the BI-to-SA prediction. Investigators have proposed that the quality of BI should be considered. Conflicted shyness could be a subtype of BI that holds more predictive power for SA (Poole et al., 2017) and dysregulated fearfulness may also signal increased risk (Buss, 2011). Moreover, we did not include biological variables in our study. Biological variables (e.g., neuroendocrine, brain structure and activity) might serve as mediators in the types of models we investigated, a feature that would complicate the models statistically (moderated mediation) but perhaps add substantial explanatory power.

**Implications of results for models of temperament-psychopathology effects**

How exactly BI relates to SA remains under debate. One hypothesis suggests that the social component of BI (i.e., shyness), which we focus on here, can be conceptualized as part of a spectrum that includes SA (Hofmann et al., 2004; Marshall & Lipsett, 1994). On the other hand, one-third of shy adults do not endorse social fears or somatic symptoms, which are two hallmarks of SA (Heiser et al., 2009). Our study’s results hold implications for the vulnerability vs. complication/scar vs. spectrum models. We must consider whether BI is a precursor (as in a spectrum model) or predisposition (as in a vulnerability model) to SA (Mumper & Klein, 2021). If BI acts as a precursor to later psychopathological anxiety, shared causal mechanisms are expected. If BI is instead a predisposition to anxiety problems, then shared causal processes would not be expected, or at most we would expect only partial overlap in causal mechanisms. The predisposition idea (vulnerability model) would suggest that individuals with high BI would be more likely to develop SA in certain contexts, but that these contexts are crucially important in the etiology of SA. That is, moderating elements typically must be present for a predisposition to give rise to a form of psychopathology with a partially independent etiology (Mumper & Klein, 2021). Of course, multiple processes can be operating, and nothing compels reality to respect our conceptual distinctions among models.

Our cross-lagged panel results support neither a strong version of the vulnerability model nor a strong version of complication/scar model. The vulnerability model would predict stronger paths from BI to later SA, and the complications/scar model would predict the opposite. Our observation of roughly equal paths for these two potentially causal associations, by default, would seem to favor a spectrum model. In other words, high SA symptoms essentially equate to extreme shyness. However, when we turn to the bivariate biometric model, support for a spectrum model dissipates. Genetic influences can be regarded as causal influences although they are distal causes. Strong versions of spectrum models require the same causal process for the precursor and the outcome. Yet, our bivariate results show that most of the genetic variance for SA is not shared with BI. Our bivariate results favor the vulnerability model, in which BI acts as a predisposition. Our bivariate biometric model was not set up to test the complication/scar model, with its opposite direction-of-effect in the BI-SA association 1.

We noted above that moderation of BI’s effect by experiential variables or parental characteristics could signal branching developmental paths toward SA. In the Essex et al. (2010) study, we also studied pathways to adolescent SA, but with a different sample, different ages of assessment, and a different analytic strategy. Essex et al. (2010) identified two paths to adolescent SA, one that was primarily evident in females and that involved higher trait-like BI present in early childhood and the other that involved early maternal stress and subsequent elevated basal cortisol levels. The overlap with our present findings is clear. We show main effects of two parenting variables (stress and over-protectiveness) with later SA, which are independent of BI. We also showed moderation of BI’s effect by several variables, three of which survive testing with the other moderators.

We conclude that hybrid models of the relationship of temperament to psychopathology, at least in the BI-to-SA domain, may be needed.

**The role of children’s sex**

Our finding of modest sex differences at various points in the analyses is unsurprising given a 30-year follow-up of shy boys and girls that showed different (non-psychopathological) manifestations of shyness through the life course (e.g., later marriage for

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1We realize that our treatment of BI in this paper is mostly framed in a risk framework. To partially correct this bias, we affirm that shy individuals can be, and often are, psychologically healthy. Although beyond the scope of this paper, an emerging literature (e.g., Schmidt & Poole, 2020) examines positive features of shyness, which is also a popular subject outside the academic sphere where shy individuals are applauded as more modest, more planful, more approachable, etc. (e.g., Susan Cain’s 2012 book, Quiet: The Power of Introverts in a World That Can’t Stop Talking).
males) (Caspi et al., 1988). Our meta-analysis of sex differences in temperament identified fearfulness as one of the traits for which females scored lower than males, but the meta-analytic effect size ($d = -0.12$) was modest and non-significant (Else-Quest et al., 2006).

Sex differences in comorbidity, course, and other features of SA are less well understood than sex differences in prevalence. However, a large epidemiological study reported that women with SA are more likely to have comorbid internalizing disorders (e.g., panic disorder, specific phobia, generalized anxiety disorder, major depressive disorder), whereas men with SA are more likely to have comorbid externalizing disorders (e.g., conduct disorder, pathological gambling, substance use disorder; Xu et al., 2012), and we did not analyze substance misuse for this report.

We emphasize that, in general, the sex effects in our analyses were small. However, as Table 6 shows, interactions with sex were an important feature of our results. Thus, at least from the dimensional perspective, sex differences are notable but not a major independent part of the explanation of our findings, given the set of other experiential measures that we incorporated.

**Implications for clinical issues**

Evidence-based child and adolescent anxiety interventions are increasingly conceptualizing the client in the context of their broader family system (Chorpita & Weisz, 2009; Ehrenreich-May et al., 2017). Clinicians regularly interact with parents and are sensitive to caregiver behaviors – such as overprotective parenting – that may contribute to the child’s avoidance behaviors, anxious thoughts, and reassurance seeking behaviors; our work provides further evidence for this practice. In fact, an online parental intervention has been developed to reduce over-protective parenting (Morgan et al., 2017). Of course, the meaning and adaptive function of parental over-protectiveness differs according to the degree of danger in the neighborhood, and racial and ethnic differences are likely correlated with these differences (Deater-Deckard, 2008).

Pediatric anxiety intake assessments typically identify parents’ experiential stress, but those evaluating the assessments are often not equipped to provide adult intervention services to stressed parents. Thus, we encourage health care systems to establish concrete pipelines that connect parents to accessible adult intervention services, such that some degree of coordination is possible.

We also emphasize that thorough assessment of peer victimization during childhood and adolescence is crucial for successful intervention at the family or school level (Graham & Bellmore, 2007; Stadler et al., 2010). To the best of our knowledge, providers inconsistently ask about peer victimization during intake and intervention, yet prioritizing communication with parents and schools about whether and how peer victimization is experienced can offer significant intervention improvements. Our findings suggest a thorough peer victimization assessment could be a helpful addition to SA treatment.

Prevention programs that increase children’s self-regulation could also mitigate genetic risk of SA and other behavioral conditions. We have recently shown, for example, that the Family Check-Up increases children’s temperamental inhibitory control across childhood, and inhibitory control mediates prevention effects on adolescent outcomes including internalizing and externalizing problems (Hentges et al., 2020). Furthermore, genetic risk for child psychopathology is ameliorated for children whose families received the intervention; those at high genetic risk in the control group had lower levels of observed childhood effortful control, whereas levels of effortful control were “repaired” for children in the intervention group (Oro et al., 2019).

**Implications for parenting and family research**

Worried or anxious parents may be more likely to exhibit overprotective parenting (Root et al., 2016). Anxious mothers may be hypervigilant to perceived threats in their child’s environment and may seek to reduce their own anxious distress by exerting control over their child’s behaviors or environment (Eysenck, 2013; Lindhout et al., 2006; Turner et al., 2003). On the other hand, relations between child reactivity and parenting behaviors may not be unidirectional, as we have examined here. Notably, parents’ overprotectiveness is unlikely to be the only parent-related variable related to SA and perhaps moderating the BI-to-SA prediction. High dismissive parenting and low supportive parenting also moderate relations between early toddler BI and anxiety symptoms from 9 to 15 years. In addition, much of the literature on overprotectiveness has mixed findings relating to child outcomes and neglects fathering behaviors (Jones et al., 2021; Moore et al., 2004). We were also limited to examining mothers’ parenting, but future work may benefit from understanding how fathers’ overprotective behaviors may exacerbate or buffer the effects of mothers’ parenting styles.

**Apparent next steps in the research**

Our results open avenues for new questions. The hint that sex might moderate the heritability of BI and BI’s heritable effect on SA raises the question of whether any of experiential measures that we examined might more strongly moderate these genetic effects. Of course, the possibility of extending the age span under study to early childhood and early adulthood is also intriguing for examining development outside the period of pubertal transition.

Although our analyses did not include biological measures, some clues to the neural processes underlying BI are emerging in the literature. For instance, children with high BI later developed anxiety symptoms that became more negatively associated with right amygdala–left dorso-lateral prefrontal cortex connectivity when attention was focused on threat (Abend et al., 2019). Including biological mediators of the effects of BI on SA are a necessary next step toward a fuller understanding of etiological processes. We pursue them elsewhere with MRI and neuroendocrine correlates of BI in children and adolescents (Adluru et al., 2017; Alisch et al., 2014; Bas-Hoogendam et al., 2022; Luo et al., 2022).

**Conclusion**

In sum, we present a comprehensive set of findings linking several moderators previously found in the literature to be significantly related to BI and SA. Though we partially replicated many previous findings, when included in a more comprehensive, robust model, only parenting stress and to a lesser extent SES significantly and independently interacted with BI to predict SA. This work, informed by both developmental and behavior-genetic models, highlights the importance of environmental moderators in models examining temperamental effects on later psychological symptoms, which will inform future designs as we continue to build the prototype for developmental models linking early temperament to emerging psychopathology.


