Review

Lung cancer incidence in patients with schizophrenia: meta-analysis

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Background

Lung cancer risk factors, like tobacco smoking, are highly prevalent in patients with schizophrenia. Whether these patients have a higher risk of lung cancer remains unknown.

Aims

We aimed to investigate whether patients with schizophrenia have a higher incidence of lung cancer compared with general population, in a meta-analysis.

Method

Eligible studies were searched from PubMed and EMBASE databases to identify cases of lung cancer in patients with schizophrenia and the general population. This meta-analysis utilised the random-effects model and prediction interval was used to calculate the heterogeneity of these eligible studies. We assessed the quality of evidence with the Grading of Recommendations Assessment, Development and Evaluation (GRADE) approach.

Results

There were 12 studies, totalling 496 265 patients, included in this meta-analysis. The data showed that the baseline schizophrenia diagnosis was not associated with any changes in lung cancer incidence in the overall population, with a standardised incidence ratio of 1.11 (95% Cl 0.90–1.37; P = 0.31), although there was a significant heterogeneity among these studies ($l^2 = 94\%$).

Moreover, there was also a substantial between-study variance with wide prediction interval values (0.47–2.64). The data were consistent for both males and females.

Conclusions

Up-to-date evidence from epidemiological studies indicates the lack of certainty about the association between schizophrenia diagnosis and lung cancer incidence.

Declaration of interest

None.

Keywords

Schizophrenia; lung cancer risk; cancer incidence; metaanalysis.

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Lung cancer is a significant health problem worldwide, accounting for 1,800,000 new lung cancer cases or 13% of newly diagnosed cancers globally in 2012.¹ Tobacco smoke is a single most significant risk factor in developing lung cancer,¹ although whether patients with schizophrenia have a higher risk of lung cancer remains unknown. Previous studies showed a higher prevalence of tobacco smoking and higher risk of lung cancer in patients with schizophrenia than in the general population.^{2–4} However, results of a previous meta-analysis did not support this notion, showing that the risk of lung cancer incidence was higher than the general population.⁵ Indeed, a previous genetic study reported that differences in p53 polymorphisms might be responsible for the genetic association between patients with schizophrenia and a potentially reduced risk of lung cancer,⁶ which further complicates the association between schizophrenia and lung cancer risk. However, results of previous cohort studies are inconclusive; that is, some studies have suggested that patients with schizophrenia have an increased risk of lung cancer,^{7,8} whereas other studies have not.⁹⁻¹⁸ Additionally, many updated analyses have been published since the previous meta-analysis, highlighting the importance of an updated meta-analysis. For example, previous meta-analyses were performed on the basis of a heterogeneity evaluation, using a conventional statistic, namely I^2 . The use of I^2 as the sole heterogeneity index has been questioned, whereas the prediction interval has been encouraged to analyse the heterogeneity index in a random-effects meta-analysis.¹⁹ Thus, in this study, we performed a meta-analysis to stratify the potential influence of tobacco smoking on lung cancer risk for patients with schizophrenia.

Method

Literature search

In this study, we searched PubMed and EMBASE databases to identify eligible studies of patients with lung cancer in both the schizophrenia and general populations. The search terms used were 'schizophrenia', 'schizophrenic', 'psychosis', 'psychoses', 'psychotic*' or 'schizo*' in combination with 'lung', and 'cancer', 'tumor', 'neoplasm' or 'carcinoma'. Two of our investigators independently searched the databases and reviewed each of publications. Any discrepancies were discussed and solved by the third investigator together with the two investigators. Kappa was calculated to reflect the interrater reliability for study inclusion/exclusion in the revised manuscript. Studies included in this meta-analysis were restricted to human studies and those published in English. Moreover, we also performed a manual search for cited reference lists in the original and review articles. The last literature search was performed on 26 June 2018.

Inclusion and exclusion criteria

Studies were included if they fulfilled all of the following criteria: (a) prospective or retrospective cohort studies without restrictions to the sample size and follow-up duration, and published in English as a full-length article; (b) enrolled adult participants (\geq 18 years of age); (c) schizophrenia was defined as the exposure of interest at baseline; (d) general population without schizophrenia were defined as a control; (e) reported lung cancer incidence during follow-up and (f) reported lung cancer incidence with adjusted standardised incidence ratios (SIRs), adjusted by age and gender at least, and the corresponding 95% confidence intervals of patients with schizophrenia relative to the general population.

Schizophrenia and lung cancer were diagnosed according to the criteria applied in all original articles. If a study included any overlapping participants, we only used the one with the larger sample sizes. Furthermore, our study only included the incidence study to avoid potential difference in outcome of lung cancer-related mortality between treatment of patients with schizophrenia and the general population.

Data extraction and quality evaluation

The following data were extracted from each study: (a) the first author name, publication years and patients' country of origin; (b) number of patients with schizophrenia at baseline and the source of the study populations; (c) follow-up durations, cases of lung cancer during follow-up and the methods applied for the identification of lung cancer cases and (d) other characteristics, such as whether tobacco smoking was adjusted when reporting the SIR outcome or whether lung cancer diagnosis was made before schizophrenia diagnosis (if this was yes, the study or participant were excluded). This meta-analysis utilised the Newcastle-Ottawa Scale for quality evaluation, according to a previous study.²⁰ This scale ranges from 1 to 9 stars, including criteria based on three aspects: selection of a study group, the comparability of the groups and the confirmation of the outcome. Moreover, we used the Grading of Recommendations Assessment, Development and Evaluation (GRADE) approach to assess the quality of the body of evidence.²¹ The GRADE methodology involves rating the initial quality of evidence for an association as high (with observational data), followed by downgrading based on five criteria (risk of bias, inconsistency, imprecision, indirectness and publication bias), and upgrading based on three criteria (large effect size, dose-response gradient, and plausible confounding). Abstracts, letters, reviews, noncohort studies and those reporting mortality data rather than incidence data were excluded from this meta-analysis.

Statistical analyses

This meta-analysis utilised the SIR as the overall measurement of the association between baseline schizophrenia and lung cancer risk for patients across all included studies. The most adequately adjusted data were then extracted and the logarithmically transformed SIRs and their corresponding s.e. were computed. Moreover, for heterogeneity evaluation, Cochrane's Q-test²² and the I^2 statistic²³ were performed and we defined the significant heterogeneity as $I^2 > 50\%$. The random-effects model was applied to combine the SIR data because this model is considered to have a more generalised result as it incorporates the potential heterogeneity.²⁴ In addition, we calculated the prediction interval based on the method provided by Borenstein et al, which was the range that the true SIR of a future study may fall in 95% of populations.¹⁹ We also performed sensitivity analyses by removing studies one at a time, to evaluate the stability of the results. Next, we utilised the predefined subgroup analyses to determine whether study characteristics that could significantly affect our data analyses and results, such as the exclusion of lung cancer incidence before schizophrenia diagnosis. We evaluated the potential publication bias by funnel plots supplemented by the Egger regression asymmetry test.²⁵ All statistical analyses were performed with RevMan version 5.1 (Cochrane Collaboration, Oxford, UK: https://community.cochrane.org/help/tools-and-software/revman-5) and Stata version 12.0 for Windows (Stata Corporation, College Station, TX).

Results

Identified studies for the meta-analysis

In this study, we searched PubMed and EMBASE databases to identify eligible studies, and the data are shown in Supplementary Figure 1 available at https://doi.org/10.1192/bjp.2019.23. In brief, we initially identified a total of 912 studies, and 881 were excluded after reviewing the titles and abstracts. We obtained 31 relevant articles for full-text review, and a further 19 articles were excluded for the following reasons: two of studies analysed patients with other mental diseases, two were duplicated reports, six were without the available outcome of lung cancer incidence, six were with mortality data only and two did not include the outcome of lung cancer incidence as the SIR. Thus, we were left with 12 studies for this metaanalysis.^{7–18} The Kappa for the interrater reliability of the process of study inclusion/exclusion was 0.89.

Study characteristics and quality evaluation

This meta-analysis included 12 studies^{7–18} totalling 496 265 patients with schizophrenia, and their baseline data are summarised in Table 1. All of these studies were retrospective database studies, including participants from the USA,8 Asia,10,13,14,16,18 Australia9 and Europe.^{7,11,12,15,17} Some of the studies included only mental health centre in-patients with schizophrenia,^{7,11,12,15} whereas other studies included patients with schizophrenia without specifying the setting.^{8-10,13,14,16-18} The inclusion and the follow-up durations spanned from 1962 to 2011. A total of 2151 lung cancer cases were diagnosed during the follow-up of patients with schizophrenia. Moreover, all studies reported the SIR of lung cancer incidence comparing patients with schizophrenia with the general population from the same country, during similar follow-up durations; however, there were only eight studies providing the SIR data separately in men and women.^{8,9,11,12,14–16,18} Only one study adjusted for tobacco smoking when presenting the SIR data,²⁶ and in seven studies, the investigators clearly stated that only lung cancer incidence after schizophrenia diagnosis was evaluated. $^{9,11,13-17}$ The Newcastle–Ottawa scores ranged from 6 to 8 for the included cohort studies and all of the included studies were database-based cohort studies. Patients with a confirmed diagnosis of schizophrenia were included, and the general population without schizophrenia were selected as controls. The follow-up durations of the included studies were between 10 and 29 years, which was adequate. However, given the large sample sizes and the registry-based nature of the included studies, poor sample controls regarding the diagnosis of schizophrenia, baseline characteristic evaluation and diagnosis of lung cancer were likely to exist. Moreover, confounding factors, such as smoking was adjusted in only one of the included studies. Following the GRADE methodology, we graded the quality of evidence for the outcome 'incidence of lung cancer' to very low because risk of bias of inconsistency, indirectness and imprecision may exist (Supplementary Table 1).

Schizophrenia and lung cancer risk in the overall populations

When combining the results from the 12 studies,^{7–18} our metaanalysis data showed that, compared with the general population from the same countries, patients with schizophrenia at baseline were not significantly associated with lung cancer risk (SIR = 1.11, 95% CI 0.90–1.37, P = 0.31; Fig. 1a), although there was a significant heterogeneity (P < 0.01 by Cochrane's Q-test, $I^2 = 94\%$) and a wide prediction interval (0.47–2.64) was detected. The data indicated an existence of substantial variance among these 12 studies. However, following the GRADE methodology, the quality of evidence was very low (Supplementary Table 1).

Table 1 Studies included in this meta-analysis												
Author, year	Country	Patient characteristics	Number of patients with schizophrenia	Comparison population	Follow-up (years)	Confirmation of lung cancer cases	Number of lung cancer cases	SIR outcomes reported	Smoking adjusted	Exclusion of cancer incidence before schizophrenia	Quality score	SIR (95% CI)
			•			0				•		
Lawrence, 2000 ⁹	Australia	Patients of mental health services	172 932	General western Australian population	1966–1995	ICD-9 classification ²⁷	856	M, F	Ν	Y	8	M: 1.10 (0.96–1.30); F: 1.10 (0.91–1.30)
Lichtermann, 2001 ⁷	Finland	In-patients or those with disability pension for schizophrenia	26 996	General Finnish population	1971–1996	Finnish Cancer Registry	106	Т	Ν	NS	7	2.17 (1.78–2.60)
Goldacre, 2005 ¹²	UK	In-patients with schizophrenia in Oxford	9649	General population	1963–1999	National Health Service data	88	Т	Ν	NS	6	1.18 (0.94–1.45)
Barak, 2005 ¹⁰	Israel	Patients at a mental health centre	3226	General population of Israel	1993–2003	National Cancer Registry of Israel	9	Т	Ν	NS	6	0.65 (0.29–1.23)
Dalton 2005 ¹¹	Denmark	In-patients with schizophrenia	22 766	General Danish population	1969–1995	Danish Cancer Registry	147	M, F	Ν	Y	8	M: 0.82 (0.65–1.02); F: 1.17 (0.90–1.49)
Grinshpoon, 2005 ¹³	Israel	Patients diagnosed with schizophrenia	26 518	General population of Israel	1962-2001	National Cancer Registry of Israel	126	M, F	Ν	Y	8	M: 1.38 (1.11–1.65); F: 0.85 (0.50–1.19)
Chou, 2011 ¹⁴	China	Patients diagnosed with schizophrenia in National Health Insurance Research Database	59 257	Age- and gender- matched individuals	2000–2008	ICD-9 classification	96	M, F, T	Ν	Y	8	0.55 (0.44–0.69)
McGinty, 2012 ⁸	USA	Patients diagnosed with schizophrenia	2315	General population from the national database	1994–2004	Data based on Surveillance Epidemiology and End Results Database	28	M, F, T	Ν	NS	6	4.70 (3.10–6.80)
Lin, 2013 ¹⁶	China	Patients diagnosed with schizophrenia	102 202	General population from the health insurance database	1995~2007	National Cancer Database	144	M, F, T	Ν	Y	8	0.81 (0.77–0.96)
Ji, 2013 ¹⁵	Sweden	In-patients with schizophrenia	59 233	General Swedish population	1965–2008	National Cancer Registry of Sweden	451	M, F, T	Ν	Y	8	0.75 (0.68–0.82)
Osborn, 2013 ¹⁷	UK	Patients diagnosed with schizophrenia	6845	Age, gender, GP and calendar time matched controls	1990–2008	Medical record– confirmed cases	67	Т	Y	Y	8	0.95 (0.65–1.41)
Raviv, 2014 ¹⁸	Israel	Male patients of a mental health centre	4326	General population of Israel	1990–2011	National Cancer Registry of Israel	33	Т, М	Ν	NS	6	1.43 (0.98–2.01)
SIR, standardised incidence ratio; M, male; F, female; N, no; Y, yes; T, total; NS, not stated; GP, general practitioner.												

(a) Study or subgroup	log(SIR)	s.e.	Weight	SIR IV, Random, 95% Cl	SIR IV, Random, 95% Cl
Lawrence 2000-M	0.0953	0.0773	7.3%	1.10 (0.95–1.28)	
Lawrence 2000-F	0.0953	0.091	7.2%	1.10 (0.92–1.31)	-
Lichtermann 2001	0.7747	0.0967	7.2%	2.17 (1.80-2.62)	-
Goldacre 2005	0.1655	0.1106	7.1%	1.18 (0.95–1.47)	
Barak 2005	-0.4308	0.3686	4.0%	0.65 (0.32–1.34)	
Dalton 2005-M	-0.1985	0.1149	7.0%	0.82 (0.65–1.03)	-
Dalton 2005-F	0.157	0.1286	6.9%	1.17 (0.91–1.51)	
Grinshpoon 2005-M	0.3221	0.1011	7.2%	1.38 (1.13–1.68)	-
Grinshpoon 2005-F	-0.1625	0.2212	5.8%	0.85 (0.55–1.31)	
Chou 2011	-0.5978	0.1148	7.0%	0.55 (0.44–0.69)	-
McGinty 2012	1.5476	0.2004	6.0%	4.70 (3.17–6.96)	
Lin 2013	-0.2107	0.0563	7.5%	0.81 (0.73–0.90)	-
Ji 2013	-0.2877	0.0478	7.5%	0.75 (0.68–0.82)	+
Osborn 2013	-0.0513	0.1975	6.1%	0.95 (0.65–1.40)	-
Raviv 2014	0.3577	0.1832	6.2%	1.43 (1.00–2.05)	
Total (95% Cl)			100.0%	1.11 (0.90–1.37)	•

Heterogeneity: Tau² = 0.15; X^2 = 225.12, d.f. = 14 (P < 0.00001); l² = 94% Test for overall effect: Z = 1.01 (P = 0.31)

0.05 0.2 20 5 Favours schizophrenic Favours control

(b) SIR SIR IV, Random, 95% Cl Study or subgroup log(SIR) IV, Random, 95% Cl s.e. Weight Lawrence 2000-M 0.0953 0.0773 11.3% 1.10 (0.95-1.28) Lawrence 2000-F 0.0953 0.091 10.9% 1.10 (0.92-1.31) Dalton 2005-M -0.1985 0.1149 10.1% 0.82 (0.65-1.03) Dalton 2005-F 0.157 0.1286 9.6% 1.17 (0.91-1.51) 0.3221 0.1011 Grinshpoon 2005-M 10.5% 1.38 (1.13-1.68) Grinshpoon 2005-F -0.1625 0.2212 0.85 (0.55-1.31) 6.5% Chou 2011 -0.5978 0.1148 10.1% 0.55 (0.44-0.69) Lin 2013 -0.2107 0.0563 11.8% 0.81 (0.73-0.90) Ji 2013 -0.2877 0.0478 12.0% 0.75 (0.68-0.82) Osborn 2013 -0.0513 0.1975 0.95 (0.65-1.40) 7.2% Total (95% Cl) 100.0% 0.92 (0.78-1.08) Heterogeneity: Tau² = 0.05; X^2 = 69.64, d.f. = 9 (P < 0.00001); l^2 = 87% 0.5 07 15 2 1 Test for overall effect: Z = 1.03 (P = 0.30) Favours schizophrenic Favours control

Fig. 1 Forest plots for the meta-analysis of the association between schizophrenia and lung cancer risk. (a) Forest plots of the overall population in this meta-analysis for association of schizophrenia with risk of developing lung cancer. (b) Forest plots of studies excluding patients with lung cancer diagnosed before the diagnosis of schizophrenia.

SIR, standardised incidence ratio: IV, inverse variance

Overall, our findings suggest that the up-to-date evidence from epidemiological studies indicates the lack of certainty about the association between schizophrenia diagnosis and lung cancer incidence. Future studies may show significantly affected lung cancer risk in patients with schizophrenia compared with the general population. Indeed, the results of our sensitivity analyses excluding one study at a time also did not significantly alter the pooled results (data not shown), suggesting good robustness of the study. Subsequent analyses of studies by excluding patients with lung cancer diagnosis before schizophrenia diagnosis^{9,11,13–17} produced similar results (SIR=0.92, 95% CI 0.78-1.08, P=0.30; Fig. 1b), with significant heterogeneity (P < 0.01 by Cochrane's Q-test, $I^2 =$ 87%; prediction interval: 0.53-1.60). The only study that adjusted tobacco smoke status when presenting the results did not support a significantly affected lung cancer incidence in patients with schizophrenia (SIR=0.95, 95% CI 0.65-1.40).

Schizophrenia and lung cancer risk stratified by gender

We then separately evaluated the risk of lung cancer in male and female patients with schizophrenia. By pooling the data from eight studies with SIR data on male patients with schizophrenia,^{8,9,11,13–16,18} our meta-analysis showed that a baseline diagnosis of schizophrenia did not significantly affect the risk of male lung cancer incidence (SIR=1.03, 95% CI 0.78-1.35, P=0.86; Fig. 2a), although there was substantial heterogeneity (P < 0.01 by Cochrane's Q-test, $I^2 = 93\%$; prediction interval: 0.39–2.73). We found that exclusion of participants with a lung cancer diagnosis before a schizophrenia diagnosis also had similar results (SIR=0.83, 95% CI 0.64–1.07, *P* = 0.15; Fig. 2b), although significant heterogeneity also occurred (P < 0.01 by Cochrane's Q-test, $I^2 =$ 92%; prediction interval: 0.34-2.06). Similarly, after pooling the results of seven studies that provided data on female patients with

(a)				SIR	SIR
Study or subgroup	log(SIR)	s.e.	Weight	IV, Random, 95% Cl	IV, Random, 95% Cl
Lawrence 2000	0.0953	0.0773	13.7%	1.10 (0.95–1.28)	-
Dalton 2005	-0.1985	0.1149	13.0%	0.82 (0.65–1.03)	
Grinshpoon 2005	0.3221	0.1011	13.3%	1.38 (1.13–1.68)	-
Chou 2011	-0.6162	0.1401	12.5%	0.54 (0.41–0.71)	
McGinty 2012	1.6094	0.311	8.4%	5.00 (2.72-9.20)	
Lin 2013	-0.3567	0.0744	13.7%	0.70 (0.61–0.81)	-
Ji 2013	-0.4005	0.0603	13.9%	0.67 (0.60–0.75)	•
Raviv 2014	0.3577	0.1832	11.5%	1.43 (1.00–2.05)	
Total (95% Cl)			100.0%	1.03 (0.78–1.35)	•
Test for overall effect: Z	2 = 1.18 (P = 0.)	86)		SIR	Favours schizophrenia Favours control
Study or subgroup	log(SIR)	s.e.	Weight	IV, Random, 95% Cl	IV, Random, 95% Cl
Lawrence 2000	0.0953	0.0773	17.3%	1.10 (0.95–1.28)	-
Dalton 2005	-0.1985	0.1149	16.1%	0.82 (0.65–1.03)	
Grinshpoon 2005	-0.3221	0.1011	16.5%	1.38 (1.13–1.68)	
Chou 2011	-0.6162	0.1401	15.1%	0.54 (0.41–0.71)	
Lin 2013	-0.3567	0.0744	17.3%	0.70 (0.61–0.81)	
Ji 2013	-0.4005	0.0603	17.7%	0.67 (0.60-0.75)	
	-0.4005	0.0000	17.17.70	0.07 (0.00 0.70)	
Total (95% Cl)	-0.4003	0.0000	100.0%	0.83 (0.64–1.07)	•

Fig. 2 Forest plots for the meta-analysis of the association between schizophrenia and lung cancer risk in males. (a) Forest plots of male patients in this meta-analysis for association of schizophrenia with risk of developing lung cancer. (b) Forest plots of studies excluding male patients with lung cancer diagnosed before the diagnosis of schizophrenia.

SIR, standardised incidence ratio; IV, inverse variance

schizophrenia,^{8,9,11,13–16} our meta-analysis showed that a baseline diagnosis of schizophrenia did not significantly affect the risk of female lung cancer incidence (SIR=1.10, 95% CI 0.85–1.41, P = 0.47; Fig. 3a), although substantial heterogeneity existed (P < 0.01 by Cochrane's Q-test, $I^2 = 87\%$; prediction interval: 0.48–2.54). After excluded participants with lung cancer diagnosis before schizophrenia diagnosis, the data were also similar (SIR=0.96, 95% CI 0.82–1.11, P = 0.55; Fig. 3b), although there was significant heterogeneity among these original studies (P = 0.02 by Cochrane's Q-test, $I^2 = 62\%$; prediction interval: 0.58–1.43).

Publication bias

We performed a funnel plot to assess publication bias and the data are shown in Supplementary Figure 2. We found that the funnel plot was visually symmetric and that the data on the Egger regression test indicated no significant publication bias occurring in these studies (P = 0.466).

Discussion

In the current study, based on the findings from previous databasederived, retrospective cohort studies, we failed to indicate a certain association between schizophrenia diagnosis and lung cancer incidence. Considerable heterogeneity exists regarding the included cohort, and the quality of the evidence of the meta-analysis was

708 https://doi.org/10.1192/bjp.2019.23 Published online by Cambridge University Press downgraded because of the potential risk of bias in inconsistency, indirectness and imprecision. Subsequent analyses demonstrated that our data were consistent in men and women, although there was significant heterogeneity underlying the pooled results of the overall population, as well as male and female participants.

Because many cancer-related risk factors, such as tobacco smoke, physical inactivity, obesity or metabolic disorders,^{2,4,5} have been shown to be prevalent in patients with schizophrenia, it has been speculated that patients with schizophrenia have an increased risk of developing many illnesses, including cancer. However, there are also studies that suggest that patients with schizophrenia may have a lower risk of cancer. The possible protective effect of schizophrenia on cancer incidence has been proposed after some early epidemiological studies showing that cancer incidence in the siblings and parents of patients with schizophrenia were lower than that in the general population.⁷ A previous meta-analysis showed that patients with schizophrenia had the same overall cancer incidence as the general populations, and that their the siblings and parents had even lower overall cancer incidence than that of the general populations.⁵ A previous meta-analysis showed that patients with schizophrenia had a slightly higher risk of lung cancer incidence (SIR = 1.31, 95% CI 1.01-1.71), but this previous study revealed that such results could be insignificant after adjusting for prevalence of tobacco smoke.⁵ In updating the previous meta-analysis, our current data showed that schizophrenia did not affect lung cancer incidence despite the fact that many lung cancer risk factors were prevalent in this population. These findings raise the possibility

(a) Study or subgroup	log(SIR)	s.e.	Weight	SIR IV, Random, 95% C	SIR Cl IV, Random, S	95% Cl
Lawrence 2000	0.0953	0.091	16.4%	1.10 (0.92–1.31)	-	
Dalton 2005	0.157	0.1286	15.2%	1.17 (0.91–1.51)		
Grinshpoon 2005	-0.1625	0.2212	11.7%	0.85 (0.55–1.31)		
Chou 2011	-0.5447	0.1831	13.2%	0.58 (0.41–0.83)		
McGinty 2012	1.5686	0.2706	10.0%	4.80 (2.82-8.16)		
Lin 2013	0.0198	0.0827	16.7%	1.02 (0.87–1.20)	+	
Ji 2013	-0.0943	0.0758	16.8%	0.91 (0.78–1.06)	-	
Total (95% Cl)			100.0%	1.10 (0.85–1.41)	•	
Heterogeneity: $Tau^2 = 0$. Test for overall effect: Z			° <0.00001);		0.05 0.2 1 Favours schizophrenic F	5 20 avours control
(1-)					·	
(b) Study or subgroup		50	Woight	SIR IV Bandom 95% (SIR SIR IV Bandom (05% CI

(d) Study or subgroup	log(SIR)	s.e. Weight	SIR IV, Random, 95%	SIR Cl IV, Random, 95% Cl
Lawrence 2000	0.0953 0	0.091 20.6%	1.10 (0.92–1.31)	
Dalton 2005	0.157 0.1	1286 15.9%	1.17 (0.91–1.51)	
Grinshpoon 2005	-0.1625 0.2	2212 8.4%	0.85 (0.55–1.31)	
Chou 2011	-0.5447 0.1	1831 10.8%	0.58 (0.41–0.83)	.
Lin 2013	0.0198 0.0	0827 21.7%	1.02 (0.87–1.20)	
Ji 2013	-0.0943 0.0	0758 22.6%	0.91 (0.78–1.06)	
Total (95% Cl)		100.0%	0.96 (0.82–1.11)	•
Heterogeneity: $Tau^2 = 0$ Test for overall effect: Z		0.5 0.7 1 1.5 2 Favours schizophrenic Favours control		

Fig. 3 Forest plots for the meta-analysis of the association between schizophrenia and lung cancer risk in females. (a) Forest plots of female patients in this meta-analysis for association of schizophrenia with risk of developing lung cancer. (b) Forest plots of studies excluding female patients with lung cancer diagnosed before the diagnosis of schizophrenia.

SIR, standardised incidence ratio; IV, inverse variance.

that schizophrenia may provide protective effects from cancer on a genetic level. Further, a previous review examined repurposing of psychiatric drugs as anti-cancer agents,²⁶ and some recent genetic studies also provided further clues for this phenomenon. For example, a previous study of Chinese patients with schizophrenia showed that genetic variants of XRCC4, which more often occurs in patients with schizophrenia, might have protective effects against colorectal cancer.²⁸ Subsequent studies proposed that microRNA 183 could also be a protective factor against solid tumours in patients with schizophrenia.²⁹ However, other studies reported shared genetic variants between schizophrenia and lung cancer, and patients with schizophrenia had a higher prevalence of tobacco use than general populations because of loss of nicotinic receptor expression,^{30,31} suggesting the complexity of the genetic influence of schizophrenia on cancer incidence. Thus, further study of cancer incidence in patients with mental illness may help us to stratify lung cancer risk factors.

Furthermore, we also calculated the prediction interval to evaluate the heterogeneity for the random effects in this meta-analysis, which should be acknowledged as a strength of our current study. Accordingly, our current data showed a considerably high between-study variance (prediction interval: 0.47–2.64), indicating that lung cancer incidence could be found to be significantly changed (either increased or decreased) in future studies comparing patients with schizophrenia with the general population. Indeed, in light of the variable quality of the included studies, the considerable heterogeneity of the meta-analysis and the wide prediction interval, the results of our study may reflect an uncertainty of the current understanding between schizophrenia and risk of lung cancer based on the findings of epidemiological studies. From this perspective, the results of our current study highlight a limitation of epidemiological studies on association of schizophrenia with lung cancer risk because of a considerable heterogeneity among these studies. With the development of methodologies in big-data research, including deep neural network analysis, the potential associations of schizophrenia with cancer risk may be better stratified according to biological functional or genetic levels. Future studies are warranted in these fields.

Limitations and future directions

Our study does have some limitations; for example, as mentioned in a previous meta-analysis,⁵ each original study might not have adjusted the lung cancer-related risk factors (like tobacco smoking) and in this case, analysis of association between schizophrenia and lung cancer risks could have confounded results. Moreover, these confounding factors remain, including body mass index, occupational information and dietary habits, and may affect association of schizophrenia with lung cancer risk. Thus, we recommend that future studies provide detailed information regarding the prevalence of associated risk factors in patients with schizophrenia and controls. In addition, our study showed a significant heterogeneity among these original studies, such as different diagnostic criteria used for schizophrenia and lung cancer, different schizophrenia treatment regimens and different methodologies to confirm lung cancer diagnosis, all of which need to be solved in future studies. Additionally, we only included studies from PubMed and EMBASE and did not include studies from other databases, those not in English or those published as conference abstracts. However, including studies from PubMed and EMBASE may have covered the majority of available literature. In addition, including conference abstracts that did not undergo peer review may potentially cause other bias. Moreover, given the large sample sizes and the registry-based nature of the included studies, poor sample controls regarding the diagnosis of schizophrenia, baseline characteristic evaluation and diagnosis of lung cancer were likely to be present in the individual cohort studies, which may weaken the validity of the comparison in this meta-analysis. Finally, our study only focused on the outcome of lung cancer incidence regardless of the histopathology of lung cancer. It is unknown whether schizophrenia contributed to incidence of small-cell or non-small-cell lung cancer incidence, which needs more precise study in future.

Overall, our current meta-analysis sheds lights on future directions in clarifying the potential association and underlying mechanisms between schizophrenia and lung cancer risk. First, studies are needed to determine whether the risk factors of lung cancer, such as tobacco smoking, obesity, substance misuse and diabetes, are potential confounding factors for the association between schizophrenia and lung cancer risk. Second, there are not many studies on the genetics and biology underlying the associations of schizophrenia with lung cancer risk. Thus, more studies are needed to elucidate the association between schizophrenia and lung cancer risk. Third, it also needs to be determined whether other characteristics in patients with schizophrenia, such as the use of antipsychotic medications, variability in patient management and disease course, or discrepancies in identifying lung cancer cases, further compound the heterogeneity of previous cohort studies. Finally, the influence of smoking on the general health status in patients with schizophrenia deserves further evaluation. Although the prevalence of tobacco smoking was significantly higher in patients with schizophrenia (approximately 80%) than that of the general population, results of our study indicated that the lung cancer risk was not higher in patients with schizophrenia.³¹ Also, smoking in schizophrenia has been hypothesized as 'self-treatment' because it may normalise sensory processing deficits in patients with schizophrenia.³⁰ Nonetheless, smoking cessation should still be encouraged in patients with schizophrenia, as no protective effect from schizophrenia on lung cancer risk has been definitively been proven, and smoking may expose these patients to risk for other smoking-related malignancies. Moreover, from a clinical perspective, smoking may lower the plasma level of antipsychotics such as clozapine, thereby hampering the treatment efficacy of these antipsychotics.³²

In conclusion, the up-to-date evidence from epidemiological studies indicates the lack of certainty about the association between schizophrenia diagnosis and lung cancer incidence. Because of the large heterogeneity of previously published cohort studies, the association between schizophrenia and lung cancer risk should be determined by further studies based on genetic and biological functional associations. These contradictory studies have prompted us to perform relevant research to investigate the mechanisms of low lung cancer incidence in smokers with schizophrenia from multiple perspectives.

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Supplementary material

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