

Research Paper

Cite this article: Foster S, Collins S, Powers D, Rickman R, Miller J, Onono M, Wekesa P, Abrego M, Ramos-Gonzalez A, Whitley K, Bradley M, Young SL, and Widen EM (2026). Higher infant dietary diversity is associated with a reduced risk of stunting in the second year among a longitudinal cohort of HIV-unexposed and HIV-exposed but uninfected Kenyan children. *Public Health Nutrition* 29: e143, 1–9. doi: [10.1017/S1368980026102912](https://doi.org/10.1017/S1368980026102912)

Received: 30 April 2025 UTC

Revised: 22 April 2026 UTC

Accepted: 18 May 2026 UTC

Keywords:

Minimum dietary diversity; Stunting; Infant growth; Child growth; Infant diet; Prospective study; Longitudinal; Linear growth

Abbreviations:

ART, Antiretroviral therapy; HCAZ, Head circumference-for-age z-score; LAZ, Length-for-age z-score; MUAC, Middle-upper arm circumference; WAZ, Weight-for-age z-score; WLZ, Weight-for-length z-score

Corresponding author:

Elizabeth M. Widen;

Email: elizabeth.widen@austin.utexas.edu

Higher infant dietary diversity is associated with a reduced risk of stunting in the second year among a longitudinal cohort of HIV-unexposed and HIV-exposed but uninfected Kenyan children

Saralyn Foster¹, Shalean Collins², Daniel Powers¹, Rachel Rickman¹, Joshua Miller³, Maricianah Onono⁴, Pauline Wekesa⁴, Marcela Abrego¹, Anahi Ramos-Gonzalez¹, Kayleen Whitley¹, Madilyn Bradley¹, Sera L. Young⁵  and Elizabeth M. Widen¹ 

¹University of Texas at Austin, USA; ²Tulane University, USA; ³The University of North Carolina at Chapel Hill, USA; ⁴Kenya Medical Research Institute, Kenya and ⁵Northwestern University, USA

Abstract

Objective: Few studies have examined whether infant dietary diversity is prospectively associated with malnutrition risk in young children. We sought to assess whether dietary diversity and animal-source food (ASF) consumption were associated with malnutrition and body composition among children younger than 2 years at high risk of malnutrition. **Design:** Longitudinal cohort of HIV-exposed uninfected (HEU) or HIV-unexposed children with repeated measures of weight, length, circumferences and skinfold thicknesses from age 9 to 23 months from which WHO z-scores were calculated. FFQ assessed infants' dietary diversity, flesh foods and egg intake. Cox proportional hazards were used to examine whether minimum dietary diversity (MDD, dietary diversity scores ≥ 5) and consumption of ASF were associated with malnutrition, defined as z-score < -2 . **Setting:** Western Kenya **Participants:** Infants (n 232; 48 % female; 49 % HEU; growth z-scores above -2 at 9 months). **Results:** At 9 months, 42 % of infants met MDD, and mean z-scores were above -1 SD. Meeting MDD at 9 months was associated with a lower risk of stunting from 9 to 23 months (HR 0.51; 95 % CI: 0.33, 0.80) but not wasting or underweight. MDD at 9 months was associated with 86 % (95 % CI 0.11, 0.51) lower risk of low subscapular skinfolds; similar trends were observed for flesh foods. **Conclusions:** In a cohort of Kenyan infants, MDD at 9 months was associated with reduced risk of stunting through the second year of life. Flesh food intake was associated with lower risk of low trunk adiposity. Improving dietary diversity during infancy may protect against early-life stunting.

© The Author(s), 2026. Published by Cambridge University Press on behalf of The Nutrition Society. This is an Open Access article, distributed under the terms of the Creative Commons Attribution-NonCommercial-NoDerivatives licence (<https://creativecommons.org/licenses/by-nc-nd/4.0/>), which permits non-commercial re-use, distribution, and reproduction in any medium, provided that no alterations are made and the original article is properly cited. The written permission of Cambridge University Press or the rights holder(s) must be obtained prior to any commercial use and/or adaptation of the article.



Child malnutrition remains a persistent global challenge^(1,2), driving increased mortality⁽³⁾ and compromising lifelong economic⁽⁴⁾, physical⁽⁵⁾ and mental health outcomes⁽⁶⁾. Globally, stunting (height or length-for-age z-score (HAZ/LAZ) < -2 SD) and wasting (weight-for-height/length (WHZ/WLZ) < -2 SD) affect approximately 22 % (150.2 million) and 6 % (42.8 million) children under 5 years, respectively⁽¹⁾. While Kenya's national undernutrition rates are lower than the global average (18.4 % stunting and 4.9 % wasting), they still fall above global targets⁽²⁾, with stunting prevalence likely higher among HIV-exposed uninfected (HEU) infants⁽⁷⁾.

Multiple factors drive undernutrition among children under 5 years in East Africa, including food insecurity⁽⁸⁾, water insecurity⁽⁹⁾ and household poverty⁽¹⁰⁾. These vulnerabilities are compounded by maternal human HIV status, which exists in a syndemic relationship with deprivation and resource insecurity^(10–13). While the overall prevalence of HIV in Kenya decreased from 5.6 % in 2010 to 3.2 % in 2023, the number of women of reproductive age living with HIV has risen, increasing the risk of HIV exposure for paediatric populations⁽¹⁴⁾.

Early-life nutrition is a critical determinant of growth that is influenced by birth weight, breast-feeding practices and complementary feeding⁽¹⁵⁾. HEU infants often have lower birth weights and poorer growth trajectories than HIV-unexposed uninfected (HUU) counterparts^(16,17). Strikingly, our prior studies from Kenya suggest that breast-feeding rates may be higher among HEU infants than HUU infants, indicating the efficacy of breast-feeding promotion⁽¹⁸⁾. Other evidence has also shown that HIV-infected women are more likely to receive intensive counselling on exclusive breast-feeding^(19,20). However, less is known about complementary feeding, the introduction of solid foods alongside breastmilk or formula, in this context⁽²¹⁾.



In low- and middle-income settings, the adequacy of complementary feeding is often suboptimal due to inadequate access to nutritious foods⁽²²⁾. Cross-sectional studies have shown that low consumption of animal-source foods (ASF) – which provide essential protein, Fe and Zn – is associated with greater risk of stunting and wasting between 6 and 24 months of age⁽²³⁾. Meeting minimum dietary diversity (MDD) (consumption of at least five of eight specific food groups) is a common indicator of complementary feeding adequacy^(24,25). Cross-sectional and longitudinal studies have linked higher MDD to reduced risks of malnutrition, yet many of these analyses use earlier definitions of dietary diversity that excluded breastmilk^(26–29) or failed to account for the unique growth trajectories of HEU infants, who may have an elevated risk of malnutrition⁽³⁰⁾.

Therefore, the objective of this study was to evaluate the association between infant dietary diversity assessed at 9 months with subsequent growth and anthropometric measures of body composition, including risk of stunting and wasting, from 9 to 23 months of age among a cohort of HEU and HUU in Western Kenya. We also evaluated whether the consumption of ASF beginning at 9 months influenced later child growth and body composition outcomes. We hypothesized that infants with higher dietary diversity scores (DDS) and ASF consumption will have a lower likelihood of future stunting or wasting, as well as higher growth and body composition z-scores, compared to those with lower DDS and no ASF intake.

Methods

Study sample

Data were from the Pith Moromo 2 study, a longitudinal cohort designed to study the health consequences of food insecurity and HIV infection during pregnancy and lactation. From 2014 to 2017, 371 pregnant women were recruited from seven Family AIDS Care and Educational Services clinics in rural and urban Nyanza, Kenya, using non-probability, stratified sampling procedures based on HIV status (known, positive and newly diagnosed) and food insecurity (low, medium and high). The study has been described in detail previously^(18,31). Briefly, pregnant women were eligible if they were ≥ 18 years of age, gestational age ≤ 30 weeks, had known HIV status and had no plans to move from the study area. Prenatal in-person visits were conducted at 6 and 8 months' gestation.

Study design

Postpartum telephone surveys for the maternal–child dyads were collected at 1 week and 6 months. Postpartum in-person dyad visits were held at approximately infant/child age 6 weeks and 3, 9, 12, 15, 18, and 23 months. This secondary analysis focuses on child visits from 9 to 23 months, as 9 months is the first visit where complementary feeding had been established. Similar to other research^(32,33), the 9-month visit was chosen as the time frame to base dietary analysis as it was the first visit following the start of complementary feeding and is late enough in the complementary feeding period where a larger variety of foods should have been introduced and accepted. The 6-month visit was a telephone visit that did not include anthropometric measurements.

Child anthropometric measures

At each in-person study visit, child body composition measurements were completed by trained staff. Weight (kg, Seca 874),

length (cm, Seca 417), head circumference (cm, non-stretchable retractable tape measure) and mid-upper arm circumference (MUAC, cm; non-stretchable retractable tape measure) were obtained. Skinfold thickness was measured (mm) at the triceps, subscapular, iliac crest and mid-thigh using Harpenden calipers (Baty International). Child weight was measured by having the mother stand on the scale, zeroing out the scale and then handing the child to the mother. Child measures were completed at least twice each visit, using a pre-specified threshold for conducting a third measure, and the mean of the measure was calculated.

Z-score conversions

Growth measurements were converted to WHO z-scores for weight (WAZ), length (LAZ), weight-for-length (WLZ), MUAC (MUACZ), triceps (TSFZ) and subscapular (SSFZ) skinfold thickness using zanthro in Stata⁽³⁴⁾. Poor growth was defined according to the WHO guidelines: (1) wasting: weight-for-length (WLZ) < -2 SD; stunting: height-for-age z-score (LAZ) < -2 SD; underweight: weight-for-age (WAZ) < -2 SD⁽³⁵⁾. A cut-off of below -2 SD was also used for skinfold thicknesses and MUAC measurements.

Dietary measures

Maternal interviews assessing child dietary intake of certain foods in the prior 24 h were completed at all visits after 6 months. The interviewer asked mothers to answer yes or no if their child had eaten a food from twenty-six different food types. For example, one food type was 'any meat, such as beef, pork, lamb, goat, chicken or duck', while another was 'baby cereal, porridge, bread, rice or other foods made from grains'. These twenty-six food types were grouped into the eight food categories recommended by the WHO and UNICEF⁽³⁶⁾. This information was used to calculate DDS (0–8) and meeting MDD – consuming at least five of the eight food categories during the previous day: breastmilk; grains, tubers, and plantains; beans, peas, lentils, nuts, and seeds; dairy products; flesh foods (e.g. fish, meat, poultry and offal); eggs; vitamin A-rich fruits and vegetables; and other fruits and vegetables. Dietary diversity is a global measure that quantifies the number of food groups out of 8 (listed above) in a diet during a specific period and is used as a potential indicator of nutritional adequacy⁽³⁷⁾. In the 2021 guidance from the WHO and UNICEF, the minimal acceptable diet was included as an additional measure of diet adequacy and is comprised of MDD, minimum meal frequency and minimum milk feeding frequency⁽³⁶⁾. Minimum meal frequency is achieved by receiving solid, semi-solid or soft foods the minimum number of times or more the previous day and minimum milk feeding frequency by having at least two milk feeds the previous day. MAD is then calculated using an established formula⁽³⁶⁾.

The Institutional Review Boards at Northwestern University, Cornell University and The University of Texas at Austin approved the study procedures. All mothers provided written informed consent.

Statistical analysis

Statistical analyses were conducted in Stata 17 using an α level of < 0.05 . To be included in the complete case analytic sample, infants had to be HIV-negative and have anthropometric measures and dietary data at 9 months as they were the outcome and predictor variables, respectively. Infants who seroconverted at any point during the study ($n = 2$) were not included, as their growth would likely be affected by this change in HIV status. Additionally,

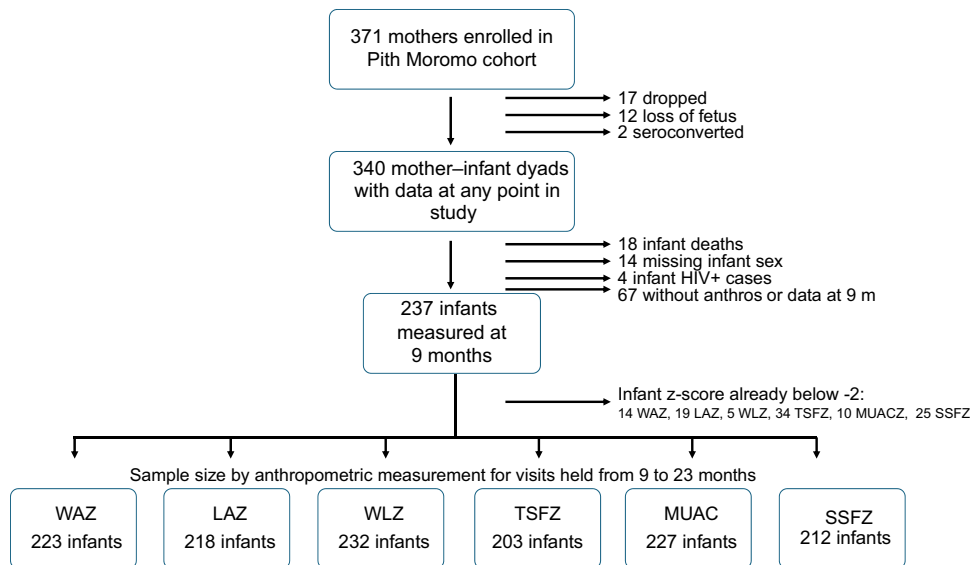


Figure 1. Flow diagram from enrollment in Pith Moromo 2 to inclusion into analytic sample for each growth measure. WAZ, weight-for-age z-score; LAZ, length-for-age z-score; WLZ, Weight-for-length z-score; TSFZ, triceps skinfold z-score; SSFZ, subscapular skinfold z-score; MUACZ, middle-upper arm circumference z-score.

those with missing sex ($n = 7$) data were not included because their z-scores could not be calculated (Figure 1).

Sensitivity analysis of those included *v.* not included in the sample showed no notable differences (Table 1, although those not included in the sample had an earlier age of food introduction (1.61 months *v.* 2.6, $P = 0.04$), compared to those included in the analysis. Independent *t* tests were conducted to assess differences between maternal–infant dyads by MDD status (dichotomised yes/no) at 9 months. Dietary diversity was examined as two different exposures: MDD (dichotomised yes/no) and continuous DDS (scale 0–8); we also assessed the presence of specific ASF types, including flesh foods (meat, fish, poultry and organ meats) and eggs.

We conducted Cox proportional hazard analysis to identify the association between time to z-score -2 and the exposure – MDD and specific food intake. The proportional hazard assumptions were tested using estat phtest in Stata. To properly fit the Cox hazards model, the infants included could not have the presence of the growth condition (z-score already below -2 SD), as they were already below the endpoint and could not be included in the hazard model. The Cox models attempt to explain how an event in time, diet at 9 months, is associated with later changes in body composition z-score. Unadjusted and adjusted Cox proportional hazard regression models assessed associations between dietary exposures and child anthropometric measures that dropped below a -2 z-score threshold. Failure time was set as the first occurrence a child reached a z-score of -2 for the anthropometric outcome, including WAZ, LAZ, WLZ, triceps skinfold z-score (TSFZ), subscapular skinfold z-score (SSFZ), MUAC z-score and head circumference z-scores (HCAZ). Additional analyses were conducted to assess whether the same associations were observed for time-varying dietary diversity among the children included in the Cox models. For this analysis, adjusted and unadjusted longitudinal regression models assessed associations between lagged dietary diversity from the prior visit (9–23 months) with growth outcomes from 12 to 23 months. The Cox model allows for a dietary measure at a point in time to be associated with the risk of failure over time, while longitudinal regression allows dynamic dietary diversity to fluctuate and follow along as growth. Lagged growth is used, as diet at 9 months does not affect growth at 9 months. Diet at 9 months will affect growth at later time points.

Model covariates included infant sex, food insecurity score, household asset index, hunger season, water insecurity and maternal HIV status. Covariates were determined based on previous research demonstrating the impact of these factors on growth in similar populations⁽¹⁸⁾. Food insecurity was calculated using the 9-Item Individual Food Insecurity Access Scale (range 0–27) at 9 months in relation to diet⁽³⁸⁾. Hunger season is a seasonal period when rainfall is minimal and thus fewer agricultural goods are produced with lower food availability. Hunger season was analysed as a time-varying binary variable aligned with study visits⁽³¹⁾. If a visit occurred from February to May, the visit was noted as ‘yes’, occurring during the hunger season. Visits from June through January were marked as not occurring during the hunger season. Household asset index included eighteen items commonly owned by households in the context and was created through item reduction via principal component analysis; higher scores indicated greater accumulated household wealth. Water insecurity was measured from a twenty-item scale previously validated in this population, evaluating water availability for drinking and other household activities as well as worry and missed opportunities due to lack of water in the prior 4 weeks⁽³⁹⁾. Additional maternal indicators such as mother’s nutritional status and disease progression were not included as covariates as maternal determinants more strongly affect early infancy growth, while infant’s dietary behaviour, activity levels and motor development more strongly influence later growth.

Results

A total of 232 children were included in the analysis (Figure 1). There were no differences in baseline or complementary feeding data between those included and those excluded in the study, except that those excluded from this analysis had an earlier introduction to complementary foods (Table 1). The mean food insecurity score was 13.0, indicating mild food insecurity (FI, range 11–14), with 45 % of households experiencing mild FI. None of the covariates had statistically significant differences between those included *v.* not included in the sample or those who met compared to those who did not meet MDD.

At the 9-month visit, 41 % of the infants met MDD. Infants who met MDD had an average DDS that was 3.3 points higher than those who did not meet (met 5.1 ± 1.0 *v.* unmet 1.8 ± 1.0). Dyads in the analytic sample who met MDD (60 %) had birth parents who

Table 1. Characteristics of Pith Moromo 2 cohort of mother–infant dyads overall, and by inclusion in analytic sample

	All <i>n</i> 340				Analytic sample <i>n</i> 232				Not included in sample [†] <i>n</i> 78				<i>P</i> -value [†]
	<i>n</i>	%	Mean	SD	<i>n</i>	%	Mean	SD	<i>n</i>	%	Mean	SD	
Maternal													
Age, years			24.8	4.8 [‡]			25.0	4.8			24.2	4.8	0.21
First pregnancy	69	22 [§]			48	20			41	28			0.15
BMI, kg/m ²			24.2	3.5			24.2	3.5			24.3	3.5	0.84
HIV-positive	154	49			115	49			39	52			0.95
Parity, first child	69	22			46	20			23	29			0.08
Food insecurity, 0–27			12.9	5.1			13.0	5.1			12.6	5.1	0.57
Household asset index			0.25	1.7			0.29	1.7			−0.16	1.9	0.77
Water insecurity [¶]			7.8	8.2			8.2	8.5			6.2	6.6	0.11
Infant													
Female	148	50			108	48			40	55			0.34
Born in hunger season	136	44			109	47			27	35			0.06
Age of food intro, months ^{††}			2.4	2.8			2.5	2.8			1.9	2.5	0.20
Exclusive BF at 6 months	67	21			53	23			14	18			0.36

[†]Comparisons between those with 9-month data and those missing 9-month data were conducted using parametric tests for continuous, normally distributed variables and non-parametric tests for non-normally distributed variables.

[‡]Mean (sd) (all such values).

[§]*n* (%) (all such values).

^{||}Household asset index score −2.3–4.9; higher score equals more assets.

[¶]20 question scale; score 0–80; mean provided from last four visits.

^{††}Data are available for only 274.

[†]Comparisons between those in and out of the analytic sample were conducted.

were younger ($P < 0.05$), experienced lower rates of depression ($P < 0.05$), had infants who were born during the hunger season ($P < 0.01$) and were longer by 23 months ($P < 0.05$) compared to those who did not meet MDD (Table 2, see online supplementary material, Supplemental Table 2). In the sample, both average LAZ and WAZ decreased over time from the 9-month to 23-month visit. For all infants, WLZ and HCAZ stayed relatively stable over time and above the median z-score. TSFZ, SSFZ and MUACZ increased from 9 months to 23 months (Figure 2, see online supplementary material, Supplemental Table 2). At 9 months, infants' WAZ score in both those meeting MDD and those not meeting was just below the median (Figure 2). LAZ score in those not meeting MDD was modestly lower than in the group not meeting MDD, but not statistically different.

Both TSFZ and MUACZ at 9 months were statistically significant between those meeting and those not meeting MDD at 9 months. Those meeting MDD had a lower TSFZ compared to those not meeting at 9 months. By 23 months, the difference was no longer statistically significant. There was no significant difference between average z-scores in any outcome for those meeting compared to not meeting MDD at 9 months.

Dietary diversity and animal-source foods with risk of adverse growth from 9 to 23 months

In Cox proportional hazard models, meeting MDD at 9 months was associated with a 51 % lower risk of stunting from 9 to 23 months and, further, 76 % and 58 % lower risk of z-scores falling below −2 SD for subscapular skinfolds z-score and MUAC z-score, respectively (Table 3, see online supplementary material, Supplemental Figure 1). In analyses using continuous DDS, each one-unit increase in DDS was associated with a 13 % reduced risk of stunting, a 24 % lower risk of

low z-score for SSFZ and a 17 % lower risk for MUACZ (Table 3). Meeting MDD and continuous DDS were not associated with the risk of falling below −2 SD for WAZ, WLZ, TSF and HCAZ. No differences were noted between WAZ, WLZ, TSFZ and HCAZ in those meeting *v.* not meeting MDD at 9 months.

Following similar patterns for dietary diversity measures, intake of any flesh foods (i.e. meat, fish, poultry or organ meats) by 9 months was associated with an 85 % reduced risk of SSFZ falling below −2. Egg intake by 9 months was not associated with risk of stunting or z-scores falling below −2 for any anthropometric measure.

Similar to the hazard models, in our sensitivity analysis with the longitudinal regression models, a one-point increase in DDS across the study was associated with an average 0.04 higher LAZ score and meeting MDD was associated with an average 0.18 z-score (Table 4, see online supplementary material, Supplemental Figure 2). Unlike in the Cox models, intake of flesh foods during the 9–23 month period was associated with 0.19 higher LAZ from 12 to 23 months. A positive association was found between DDS and SSFZ, while a negative association was found between flesh food intake and WLZ (Table 4, see online supplementary material, Supplemental Figure 3). Finally meeting MDD was associated with 0.13 higher HCAZ score. In a sensitivity analysis including all infants with a 9-month visit, regardless of z-score, regression models demonstrated similar associations for all anthropometric outcomes as the primary models shown in the paper.

Discussion

In this cohort of Kenyan children who were HIV-exposed, uninfected or HIV-unexposed, we sought to examine the association between dietary diversity and ASF intake at 9 months with

Table 2. Characteristics of mother–infant dyads by MDD included in analytic sample, *n* 232

	Below MDD [†] (<i>n</i> 134)				Met MDD (<i>n</i> 98)			
	<i>n</i>	%	Mean	SD	<i>n</i>	%	Mean	SD
Maternal								
Age, years			24.5	4.6			25.8	5.0*
BMI, kg/m ²			23.9	3.1			24.5	3.9
HIV-positive	73	47			55	53		
Parity, first child	27	20			19	19		
Food insecurity score			13.5	5.2			12.2	4.8
Depression [‡]	52	33			20	20*		
Household asset index			0.11	1.5			0.54	1.8
Water insecurity			8.3	7.5			8.0	9.8
Infant								
Male	71	48			56	56		
Exclusively breastfed, 3 months	79	51			63	61		
Any breastfed, 9 months	125	85			96	96		
Age at 9-month visit			9.3	0.7			9.4	0.4
DDS [§] at 9-month visit			1.8	1.0			5.1	1.0 [^]
Hunger Season Birthday	58	37			59	57**		
Flesh food at 9-month visit	6	4			66	66 [^]		
Eggs at 9-month visit	8	6			49	49 [^]		

[†]MDD, minimum dietary diversity: below MDD (DDS ≤ 5); met MDD (DDS > 5).

[‡]Depression measured by CESD, scale 0–60.

[§]DDS, Dietary Diversity Score, range 0–8; z-scores: Testing for differences by MDD.

P* < 0.05, *P* < 0.01, [^]*P* < 0.001.

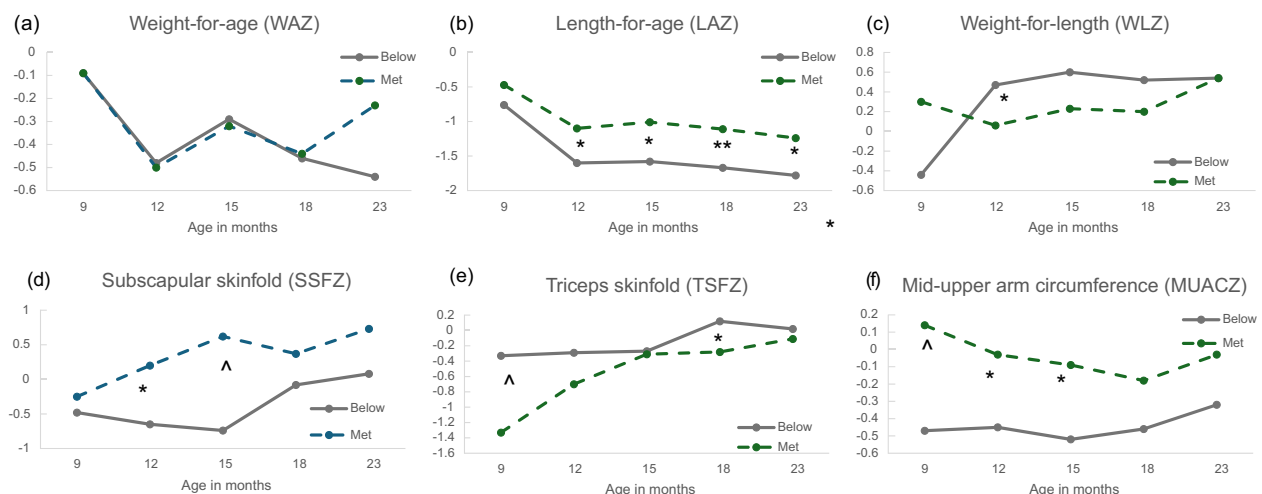


Figure 2. Mean z-scores of children at visits from 9 to 23 months of age in Pith Moromo 2, by minimum dietary diversity (MDD) status (met, MDD ≥ 5 or below, MDD < 5) for (a) weight-for-age z-score, (b) length-for-age z-score, (c) weight-for-length z-score, (d) subscapular skinfold thickness z-score, (e) triceps skinfold z-score and (f) mid-upper arm circumference z-score. Not shown head circumference z-score with no significant differences between the groups; **P* < 0.05, ***P* < 0.01, [^]*P* < 0.001.

subsequent growth in the second year of life. Although most of the children in the study had inadequate growth during the first 2 years, those who met MDD had a lower risk of stunting, low trunk adiposity and low MUAC, reflecting a reduced risk of impaired linear growth, low central adiposity and smaller upper arm composition, compared to those who did not meet MDD. Using

longitudinal regression models, we also examined whether dietary diversity over time was associated with growth and found that dietary diversity over time was positively associated with subsequent LAZ and SSFZ, suggesting higher dietary diversity may have persistent associations across the second year of life on linear growth and subcutaneous truncal adiposity. Together, these

Table 3. Associations between dietary diversity and dietary exposures with z-scores falling below -2 sd from 9–23 months, Pith Moromo 2

	Cox proportional hazard models							
	Met MDD		DDS (0–8)		Flesh foods		Egg	
	HR	95 % CI	HR	95 % CI	HR	95 % CI	HR	95 % CI
WAZ (<i>n</i> 223) 22 failures	0.87	0.36, 2.12	0.91	0.74, 1.13	0.84	0.29, 2.42	0.81	0.27, 2.44
LAZ (<i>n</i> 218) 96 failures	0.51	0.33, 0.80**	0.87	0.79, 0.96**	0.60	0.35, 1.00	0.89	0.54, 1.47
WLZ (<i>n</i> 232) 10 failures	0.94	0.25, 3.56	0.83	0.60, 1.14	1.05	0.23, 4.72	–	
TSFZ (<i>n</i> 199) 41 failures	0.92	0.47, 1.78	0.95	0.81, 1.11	1.19	0.57, 2.46	0.86	0.37, 1.99
SSFZ (<i>n</i> 212) 48 failures	0.24	0.11, 0.51^	0.76	0.66, 0.88^	0.13	0.04, 0.45**	0.83	0.41, 1.68
MUACZ (<i>n</i> 227) 29 failures	0.42	0.18, 0.98*	0.83	0.69, 0.99*	0.35	0.12, 1.01	0.87	0.35, 2.15
HCAZ (<i>n</i> 230) 6 failures	1.13	0.21, 6.22	0.98	0.68, 1.40	0.62	0.58, 6.72	0.88	0.09, 8.24

MDD, minimum dietary diversity; DDS, dietary diversity score; WAZ, weight-for-age z-score; LAZ, length-for-age z-score; WLZ, Weight-for-length z-score; TSFZ, triceps skinfold z-score; SSFZ, subscapular skinfold z-score; MUACZ, middle-upper arm circumference z-score; HCAZ, head circumference-for-age z-score.

Results are presented as hazard ratios, relative to 1 with values above 1 indicating increased hazard and values below 1 indicating lower hazard. Models adjusted for maternal HIV status, hunger season, food insecurity, infant sex, household asset index and water insecurity. Children excluded if 9-month data not available or already below cut-off. Analysis time: infant age in month. *P*-value: * < 0.05, ** < 0.01, ^ < 0.001.

Table 4. Associations between lagged dietary diversity and dietary exposures with continuous growth measures from 9 to 23 months, Pith Moromo 2

	Longitudinal linear regression (results are relative to 0 with numbers above 0 indicating higher values)							
	Met MDD		DDS (0–8)		Flesh foods		Egg	
	β	95 % CI	β	95 % CI	β	95 % CI	β	95 % CI
WAZ (<i>n</i> 223)	0.09	-0.05, 0.23	0.001	-0.03, 0.03	0.01	-0.12, 0.15	0.03	-0.11, 0.17
LAZ (<i>n</i> 218)	0.18	0.004, 0.37*	0.04	0.01, 0.08*	0.19	0.02, 0.36*	0.05	-0.12, 0.21
WLZ (<i>n</i> 232)	-0.12	-0.32, 0.09	-0.03	-0.07, 0.01	-0.21	-0.40, -0.19*	-0.04	-0.23, 0.15
TSFZ (<i>n</i> 199)	-0.22	-0.48, 0.28	-0.005	-0.05, 0.04	-0.02	-0.27, 0.23	-0.07	-0.32, 0.18
SSFZ (<i>n</i> 212)	-0.15	-0.46, 0.15	0.07	0.01, 0.13*	-0.08	-0.38, 0.21	0.06	-0.22, 0.35
MUACZ (<i>n</i> 227)	0.09	-0.08, 0.25	0.008	-0.02, 0.04	0.04	-0.11, 0.20	0.11	-0.04, 0.27
HCAZ (<i>n</i> 230)	0.13	0.01, 0.25*	0.005	-0.18, 0.03	0.06	-0.06, 0.19	0.11	-0.01, 0.23

MDD, minimum dietary diversity; DDS, dietary diversity score; WAZ, weight-for-age z-score; LAZ, length-for-age z-score; WLZ, Weight-for-length z-score; TSFZ, triceps skinfold z-score; SSFZ, subscapular skinfold z-score; MUACZ, middle-upper arm circumference z-score; HCAZ, head circumference-for-age z-score.

Results are presented as β coefficients, relative to 0 with values above 0 indicating higher scores. Each value is showing an estimate from an individual regression model. Models are adjusted for maternal HIV status, hunger season, food insecurity, infant sex, household asset index and water insecurity.

P-value: * < 0.05.

findings support some of our hypotheses that meeting MDD may support more optimal growth.

Minimum dietary diversity and anthropometric changes

Interestingly, in the Cox models, we found that MDD was not associated with HCAZ, WLZ, WAZ or TSFZ, but it was associated with lower risk of z-score falling below -2 for LAZ, SSFZ and MUACZ. While overall linear growth (LAZ) in the sample was faltering, relative size (WLZ), weight and upper arm skinfolds remained relatively stable within 1 SD, which may be why we observed no association between WLZ and WAZ with any of the dietary exposures. Our longitudinal regression models found an inverse association between lagged flesh food intake and WLZ, with intake of flesh foods being associated with lower weight relative to length z-scores. One reason could be children who had flesh food intake could potentially be longer in length but have similar weight to those without flesh food intake reported during the interview. These children may appear thinner but, as evidenced by our results in Table 3, still maintain a z-score within acceptable

boundaries. Interestingly, both our Cox and longitudinal models show a positive association between flesh food intake and LAZ. Similar findings were also found by Duong *et al.* on associations with dietary diversity and linear growth but not ponderal (weight (kg)/height (m)) or relative size in children up to 25 months of age⁽⁴⁰⁾.

Lower dietary diversity has been associated with stunting in both cross-sectional and prospective studies^(29,41,42). Similar to this study, in a cohort study of Tanzanian infants by Kamenju *et al.*, a higher DDS from 6 to 24 months protected against stunting and underweight at 24 months⁽²⁹⁾; yet, body composition was not assessed, so body fat, another indicator of growth and nutritional risk, could not be evaluated. In our findings, meeting MDD at 9 months was associated with both a decreased risk of z-score -2 (indicative of malnutrition) for MUACZ and SSFZ, suggesting that both adiposity and muscle mass may be improved by the early introductions to a variety of foods.

Unlike our findings, others have found associations between dietary diversity with wasting and underweight in cross-sectional research, which is unable to disentangle temporality between

associations^(26,43). Aboagye *et al.* used cross-sectional survey data in children aged 6–23 months in sub-Saharan Africa and found adequate MDD was associated with a 13 % lower likelihood of wasting and a 17 % lower likelihood of being underweight⁽²⁶⁾. Motadi *et al.* found, in their cross-sectional study of 273 children aged 3–4 years from South Africa, that weight-for-age but not weight-for-length was associated with DDS⁽⁴³⁾. However, a recent systematic review including both cross-sectional and longitudinal studies found that current evidence indicates possible associations between dietary diversity and stunting, which aligns with our findings; however, this association for wasting and underweight remained uncertain⁽²⁷⁾.

Flesh food intake and anthropometric changes

Flesh foods were consumed by two-thirds of those meeting MDD but less than 10 % of those not meeting MDD. Flesh food consumption at 9 months was also associated with lower risk of SSFZ falling below -2 SD growth in our analysis, indicating that intake of flesh foods may also be associated with adiposity possibly by reducing fat wasting. Few studies have evaluated skinfold thickness measures with dietary diversity. However, these important measures can help identify how weight and subcutaneous adipose tissue are distributed and how they relate to health. Our study is likely the first to report outcomes on skinfold thickness associations with dietary diversity in children aged 9 to 23 months in a population of HEU and HUU children.

Egg intake and anthropometric changes

Our findings that egg consumption at 9 months was not associated with the risk of malnutrition up to 23 months are in contrast to Pasqualino *et al.*, who found a difference in higher egg intake from infant age 6 to 12 months with increased WAZ at infant age 12 months in Bangladesh⁽⁴⁴⁾. A randomised control trial of 500 South African infants enrolled at age 6 to 9 months with the provision of one egg per day for 6 months *v.* no egg found no differences in LAZ, WAZ or WLZ between the two groups at child age 12–15 months⁽⁴⁵⁾. Our longitudinal regression models also found no association between lagged egg intake over time and anthropometric measurements across the 9–23 month period. Differences in outcomes from the previous research compared to ours could be related to portion size and frequency of intake, neither of which we were able to capture.

While specific foods provide nutrients necessary for human growth and survival (such as the Fe and Zn in meats), other foods can also provide these nutrients. The ability of multiple foods to provide essential nutrients is likely one reason we observed associations between overall dietary diversity and the risk of stunting, subscapular skinfold thickness, and MUACz but not with specific food groups.

Our findings highlight that early DD is important for supporting linear growth. Yet, research has shown that improving access to foods does not always change dietary diversity⁽⁴⁶⁾. The Supporting Healthy Mothers study, a feeding support programme with education and monetary support in Kenya for women with HIV, changed the nutrition behaviour of exclusive breast-feeding for 3 months⁽⁴⁷⁾. Community-based programmes and other policies and interventions should incorporate nutrition components addressing dietary diversity practices to not only improve access but also dietary diversity and growth.

This study has limitations that should be noted. One limitation is that we did not collect information on meal frequency, so we

were unable to calculate minimal meal frequency and MAD. In addition, due to the nature of the dietary diversity questions, the amounts of each type of food eaten were unknown, so we could not calculate estimated energy intake and therefore cannot know if children with greater dietary diversity and more favourable eating profiles were eating more food or just higher-quality food. To address this limitation, we performed longitudinal regressions to evaluate the relationship of dietary diversity with growth from 9 to 23 months. Moreover, infant dietary diversity is validated as a tool for measuring nutritional exposures in this context, and it allowed us to contextualise how different dietary components were associated with growth⁽²⁴⁾. While skinfolds are a proxy for adiposity in children, there are limitations to the measures such as difficulty in measuring squirming infants and the influence of hydration status or illness. Due to the small sample size, subgroup analysis comparing exposed *v.* unexposed infants was not conducted. Finally, we acknowledge that the missingness of data at 9 months and subsequent visits may underreport our low growth measures. Due to Cox model parameters, our analysis excluded infants who had suboptimal growth at 9 months; thus, our findings are generalisable to infants with similar growth parameters. Other limitations of Cox models are low event rates, although our sample had adequate events, and competing risks, which may bias results.

Nevertheless, this paper has several strengths, including the prospective longitudinal design and the frequency of repeated anthropometric measurements across the second year of life. Another strength is that almost half of the cohort was HEU infants, allowing us to evaluate growth in the first two years in this group of infants who are understudied at higher risk of poor growth⁽¹⁸⁾ and who are increasing part of Kenya's infant population. Cox proportional hazards have not been a common method to evaluate diet in relation to growth in infant populations, and our paper is one of the first to examine associations between dietary diversity and time to event malnutrition, as a measure of nutritional health. Abebe *et al.* used Cox modelling in 476 Ethiopian infants and children to estimate the time to recovery from malnutrition (returning to a z-score > -2)⁽⁴⁸⁾. Tewoldie *et al.* used time to full enteral feeding in a low birth weight infant population⁽⁴⁹⁾. Similar to our research, Ejigu *et al.* used Cox hazards to evaluate factors associated with stunting and underweight but did not assess sixteen dietary measures⁽⁵⁰⁾. In prospective, longitudinal studies like ours, Cox proportional hazard models should be considered as a way to understand early-life predictors of aberrant growth before they occur.

In conclusion, in our longitudinal cohort of HEU and HUU children, meeting MDD at 9 months was associated with reduced risk of stunting through 23 months, compared to those not meeting MDD. Improving complementary feeding by increasing dietary diversity and supporting the intake of flesh food throughout the first two years may help reduce the incidence of stunting and support optimal growth in HEU and HUU children. Our results highlight the need for early-life complementary feeding interventions to support children living in low- and middle-income countries. Policies that increase access to foods and other resources to increase dietary diversity are needed to mitigate the risk of stunting and support healthy growth. Our findings further suggest that we must continue to focus on improving the feeding environment in lower- and middle-income countries to improve infant and global health, especially in vulnerable populations exposed to HIV.

Supplementary material. For supplementary material accompanying this paper visit <https://doi.org/10.1017/S1368980026102912>

Acknowledgements. The authors are thankful for the support and collaboration of the Kenya Medical Research Institute (KEMRI); the Director General of KEMRI and the Deputy Director of the Centre for Microbiology Research (CMR); the Pith Moromo and Pii Ngima participants; the study nurses Joy China, Joyce Bonke and Tobias Odwar; the study trackers Benter Ogwana, Teresa Owade and Sarah Obaje and the anthropologist Patrick Mbullo.

Financial support. Research activities and SLY were supported by the National Institute of Mental Health (NIH/NIMH K01MH098902; R21MH108444). EMW was supported by the Eunice Kennedy Shriver National Institute of Child Health & Human Development (K99/R00 HD086304 and P2CHD042849), the National Institute of Diabetes and Digestive and Kidney Diseases (NIH/NIDDK T32DK091227; T32DK007559) and an unrestricted fellowship to support research in maternal and child health from PepsiCo Global R + D. The content is solely the responsibility of the authors and does not necessarily represent the official views of the National Institutes of Health.

Conflict of interest. The authors declare no conflicts of interest.

Authorship. E.M.W., M.O. and S.Y. designed the PM study; S.M.C., P.W. and J.D.M. conducted the research; S.F.F., S.Y. and E.M.W. designed the secondary analysis; S.F.F., R.R.R., D.P. and E.M.W. analysed the data; S.F.F., S.M.C., J.D.M., D.P., S.Y. and E.M.W. interpreted the findings; S.F.F., S.Y. and E.M.W. wrote the paper. E.M.W. has primary responsibility for final content. All authors read and approved the final manuscript.

Ethics of human subject participation. This study was conducted according to the guidelines laid down in the Declaration of Helsinki, and all procedures involving research study participants were approved by the Institutional Review Boards at Cornell University, Northwestern University, and the Kenya Medical Research Institute. Written informed consent was obtained from all mothers.

References

- United Nations Children's Fund (UNICEF), World Health Organization, International Bank for Reconstruction and Development/The World Bank (2025) Levels and trends in child malnutrition: UNICEF / WHO / World Bank Group Joint Child Malnutrition Estimates. Key findings of the 2025 edition. Geneva: World Health Organization.
- United Nations Children's Fund (UNICEF), World Health Organization (WHO), International Bank for Reconstruction and Development/The World Bank (2023) Levels and trends in child malnutrition: UNICEF / WHO / World Bank Group Joint Child Malnutrition Estimates: Key findings of the 2023 edition. New York: UNICEF and WHO.
- Kassaw A, Amare D, Birhanu M, et al. (2021) Survival and predictors of mortality among severe acute malnourished under-five children admitted at Felege-Hiwot comprehensive specialized hospital, northwest, Ethiopia: a retrospective cohort study. *BMC Pediatr* **21**, 176.
- Akseer N, Tasic H, Nnachebe Onah M, et al. (2022) Economic costs of childhood stunting to the private sector in low- and middle-income countries. *EClinicalMedicine* **45**, 101320.
- Soni A, Fahey N, Bhutta ZA, et al. (2021) Early childhood undernutrition, preadolescent physical growth, and cognitive achievement in India: a population-based cohort study. *PLoS Med* **18**, e1003838.
- Roger K, Vannasing P, Tremblay J, et al. (2022) Impact of early childhood malnutrition on adult brain function: an evoked-related potentials study. *Front Hum Neurosci* **16**, 884251.
- Neary J, Langat A, Singa B, et al. (2022) Higher prevalence of stunting and poor growth outcomes in HIV-exposed uninfected than HIV-unexposed infants in Kenya. *Aids* **36**, 605–610.
- FAO, IFAD, UNICEF, WFP and WHO. (2025) *The State of Food Security and Nutrition in the World 2025 – Addressing High Food Price Inflation for Food Security and Nutrition*. Rome. <https://doi.org/10.4060/cd6008en>
- Krumdieck NR, Collins SM, Wekesa P, et al. (2016) Household water insecurity is associated with a range of negative consequences among pregnant Kenyan women of mixed HIV status. *J Water Health* **14**, 1028–1031.
- Kalichman SC (2023) Ending HIV hinges on reducing poverty. *AIDS Behav* **27**, 1–3.
- The Lancet HIV (2020) The syndemic threat of food insecurity and HIV. *Lancet HIV* **7**, e75.
- Boateng GO, Workman CL, Miller JD, et al. (2022) The syndemic effects of food insecurity, water insecurity, and HIV on depressive symptomatology among Kenyan women. *Soc Sci Med* **295**, 113043.
- Workman CL & Ureksoy H (2017) Water insecurity in a syndemic context: understanding the psycho-emotional stress of water insecurity in Lesotho, Africa. *Soc Sci Med* **179**, 52–60.
- Joint United Nations Programme on HIV/AIDS (2024) UNAIDS Data 2024. Geneva: Joint United Nations Programme on HIV/AIDS.
- Adair LS & Guilkey DK (1997) Age-specific determinants of stunting in Filipino children. *J Nutr* **127**, 314–320.
- Alvarez GG, Miller JD, Santoso MV, et al. (2021) Prevalence and covariates of food insecurity across the first 1000 days among women of mixed HIV status in Western Kenya: a longitudinal perspective. *Food Nutr Bull* **42**, 319–333.
- Tiwari R, Singa BO, Lihanda P, et al. (2025) Differences in growth trajectories in breastfed HIV-exposed uninfected and HIV-unexposed infants in Kenya: an observational cohort study. *PLoS Med* **22**, e1004781.
- Rickman RR, Lane CE, Collins SM, et al. (2023) Body composition trajectories during the first 23 months of life differ by HIV exposure among infants in Western Kenya: a prospective study. *J Nutr* **153**, 331–339.
- Okanda J, Otieno G, Kinuthia J, et al. (2018) Higher likelihood of 6-months exclusive breastfeeding among HIV infected than uninfected mothers: a household survey in Kenya. *Int Breastfeed J* **13**, 51.
- Nabakwe EC, Egesah O & Kiverenge-Ettinger GA (2022) Maternal and health care workers' perspectives on exclusive breastfeeding in the context of maternal HIV infection, in Busia county, western Kenya: a mixed methods cross-sectional survey. *Int Breastfeed J* **17**, 17.
- Grote V, Theurich M, Luque V, et al. (2018) Complementary feeding, infant growth, and obesity risk: timing, composition, and mode of feeding. *Nestle Nutr Inst Workshop Ser* **89**, 93–103.
- Sokhela H, Govender L & Siwela M (2023) Complementary feeding practices and childhood malnutrition in South Africa: the potential of Moringa Oleifera leaf powder as a fortificant: a narrative review. *Nutrients* **15**, 2011.
- Tadele TT, Gebremedhin CC, Markos MU, et al. (2022) Stunting and associated factors among 6–23 month old children in drought vulnerable kebeles of Demba Gofa district, southern Ethiopia. *BMC Nutr* **8**, 9.
- Heidkamp RA, Kang Y, Chimanya K, et al. (2020) Implications of updating the minimum dietary diversity for children indicator for tracking progress in the Eastern and Southern Africa region. *Curr Dev Nutr* **4**, nzaa141.
- Food and Agriculture Organization (2010) Guidelines for Measuring Household and Individual Dietary Diversity. Rome, Italy.
- Aboagye RG, Seidu AA, Ahinkorah BO, et al. (2021) Dietary diversity and undernutrition in children aged 6–23 months in Sub-Saharan Africa. *Nutrients* **13**, 3431.
- Molani Gol R, Kheirouri S & Alizadeh M (2022) Association of dietary diversity with growth outcomes in infants and children aged under 5 years: a systematic review. *J Nutr Educ Behav* **54**, 65–83.
- Gebremariam T, Amare D, Ayalew T, et al. (2022) Determinants of severe acute malnutrition among children aged 6–23 months in Bahir Dar city public hospitals, Northwest Ethiopia, 2020: a case control study. *BMC Pediatr* **22**, 296.
- Kamenju P, Liu E, Hertzmark E, et al. (2017) Nutritional status and complementary feeding among HIV-exposed infants: a prospective cohort study. *Matern Child Nutr* **13**, e12358.
- Lane CE, Widen EM, Collins SM, et al. (2020) HIV-exposed, uninfected infants in Uganda experience poorer growth and body composition trajectories than HIV-unexposed infants. *J Acquir Immune Defic Syndr* **85**, 138–147.
- Widen EM, Tsai I, Collins SM, et al. (2019) HIV infection and increased food insecurity are associated with adverse body composition changes among pregnant and lactating Kenyan women. *Eur J Clin Nutr* **73**, 474–482.
- Andrade EDO, Rebouças AS, Filho JQ, et al. (2022) Evolution of infant feeding practices in children from 9 to 24 months, considering complementary feeding indicators and food processing: results from the Brazilian cohort of the MAL-ED study. *Matern Child Nutr* **18**, e13413.

33. Owais A, Schwartz B, Kleinbaum DG, *et al.* (2016) Minimum acceptable diet at 9 months but not exclusive breastfeeding at 3 months or timely complementary feeding initiation is predictive of infant growth in rural Bangladesh. *PLoS One* **11**, e0165128.
34. Vidmar SI, Cole TJ & Pan H (2013) Standardizing anthropometric measures in children and adolescents with functions for egen: update. *Stata J* **13**, 366–378.
35. WHO Multicentre Growth Reference Study Group (2006) WHO Child Growth Standards based on length/height, weight and age. *Acta Paediatrica* **450**, Supplement 76–85.
36. World Health Organization and the United Nations Children's Fund (2021) Indicators for Assessing Infant and Young Child Feeding Practices: Definitions and Measurement Methods. Geneva: World Health Organization and the United Nations Children's Fund (UNICEF).
37. Moursi MM, Arimond M, Dewey KG, *et al.* (2008) Dietary diversity is a good predictor of the micronutrient density of the diet of 6- to 23-month-old children in Madagascar. *J Nutr* **138**, 2448–2453.
38. Natamba BK, Kilama H, Arbach A, *et al.* (2015) Reliability and validity of an individually focused food insecurity access scale for assessing inadequate access to food among pregnant Ugandan women of mixed HIV status. *Public Health Nutr* **18**, 2895–2905.
39. Boateng GO, Collins SM, Mbullo P, *et al.* (2018) A novel household water insecurity scale: procedures and psychometric analysis among postpartum women in western Kenya. *PLoS One* **13**, e0198591.
40. Duong C, Young MF, Nguyen PH, *et al.* (2023) Temporal dietary diversity patterns are associated with linear growth but not ponderal growth in young children in rural Vietnam. *J Nutr* **153**, 3083–3091.
41. Chalashika P, Essex C, Mellor D, *et al.* (2017) Birthweight, HIV exposure and infant feeding as predictors of malnutrition in Botswanan infants. *J Hum Nutr Diet* **30**, 779–790.
42. Park JJH, Harari O, Siden E, *et al.* (2019) Interventions to improve linear growth during complementary feeding period for children aged 6–24 months living in low- and middle-income countries: a systematic review and network meta-analysis. *Gates Open Res* **3**, 1660.
43. Motadi SA, Zuma MK, Freeland-Graves JH, *et al.* (2023) Dietary diversity and nutritional status of children attending early childhood development centres in Vhembe District, Limpopo province, South Africa. *J Nutr Sci* **12**, e92.
44. Pasqualino MM, Shaikh S, Hossain MI, *et al.* (2024) An egg intervention improves ponderal but not linear growth among infants 6–12 months of age in rural Bangladesh. *J Nutr* **154**, 2290–2299.
45. Ricci H, Faber M, Ricci C, *et al.* (2023) Effects of egg as an early complementary food on growth of 6- to 9-month-old infants: a randomised controlled trial. *Public Health Nutr* **27**, e1.
46. Nicastro TM, Mocello AR, Weke E, *et al.* (2025) Effect of a climate-smart intervention on agriculture and nutrition of people with HIV. *Aids* **39**, 1650–1655.
47. Maltby AE, Odhiambo BC, Nyaura M, *et al.* (2023) Feasibility, acceptability and lessons learned from an infant feeding intervention trial among women living with HIV in western Kenya. *BMC Public Health* **23**, 1930.
48. Abebe A, Simachew Y & Delbiso TD (2023) Effect of ready-to-use therapeutic foods on time to recovery among children with severe acute malnutrition in Ethiopia: a prospective cohort study. *BMC Pediatr* **23**, 340.
49. Tewoldie MT, Girma M & Seid H (2022) Determinants of time to full enteral feeding achievement among infants with birth weight 1000–2000 g admitted to the neonatal intensive care unit of public hospitals in Hawassa city, Sidama region Ethiopian, 2019: a retrospective cohort study. *PLoS One* **17**, e0271963.
50. Ejigu Y, Magnus JH, Sundby J, *et al.* (2020) Differences in growth of HIV-exposed uninfected infants in Ethiopia according to timing of in-utero antiretroviral therapy exposure. *Pediatr Infect Dis J* **39**, 730–736.