On Errors in Determinants

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§1. Introduction. Care is needed in dealing with determinants whose elements are subject to experimental error, particularly when a determinant itself is small compared with its first minors. For, as these examples show, a relatively tiny error in one element may be responsible for a large error in the determinant :

$$\begin{vmatrix} -73 & 78 & 24 \\ 92 & 66 & 25 \\ -80 & 37 & 10 \end{vmatrix} = 1; \quad \begin{vmatrix} -73 & 78 & 24 \\ 92 & 66 & 25 \\ -80 & 37 & 10 \cdot 01 \end{vmatrix} = -118 \cdot 94.$$

Consider a determinant Δ , of order *n*, with elements a_i $(i = 1, \ldots, n^2)$. Let the actual error in each element a_i be e_i , and let the resultant error in Δ be *E*. In §§ 2-4, I shall find expressions for (i) *E* in terms of e_i ; (ii) the maximum range of *E* in terms of the ranges of e_i ; (iii) the probability distribution P(E) of *E* in terms of the probability distributions $p_i(e_i)$ of e_i , assumed independent and (for simplicity) symmetrical. In §§ 5-7, I shall investigate corresponding results for the quotient of two determinants which are identical except for one row or column. Such quotients are of importance in practical work, occurring in the solution of a set of simultaneous linear algebraic equations. In §8, the method is applied to an arbitrary function whose arguments are subject to error.

To apply the formulae to numerical determinants, it is necessary for first approximations to calculate the complete set of first minors; the second, third, etc., minors are required for closer approximations. The calculation of these minors is very laborious in the case of a determinant of large order. It may be pointed out, however, that all the first minors can be found in the course of evaluating a determinant and its adjugate by extant methods, for example that¹ of T. Smith, *Phil. Mag.* (7), **3** (1927), 1007.

¹ There is also a method suitable for use with a calculating machine, due to Dr Aitken, and as yet unpublished.

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§2. Actual error in a determinant. Let $A_i, A_{ij}, A_{ijk}, \ldots$ denote the first, second, third, \ldots minors of Δ corresponding to the elements indicated by the suffixes. An error e_i in the term a_i will cause an error $e_i A_i$ in Δ ; errors e_i, e_j in the terms a_i, a_j will together cause an error $(e_i A_i + e_j A_j + e_i e_j A_{ij})$ in Δ ; and so on. Hence in general:

$$E = \sum e_i A_i + \sum e_i e_j A_{ij} + \sum e_i e_j e_k A_{ijk} + \ldots, \qquad (1)$$

the suffixes in each summation running from 1 to n^2 . (This result follows at once from Taylor's Theorem, $A_i, A_{ij}, A_{ijk}, \ldots$ being partial derivatives of Δ .) It is to be observed that none of the errors occur squared or to higher powers in (1).

§3. Range of error. If $|e_i| \leq \epsilon_i$, we deduce from (1):

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$$|E| \leqslant \Sigma \epsilon_i |A_i| + \Sigma \epsilon_i \epsilon_j |A_{ij}| + \Sigma \epsilon_i \epsilon_j \epsilon_k |A_{ijk}| + \dots, \qquad (2)$$

vertical bars denoting absolute values. If we put $\epsilon_i = \epsilon$, and neglect ϵ^2 , (2) becomes

$$|E| \leqslant \epsilon \Sigma |A_i|. \tag{3}$$

Thus $\Sigma |A_i|$ may be regarded as a measure of the sensitivity of a determinant to small errors whose squares may be neglected.

As an application of these formulae, consider the first determinant in §1, and suppose that the range of error in each element is $\pm \frac{1}{2}$. We find $\Sigma |A_i| = 33099$, $\Sigma |A_{ij}| = 2\Sigma |a_i| = 970$, and (for any third order determinant) $\Sigma |A_{ijk}| = 6$. Hence (3) gives ± 16550 as a first approximation to the range of error in Δ . More accurately, (4) gives:

$$|E| \leq \frac{1}{2} \cdot 33099 + \frac{1}{4} \cdot 970 + \frac{1}{8} \cdot 6 < 16793.$$

If we suppose that the determinant has occurred in some practical calculation, and that the elements as given are only rough first approximations (with $\epsilon = \frac{1}{2}$), then the important question arises: To what further degree of accuracy η should the elements be evaluated, in order that the error in Δ may not exceed a given limit, say for example 0.2?

We have in general, when the range of error in each element is $\pm \eta$,

 $|E| \leqslant \eta \Sigma |B_i| + \eta^2 \Sigma |B_{ij}| + \ldots,$

where B_i , B_{ij} , are the *correct* values of the minors to which A_i , A_{ij} , are approximations.

Now

$$egin{aligned} &A_i-B_i|= ext{error in minor of }a_i\ &\leqslant \epsilon \sum\limits_{j} |A_{ij}|+\epsilon^2 \sum\limits_{ik} |A_{ijk}|+\ldots, \end{aligned}$$

and hence

$$|B_i| \leqslant |A_i| + \epsilon \sum_j |A_{ij}| + \epsilon^2 \sum_{jk} |A_{ijk}| + \ldots$$

Summing for $i = 1, \ldots, n$,

$$\Sigma |B_i| \leqslant \Sigma |A_i| + \epsilon \Sigma |A_{ij}| + \epsilon^2 \Sigma |A_{ijk}| + \dots$$
. Similarly,

$$\Sigma \left| \left| B_{ij} \right| \leqslant \Sigma \left| \left| A_{ij} \right| + \epsilon \Sigma \left| \left| A_{ijk} \right| + \ldots
ight|$$

Hence

$$egin{aligned} E &| \leqslant \eta \left[\Sigma \left| A_i
ight| + \epsilon \left. \Sigma \left| A_{ij}
ight| + \epsilon^2 \left. \Sigma \left| A_{ijk}
ight| + \ldots
ight]
ight. \ &+ \eta^2 \left[\Sigma \left| A_{ij}
ight| + \epsilon \left. \Sigma \left| A_{ijk}
ight| + \ldots
ight]
ight. \end{aligned}$$

This formula gives the range of error in a determinant when the elements are correct to within $\pm \eta$, in terms of the minors in a first approximation in which the elements are correct to within $\pm \epsilon$.

Let us apply this to the determinant already considered, for which

$$\Sigma \left| \left| A_i
ight| + \epsilon \left. \Sigma \left| \left| A_{ij}
ight| + \epsilon^2 \left. \Sigma \left| \left| A_{ijk}
ight|
ight| = 33585^{\circ}5.$$

Putting $\eta = .000005$, we find |E| < 0.17. Thus, in order that the error in Δ should not exceed 0.2, it is sufficient to evaluate each element to 5 decimal places.

§4. Probability Distribution. Denote the probability distribution of e_i by $p_i(e_i)$, and the r^{th} moment of this function by

$$m_{ri} \equiv \int_{-\infty}^{\infty} e_i^r p_i(e_i) de_i. \qquad (4)$$

The functions p_i being given, we can assume that these moments m_{ri} are known; in terms of them we can calculate the r^{th} moment M_r of P(E). We have in fact:

$$M_r \equiv \int_{-\infty}^{\infty} E^r P(E) dE = \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} E^r \prod_{i=1}^{n^z} [p_i(e_i) de_i].$$
 (5)

On the right hand side of (5), E^r is to be interpreted as a function of e_i , by means of (1). Consequently, for any value of r, we can expand E^r and integrate term by term, making use of (4), and of the conditions $\int_{-\infty}^{\infty} p_i de_i = 1$.

For example,

 $E^{2} = \sum A_{i}^{2} e_{i}^{2} + 2 \sum A_{i} A_{j} e_{i} e_{j} + 2 \sum A_{i} A_{ij} e_{i}^{2} e_{j} + 2 \sum A_{i} A_{jk} e_{i} e_{j} e_{k} + \dots$

On integration this gives

 $\Sigma A_i^2 m_{2i} + 2\Sigma A_i A_j m_{1i} m_{1j} + 2\Sigma A_i A_{ij} m_{2i} m_{1j} + 2\Sigma A_i A_{jk} m_{1i} m_{1j} m_{1k} + \dots$ as the value of M_2 .

The given functions have been assumed to be symmetrical; *i.e.* $p_i(e_i) = p_i(-e_i)$; consequently the odd moments m_{1i} , m_{3i} , all vanish. We need, then, only take account of those terms of E^r which contain only *even* powers of the errors e_i ; all other terms vanish on integration. We then find

$$\begin{split} & M_1 = 0, \qquad (6) \\ & M_2 = \sum A_i^2 \, m_{2i} + \sum A_{ij}^2 \, m_{2i} \, m_{2j} + \sum A_{ijk}^2 \, m_{2i} \, m_{2j} \, m_{2k} + \dots, \qquad (7) \\ & M_3 = 6 \sum A_i \, A_j \, A_{ij} \, m_{2i} \, m_{2j} + 6 \sum A_{ij} \, A_{jk} \, A_{ik} \, m_{2i} \, m_{2j} \, m_{2k} + 6 \sum A_i A_{jk} \, A_{ijk} \, m_{2i} \, m_{2j} \, m_{2k} + . \\ & M_4 = \sum A_i^4 \, m_{4i} + 6 \sum A_i^2 \, A_j^2 \, m_{2i} \, m_{2j} + \dots, \end{split}$$

and so on.

Each of these expressions has only a finite number of terms, but they rapidly become very complicated. However, if the original errors are known to be sufficiently small, *i.e.* if p_i are practically zero except where e_i are small, then the moments m_{ri} are also small quantities; we can therefore approximate to M_r successfully with only a few terms, and the succeeding moments M_r will rapidly diminish in order of magnitude.

With a change of notation, equation (7) can be rewritten:

$$S^{2} = \sum A_{i}^{2} \sigma_{i}^{2} + \sum A_{ij}^{2} \sigma_{i}^{2} \sigma_{j}^{2} + \sum A_{ijk}^{2} \sigma_{i}^{2} \sigma_{j}^{2} \sigma_{k}^{2} + \dots$$
(8)

giving the standard deviation $S (= \sqrt{M_2})$ of the determinant in terms of the standard deviations $\sigma_i (= \sqrt{m_{2i}})$ of the elements.

Having thus found M_1, M_2, M_3, \ldots , we can approximate to P(E) in one of the usual ways. For instance, if we assume that the errors e_i follow the normal law of distribution,

$$p_i\left(e_i
ight) = rac{1}{\sigma_i \sqrt{2\pi}} \exp\left(- e_i^2/2\sigma_i^2
ight)$$

(or, more generally, that they follow a symmetrical law of Charlier's Type A), then the law for the determinant must necessarily be of Charlier's Type A. We can thus assume

$$P(E) = \left[1 + A \frac{d^3}{dE^3} + B \frac{d^4}{dE^4} + \dots \right] \frac{1}{S\sqrt{2\pi}} e^{-E^2/2S^2}.$$

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S is given by (8), and the other constants by:

$$A = -rac{M_3}{3!}, \ B = rac{M_4 - 3S^4}{4!}, \ etc.,$$

these results being found by comparing the moments of the function with M_2 , M_3 ,

To a first approximation,

$$S^2 = \sum A_i^2 \sigma_i^2; A, B, \ldots = 0.$$

To a second approximation,

$$S^{2} = \sum A_{i}^{2} \sigma_{i}^{2} + \sum A_{i}^{2} A_{j}^{2} \sigma_{i}^{2} \sigma_{j}^{2};$$

$$A = -\sum A_{i} A_{j} A_{ij} \sigma_{i}^{2} \sigma_{j}^{2};$$

$$4! B = \sum A_{i}^{4} m_{4i} + 6 \sum A_{i}^{2} A_{j}^{2} \sigma_{i}^{2} \sigma_{j}^{2} - 3 (\sum A_{i}^{4} \sigma_{i}^{4} + 2 \sum A_{i}^{2} A_{j}^{2} \sigma_{i}^{2} \sigma_{j}^{2})$$

$$= 0,$$

since for normal distributions $m_{4i} = 3\sigma_i^4$.

For practical purposes a more important case arises when the elements have been calculated correct to x places of decimals, the last figure being forced, *i.e.* the last digit retained is increased by 1 when the first digit not retained is 5, 6, 7, 8 or 9. We may take it that all errors between $\pm \frac{1}{2} 10^{-x}$ are equally likely, and that none exceed these limits. Remembering that $\int_{-\infty}^{\infty} p(e) de$ must be unity, we have

$$p(e) = 10^{x}, \quad \text{if } -\frac{1}{2} 10^{-x} \leqslant e < \frac{1}{2} 10^{-x}, \\ = 0, \qquad \text{otherwise}; \end{cases}$$

the same law of distribution applies to all the elements.

Then we have

$$m_{1} = m_{3} = \dots = 0;$$

$$m_{2} = \int_{-\frac{1}{2}10^{-x}}^{\frac{1}{2}10^{-x}} e^{2} \cdot 10^{x} de = \frac{1}{12 \cdot 10^{2x}};$$

$$m_{4} = \int_{-\frac{1}{2}10^{-x}}^{\frac{1}{2}10^{-x}} e^{4} \cdot 10^{x} de = \frac{1}{80 \cdot 10^{4x}};$$

$$M_{1} = 0;$$

$$M_{2} = \frac{\sum A_{i}^{2}}{12 \cdot 10^{2x}} + \frac{\sum A_{ij}^{2}}{144 \cdot 10^{4x}} + \dots;$$

$$M_{3} = \frac{\sum A_{i} A_{j} A_{ij}}{24 \cdot 10^{4x}} + \frac{\sum A_{i}^{2} A_{j}^{2}}{24 \cdot 10^{4x}} + \dots;$$

Hence, neglecting terms involving 10^{-6x} , we find:

$$P(E) = \left[1 - \frac{\sum A_i A_j A_{ij}}{144 \cdot 10^{4x}} \frac{d^3}{dE^3} - \frac{\sum A_i^4}{288 \cdot 10^{5x}} \frac{d^4}{dE^4}\right] \frac{1}{S\sqrt{2\pi}} e^{-B^2/2\delta^2}, \quad (9)$$

where

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$$S^2 = rac{\sum A_i^2}{12 \cdot 10^{2x}} + rac{\sum A_{ij}^2}{144 \cdot 10^{4x}}.$$
 (10)

As an application of these formulae, consider the determinant:

$\sqrt{7}$	$\sqrt{2}$	$\sqrt{2}$	
$\sqrt{2}$	$\sqrt{11}$	$\sqrt{3}$,
$\sqrt{3}$	$\sqrt{5}$	$\sqrt{6}$]

the value of which, to five places, is 6.93899. Evaluating each element to two places, we get:

2.65	1.41	1.41	Ì
1.41	3.32	1.73	,
1.73	2.24	2.42	

the value of which, to five places, is 6.98983. For the latter determinant we find (correct to the number of places given in each case):

$$\begin{split} \Sigma |A_i| &= 26 \cdot 8, \\ \Sigma |A_i^2| &= 111 \cdot 9, \\ \Sigma |A_{ij}^2| &= 78, \\ \Sigma |A_i A_j A_{ij} = 287, \\ \Sigma |A_i^4| &= 1078, \\ \epsilon &= \cdot 005. \end{split}$$

By (3), the maximum range of error is \pm 135. By (10), the standard deviation (to a first approximation) is 031. These results may be compared with the actual error 051. The second approximation to the probability distribution is given by:

$$S = .0314, A = -.0002, B = -.0004.$$

§5. Error in quotient of determinants. Consider the quotient $\frac{\Delta_1}{\Delta_2}$ of the determinants¹:

 $\Delta_1 = | a_i a_{i1} a_{i2} \dots a_{i_{n-1}} |, \quad (i = 1 \dots n)$ $\Delta_2 = | b_i a_{i1} a_{i2} \dots a_{i_{n-1}} |.$

¹ The notation means that the determinants have these as i^{th} rows. The determinants are identical except in their first columns. As pointed out in §1, such quotients occur in the solution of a set of n simultaneous linear equations.

Suppose that the elements a_i , b_i , a_{ij} are subject to errors e_i , f_i , e_{ij} ; and that the resultant errors in Δ_1 , Δ_2 , $\frac{\Delta_1}{\Delta_2}$ are E_1 , E_2 , E. Form the determinant

where
$$\Delta_3 = |c_i a_{i1} a_{i2} \dots a_{i|n-1}|,$$

 $c_1 = a_i - Xb_i,$
 $X = ext{the calculated value of } \frac{\Delta_1}{\Delta_2};$
so that $\Delta_3 = ext{the calculated value of } \Delta_1 - X\Delta_2 = 0.$ (11)

The correct value of $\Delta_1 - X \Delta_2$ is, however, not zero.

Let
$$g_i = \operatorname{error} \operatorname{in} c_i = e_i - Xf_i,$$

 $E_3 = \operatorname{error} \operatorname{in} \Delta_3 = E_1 - XE_2,$
 $A_i = \operatorname{minor} \operatorname{of} a_i \text{ in } \Delta_1$
 $= \operatorname{minor} \operatorname{of} b_i \text{ in } \Delta_2$
 $= \operatorname{minor} \operatorname{of} c_i \text{ in } \Delta_3,$
 $A_{ij} = \operatorname{minor} \operatorname{of} a_{ij} \text{ in } \Delta_1,$
 $B_{ij} = \operatorname{minor} \operatorname{of} a_{ij} \text{ in } \Delta_2,$
 $C_{ij} = \operatorname{minor} \operatorname{of} a_{ij} \text{ in } \Delta_3 = A_{ij} - XB_{ij}.$

It follows from (11) that the minors of any one row of Δ_3 are proportional to the minors of the first row, *i.e.*

$$C_{ij} = \frac{A_i C_{1j}}{A_1}.$$
 (12)

Using (1), we have as first approximations,

$$egin{aligned} &E_2 = \Sigma f_i \, A_i + \Sigma \, e_{ij} \, B_{ij}, \ &E_3 = \Sigma \, g_i \, A_i + \Sigma \, e_{ij} \, C_{ij}. \end{aligned}$$

Correct to the second order,

$$egin{aligned} E &= rac{\Delta_1 + E_1}{\Delta_2 + E_2} - rac{\Delta_1}{\Delta_2} \ &= rac{\Delta_2 E_1 - \Delta_1 E_2}{\Delta_2 (\Delta_2 + E_2)} \ &= rac{E_1 - X E_2}{\Delta_2} \left(1 + rac{E_2}{\Delta_2}
ight)^{-1} \ &= rac{E_3}{\Delta_2} \left(1 - rac{E_2}{\Delta_2}
ight). \end{aligned}$$

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Thus, to the first order,

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$$E = \frac{1}{\Delta_2} [\Sigma (e_i - X f_i) A_i + \Sigma e_{ij} C_{ij}].$$
⁽¹³⁾

To the second order,

$$E = \frac{E_3}{\Delta_2} - \frac{1}{\Delta_2^2} \left[\sum f_i A_i + \sum e_{ij} B_{ij} \right] \left[\sum \left(e_i - X f_i \right) A_i + \sum e_{ij} C_{ij} \right], \quad (14)$$

$$E^{2} = \frac{1}{\Delta_{2}^{2}} [\Sigma (e_{i} - Xf_{i}) A_{i} + \Sigma e_{ij} C_{ij}]^{2}.$$
(15)

§6. Range of error in quotient of determinants. Suppose

 $|e_i|\leqslant\epsilon_i, |f_i|\leqslant\zeta_i, |e_{ij}|\leqslant\epsilon_{ij}.$

From (13), the range of E is given to a first approximation by:

$$ig| E ig| \leqslant rac{1}{ig| \Delta_2 ig|} \, [\Sigma \left(\epsilon_i + ig| X ig| \zeta_i
ight) ig| A_i ig| + \Sigma \, \epsilon_{ij} \, |C_{ij}| \,].$$

If the given ranges are all equal, *i.e.*, $\epsilon_i = \zeta_i = \epsilon_{ij} = \epsilon$, we have, using (12)

$$E \mid \leq \frac{\epsilon}{|\Delta_2|} \left[(1+|X|) \Sigma |A_i| + \Sigma \left| \frac{A_i C_{1j}}{A_1} \right| \right]$$
$$\leq \frac{\epsilon}{|\Delta_2 A_1|} \Sigma A_i \cdot \left[(1+|X|) |A_1| + \Sigma |C_{1j}| \right].$$
(16)

This can be expressed in the form

$$|\,E\,|\,\leqslant\epsilon\,\,{1+|\,X\over |\,\Delta_2\,A_1\,|}|\,S_1\,S_2$$
 ,

where $S_1 = \text{sum of absolute values of first minors of first column of } D$, $S_2 = \text{sum of absolute values of first minors of first row of } D$, D being the determinant $\left| \frac{a_i - Xb_i}{1 + |X|} a_{i1} a_{i2} \dots a_{in-1} \right|$.

§7. Probability distribution for quotient of determinants. Let the standard deviations of e_i , f_i , e_{ij} be σ_i , τ_i , σ_{ij} , it being assumed that the first moments are zero. Let $M_1, M_2, M_3 \ldots$ be the moments of the probability distribution P(E) of E. Using (14) and (15), proceeding as in §4, and remembering that in accordance with (6) the

first moment of E_3 is zero, we get, correct to the second order in the deviations:

$$egin{aligned} &M_1 = rac{1}{\Delta_2^2} [X \ \Sigma \ au_i^2 A_i^2 - \Sigma \ \sigma_{ij}^2 \ B_{ij} \ C_{ij}], \ &M_2 = rac{1}{\Delta_2^2} [\Sigma \ (\sigma_i^2 + \ X^2 \ au_i^2) \ A_i^2 + \Sigma \ \sigma_{ij}^2 \ C_{ij}^2], \ &M_3 = \ \ldots \ = 0. \end{aligned}$$

Thus, assuming that the required function is of Charlier's Type A, we have as a first approximation:

$$P(E) = \frac{1}{S\sqrt{2\pi}}e^{-(E-a)^2/2S^2},$$

where

$$egin{aligned} a &= M_1 = \left[X \ \Sigma \ au_i^2 A_i^2 - \Sigma \ \sigma_{ij}^2 \ B_{ij} \ C_{ij}
ight] / \Delta_2^2, \ S^2 &= M_2 - M_1^2 = \left[\Sigma \left(\sigma_i^2 + X^2 \ au_i^2
ight) A_i^2 + \Sigma \ \sigma_{ij}^2 \ C_{ij}^2
ight] / \Delta_2^2 \end{aligned}$$

The second approximation to P(E) involves many complicated summations.

If the given standard deviations are all equal, *i.e.* $\sigma_i = \tau_i = \sigma_{ij} = \sigma$, we have, using (12):

$$a = \frac{\sigma^2}{\Delta_2^2} [X \Sigma A_i^2 - \frac{1}{A_1} \Sigma B_{ij} A_i C_{1j}], \qquad (17)$$

$$S^{2} = \frac{\sigma^{2}}{\Delta_{2}^{2} A_{1}^{2}} \Sigma A_{i}^{2} . [(1 + X^{2}) A_{1}^{2} + \Sigma C_{1j}^{2}].$$
(18)

As in $\S6$, we note that (18) can be written

$$S^2 = \sigma^2 rac{1+X^2}{\Delta_2^2 \, A_1^2} \, S_3 \, S_4,$$

where $S_3 = \text{sum of squares of first minors of first column of } D'$,

 $S_4 =$ sum of squares of first minors of first row of D',

$$D' ext{ being the determinant } \left| rac{a_i - X b_i}{(1+X^2)^{rac{1}{2}}} \, a_{i1} \, a_{i2} \, \dots \, a_{i\,n-1}
ight|.$$

As an application of these formulae, consider the value of X found from the equations

$$\begin{split} X\sqrt{7} + Y\sqrt{2} &+ Z\sqrt{2} = \sqrt{17}, \\ X\sqrt{3} + Y\sqrt{11} + Z\sqrt{3} = -\sqrt{2}, \\ X\sqrt{2} + Y\sqrt{5} &+ Z\sqrt{6} = -\sqrt{3}. \end{split}$$

To four places, the value is X = 3.1468. Let us however approximate by evaluating each coefficient to two places. We then have:

$$\Delta_1 = egin{bmatrix} 4\cdot 12 & 1\cdot 41 & 1\cdot 41 \ -1\cdot 41 & 3\cdot 32 & 1\cdot 73 \ -1\cdot 73 & 2\cdot 24 & 2\cdot 45 \ \end{bmatrix} = 21\cdot 842244,$$
 $\Delta_2 = egin{bmatrix} 2\cdot 65 & 1\cdot 41 & 1\cdot 41 \ 1\cdot 41 & 3\cdot 32 & 1\cdot 73 \ 1\cdot 73 & 2\cdot 24 & 2\cdot 45 \ \end{bmatrix} = 6\cdot 989832.$

The following results are correct to the number of places given in each case:

$$\begin{split} X &= \frac{\Delta_1}{\Delta_2} = 3.1249, \\ \Delta_3 &= \begin{vmatrix} -4.16 & 1.41 & 1.41 \\ -5.82 & 3.32 & 1.73 \\ -7.14 & 2.24 & 2.45 \end{vmatrix} = .01, \\ A_1 &= 4.26, \\ \Sigma \mid A_i \mid = 6.8, \\ \Sigma \mid A_i^2 &= 23.25, \\ \Sigma \mid B_{ij} \mid A_i \mid C_{1j} = -264, \\ (1 + \mid X \mid) \mid A_1 \mid + \Sigma \mid C_{1j} \mid = 30, \\ X \sum A_i^2 - \frac{1}{A_1} \sum B_{ij} \mid A_i \mid C_{1j} = 134, \\ (1 + X^2) \mid A_1^2 + \sum C_{1j}^2 = 313, \\ \epsilon &= .005, \\ \sigma^2 &= \frac{1}{12 \cdot 10^4}. \end{split}$$

From (16), (17) and (18) we deduce that the range of error is $\pm .035$, and that a = .00002, S = .008. The actual error is .022.

§8. Error in an arbitrary function. The method of §4 may also be used to determine the probability distribution of error in an arbitrary function F (instead of the determinant Δ) of the quantities a_i . Equation (1) is then formally the same, but, since $F_{ii} = \partial^2 F / \partial a_i^2 \pm 0$, it includes terms with higher powers of the errors e_i . The following are the corresponding results, suffixes of F indicating partial derivatives.

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If the given laws of distribution are symmetrical,

$$egin{aligned} &M_1 = rac{1}{2} \sum F_{ii} \, m_{2i} + rac{1}{24} \sum F_{iiii} \, m_{4i} + rac{1}{4} \sum F_{iijj} \, m_{2i} \, m_{2j} + \, \ldots \, ; \ &M_2 = \sum F_i^2 m_{2i} + rac{1}{2} \sum F_{ii}^2 \, m_{4i} + \sum \left(F_{ii} \, F_{jj} + F_{ij}^2
ight) \, m_{2i} \, m_{2j} + \, \ldots \, ; \ &M_3 = rac{3}{2} \sum F_i^2 F_{ii} \, m_{4i} + \left(rac{3}{2} \sum F_i^2 F_{jj} + 6 \sum F_i \, F_j \, F_{ij}
ight) \, m_{2i} \, m_{2j} + \, \ldots \, ; \ &M_4 = \sum F_i^4 \, m_{4i} + 6 \sum F_i^2 F_j^2 m_{2i} \, m_{2j} + \, \ldots \, \, (i \pm j). \end{aligned}$$

If the given laws of distribution are *normal*, the required function will be of the form

$$P(E) = \left[1 + A \frac{d^3}{dE^3} + B \frac{d^4}{dE^4} + \dots \right] \frac{1}{S\sqrt{2\pi}} e^{-(B-a)^2/2S^2}.$$

The constants are given by

$$a = M_1, S^2 = M_2 - M_1^2, A = -\frac{M_3}{6} + \frac{M_1M_2}{2} - \frac{M_1^3}{3}, B = \frac{M_4}{24} + \frac{M_1^4}{12} - \frac{M_2^2}{8} + M_1A.$$

To a first approximation

$$a = \frac{1}{2} \sum F_{ii} \sigma_i^2; \quad S^2 = \sum F_i^2 \sigma_i^2; \quad A, B, \ldots = 0.$$

To a second approximation

$$a = \frac{1}{2} \sum F_{ii} \sigma_i^2 + \frac{1}{8} \sum F_{iii} \sigma_i^4 + \frac{1}{4} \sum F_{iij} \sigma_i^2 \sigma_j^2;$$

$$S^2 = \sum F_i^2 \sigma_i^2 + \frac{1}{2} \sum F_{ii}^2 \sigma_i^4 + \sum F_{ij}^2 \sigma_i^2 \sigma_j^2;$$

$$A = -\frac{1}{2} \sum F_{ii} F_i^2 \sigma_i^4 - \sum F_{ij} F_i F_j \sigma_i^2 \sigma_j^2;$$

$$B, \ldots = 0.$$

If the arguments of F have been calculated correct to x decimal places, the last figure being forced, then the second approximation is given by:

$$a = \frac{\sum F_{ii}}{24 \cdot 10^{2x}} + \frac{\sum F_{iiii}}{1920 \cdot 10^{4x}} + \frac{\sum F_{iiji}}{576 \cdot 10^{4x}};$$

$$S^{2} = \frac{\sum F_{i}^{2}}{12 \cdot 10^{2x}} + \frac{\sum F_{ii}^{2}}{720 \cdot 10^{4x}} + \frac{\sum F_{ii} F_{jj}}{288 \cdot 10^{4x}} + \frac{\sum F_{ij}^{2}}{144 \cdot 10^{4x}};$$

$$A = -\frac{\sum F_{ii} F_{i}^{2}}{720 \cdot 10^{4x}} - \frac{\sum F_{ij} F_{i} F_{j}}{144 \cdot 10^{4x}};$$

$$B = -\frac{\sum F_{i}^{4}}{2880 \cdot 10^{4x}}.$$

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