Body cell mass index in children: interpretation of total body potassium results

Alexia J. Murphy* and Peter S. W. Davies

Children’s Nutrition Research Centre, Discipline of Paediatrics and Child Health, Royal Children’s Hospital, University of Queensland, Herston, Queensland 4029, Australia

(Received 20 August 2007 – Revised 12 November 2007 – Accepted 4 December 2007 – First published online 15 January 2008)

Body cell mass (BCM) is a valuable measure of functional nutritional status in children. As BCM is related to body size, it is essential that BCM is adjusted for stature when interpreting BCM data in children. Our aim was to examine the relationship between height and BCM in healthy children to determine the power by which height should be raised to adjust BCM for stature. This cross-sectional study calculated BCM by ⁴⁰K counting in 146 healthy children aged between 5 and 18 years. The relationship between BCM and height was explored using log–log regression. The present results demonstrate that the power by which height should be raised to adjust for BCM in females is 2.39 (SE 0.09) and for males is 2.92 (SE 0.10). A simplified sex-specific version of the index, BCM/height².⁵ for females and BCM/height² for males, was found to be statistically valid and numerically convenient, with the proportion of variation that could be attributed to height being less than 2 %. The present study shows that there is a difference in the relationship between height and BCM for males and females and that BCM can be adjusted in children using the BCM index of BCM/height².⁵ for females and BCM/height² for males.

Abbreviations: BCM, body cell mass; BCMI, body cell mass index; FFM, fat-free mass; FM, fat mass; HT, height; lnBCM, log of body cell mass; lnHT, log of height; TBK, total body K.

* Corresponding author: Dr Alexia J. Murphy, fax +61 7 3346 4684, email alexia.murphy@uq.edu.au

Material and methods

Subjects

The study sample was 146 healthy Caucasian children between the ages of 5 and 18 years who were controls in a...
study or recruited as part of a normative study. Subjects were recruited from schools, workplace newsletters and sporting groups. The study protocol was approved by the Royal Children’s Hospital Ethics Committee. Written consent was obtained from all parents and the children over age 12 years, while verbal assent was obtained for children under age 12 years.

**Measurements**

All measurements were taken in the Body Composition Laboratory at the Royal Children’s Hospital. Body weight was measured to the nearest 0·05 kg using calibrated digital scales (Tanita BWB-600; Wedderburn Scales, Brisbane, Qld, Australia) and HT was measured to the nearest 0·1 cm using a wall-mounted stadiometer (Holtain Instruments Ltd, Crymych, Dyfed, UK). BMI was calculated as weight divided by HT squared and BMI Z-scores were calculated using the LMS values published by the 2000 Centers for Disease Control and Prevention LMS values (10).

As 98 % of TBK is located in the BCM, BCM can be estimated from a known TBK. TBK analysis was performed using a shadow shield whole-body counter (Accuscan, Canberra Industries, Boston, MA, USA), which contains three sodium iodide crystal scintillation detectors arranged above a scanning bed. The crystals detect the 1·46 MeV γ rays emitted by the 40K in the body. A fixed proportion of the body’s K occurs as the natural isotope 40K; thus TBK can be determined from the 40K scan.

The measurement of a subject’s TBK required the subject to lie supine on a bed that is moved under the detectors. Two 1067 s scans were performed for each subject to check reproducibility. Background and sensitivity checks were considered in each measurement, with TBK reported in grams. BCM was calculated from TBK using the equation of Wang et al. (10).

\[
BCM (kg) = \frac{(TBK(g) \times 9·20)}{39·1}.
\]

**Statistical analysis**

The relationship between BCM and HT was investigated using ln–ln regression analysis. For each sex, the ln of BCM (lnBCM) was regressed on the ln of HT (lnHT), and the gradient (p) of the regression equation was determined. The index BCM/HT\(^p\) will then have zero correlation with HT, therefore representing a size-independent index of BCM.

A power will be more clinically useful if it is easy to apply, therefore we rounded p to the closest half decimal, ensuring that this value was within the standard error of p. To determine the implication of a correlation between BCM and HT\(^p\) using this adjustment, the proportion of variation in BCM that could be attributed to HT was calculated. The statistical package Minitab (version 8.21; Minitab, Inc., State College, PA, USA) was used for statistical analysis.

**Results**

The final study sample consisted of 146 healthy children (seventy-two females and seventy-four males) between the ages of 5·2 and 17·9 years. The subject characteristics are given in Table 1. The male and female subjects were similar in ages and representative of a healthy paediatric population with BMI Z-scores close to zero.

The statistics for the regression of lnBCM on lnHT are given in Table 2. The regression equation for females was lnBCM (kg) = 2·354 lnHT (m) + 1·881 and for males was lnBCM (kg) = 2·927 lnHT (m) + 1·724. The values of p were 2·35 (SE 0·09) and 2·93 (SE 0·10) for the female and male groups respectively. The closest half decimal to which these p values were rounded was 2·5 for females and 3·0 for males. If the BCM indices (BCM/HT\(^p\)) for females and BCM/HT\(^{3·0}\) for males were applied, then the proportion of variation in these indices that could be attributed to HT was 2 % for females and 1 % for males. The BCM was plotted against age for males and females in Fig. 1 to demonstrate that effect of stature on BCM was essentially eliminated using the BCMI. The regression equation was BCMI = 0·033age + 5·0386 for males and BCMI = −0·0176age + 6·4554 for females.

**Discussion**

The aim of the present study was to examine the relationship between BCM and HT in healthy children to determine the optimal adjustment of BCM for HT. The study shows that BCM can be adjusted in children using the BCMI of BCM/HT\(^{2·5}\) for females and BCM/HT\(^{3·0}\) for males. The regression equations for females and males both gave values for p close to 2·5 and 3, but not equal to this value. By using the numerically convenient and statistically valid HT adjustment of 2·5 for females and 3 for males, the amount of variation that is not accounted for by this adjustment is minimal, being less than 2 %. Thus there is no benefit to be gained from using the specific power values determined.

Our findings show that there is a difference between males and females in the relationship between BCM and HT. It is clear from the present study that the same power correction for HT can not be used in males and females and using the same BCMI would introduce significant variation not accounted for in HT of up to 40 %. Flynn et al. (11) examined the regression

**Table 1. Characteristics of the study population**

<table>
<thead>
<tr>
<th></th>
<th>Females (n 72)</th>
<th>Males (n 74)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Mean</td>
<td>SD</td>
<td>Mean</td>
</tr>
<tr>
<td>Age (years)</td>
<td>11.7</td>
<td>3.6</td>
</tr>
<tr>
<td>Height (cm)</td>
<td>146.4</td>
<td>19.8</td>
</tr>
<tr>
<td>Weight (kg)</td>
<td>40.90</td>
<td>15.45</td>
</tr>
<tr>
<td>BMI (kg/m(^2))</td>
<td>18.2</td>
<td>2.9</td>
</tr>
<tr>
<td>BMI Z-score</td>
<td>−0.05</td>
<td>0.83</td>
</tr>
<tr>
<td>Total body K (g)</td>
<td>71</td>
<td>23</td>
</tr>
<tr>
<td>Body cell mass (kg)</td>
<td>16.6</td>
<td>5.4</td>
</tr>
</tbody>
</table>

**Table 2. Regression of log of body cell mass on log of height**

<table>
<thead>
<tr>
<th></th>
<th>Females (n 72)</th>
<th>Males (n 74)</th>
</tr>
</thead>
<tbody>
<tr>
<td>p</td>
<td>2.35</td>
<td>2.93</td>
</tr>
<tr>
<td>Intercept</td>
<td>1.88</td>
<td>1.72</td>
</tr>
<tr>
<td>SE of gradient</td>
<td>0.09</td>
<td>0.10</td>
</tr>
<tr>
<td>95 % CI</td>
<td>2.18, 2.53</td>
<td>2.73, 3.12</td>
</tr>
</tbody>
</table>

*Gradient of the regression equation.*
of TBK on HT in 462 children. Their findings support the present result of a sex difference, showing that above 135 cm, about age 10 years, girls have less K per cm of HT than boys do. A sex difference is to be expected in the relationship between HT and BCM to account for the variation in composition of weight gain seen during childhood and adolescence.

No other studies, to the best of our knowledge, have examined the adjustment of BCM for HT in children. However, the relationship between BCM and size has been investigated recently in women of different weights. Wells et al. (8) found that the power adjustment of two was suitable for use in healthy and malnourished women. The HT adjustment of 2 has also been used in previous studies of adults (12,13).

From our findings it is evident that the power adjustment of 2 is not suitable for use in children. For both sexes the power of 2 falls beyond the 95 % CI, and if the power of 2 is applied, the proportion of variation not accounted for by HT increases to 21 % for females and 55 % for males.

Previous studies have used these sex-specific BCMI to represent nutritional status in children with cystic fibrosis and those undergoing bone marrow transplant (14,15). The use of these paediatric sex-specific BCMI will allow more accurate analysis of BCM data and expression of nutritional status between groups and individuals of children with different heights, especially those with clinical conditions and healthy controls. The clinical application of the present study’s findings will also ensure that BCM will be able to monitor the longitudinal response of children to nutritional support, eliminating the effect of growth in stature.

The present study shows for the first time the normalisation for body size of BCM in children. By using the BCMI for healthy children of BCM/HT2·5 for females and BCM/HT3 for males, the effect of stature on BCM can be essentially eliminated. It is vital to adjust body compartment data for stature in children and thus this finding has significant clinical and academic significance for the reporting, monitoring and comparison of BCM data in children.

Acknowledgements

The authors wish to thank the Royal Children’s Hospital Foundation for their financial support of the Children’s Nutrition Research Centre. The authors have no personal, commercial, political, academic or financial conflicts of interest in this research. A. J. M. was responsible for the subject testing, research design and writing of the manuscript. P. S. W. D. conducted the statistical analysis and significantly contributed to the research design, interpretation of results and review of the manuscript.

References