

ON THE NONCENTRAL DISTRIBUTIONS OF THE SECOND LARGEST ROOTS OF THREE MATRICES IN MULTIVARIATE ANALYSIS⁽¹⁾

BY
SABRI AL-ANI

1. Introduction and Summary. The central distribution of the second largest (smallest) root following the Fisher-Girshick-Hsu-Roy distribution under certain null-hypothesis has been derived in series form by Pillai and Al-Ani [6]. In this paper the noncentral distributions of the second largest roots in the MANOVA situation, the canonical correlation, and equality of two covariance matrices are obtained. Further, the distribution of the second largest root of the covariance matrix is obtained as a limiting case. The largest root and its noncentral distributions have been considered already by Pillai and Sugiyama [7] for the situations stated above. However, in the present paper, the joint densities of the largest and the second largest roots are derived in all the above cases from which the distributions of the largest roots can be obtained, although in more elaborate forms.

2. Noncentral distribution of the second largest root in the MANOVA case. Let \mathbf{X} be a $p \times n_1$ matrix variate ($p \leq n_1$) and \mathbf{Y} a $p \times n_2$ matrix variate ($p \leq n_2$) and the columns be all independently normally distributed with covariance matrix Σ , $E\mathbf{X} = \mathbf{M}$ and $E\mathbf{Y} = 0$. Then it is known that $\mathbf{X}\mathbf{X}' = \mathbf{S}_1$ is noncentral Wishart with n_1 degrees of freedom and $\mathbf{Y}\mathbf{Y}' = \mathbf{S}_2$ is central Wishart with n_2 degrees of freedom and the covariance matrix Σ , respectively.

Let $0 < l_1 < l_2 < \dots < l_p < 1$ be the characteristic roots of $\mathbf{S}_1\mathbf{S}_2^{-1}$; then the joint density function of l_1, \dots, l_p is given by Constantine [1],

$$(2.1) \quad c(p, n_1, n_2) \exp(\text{tr} - \Omega) |\mathbf{L}|^m |\mathbf{I} - \mathbf{L}|^n \prod_{i>j} (l_i - l_j) \sum_{\kappa=0}^{\infty} \sum_{\kappa} \frac{(\nu)_{\kappa} c_{\kappa}(\Omega) c_{\kappa}(\mathbf{L})}{(n_1/2)_{\kappa} c_{\kappa}(\mathbf{I}) k!},$$

where Ω is the noncentrality matrix, $\frac{1}{2}\mathbf{M}'\Sigma^{-1}\mathbf{M}$, and $\mathbf{L} = \text{diag}(l_1, \dots, l_p)$ and $c(p, n_1, n_2) = \pi^{p^2/2} \Gamma_p(\nu/2) / \Gamma_p(p/2) \Gamma_p(n_1/2) \Gamma_p(n_2/2)$, $m = \frac{1}{2}(n_1 - p - 1)$, $n = \frac{1}{2}(n_2 - p - 1)$, $\nu = \frac{1}{2}(n_1 + n_2)$ and $c_{\kappa}(\mathbf{L})$ are zonal polynomials defined in [3]. Consider the transformation $q_i = l_i/l_{p-1}$, $i = 1, \dots, p-2$, and decompose $c_{\kappa}(\mathbf{L}) = \sum_{\tau, \nu} a_{\tau, \nu} l_p^{k_2} c_{\nu}(\mathbf{L}_1)$ where $\mathbf{L}_1 = \text{diag}(l_1, \dots, l_{p-1})$ and the summation is over the partitions τ of k_1 and ν of k_2 such that $k_1 + k_2 = k$, and κ is the partition of k , and $a_{\tau, \nu}$ are constants defined

Received by the editors November 28, 1968.

⁽¹⁾ This research was supported by the NSF Grant GP 7663, and this paper was written when the author was at Purdue University.

in [2]. Then the joint distribution of $q_1, \dots, q_{p-2}, l_{p-1}, l_p$ can be written in the form

$$(2.2) \quad q(l_{p-1}, l_p) | \mathbf{Q} |^m | \mathbf{I} - \mathbf{Q} | | \mathbf{I} - l_{p-1} \mathbf{Q}_1 |^n | \mathbf{I} - (l_{p-1} | l_p) \mathbf{Q}_1 | \\ \times \prod_{i>j} (q_i - q_j) \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(v)_{\kappa} c_{\kappa}(\Omega)}{k! c_{\kappa}(\mathbf{I})(n_1/2)_{\kappa}} \sum_{\tau, \nu} a_{\tau, \nu} l_p^{k_1} l_{p-1}^{k_2} c_{\nu}(\mathbf{Q}_1),$$

where $\mathbf{Q} = \text{diag}(q_1, \dots, q_{p-2})$, $\mathbf{Q}_1 = \text{diag}(q_1, \dots, q_{p-2}, 1)$, and $q(l_{p-1}, l_p) = c(p, n_1, n_2) \exp \text{tr} - \Omega \cdot l_p^{m(p-1) + 1/2(p-2)(p+1)} l_p^{m+p-1} (1-l_p)^n$. By expanding $| \mathbf{I} - l_{p-1} \mathbf{Q}_1 |^n$ as well as $| \mathbf{I} - (l_{p-1} | l_p) \mathbf{Q}_1 |$ and using the results from Khatri and Pillai [5] for multiplication of zonal polynomials, we write (2.2) in the form

$$(2.3) \quad q(l_{p-1}, l_p) | \mathbf{Q} |^m | \mathbf{I} - \mathbf{Q} | \prod_{i>j} (q_i - q_j) \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(v)_{\kappa} c_{\kappa}(\Omega)}{k! (n_1/2)_{\kappa} c_{\kappa}(\mathbf{I})} \sum_{s=0}^{\infty} \sum_{\eta} \{ (-n)_{\eta} l_p^s / s! \} \\ \times \sum_{l=0}^{p-2} (c(l) l_p^l / l! l_p!) \sum_{\tau, \nu} a_{\tau, \nu} l_p^{k_1} l_{p-1}^{k_2} \sum_{\delta} g_{\delta}^l c_{\delta}(\mathbf{Q}_1),$$

where η and δ are the partitions of s and $l+s+k_2$ respectively such that $\eta = (\eta_1, \dots, \eta_p)$ and $\delta = (\delta_1, \dots, \delta_p)$ where $s = \sum_{i=1}^p \eta_i$, $l+s+k_2 = \sum_{i=1}^p \delta_i$, g_{δ}^l are constants defined in [5] and $c(l) = (-1)^l (2l)! / [(l!)^2 2^l \chi_{[21^l]}(1)]$, where $\chi_{[21^l]}$ is the degree of the representation [21^l] of the symmetric group on $2l$ symbols, and such that $\chi_{[k]}(1) = k! \prod_{i<j} (k_i - k_j - i + j) / \prod_{i=1}^k (k_i + p - i)!$ and $\kappa = (k_1 \geq k_2 \geq \dots \geq k_p \geq 0)$. Now integrate (2.3) with respect to $0 \leq q_1 \leq q_2 \leq \dots \leq q_{p-2} < 1$ by the use of the lemma in [8], we get the joint density function of l_{p-1}, l_p in the form

$$(2.4) \quad (\Gamma_{p-1}((p-1)/2) / \pi^{(p-1)^2/2}) \Gamma_{p-1}(p/2) q(l_{p-1}, l_p) \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(v)_{\kappa} c_{\kappa}(\Omega)}{k! (n_1/2)_{\kappa} c_{\kappa}(\mathbf{I})} \\ \times \sum_{s=0}^{\infty} \sum_{\eta} \{ (-n)_{\eta} / s! \} \sum_{l=0}^{p-2} \{ c(l) / l! l_p! \} \sum_{\tau, \nu} a_{\tau, \nu} l_p^{k_1} l_{p-1}^{s+l+k_2} \sum_{\delta} g_{\delta}^l c_{\delta}(\mathbf{I}_{p-1}) \\ \times ((n_1 - 1)(p - 1) / 2 + s + l + k_2) (\Gamma_{p-1}((n_1 - 1) / 2, \delta) / \Gamma_{p-1}((n_1 + p - 1) / 2, \delta)).$$

Further, integrate (2.4) with respect to l_p , then the density function of l_{p-1} can be written

$$(2.5) \quad c_1(p, n_1, n_2) \exp(\text{tr} - \Omega) l_p^{m(p-1) + (p-2)(p+1)/2} \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(v)_{\kappa} c_{\kappa}(\Omega)}{k! (n_1/2)_{\kappa} c_{\kappa}(\mathbf{I}_p)} \\ \times \sum_{s=0}^{\infty} \sum_{\eta} \{ (-n)_{\eta} / s! \} \sum_{l=0}^{p-2} (c(l) / l!) \sum_{\tau, \nu} a_{\tau, \nu} l_p^{s+l+k_2} B(l_{p-1}, 1; m+p+k_1-l; n+1) \\ \times \sum_{\delta} g_{\delta}^l c_{\delta}(\mathbf{I}_{p-1}) ((n_1 - 1)(p - 1) / 2 + s + l + k_2) \\ \times (\Gamma_{p-1}((n_1 - 1) / 2, \delta) / \Gamma_{p-1}((n_1 + p - 1) / 2, \delta))$$

where

$$B(a, b; c, d) = \int_a^b x^{c-1} (1-x)^{d-1} dx; \\ c_1(p, n_1, n_2) = \pi^{p-1} \Gamma_p(v/2) \Gamma_{p-1}((p-1)/2) / \Gamma_p(n_1/2) \Gamma_p(n_2/2).$$

It may be pointed out that the density function of the largest root can be obtained from (2.4) by integrating it with respect to l_{p-1} over the range $0 < l_{p-1} < l_p$, however a simpler form has been given in [7].

Let $\Omega = 0$ in (2.5) then the central case is of the form

$$\begin{aligned}
 & c_1(p, n_1, n_2) l_{p-1}^{m(p-1) + (p-2)(p+1)/2} \sum_{s=0}^{\infty} \sum_{\eta} \{(-n)_{\eta}/s!\} \sum_{l=0}^{p-2} \{c(l)/l!\} l_{p-1}^{s+1} \\
 (2.6) \quad & \times B(l_{p-1}, 1; m+p-l; n+1) \sum_{\delta} g_{\delta}^i c_{\delta}(\mathbf{I}_{p-1}) ((n_1-1)(p-1)/2 + s + l) \\
 & \times \Gamma_{p-1}((n_1-1)/2, \delta) / \Gamma_{p-1}((n_1+p-1)/2, \delta)
 \end{aligned}$$

where δ is the partition of $l+s$, (2.6) has been obtained by Pillai and Al-Ani [6].

3. The distribution of the second largest root in the canonical correlation case.

Let the columns of $\begin{pmatrix} \mathbf{X}_1 \\ \mathbf{X}_2 \end{pmatrix}$ be n independent normal $(p+q)$ -dimensional variates ($p \leq q$) with zero means and covariance matrix

$$\Sigma = \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma'_{12} & \Sigma_{22} \end{pmatrix}.$$

Let $\mathbf{R} = \text{diag}(r_1, r_2, \dots, r_p)$, where r_1^2, \dots, r_p^2 are the characteristics roots of the equation

$$|\mathbf{X}_1 \mathbf{X}'_2 (\mathbf{X}_2 \mathbf{X}'_2)^{-1} \mathbf{X}_2 \mathbf{X}'_1 - r^2 \mathbf{X}_1 \mathbf{X}'_1| = 0$$

and also $\mathbf{P} = \text{diag}(\rho_1, \rho_2, \dots, \rho_p)$ where $\rho_1^2, \dots, \rho_p^2$ are the characteristics roots of the equation

$$|\Sigma_{12} \Sigma_{22}^{-1} \Sigma'_{12} - \rho^2 \Sigma_{11}| = 0.$$

Then, the density function of r_1^2, \dots, r_p^2 is given by Constantine [1] in the following form

$$\begin{aligned}
 (3.1) \quad & c(n, p, q) |\mathbf{I} - \mathbf{P}^2|^{n/2} |\mathbf{R}^2|^{(q-p-1)/2} |\mathbf{I} - \mathbf{R}^2|^{(n-p-q-1)/2} \prod_{i>j} (r_i^2 - r_j^2) \\
 & \times \sum_{\kappa=0}^{\infty} \sum_{\kappa} \frac{(n/2)_{\kappa} (n/2)_{\kappa} c_{\kappa}(\mathbf{R}^2) c_{\kappa}(\mathbf{P}^2)}{(q/2)_{\kappa} \kappa! c_{\kappa}(\mathbf{I}_p)}
 \end{aligned}$$

where

$$c(n, p, q) = \frac{\Gamma_p(n/2) \pi^{p^2/2}}{\Gamma_p(q/2) \Gamma_p((n-q)/2) \Gamma_p(p/2)}.$$

By using the same transformation, namely $q_i = r_i^2 / r_{p-1}^2, i = 1, \dots, p-2$ and the same method as in §2, the joint density function of r_{p-1}^2, r_p^2 can be shown to have the following form

$$\begin{aligned}
 & c_1(n, p, q) |\mathbf{I} - \mathbf{P}^2|^{n/2} (r_{p-1}^2)^{((q-p-1)(p-1) + (p-2)(p+1))/2} (r_p^2)^{(q+p-3)/2} \\
 & \times (1 - r_p^2)^{(n-p-q-1)/2} \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(n/2)_{\kappa} (n/2)_{\kappa} c_{\kappa}(\mathbf{P}^2)}{(q/2)_{\kappa} k! c_{\kappa}(\mathbf{I}_p)} \sum_{s=0}^{\infty} \sum_{\eta} \frac{((p+q+1-n)2)_{\eta}}{s!} \\
 (3.2) \quad & \times \sum_{l=0}^{p-2} \{c(l)/l!(r_p^2)^l\} \sum_{\tau, \nu} a_{\tau, \nu} (r_p^2)^{k_1} (r_{p-1}^2)^{s+l+k_2} \sum_{\delta} g_{\delta}^{\delta} c_{\delta}(\mathbf{I}_{p-1}) \\
 & \times ((q-1)(p-1)/2 + s + l + k_2) (\Gamma_{p-1}((q-1)/2, \delta)) / \Gamma_{p-1}((q+p-1)/2, \delta),
 \end{aligned}$$

where $c_1(n, p, q) = \pi^{p-1} \Gamma_{p-1}((p-1)/2) \Gamma_p(n/2) / \Gamma_p(q/2) \Gamma_p((n-q)/2)$. Now, integrate (3.2) with respect to r_p^2 then the density function of r_{p-1}^2 can be written in the form

$$\begin{aligned}
 & c_1(n, p, q) |\mathbf{I} - \mathbf{P}^2|^{n/2} (r_{p-1}^2)^{((q-p-1)(p-1) + (p-2)(p+1))/2} \\
 & \times \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(n/2)_{\kappa} (n/2)_{\kappa} c_{\kappa}(\mathbf{P}^2)}{(q/2)_{\kappa} k! c_{\kappa}(\mathbf{I}_p)} \sum_{s=0}^{\infty} \sum_{\eta} \frac{(p+q+1-n)/2)_{\eta}}{s!} \sum_{l=0}^{p-2} \{c(l)/l!\} \\
 (3.3) \quad & \times \sum_{\tau, \nu} a_{\tau, \nu} (r_{p-1}^2)^{s+l+k_2} B(r_{p-1}^2, 1; (q+p-1)/2 + k_1 - l; (n-p-q+1)/2) \\
 & \times \sum_{\delta} g_{\delta}^{\delta} c_{\delta}(\mathbf{I}_{p-1}) ((q-1)(p-1)/2 + s + l + k_2) (\Gamma_{p-1}((q-1)/2, \delta)) / \Gamma_{p-1} \\
 & \times ((q+p-1)/2, \delta).
 \end{aligned}$$

4. Noncentral distribution of the second largest root of $\mathbf{S}_1 \mathbf{S}_2^{-1}$. Let \mathbf{S}_1 and \mathbf{S}_2 be independently distributed as Wishart $W(n_1, p, \mathbf{\Sigma}_1)$ and $W(n_2, p, \mathbf{\Sigma}_2)$, respectively. Let the characteristics roots of $\mathbf{S}_1 \mathbf{S}_2^{-1}$ and $\mathbf{\Sigma}_1 \mathbf{\Sigma}_2^{-1}$ be denoted c_i and $\lambda_i, i = 1, \dots, p$ and such that $0 < c_1 < c_2 < \dots < c_p < \infty$ and $0 < \lambda_1 < \lambda_2 < \dots < \lambda_p < \infty$.

Let $g_i = \delta c_i / (1 + \delta c_i), i = 1, \dots, p; \delta > 0$ and $\mathbf{G} = \text{diag}(g_1, \dots, g_p)$, and $\mathbf{W} = \text{diag}(\lambda_1, \dots, \lambda_p)$, then the distribution of g_1, \dots, g_p is given by Khatri [4] in the following form

$$(4.1) \quad c(p, m, n) |\delta \mathbf{W}|^{-in_1} |\mathbf{G}|^m |\mathbf{I} - \mathbf{G}|^n \prod_{i>j} (g_i - g_j) \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(\nu)_{\kappa} c_{\kappa}(\mathbf{I} - (\delta \mathbf{W})^{-1}) c_{\kappa}(\mathbf{G})}{k! c_{\kappa}(\mathbf{I})},$$

where m, n , and ν as defined in §2. Then, as before, we can obtain the joint density function of g_{p-1} and g_p in the following form

$$\begin{aligned}
 & c_1(p, n_1, n_2) |\delta \mathbf{W}|^{-in_1} g_p^{m(p-1) + \frac{1}{2}(p-2)(p+1)} g_p^{m+p-1} (1 - g_p)^n \\
 & \times \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(\nu)_{\kappa} c_{\kappa}(\mathbf{I} - (\delta \mathbf{W})^{-1})}{k! c_{\kappa}(\mathbf{I})} \sum_{s=0}^{\infty} \sum_{\eta} \{(-n)_{\eta} / s!\} \sum_{l=0}^{p-2} c(l)/l! g_p^l \\
 (4.2) \quad & \times \sum_{\tau, \nu} a_{\tau, \nu} g_p^{k_1} g_{p-1}^{s+l+k_2} \sum_{\delta} g_{\delta}^{\delta} c_{\delta}(\mathbf{I}_{p-1}) ((n_1-1)(p-1)/2 + s + l + k_2) \\
 & \times (\Gamma_{p-1}((n_1-1)/2, \delta)) / \Gamma_{p-1}((n_1+p-1)/2, \delta).
 \end{aligned}$$

Now, integrate (4.2) with respect to g_p , the density function if g_{p-1} can be written in the following form

$$\begin{aligned}
 & c_1(p, n_1, n_2) |\delta \mathbf{W}|^{-\frac{1}{2} n_1} g_{p-1}^{m(p-1) + (p-2)(p+1)/2} \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(p)_{\kappa} c_{\kappa}(\mathbf{I} - (\delta \mathbf{W})^{-1})}{k! c_{\kappa}(\mathbf{I})} \\
 & \times \sum_{s=0}^{\infty} \sum_{\eta} \{(-n)_{\eta} / s!\} \sum_{l=0}^{\infty} (c(l) / l!) \sum_{\tau, \nu} a_{\tau, \nu} g_{p-1}^{s+l+k_2} B(g_{p-1}, 1; \\
 & \times m+p+k_1-l; n+1) \sum_{\delta} g_{\delta}^{\delta} c_{\delta}(\mathbf{I}_{p-1}) ((n_1-1)(p-1)/2 + s+l+k_2) \\
 & \times (\Gamma_{p-1}((n_1-1)/2, \delta) / \Gamma((n_1+p-1)/2, \delta)).
 \end{aligned}$$

5. The distribution of the second largest root of a covariance matrix. The distribution of the characteristics roots, $0 < \omega_1 \leq \dots \leq \omega_p < \infty$, of $\mathbf{X}\mathbf{X}'$ depends only upon the characteristic roots of $\mathbf{\Sigma}$ and can be given in the form (James [3])

$$(5.1) \quad k(p, n) |\mathbf{\Sigma}|^{-\frac{1}{2} n} |\mathbf{W}_1|^m \{ \exp(-\frac{1}{2} \text{tr } \mathbf{W}_1) \} \prod_{i>j} (\omega_i - \omega_j) {}_0F_0(\frac{1}{2}(\mathbf{I}_p - \mathbf{\Sigma}^{-1}), \mathbf{W}_1)$$

where $k(p, n) = \pi^{\frac{1}{2} p^2} / 2^{\frac{1}{2} p n} \Gamma_p(n/2) \Gamma_p(p/2)$.

It may be pointed out that the form (5.1) can also be viewed as a limiting form of (4.1), when $n_2 \rightarrow \infty$.

However, by methods similar to those in the previous sections, the density function of the second largest root $\gamma_{p-1} = \omega_{p-1} / 2$ can be written in the form

$$\begin{aligned}
 & k_1(p, n) |\mathbf{\Sigma}|^{-\frac{1}{2} n} \gamma_{p-1}^{m(p-1) + (p-2)(p+1)/2} e^{-\gamma_{p-1}} \sum_{k=0}^{\infty} \sum_{\kappa} \frac{c_{\kappa}(\mathbf{I} - \mathbf{\Sigma}^{-1})}{k! c_{\kappa}(\mathbf{I})} \sum_{\tau, \nu} a_{\tau, \nu} \\
 (5.2) \quad & \times \sum_{s=0}^{\infty} \sum_{\eta} \frac{(-1)^s}{s!} \sum_{l=0}^{p-2} c(l) \gamma_{p-1}^{s+l+k_2} \sum_{i=0}^{k_2} \sum_{\delta} b_{\delta, \nu} \sum_{\mu} g_{\delta, \nu}^{\mu} c_{\mu}(\mathbf{I}_{p-2}) [\Gamma_{p-2} \\
 & \times ((n-2)/2, \mu) / \Gamma_{p-2}((n+p)/2, \mu)] [I(\gamma_{p-1}, \infty; m+p+k_1-j) - \gamma_{p-1} \\
 & \times I(\gamma_{p-1}, \infty; m+p+k_1-j-1)],
 \end{aligned}$$

where $b_{\delta, \nu}$ are constants defined in [5], δ and μ are the partitions of i and $i+l+s$ respectively.

$$k_1(p, n) = k(p, n) \Gamma_{p-2}((p-2)/2) \Gamma_{p-2}((p+1)/2) / \pi^{(p-2)^2/2},$$

and

$$I(a, b; c) = \int_a^b x^{c-1} e^{-x} dx.$$

ACKNOWLEDGEMENT. The author wishes to express his very sincere thanks to Professor K. C. S. Pillai of Purdue University for suggesting the problem and his generous discussion and guidance during the preparation of the paper.

REFERENCES

1. A. G. Constantine, *Some non-central distribution problems in multivariate analysis*, Ann. Math. Statist. **34** (1963), 1270-1285.

2. T. Hayakawa, *On the distribution of the maximum latent root of a positive definite symmetric random matrix*, Ann. Inst. Statist. Math. **19** (1967), 1–17.
3. A. T. James, *The distribution of the latent roots of the covariance matrix*, Ann. Math. Statist. **31** (1960), 151–158.
4. C. G. Khatri, *Some distribution problems connected with the characteristic root of $S_1 S_2^{-1}$* , Ann. Math. Statist. **38** (1967), 944–948.
5. C. G. Khatri, and K. C. S. Pillai, *On the non-central distribution of two test criteria in multivariate analysis of variance*, Ann. Math. Statist. **39**, 1968 (to appear).
6. K. C. S. Pillai and S. Al-Ani, *On some distribution problems concerning characteristic roots and vectors in multivariate analysis*. Mimeo. Series **123**, Dept. of Statistics, Purdue University (1967).
7. K. C. S. Pillai and T. Sugiyama, *Non-central distributions of the largest latent roots of three matrices in multivariate analysis*, Mimeo. Series **129**, Dept. of Statistics, Purdue University, 1967.
8. T. Sugiyama, *On the distribution of the largest latent root of the covariance matrix*, Ann. Math. Statist. **38** (1967), 1148–1151.

UNIVERSITY OF CALGARY,
CALGARY, ALBERTA