Meta-Analysis of Aldehyde Dehydrogenase 2 Gene Polymorphism and Alzheimer’s Disease in East Asians

Pan-Pan Hao, Yu-Guo Chen, Jia-Li Wang, Xing Li Wang, Yun Zhang

ABSTRACT: Background: The association of genetic polymorphism of mitochondrial aldehyde dehydrogenase 2 (ALDH2) and Alzheimer’s disease (AD) has been controversial and has been investigated only in several small-sample studies. In the present study, we performed a meta-analysis to evaluate the cross-sectional association of ALDH2 variants and AD risk in East Asian populations. Methods: Trials were retrieved through MEDLINE, EMBASE, J-STAGE and the China National Knowledge Internet databases (from January 1, 1994 to November 1, 2010) without any restriction on language. Data were abstracted by a standardized protocol. Results: We found four studies of 821 AD patients and 1380 healthy controls that qualified for the analysis. The variant ALDH2 genotype GA/AA was not associated with increased AD risk (odds ratio (OR) = 1.35; 95% confidence interval (CI) = 0.75–2.42; p = 0.32), even after stratification for the status of apolipoprotein E epsilon 4 allele. However, in the subgroup analyses, the association was significant for men (OR = 1.72; 95% CI = 1.10–2.67; p = 0.02). Conclusions: This study adds to the evidence that ALDH2 GA/AA genotype increases the risk of AD among East Asian men, although the effect size is moderate.

RÉSUMÉ: Meta-analyse portant sur la relation entre le polymorphisme du gène ALDH2 et la maladie d’Alzheimer chez des populations de l’Asie de l’Est. Contexte : L’association entre le polymorphisme du gène de l’aldéhyde déshydrogénase 2 (ALDH2) et la maladie d’Alzheimer (MA) est controversée et elle a été examinée seulement au cours d’études faites chez un petit nombre de sujets. Nous avons effectué une méta-analyse pour évaluer l’association transversale de variations dans le gène de l’ALDH2 et le risque de MA dans des populations de l’Asie de l’Est. Méthode : Nous avons identifié les études dans MEDLINE, EMBASE, J-STAGE et dans les bases de données de China National Knowledge Internet, du 1er janvier 1994 au 1er novembre 2010, sans restriction quant à la langue de publication. Les données ont été recueillies selon un protocole standardisé. Résultats : Nous avons identifié 4 études portant sur un total de 821 patients atteints de MA et 1380 témoins en bonne santé, qui rencontraient les critères établis pour l’analyse. La variante génotypique GA/AA du gène de l’ALDH2 n’était pas associée à un risque accru de MA (RC = 1.35; IC à 95% : 0.75 à 2.42; p = 0.32), même après ajustement pour la présence de l’allèle epsilon 4 du gène de l’apolipoprotéine E. Cependant, dans les analyses de sous-groupes, l’association était significative chez les hommes (RC = 1.72 ; IC à 95% : 1.10 à 2.67 ; p = 0.02). Conclusions : Cette étude apporte la notion que le génotype GA/AA du gène de l’ALDH2 augmente le risque de MA chez les hommes de l’Asie de l’Est, mais que cet effet est modéré.


Alzheimer’s disease (AD) is an irreversible, progressive neurodegenerative disease. It has been proved that apolipoprotein E epsilon 4 allele (APOE ε4) is associated with increased susceptibility to AD.1,2 However, the cause of AD has not been fully elucidated. The discovery of additional genetic factors is needed for clinical diagnoses and therapy of AD. Mitochondrial acetaldehyde dehydrogenase 2 (ALDH2), a polymorphic enzyme responsible for the oxidation of acetaldehyde to acetate, is encoded by the ALDH2 gene on chromosome 12. The ALDH2 gene is composed of 13 exons. Exon 12 contains a G-to-A missense mutation resulting in a substitution of lysine for glutamic acid at amino acid position 504 of ALDH2 (known as Glu504Lys). Hence, two ALDH2 alleles encode the active and inactive subunits (Glu504 and Lys504, also called ALDH2*1 and ALDH2*2, respectively) and three combinations, namely, *1/*1 (wide-type homozygote), *1/*2 (heterozygote) and *2/*2 (mutant homozygote). The ALDH2*2 allele is associated with reduced ALDH2 activity and thus, with an accumulation of acetaldehyde and 4-hydroxy-2-nonenal (HNE), which have been proposed to be potentially linked with AD.3,5 Therefore ALDH2*2 could have a role in the pathogenesis of AD. This hypothesis is important to genetic research into AD, especially in East Asian populations, because the ALDH2*2 allele is rare in Caucasians (lower than 5%) but is widely prevalent among East Asians (30%–50%).6,7

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Although results from two Western studies suggest that increased ALDH activity might be a protective response of cerebral cortex and putamen to AD by detoxifying acetaldehyde and HNE,5,8 no investigation of the association of ALDH2 gene polymorphism with AD in non-Asian populations, in which the ALDH2*2 allele is rare, has been reported until now. Two case-control studies9,10 found that ALDH2 gene polymorphism might be a risk factor for AD and be dependent on APOE ε4 status in East Asian populations. However, several other studies11-13 did not obtain the same results. The reason for the discrepancy is unclear but may be related to low statistical power or to differences in APOE ε4 status. Thus, we conducted a meta-analysis to investigate the possible association of ALDH2 genetic variation and AD in East Asian populations. Meta-analysis is a commonly used analytical tool to answer clinical questions that are not addressed by individual studies; its advantages include increased statistical power, large sample size, wide population coverage, and cost-effectiveness.

METHODS

Selection criteria

Studies were included in the meta-analysis if they (i) investigated the association of ALDH2 Glu504Lys polymorphism and AD risk in an East Asian population; (ii) were retrospective or nested case-control studies of a hospital- or population-based design; and (iii) contained sufficient information for genotype counts for estimating the odds ratio (OR) and its corresponding 95% confidence interval (CI). Studies not meeting these criteria, animal studies, and studies without sufficient data were excluded from the analysis. When two papers reported the same study, only one publication with more complete and recent data was included in the analysis.

Search source and strategy

Publications were identified via the search engines of MEDLINE, EMBASE, J-STAGE and the China National Knowledge Internet databases (from January 1, 1994 to November 1, 2010) without any restriction on language. The following keywords were used for search: (aldehyde dehydrogenase 2) OR (ALDH2)) AND ((Alzheimer’s disease), OR (AD), OR (dementia)). The reference lists of all retrieved articles were checked for other relevant publications, and studies without sufficient data were excluded from the analysis. The following information was extracted using a standardized protocol and reporting form: first author’s last name, year of publication, number of subjects in each category, counts of subjects with different genotypes among AD patients and controls, clinical characteristics of each study population, study design, and ethnicity of the population studied. Absolute numbers were recalculated when percentages were reported. Furthermore, information on Hardy–Weinberg equilibrium (HWE) was collected or calculated manually if missing.

The literature search and data extraction were undertaken independently and blindly by two authors (PPH and YGC) using a standardized approach. Any discrepancies were resolved by consensus.

Statistical analysis

RevMan 5.0.24 software, developed by the Cochrane Collaboration (http://www.cc-ims.net/revman, accessed on June 16, 2010), was used for the meta-analysis. The HWE was assessed by the chi-square test for studies if needed. In general, the fixed-effects model was used in the absence of between-study heterogeneity (the Cochran’s Q and I² statistics); otherwise the random-effects model was used.14 In this meta-analysis, we implemented the random-effects model only to coordinate the individual effect-size estimates because with a fixed-effects model, only sampling error contributes to the differences between the observed effect-size estimates across individual studies.15 In contrast, two sources of variance coexist in a random-effects model: the sample error and between-study heterogeneity. Given the ubiquitous nature of heterogeneity between studies, a random-effects model is appropriate.

In addition, sensitivity or subgroup analyses were conducted to seek more narrowly drawn subsets of the studies by removing an individual study each time or studies with similar features such as deviations from HWE to assess individual effects. Finally, we assessed publication bias using a funnel plot and the fail-safe number (Nfs) with the significance set at 0.05 for each meta-comparison. If the calculated Nfs value was smaller than the number of observed studies, the meta-analysis results might show risk of publication bias. We calculated the Nfs according to the formula \[ N_{fs} = \frac{(\Sigma Z)^2}{1.64^2} \] (k is the number of articles included in the meta-analysis).

RESULTS

Study Characteristics

The initial search with the electronic search strategy described above and manual search yielded 226 potential literature citations, of which 218 studies were excluded after scanning the titles and abstracts (Figure 1). The two authors searched the full articles for the remaining eight abstracts and excluded four articles (the full-text of one abstract was not available, two articles had insufficient data for analysis, and two articles used the same subjects11,12). Only four original articles8-11 satisfied our selection criteria and were included in the analysis, with 821 AD patients and 1380 healthy controls. The interobserver agreement for the study selection was excellent (k = 0.92). All these studies were full papers published in English from 2000 to 2010. According to the study design, the Kim et al study11 involved a population-based design, and the remaining three a hospital-based design. Of these four populations, two (50%) were from mainland China,9,13 one (25%) was from Japan,10 and one (25%) was from South Korea.11

The detailed characteristics of the four eligible studies are listed in Table 1. The percentages of ALDH2*2 carriers in AD patients ranged from 25% to 71.3% and that in controls from 28.7% to 46%. Except for the Kamino et al study10 (p = 0.04 for controls), the genotype distributions for both patients and controls were in HWE (p > 0.05).
Figure 1: Flow chart of the systematic search process.

Table 1: Detailed characteristics of eligible studies

<table>
<thead>
<tr>
<th>Publication</th>
<th>Kamino et al.\textsuperscript{10}</th>
<th>Kim et al.\textsuperscript{11}</th>
<th>Wang et al.\textsuperscript{9}</th>
<th>Zhou et al.\textsuperscript{13}</th>
</tr>
</thead>
<tbody>
<tr>
<td>Sample size, n</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Patients</td>
<td>447</td>
<td>80</td>
<td>188</td>
<td>106</td>
</tr>
<tr>
<td>Controls</td>
<td>447</td>
<td>610</td>
<td>223</td>
<td>100</td>
</tr>
<tr>
<td>ALDH2*2 carriers, n (%)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Patients</td>
<td>215 (48.1)</td>
<td>20 (25.0)</td>
<td>134 (71.3)</td>
<td>41 (38.7)</td>
</tr>
<tr>
<td>Controls</td>
<td>167 (37.4)</td>
<td>175 (28.7)</td>
<td>99 (44.4)</td>
<td>46 (46.0)</td>
</tr>
<tr>
<td>p values for Hardy-Weinberg equilibrium test</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Patients</td>
<td>0.62</td>
<td>0.24</td>
<td>0.81</td>
<td>0.09</td>
</tr>
<tr>
<td>Controls</td>
<td>0.04</td>
<td>0.88</td>
<td>0.72</td>
<td></td>
</tr>
<tr>
<td>Age, mean (SD) years</td>
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<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Patients</td>
<td>76.7 (6.9)</td>
<td>72.8 (5.7)</td>
<td>77.4 (6.4)</td>
<td>71.6 (6.6)</td>
</tr>
<tr>
<td>Controls</td>
<td>77.1 (7.2)</td>
<td>72.2 (5.3)</td>
<td>69.2 (7.7)</td>
<td></td>
</tr>
<tr>
<td>Sex, male/female</td>
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<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Patients</td>
<td>123/324</td>
<td>282/408</td>
<td>92/96</td>
<td>41/65</td>
</tr>
<tr>
<td>Controls</td>
<td>123/324</td>
<td>138/85</td>
<td></td>
<td>41/59</td>
</tr>
<tr>
<td>Study design</td>
<td>Hospital-based</td>
<td>Population-based</td>
<td>Hospital-based</td>
<td>Hospital-based</td>
</tr>
<tr>
<td>Ethnicity</td>
<td>Japanese</td>
<td>South Koreans</td>
<td>Chinese</td>
<td>Chinese</td>
</tr>
</tbody>
</table>
Main meta-analysis results

Under the random-effects model, we found no association of \textit{ALDH2}*2 allele and increased AD risk (OR = 1.35; 95% CI = 0.75–2.42; \(p = 0.32\)) (Figure 2). Nevertheless, of note, the initial studies for studies by Wang et al\(^9\) (OR = 3.11; 95% CI = 2.06–4.69) and Kamino et al\(^10\) (OR = 1.55; 95% CI = 1.19–2.03) showed a significantly increased risk of \textit{ALDH2}*2 allele present versus \textit{ALDH2}*2 allele absent for susceptibility to AD, whereas others consistently showed the opposite associations. Furthermore, the meta-analysis confirmed a significant association of presence of \textit{APOE} \(\varepsilon4\) and increased AD risk (OR = 2.98; 95% CI = 1.84–4.84; \(p < 0.00001\)) (Figure 3).

Sensitivity/subgroup analyses

In the study by Kamino et al\(^10\), \textit{ALDH2} genotype frequencies in the control group deviated from HWE (\(p = 0.04\)). The Kim et al study\(^11\) was population-based and HWE in either group was not assessed because of insufficient information for genotype counts. However, exclusion of each of the above two studies generated no materially changes for the pooled ORs (OR = 1.26, 95% CI = 0.47–3.33, \(p = 0.64\); OR = 1.56, 95% CI = 0.79–3.08, \(p = 0.20\), respectively).

Only two articles\(^9,13\) were available in the subgroup analysis by gender and three\(^9,10,13\) by \textit{APOE} \(\varepsilon4\) status. Results of subgroup analyses for the estimates are in Table 2. The between-study heterogeneity was explained in part by gender ratio. The subgroup analyses still revealed no significant association of \textit{ALDH2} genotype and AD risk after adjustment for \textit{APOE} \(\varepsilon4\) status (\(p = 0.07\) for carriers and \(p = 0.22\) for non-carriers); however, the association was significant for men after stratification by sex (OR = 1.72; 95% CI = 1.10–2.67; \(p = 0.02\)).

Publication bias

The funnel plot assessing the publication bias of all four studies is shown in Figure 4. The plot shows a degree of asymmetry possibly consistent with “small study” bias. We calculated the \(N_{\alpha=0.05}\) for both \textit{ALDH2}*2 and \textit{APOE} \(\varepsilon4\) being present. The \(N_{\alpha=0.05}\) values for both comparisons were greater than the number of studies included in the meta-analysis.
To our knowledge, this is the first meta-analysis examining the association of ALDH2 Glu504Lys polymorphism and AD in East Asians. In view of the inconsistencies of present results of studies regarding this association with AD, our meta-analysis proved to be a more powerful approach in estimating the true effect size than a single study. Our stratified analyses suggested that a gender difference might account for the inconsistency of findings across studies, and we found a significant association of presence of $\text{ALDH2}^{\ast 2}$ and AD risk in East Asian men. However, the association did not differ substantially by $\text{APOE} \varepsilon 4$ status so was unlikely to have been affected by $\text{APOE} \varepsilon 4$.

Previous studies\(^\text{16-18}\) have suggested that alcohol consumption might affect the development of AD by producing the neurotoxic substance acetaldehyde, which might alter neurons and transmitting molecules such as muscarinic cholinergic receptor and serotonin in the brain. Mitochondrial ALDH2 metabolizes acetaldehyde produced from ethanol into acetate. The ALDH2 gene is the strongest genetic factor influencing alcohol drinking behavior and relates to the risk of alcoholism, because carriers of the $\text{ALDH2}^{\ast 2}$ allele (rare in populations except Asians) have a low tolerance of alcohol.\(^\text{19-21}\) In addition, oxidative stress may underlie age-dependent memory loss and cognitive decline. A reactive intermediate

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**DISCUSSION**

Table 2: Subgroup analyses of association of ALDH2 genotype and Alzheimer’s disease

<table>
<thead>
<tr>
<th>Category</th>
<th>Studies, n</th>
<th>References</th>
<th>AD, n</th>
<th>OR (95% CI)</th>
<th>p(^\dagger)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Adjustment by gender</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Male</td>
<td>2</td>
<td>10,13</td>
<td>163</td>
<td>1.72 (1.10–2.67)</td>
<td>0.02</td>
</tr>
<tr>
<td>Female</td>
<td>2</td>
<td>10,13</td>
<td>390</td>
<td>0.91 (0.33–2.49)</td>
<td>0.86</td>
</tr>
<tr>
<td>Adjustment by APOE $\varepsilon 4$ status</td>
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<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>APOE $\varepsilon 4$ carriers</td>
<td>3</td>
<td>9,10,13</td>
<td>296</td>
<td>1.58 (0.96–2.61)</td>
<td>0.07</td>
</tr>
<tr>
<td>APOE $\varepsilon 4$ non-carriers</td>
<td>3</td>
<td>9,10,13</td>
<td>445</td>
<td>1.67 (0.74–3.80)</td>
<td>0.22</td>
</tr>
</tbody>
</table>

**Figure 4:** Funnel plot of publication bias for the association of ALDH2 polymorphism with risk of Alzheimer’s disease.

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AD: Alzheimer’s disease; OR: odds ratio; CI: confidence interval; $\text{APOE} \varepsilon 4$: apolipoprotein E epsilon 4 allele.

\(^\dagger\) The p-value of OR determined by the Z-test. Statistical significance was set at p-value 0.05.
generated by lipid peroxidation, HNE, is known to accumulate in the brain in neurodegenerative disease.\textsuperscript{22} Ohsawa et al\textsuperscript{13} found that ALDH2 detoxifies HNE by oxidizing its aldehyde group, and transgenic mice with low ALDH2 activity exhibited an age-dependent neurodegeneration accompanying memory loss. Furthermore, Bai et al\textsuperscript{24} found that overexpressed $ALDH2$ gene might moderate HNE-induced neuronal death by regulating caspase-3 protein and reactive oxygen species in cultured hippocampal neurons. Thus the $ALDH2$ polymorphism may represent a desirable candidate for genetic risk factors for cognitive impairment and dementia among older East Asians. However, several small-sample investigations of the association of the mutant $ALDH2$ genotype with AD in East Asians showed conflicting results.\textsuperscript{9-13}

In most Asian countries, cultural norms and social and attitudinal factors lead to much more frequent alcohol abuse in men than women,\textsuperscript{25,26} which might be the main reason that East Asian men carrying the $ALDH2^*2$ allele were more likely to have subsequent AD diagnosis in our stratified analyses. Thus abstinence or mild drinking might be helpful to reduce the risk of cognitive impairment, dementia, and AD in this population. Evidence of behavioural and physiological interactions suggests that tobacco and alcohol use may not only individually affect AD, but may also modify each other’s effects, with smoking reducing the risk of AD among drinkers.\textsuperscript{27}

In a recent simulation study,\textsuperscript{28} case-control status was determined by two interacting polymorphisms with heritabilities ranging from 0.025 to 0.4 with replication sample sizes ranging from 400 to 1600 individuals. This study showed that the power to replicate the statistically significant independent main effect of one polymorphism could drop dramatically with a change of allele frequency of less than 0.1 at a second interacting polymorphism, and differences in allele frequency could result in a reversal of allelic effects whereby a protective allele becomes a risk factor in replication studies, which seems to be a possible explanation for divergent results between the present initial study and subsequent replication. It is thus reasonable to speculate that if involved, the impact of $ALDH2$ polymorphism on AD might fail to replicate and should be checked for interactions with other polymorphisms, particularly when samples are collected from groups with distinct ethnic backgrounds or different geographic regions.

Although our sample size of more than 2000 subjects is not small, it may not be large enough to detect genes that contribute to AD through small effects. As well, $ALDH2$ Glu504Lys polymorphism might be linked to a causal variant or others within or near the $ALDH2$ gene to produce the final disease phenotype. Therefore, conclusions are premature until a large, well-performed study confirms or refutes our results. $ALDH2^*2$ is negatively associated with alcohol use. If people with a combination of alcohol use and dementia are under-represented in clinical settings or if they are less likely to be classified as having AD, then $ALDH2^*2$ frequencies may be exaggerated in clinical samples of AD patients as compared with controls. The $ALDH2^*2$ allele frequencies among the four Asian populations ranged from 15.7% to 35.3%. If $ALDH2^*2$ is truly associated with AD, the different frequencies of $ALDH2^*2$ among different populations might obscure the associations. Level of education can modulate the risk of dementia and may modify the effect of biological risk factors on incidence of AD.\textsuperscript{29,30} We were not sure of the educational levels of the four populations included in our analyses, but differences in educational attainment might underlie heterogeneous findings.

Meta-analysis, as a quantitative approach to combine results from similar studies, has earned a crucial position in providing useful information for evidence-based medicine and health care decision making. However, there are several possible limitations in a meta-analysis. Such as: (i) not all the relevant studies are included, which might limit the generalization of the results and lead to a biased result; and (ii) as no blinding measures are devised regarding the names of authors or journals, it is possible for the reviewers to be subjective.\textsuperscript{31} Although two authors (PPH and YGC) independently and blindly searched and selected as many relevant articles as possible via four databases, our meta-analysis contains some inevitable limitations. First, the cross-sectional nature of our included studies precludes comments on causality. Second, we could not retrieve information for various confounding factors considered effective modulators for the development of AD and should be considered in the analyses. This analysis did not consider alcohol-related histories and the educational levels of the subjects, which might mediate or modify the association. To our knowledge, only the study by Kim et al\textsuperscript{11} took into account alcohol consumption, and no significant association was found between $ALDH2$ genotype and AD in either drinkers (OR = 0.39; 95% CI = 0.05–3.07; p = 0.37) or nondrinkers (OR = 0.83; 95% CI = 0.47–1.46; p = 0.51). Finally, we focused only on $ALDH2$ Glu504Lys polymorphism and did not evaluate other single-nucleotide polymorphisms in $ALDH2$. The potential role of Glu504Lys polymorphism may be diluted or masked by other gene–gene or gene–environment interactions.

**Conclusions**

In summary, our study expands the previous findings of AD by showing no significant association of presence of $ALDH2^*2$ and increased risk of AD in East Asians, either before or after stratification for $APOE\varepsilon4$ status. However, among East Asian men, $ALDH2\textsuperscript{GA/AA}$ genotype increases AD risk, although the effect size is moderate. Additional cross-sectional or longitudinal studies with large sample sizes and in different populations examining gene–gene or gene–environment interactions, as well as studies seeking to provide biological or clinical validations of our results, are warranted.

**Acknowledgments**

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