Polymorphisms in the methylenetetrahydrofolate reductase (MTHFR) gene, intakes of folate and related B vitamins and colorectal cancer: a case–control study in a population with relatively low folate intake

Linda Sharp1,2*, Julian Little1,3, Nigel T. Brockton1,4, Seonaidh C. Cotton1,5, Lindsey F. Masson1,5, Neva E. Haites1 and Jim Cassidy1,6

1Department of Medicine and Therapeutics, University of Aberdeen, Polwarth Building, Foresterhill, Aberdeen AB25 2ZD, UK
2National Cancer Registry Ireland, Elm Court, Boreennanma Road, Cork, Republic of Ireland
3Canada Research Chair in Human Genome Epidemiology, Department of Epidemiology and Community Medicine, University of Ottawa, Ottawa, Canada
4Maternal and Child Health Sciences, University of Dundee, Ninewells Hospital and Medical School, Dundee DD1 9SY, UK
5Department of Public Health, University of Aberdeen, Polwarth Building, Foresterhill, Aberdeen AB25 2ZD, UK
6Beatson Oncology Centre, Western Infirmary, Glasgow G11 6NT, UK

(Received 16 October 2006 – Revised 21 May 2007 – Accepted 15 June 2007)

Folate is key in one-carbon metabolism, disruption of which can interfere with DNA synthesis, repair, and methylation. Efficient one-carbon metabolism requires other B vitamins and the optimal activity of enzymes including 5,10-methylenetetrahydrofolate reductase (MTHFR). We report a population-based case–control study of folate intake, related dietary factors and MTHFR polymorphisms (C677T, A1298C) and colorectal cancer in a population with relatively high colorectal cancer incidence and relatively low folate intake. A total of 264 cases with histologically confirmed incident colorectal cancer and 408 controls participated. There was no clear trend in risk with reported intakes of total, or dietary, folate, riboflavin, vitamin B12 or vitamin B6, nor were there interactions between folate intake and the other B vitamins or alcohol. For C677T, risk decreased with increasing variant alleles (multivariate OR for CT v. CC = 0.77 (95 % CI 0.52, 1.16); OR for TT v. CC = 0.62 (95 % CI 0.31, 1.24)), which, although not statistically significant, was consistent with previous studies. For A1298C, compared with AA subjects, CC subjects had modest, non-significant, reduced risk (multivariate OR = 0.81 (95 % CI 0.45, 1.59)). There were significant interactions between total folate and C677T (P = 0.029) and A1298C (P = 0.025), and total vitamin B6 and both polymorphisms (C677T, P = 0.016; A1298C, P = 0.033), although the patterns observed differed from previous studies. Seen against the setting of low folate intake, the results suggest that the role of folate metabolism in colorectal cancer aetiology may be more complex than previously thought. Investigation of particular folate vitamers (for example, tetrahydrofolate, 5,10-methylenetetrahydrofolate) may help clarify carcinogenesis pathways.

Folate: MTHFR: Colorectal cancer: One-carbon metabolism

Folate, and its synthetic form, folic acid, is key in one-carbon metabolism, the disruption of which can interfere with DNA synthesis, repair and methylation. Low folate status or, more generally, low dietary methyl status (a combination of folate, methionine, alcohol and other B vitamins) may promote carcinogenesis. The mechanisms by which low folate could increase risk of malignancy include: (1) DNA hypomethylation and inappropriate activation of oncogenes1; and/or (2) uracil misincorporation during DNA repair and synthesis, leading to DNA strand breaks, chromosome damage and, eventually, malignant transformation2. Higher reported folate intake has been associated with reduced colorectal cancer risk3. However, most studies have been conducted in populations where intake is relatively high, and a substantial proportion comes in the form of folic acid, either from supplements or through fortified foods; folic acid is more bioavailable than natural folates4.

Efficient one-carbon metabolism also requires riboflavin, vitamin B6, vitamin B12 and methionine. This raises the possibility that these factors, either on their own account, or by acting together with folate, might influence colorectal cancer risk. However, the available evidence is limited and/or inconsistent5.

One-carbon metabolism further requires the optimal activity of multiple enzymes, including 5,10-methylenetetrahydrofolate reductase (MTHFR), which directs the folate pool towards methylation or DNA repair. Several polymorphisms have been reported in the MTHFR gene, but only two have been investigated in relation to colorectal neoplasia – C677T and A1298C6. The 677T variant lowers enzyme activity in vitro7.
and has been associated with decreased plasma folate and vitamin B₁₂ levels and raised homocysteine⁹,⁸. In vitro studies suggest that low folate status may be required for this polymorphism to affect enzyme activity in vivo⁹,₁⁰. The functional impact of the A1298C polymorphism is less clear⁹,₁¹. In vitro enzyme activity appears to be reduced in homozygous variants and compound heterozygotes, but the in vivo effects remain to be determined.

Most, but not all, previous studies have found reduced colorectal cancer risk in homozygotes for the variant C677T allele⁵,₁⁴. A1298C has been less extensively investigated, and results have been inconsistent⁵,₁⁵. The inconsistencies may be due, in part, to differences in genotype distributions and folate intakes between populations.

The Scottish diet is distinguished by low vegetable intake¹⁶. This suggests folate levels may be relatively low, a view confirmed by the National Food Survey which revealed geographical variations in blood folate across Britain, with the lowest levels in Scotland¹⁷. The high alcohol intake¹⁸ and smoking prevalence¹⁹ may further compromise folate and, more generally, methyl status in the Scottish population. Furthermore, colorectal cancer incidence in Scotland is in the upper quarter of rates observed worldwide²⁰.

We undertook a population-based case–control study in the north-east of Scotland to investigate whether intakes of folate and related dietary factors, and MTHFR polymorphisms, were associated with colorectal cancer risk in a population with a relatively low folate intake.

**Materials and methods**

**Ascertainment of cases and controls**

Eligible cases were resident in the Grampian Health Board region and diagnosed with a histological confirmed first primary invasive tumour of the colon or rectum between September 1998 and February 2000. They were identified each month from the computerised records of Aberdeen Royal Infirmary pathology laboratory, which provides the centralised service for Grampian.

Controls were selected from the Grampian Community Health Index, an inventory of everyone registered with a general practitioner. The Grampian Community Health Index has a high level of completeness for the Grampian population²¹. Each month, a pool of potential controls was selected at random, frequency matched to cases on age and sex. Individuals who declined to take part were replaced.

The study was approved by the Joint Ethical Committee of the University of Aberdeen and Grampian Health Board (Scotland, UK). All participants gave informed consent.

**Assessment of diet and other exposures**

With the permission of their general practitioner, subjects were recruited by post. They completed a questionnaire comprising a 150-item semi-quantitative FFQ²² plus questions on a range of socio-demographic and lifestyle factors relevant to colorectal cancer aetiology. The FFQ had been validated in the local population²³ and included questions on dietary supplement use in the reference period (approximately 1 year before completion).

**Genotyping**

Subjects provided a mouthwash sample. DNA was extracted from exfoliated buccal cells using the Elucigene CF12 protocol (Zeneca Diagnostics, Abingdon, Oxon, UK) and amplified by PCR. The amplification products were digested with Hinfl for C677T²⁴ and MboII for A1298C²⁴. Resultant products were separated by electrophoresis through 3 % (Metaphor BioProducts) gels. Bands were visualised by ethidium bromide staining and UV transillumination.

Analyses were blind to case–control status. Each batch contained a negative (no template DNA) and two positive controls. Analyses were repeated if either amplification or digestion failed. Gels were double-read by two individuals and assays repeated for any samples considered ambiguous. Genotyping was also repeated, blind to the original results, for a 10 % random selection (n 67) of samples; no differences were found compared with the originally assigned genotypes.

**Statistical analysis**

The dietary data were converted into estimated nutrient intakes using the UK national food composition tables, taking account of cereal fortification²⁵. For seven subjects (three cases and four controls) the FFQ was not completed or was very incomplete. A further twelve subjects (six cases and six controls) were excluded on the basis of implausible total energy intakes (<3347 kJ/d for men or <2092 kJ/d for women²⁶ or more than three standard deviations above mean intake in men and women separately). Protein intake was used as a surrogate for methionine. Dietary intakes were adjusted for total energy using the nutrient residual method²⁶. For folate, vitamin B₆, vitamin B₁₂ and riboflavin, the primary analyses related to total intake, i.e. the sum of dietary (energy adjusted) and supplement intakes; secondary analyses related to dietary intake. Subjects were grouped into intake tertiles (for alcohol) or quartiles (other variables), based on the combined case and control distribution²⁷, with the baseline group comprising non-consumers (alcohol) or the lowest quartile. For methyl content, the ‘low’ group comprised low total folate and protein (< median) and high alcohol (upper two tertiles) and the ‘high’ group had high total folate and protein (≥ median) and low alcohol (no intake or lowest tertile).

For C677T and A1298C, separate and compound genotype was assigned. Hardy–Weinberg equilibrium was assessed in controls using Pearson’s χ². Association analyses were done for individual genotypes, and combining carriers of the variant alleles, taking homozygotes for the common allele as the reference group. In addition, to explore the impact of one variant in the absence of the other, risk estimates were computed for (i) C677T genotype among 1298AA carriers and (ii) A1298C genotypes among 677CC carriers. Logistic regression was used to compute adjusted and multivariate odds ratios (OR) and 95 % CI in Stata 8 (StataCorp LP, College Station, TX, USA). Adjusted OR were adjusted for sex, age and, for the dietary variables, total energy intake. Multivariate OR were further adjusted for potential confounders; factors which made a significant contribution (likelihood ratio test P<0.1) were retained in
the model. Trend tests were used to investigate dose–response across dietary intakes or by number of variant alleles. Effect modification was explored by computing OR for combinations of (i) intake variables and (ii) intake and genotype. The test for interaction was the change in deviance ($-2 \times \log$ likelihood) between a main-effects model and one including a cross-product term. All multivariate models were checked for adequate fit.

**Results**

A total of 264 cases (62% of those eligible) and 408 controls (61%) participated. Of the 264 cases, 189 cases had colon cancer and seventy-five rectal cancer (Table 1). Cases tended to be older than controls, and a higher proportion had a first-degree family history of colorectal cancer and seventy-five rectal cancer (Table 1). Cases included 200 (thirty-eight controls (9·4%) and twenty-five cases (9·6%)) who reported taking a supplement that contained folic acid. The majority (fifty-three out of sixty-three) took supplements that were men.

Median total folate intake was 299·7 (interquartile range 261–332) g/d. Of the subjects, 9·5% (thirty-eight controls (9·4%) and twenty-five cases (9·6%)) reported taking a supplement that contained folic acid. The majority (fifty-three out of sixty-three) took supplements that included 200 µg folic acid.

For total folate intake, the multivariate OR exceeded unity for quartiles 2–4, but there was no clear trend (Table 2).

<table>
<thead>
<tr>
<th>Table 1. Characteristics of study participants by case–control status (Numbers and percentages)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Factor</td>
</tr>
<tr>
<td>Sex</td>
</tr>
<tr>
<td>Female</td>
</tr>
<tr>
<td>Male</td>
</tr>
<tr>
<td>Age (years)†</td>
</tr>
<tr>
<td>&lt; 55</td>
</tr>
<tr>
<td>55–64</td>
</tr>
<tr>
<td>65–74</td>
</tr>
<tr>
<td>75 and older</td>
</tr>
<tr>
<td>Tumour location</td>
</tr>
<tr>
<td>Colon</td>
</tr>
<tr>
<td>Rectum</td>
</tr>
<tr>
<td>First-degree family history of colorectal cancer</td>
</tr>
<tr>
<td>No</td>
</tr>
<tr>
<td>Yes</td>
</tr>
<tr>
<td>Social class‡</td>
</tr>
<tr>
<td>I and II (non-manual)</td>
</tr>
<tr>
<td>III (skilled manual)</td>
</tr>
<tr>
<td>IV and V (unskilled manual)</td>
</tr>
<tr>
<td>Country of birth</td>
</tr>
<tr>
<td>UK or Ireland</td>
</tr>
<tr>
<td>Elsewhere</td>
</tr>
<tr>
<td>Smoking status§</td>
</tr>
<tr>
<td>Never smoked</td>
</tr>
<tr>
<td>Ex-smoker</td>
</tr>
<tr>
<td>Current smoker</td>
</tr>
</tbody>
</table>

* Percentage of those subjects with complete data for a particular factor.
† Age at diagnosis for cases and at pseudo date of diagnosis for controls (middle of month in which cases diagnosed).
‡ Based on occupation at time questionnaire completed, or last occupation before retirement: for men and single women, based on own occupation; for women who were married, living as married, or widowed, based on occupation of husband or partner.
§ At approximately 1 year before the completion of the questionnaire.

A similar pattern was observed for total riboflavin. There was a borderline significant association between protein intake and risk (global $P=0·055$; $P$ (linear trend)$=0·066$), due to a one-third reduction in risk in the highest intake group. Compared with non-drinkers, disease risk was slightly, but non-significantly, raised for all categories of alcohol intake. The association was strongest in males (multivariate OR for T1 v. none $=1·58$ (95% CI 0·60, 4·15); OR for T2 v. none $=1·57$ (95% CI 0·65, 3·80); OR for T3 v. none $=1·72$ (95% CI 0·73, 4·04)). There was a suggestion of increasing risk with decreasing dietary methyl content, although the risk estimates were not statistically significant. There were no associations between total vitamin B12 or vitamin B6 intakes and colorectal cancer.

Results of analyses restricted to dietary intake of folate, vitamin B12, vitamin B6, and riboflavin were similar to those for total intake, with slightly attenuated risk estimates (data not shown).

There was no evidence of interactions between total folate intake and total intake of vitamin B12, vitamin B6, or riboflavin, nor did particular combinations of alcohol and total folate influence risk (data not shown). Results were unchanged when analyses were repeated for dietary intake (data not shown).

The C677T genotype distribution among controls conformed with Hardy–Weinberg equilibrium ($x^2 = 0·00; P = 0·999$), but the A1298C distribution did not ($x^2 = 6·25; P < 0·05$); there was an excess of homozygous subjects (AA and CC) and a deficit of heterozygotes. For C677T, risk decreased with increasing number of variant alleles (multivariate OR for CT v. CC $=0·77$ (95% CI 0·52, 1·16); OR for TT v. CC $=0·62$ (95% CI 0·31, 1·24)) but not significantly ($P$ (linear trend)$=0·103$; Table 3). For A1298C, the risk for homozygous variant subjects was modestly, but not significantly, reduced (OR for CC v. AA $=0·81$ (95% CI 0·45, 1·49)). As regards compound genotype, no subjects were homozygous for both variants, while four cases (17%) and eleven controls (28%) had three variant alleles. Compared with those with no variant alleles, individuals who were double heterozygous (OR $=0·85$ (95% CI 0·41, 1·75)) or homozygous variant (OR $=0·71$ (95% CI 0·35, 1·45)) had slightly reduced risk. Among individuals who were 1298AA, multivariate OR for the C677T genotypes were: 1·00, 0·82, 0·77 (95% CI 0·41, 1·58) and 0·75 (95% CI 0·33, 1·73).

There was a significant interaction between C677T and total folate (Table 4; $P$ (interaction)$=0·029$). Compared with CC individuals with low (< median) intake, risk was significantly reduced in carriers of the variant allele with low intake (OR $=0·47$ (95% CI 0·27, 0·83)), with a less pronounced risk reduction in those with high intake. A significant interaction was found between total vitamin B6 intake and C677T (P (interaction)$=0·016$), which followed a similar pattern to that for total folate. Repeating the analyses for dietary folate and vitamin B6 gave less pronounced risk estimates (data not shown).

There was a suggestion that C677T modified the associations between protein and alcohol intakes and colorectal cancer, although the tests for interaction were not statistically
significant. There were no interactions between C677T and either vitamin B$_{12}$ or riboflavin.

A significant interaction was observed between total folate and A1298C (Table 5; $P$(interaction)$=0·025$). The group at lowest risk comprised homozygotes for the wild-type allele with low intake. Compared with this group, those with the AC/CC genotype and low intake or AA genotype and high intake had significantly raised risk (OR$=1·75$ (95% CI $1·00$, $3·10$) and OR$=1·91$ (95% CI $1·05$, $3·50$) respectively). Similar patterns were apparent for dietary folate (data not shown; $P$(interaction)$=0·060$) and total (Table 5; $P$(interaction)$=0·033$) and dietary vitamin B$_{6}$ (data not shown; $P$(interaction)$=0·053$). A1298C did not modify the relationships between colorectal cancer and vitamin B$_{12}$ or

Table 2. Association between intake variables and colorectal cancer
(Numbers and percentages of subjects, $P$ values for likelihood ratio tests, odds ratios and 95% confidence intervals)

<table>
<thead>
<tr>
<th></th>
<th>Controls</th>
<th>Cases</th>
<th>Adjusted OR*</th>
<th>95% CI</th>
<th>Multivariate OR†</th>
<th>95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>Total folate ($\mu$g/d)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Q1, $\leq$ 263·9</td>
<td>102</td>
<td>25·6</td>
<td>26·4</td>
<td>1·0</td>
<td>–</td>
<td>1·0</td>
</tr>
<tr>
<td>Q2, 264·0–299·6</td>
<td>90</td>
<td>22·6</td>
<td>73</td>
<td>1·66</td>
<td>0·94, 2·66</td>
<td>1·87, 3·22</td>
</tr>
<tr>
<td>Q3, 299·7–348·5</td>
<td>97</td>
<td>24·4</td>
<td>66</td>
<td>1·48</td>
<td>0·92, 2·38</td>
<td>0·77, 1·77</td>
</tr>
<tr>
<td>Q4, $\geq$ 348·6</td>
<td>109</td>
<td>27·4</td>
<td>54</td>
<td>1·05</td>
<td>0·65, 1·70</td>
<td>1·37, 1·92</td>
</tr>
<tr>
<td>Global $P$</td>
<td></td>
<td></td>
<td>0·087</td>
<td>0·098</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Trend $P$</td>
<td></td>
<td></td>
<td>0·982</td>
<td>0·309</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Total vitamin B$_{12}$ ($\mu$g/d)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Q1, $\leq$ 5·25</td>
<td>100</td>
<td>25·1</td>
<td>64</td>
<td>1·0</td>
<td>–</td>
<td>1·0</td>
</tr>
<tr>
<td>Q2, 5·26–6·46</td>
<td>97</td>
<td>24·4</td>
<td>66</td>
<td>1·21</td>
<td>0·75, 1·94</td>
<td>1·21, 2·05</td>
</tr>
<tr>
<td>Q3, 6·46–7·97</td>
<td>99</td>
<td>24·9</td>
<td>64</td>
<td>1·07</td>
<td>0·67, 1·72</td>
<td>1·21, 2·04</td>
</tr>
<tr>
<td>Q4, $\geq$ 7·98</td>
<td>102</td>
<td>25·6</td>
<td>61</td>
<td>0·77</td>
<td>0·63, 1·63</td>
<td>0·95, 1·62</td>
</tr>
<tr>
<td>Global $P$</td>
<td></td>
<td></td>
<td>0·858</td>
<td>0·729</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Trend $P$</td>
<td></td>
<td></td>
<td>0·258</td>
<td>0·866</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Total vitamin B$_{6}$ (mg/d)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Q1, $\leq$ 2·29</td>
<td>94</td>
<td>23·6</td>
<td>70</td>
<td>1·0</td>
<td>–</td>
<td>1·0</td>
</tr>
<tr>
<td>Q2, 2·30–2·61</td>
<td>97</td>
<td>24·4</td>
<td>66</td>
<td>1·17</td>
<td>0·73, 1·87</td>
<td>1·03, 2·00</td>
</tr>
<tr>
<td>Q3, 2·62–3·03</td>
<td>103</td>
<td>25·9</td>
<td>60</td>
<td>1·08</td>
<td>0·67, 1·74</td>
<td>1·24, 2·12</td>
</tr>
<tr>
<td>Q4, $\geq$ 3·04</td>
<td>104</td>
<td>26·1</td>
<td>59</td>
<td>1·09</td>
<td>0·62, 1·59</td>
<td>1·07, 2·02</td>
</tr>
<tr>
<td>Global $P$</td>
<td></td>
<td></td>
<td>0·894</td>
<td>0·862</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Trend $P$</td>
<td></td>
<td></td>
<td>0·891</td>
<td>0·661</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Total riboflavin (mg/d)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Q1, $\leq$ 1·87</td>
<td>106</td>
<td>26·6</td>
<td>58</td>
<td>1·0</td>
<td>–</td>
<td>1·0</td>
</tr>
<tr>
<td>Q2, 1·88–2·19</td>
<td>96</td>
<td>24·1</td>
<td>67</td>
<td>1·36</td>
<td>0·84, 2·18</td>
<td>1·43, 3·26</td>
</tr>
<tr>
<td>Q3, 2·14–2·48</td>
<td>98</td>
<td>24·6</td>
<td>65</td>
<td>1·30</td>
<td>0·81, 2·08</td>
<td>1·63, 2·79</td>
</tr>
<tr>
<td>Q4, $\geq$ 2·49</td>
<td>98</td>
<td>24·6</td>
<td>65</td>
<td>1·32</td>
<td>0·82, 2·12</td>
<td>1·44, 2·47</td>
</tr>
<tr>
<td>Global $P$</td>
<td></td>
<td></td>
<td>0·545</td>
<td>0·312</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Trend $P$</td>
<td></td>
<td></td>
<td>0·513</td>
<td>0·165</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Protein (g/d)‡</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Q1, $\leq$ 85·2</td>
<td>97</td>
<td>24·4</td>
<td>65</td>
<td>1·0</td>
<td>–</td>
<td>1·0</td>
</tr>
<tr>
<td>Q2, 85·3–93·6</td>
<td>81</td>
<td>20·4</td>
<td>80</td>
<td>1·61</td>
<td>1·00, 2·58</td>
<td>1·43, 2·43</td>
</tr>
<tr>
<td>Q3, 93·6–102·9</td>
<td>105</td>
<td>26·4</td>
<td>61</td>
<td>1·07</td>
<td>0·67, 1·73</td>
<td>0·92, 1·57</td>
</tr>
<tr>
<td>Q4, $\geq$ 103·0</td>
<td>115</td>
<td>28·9</td>
<td>49</td>
<td>0·78</td>
<td>0·48, 1·26</td>
<td>0·67, 1·96</td>
</tr>
<tr>
<td>Global $P$</td>
<td></td>
<td></td>
<td>0·026</td>
<td>0·055</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Trend $P$</td>
<td></td>
<td></td>
<td>0·143</td>
<td>0·066</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Alcohol (g/d)$§$</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>None</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>T1, $\leq$ 3·91</td>
<td>103</td>
<td>25·9</td>
<td>72</td>
<td>1·02</td>
<td>0·61, 1·70</td>
<td>1·22, 2·23</td>
</tr>
<tr>
<td>T2, 3·94–11·99</td>
<td>106</td>
<td>28·6</td>
<td>69</td>
<td>1·20</td>
<td>0·73, 1·97</td>
<td>1·26, 2·25</td>
</tr>
<tr>
<td>T3, $\geq$ 12·00</td>
<td>112</td>
<td>28·1</td>
<td>63</td>
<td>1·14</td>
<td>0·67, 1·93</td>
<td>1·22, 2·23</td>
</tr>
<tr>
<td>Global $P$</td>
<td></td>
<td></td>
<td>0·876</td>
<td>0·877</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Trend $P$</td>
<td></td>
<td></td>
<td>0·510</td>
<td>0·553</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Methyl content§</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>High</td>
<td>57</td>
<td>14·3</td>
<td>27</td>
<td>1·0</td>
<td>–</td>
<td>1·0</td>
</tr>
<tr>
<td>Intermediate</td>
<td>286</td>
<td>71·9</td>
<td>188</td>
<td>1·45</td>
<td>0·86, 2·43</td>
<td>1·36, 2·40</td>
</tr>
<tr>
<td>Low</td>
<td>55</td>
<td>13·8</td>
<td>40</td>
<td>1·59</td>
<td>0·83, 3·06</td>
<td>1·54, 3·19</td>
</tr>
<tr>
<td>Global $P$</td>
<td></td>
<td></td>
<td>0·306</td>
<td>0·474</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Trend $P$</td>
<td></td>
<td></td>
<td>0·175</td>
<td>0·250</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Q, quartile; T, tertile; NSAID, non-steroidal anti-inflammatory drugs.
* Adjusted for age, sex and total energy.
†All models adjusted for sex, age, total energy, physical activity, family history of colorectal cancer, regular use of any NSAID, sex × NSAID interaction term; model for protein also adjusted for type of dietary supplement; model for alcohol also adjusted for type of dietary supplement and protein.
‡Protein from food only.
§High methyl content is high total folate, high protein and either no alcohol intake or intake in T1; low methyl status is low total folate, low protein and alcohol in upper two tertiles; intermediate is all other combinations of total folate, protein and alcohol. High folate or protein is intake in upper two quartiles; low folate or protein is intake in lower two quartiles.
riboflavin, either for total (Table 5) or dietary intake (data not shown). There were no significant interactions between A1298C and protein or alcohol.

Discussion

Strengths and limitations

The major strength of the study was the population basis, with both cases and control recruited from population-based sampling frames. The participation rate was similar to another UK study in which contact was made by post30. Other than refusal, the main reasons for non-participation among cases were death (15; 3·5 %), primarily because subjects were deemed to be too ill to approach. For controls, the general practitioner refused permission to approach sixteen individuals and a further eight had died (3·6 % in total). Eligible individuals and a further eight had died (3·6 % in total). Eligible

Table 3. Association between methylenetetrahydrofolate reductase (MTHFR) genotype and colorectal cancer (Numbers and percentages of subjects, \( P \) values for likelihood ratio tests, odds ratios and 95 % confidence intervals)

<table>
<thead>
<tr>
<th>Controls</th>
<th>Cases</th>
</tr>
</thead>
<tbody>
<tr>
<td>( n )</td>
<td>( % )</td>
</tr>
<tr>
<td>C677T‡</td>
<td>170</td>
</tr>
<tr>
<td>Homozygous wild type (CC)</td>
<td>170</td>
</tr>
<tr>
<td>Heterozygous (CT)</td>
<td>177</td>
</tr>
<tr>
<td>Homozygous variant (TT)</td>
<td>47</td>
</tr>
<tr>
<td>Global ( P )</td>
<td>0·603</td>
</tr>
<tr>
<td>Trend ( P )</td>
<td>0·329</td>
</tr>
<tr>
<td>Heterozygous/homozygous variant (CT/TT)§</td>
<td>224</td>
</tr>
<tr>
<td>Global ( P )</td>
<td>0·441</td>
</tr>
<tr>
<td>A1298C¶</td>
<td>177</td>
</tr>
<tr>
<td>Homozygous wild type (AA)</td>
<td>177</td>
</tr>
<tr>
<td>Heterozygous (AC)</td>
<td>157</td>
</tr>
<tr>
<td>Homozygous variant (CC)</td>
<td>60</td>
</tr>
<tr>
<td>Global ( P )</td>
<td>0·100</td>
</tr>
<tr>
<td>Trend ( P )</td>
<td>0·391</td>
</tr>
<tr>
<td>Heterozygous/homozygous variant (AC/CC)§</td>
<td>217</td>
</tr>
<tr>
<td>Global ( P )</td>
<td>0·879</td>
</tr>
<tr>
<td>Compound genotype¶</td>
<td>43</td>
</tr>
<tr>
<td>Homozygous wild type (CC and AA)</td>
<td>43</td>
</tr>
<tr>
<td>Single heterozygous (CT or AC)</td>
<td>166</td>
</tr>
<tr>
<td>Double heterozygous (CT and AC)</td>
<td>74</td>
</tr>
<tr>
<td>Homozygous variant (CC and AA)</td>
<td>43</td>
</tr>
<tr>
<td>Global ( P )</td>
<td>0·176</td>
</tr>
<tr>
<td>Trend ( P )</td>
<td>0·076</td>
</tr>
<tr>
<td>Up to one variant allele**</td>
<td>209</td>
</tr>
<tr>
<td>Two or more variant alleles††</td>
<td>178</td>
</tr>
<tr>
<td>Global ( P )</td>
<td>0·078</td>
</tr>
</tbody>
</table>

NSAID, non-steroidal anti-inflammatory drugs.

* Adjusted for sex, age, family history of colorectal cancer, physical activity, regular use of any NSAID, sex \( \times \) NSAID interaction term, total energy intake and type of dietary supplements.

† Adjusted for age and sex.

‡ Three cases did not provide a mouthwash sample; genotyping failed for ten cases and fourteen controls.

§ OR for carriers of variant allele v. homozygous wild-types.

‖ Three cases did not provide a mouthwash sample; genotyping failed for sixteen cases and fourteen controls.

§§ Includes homozygous wild type for both polymorphisms or heterozygote for either polymorphism.

†† Includes heterozygotes for both polymorphisms or homozygous variant for either; OR for two or more variant alleles v. zero or one variant allele.

risk factors resulted in associations consistent with previous evidence32.

Of the controls, 12 % were 677TT homozygotes, consistent with the frequency in other white and northern European populations33. Of the controls, 15 % carried the 1298 CC genotype, slightly higher than the prevalence in most European studies34, but consistent with another study in the same area (18 % CC)35. The A1298C genotype frequencies were not in Hardy–Weinberg equilibrium. Since subjects were unaware of the study hypotheses or their genotype, differential participation by genotype seems unlikely. Moreover, the genotyping followed rigorous quality-control measures, which should help guard against systematic errors35. Other potential explanations for a departure from Hardy–Weinberg equilibrium include non-random mating, genetic drift and chance.

The FFQ was developed and extensively validated in the local area. For folate, alcohol and riboflavin, high levels of agreement (rank correlation coefficients 0·55–0·79) were found when comparing questionnaire responses with 4 d weighed records36.

The case–control design is potentially susceptible to the effects of recall bias in the assessment of lifestyle (but not
Unlike most previous studies, we did not find an inverse relationship between total folate intake and colorectal cancer. In contrast, compared with the lowest intake quartile, risk estimates for the other quartiles exceeded unity, and followed a bell-shaped pattern. A similar pattern has been reported between plasma folate and colorectal cancer in a prospective study from Sweden. When our analysis was restricted to subjects who did not use folic acid-containing supplements, the pattern persisted (multivariate OR: 1.0, 1.86 (95 % CI 1.07, 3.22), 1.82 (95 % CI 1.03, 3.19) and 0.96 (95 % CI 0.51, 1.79)). Although some previous studies have found the inverse relationship to be stronger for colon than rectal tumours, when we stratified by site, there was no clear evidence of a trend with colon or rectal cancer. Although there may be substantial error in assessment of dietary folate intakes and supplement use, the resulting misclassification is most probably random, and would be unlikely to produce the observed result.

In further analyses, in women we found a modest, albeit non-significant, reduced risk in the highest total folate group (age- and sex-adjusted OR for quartile 4 v. quartile 1 = 0.73 (95 % CI 0.35, 1.52)). Several other case–control studies found that the association between colorectal neoplasia and folate (as intake or from supplements) was only evident, or was stronger, amongst females and folate-containing supplements. These findings raise the possibility that folate status (or bioavailable folate, at least) may be genetic) risk factors. Other than this, the main limitation was limited power both for main effects and, in particular, for interactions. As regards main effects, for a genetic variant occurring in 12 % of the population (i.e. MTHFR TT) the study had 80 % power to detect an OR of 0.45 or less (α = 0.05; two-sided test); this OR is in the region of the expected effect sizes for most previous studies. It should be borne in mind that a relatively large number of statistical tests were conducted, and a proportion of positive results would be expected by chance.

Ethnic group was not assessed, but only 3 % of subjects were born outside the UK or Ireland (78 % born in Scotland; 17 % elsewhere in the UK or Ireland). In the 2001 census for Scotland, 98 % of the population described themselves as ‘white’ (Scottish, British, Irish or other) and 2 % as southeast Asian, Chinese or black. Thus it seems unlikely that the results were adversely affected by population stratification, which needs to be quite extreme to have a major impact.
higher for females than for males for the same intake; we are not, however, aware of any evidence confirming this. There may also be a differential effect by sex on folate status of other factors influencing dietary methyl content. For example, men are more likely to drink alcohol, and to consume greater quantities of red meat than women; in the present study, 74% of females were more likely to drink alcohol, and to consume greater quantities of red meat than males.23. Alternatively, folate intake estimates among women may be subject to less measurement error. In the validation of the FFQ used in the present study, for most nutrients the agreement between questionnaire estimates and weighed records was higher for females than males.23.

The median dietary folate intake among controls (295 μg/d) was close to population estimates (average intake from food and drink 274 μg/d). In the USA, by contrast, a large proportion of the population meets the recommended intake of 400 μg/d.22. The proportion of controls taking folic acid-containing supplements (<10%) is typical of the UK population, but considerably lower than in the USA.48,54. The range of total intake in the present study was relatively narrow; cut-off points for the lowest and highest quartiles were 264 and 349 μg/d. This contrasts with several previous studies in which the cut-off point for the highest quantile was > 600 μg/d.22,55–57. Moreover, in most previous studies, including those where average intake was lower than in the USA,58,59, there was greater variability in intake than in the present study. It is noteworthy that in the prospective study in Sweden which reported a bell-shaped relationship between plasma folate and colorectal cancer,39 – similar to our finding in Sweden which reported a bell-shaped relationship between plasma folate and colorectal cancer,39 similar to our finding in the present study. It is noteworthy that in the prospective study in Sweden which reported a bell-shaped relationship between plasma folate and colorectal cancer,39– similar to our finding in the present study, it is noteworthy that in the prospective study in Sweden which reported a bell-shaped relationship between plasma folate and colorectal cancer39 – similar to our finding in the present study, it is noteworthy that in the prospective study in Sweden which reported a bell-shaped relationship between plasma folate and colorectal cancer39 – similar to our finding in the present study.
High folate intake was confined to those with low methionine intake, and we found no interaction between folate and protein (as a marker for methionine). There is tension in the previous evidence, however, in that the pattern of the folate–methionine interaction is not consistent with what would be expected based on the low-methyl diet hypothesis: according to this, the combination of low intakes of folate and methionine together with high alcohol should increase disease risk. The modest raised risk we observed for intermediate and low dietary methyl content is broadly consistent with previous studies. The extent to which the effect of the ‘methyl diet’ on disease risk exceeds that of folate alone remains unclear, however.

Effects of genotype

In most previous studies, carriers of three variant alleles, if they have occurred at all, have been rare. Our frequencies of 2.8% among controls and 1.7% among cases are quite high but are, of course, based on very small numbers of individuals. The high frequency, which appears to occur mainly in UK and Canadian populations, may be due to a founder effect.

The moderate risk reduction, and dose–response, associated with the 677TT genotype, although not statistically significant, is compatible with previous studies. As regards A1298C, eight of eleven studies found lower risk for the variant allele, although not all reached significance and an effect was not seen in all subgroups. Our modest, non-significant, reduced risk in CC subjects is consistent with this.

In common with previous investigators, we found a significant interaction between folate and C677T. However, the pattern of interaction differed from that reported previously in part, to differences in the folate main effect. In addition, for reasons of statistical power, we combined CT and TT genotypes in the analysis. This would have concealed the low risk in TT subjects with high folate (or dietary methyl content) that has been reported elsewhere.

Although we observed a significant folate–A1298C interaction, previous studies have been inconsistent. There is linkage disequilibrium between C677T and A1298C. In the present study, sixty of the sixty-eight individuals with the 677TT genotype also carried 1298AA, while seventy-nine of the eighty-seven who were 1298CC also carried 677CC. This meant that there was considerable overlap between the sub-groups at lowest risk in the two interaction analyses (C677T: low folate and CT/TT genotype; A1298C: low folate and AA genotype), which explains, in part, the different patterns of interaction observed for the two polymorphisms. More generally, differences in patterns of linkage disequilibrium between populations may contribute to heterogeneity between studies.

The less consistent findings for A1298C, than C677T, regarding colorectal cancer could also be because A1298C appears to have less strong functional effects, although it should be noted that the functional evidence is limited and the in vitro and in vivo impact of the variant allele is not fully resolved. In analyses of 199 controls from the present study, plasma and erythrocyte folate (measured by the micro-biological assay) levels were significantly lower, and plasma homocysteine levels significantly higher, in 677TT than CC individuals, while the A1298C variant was not associated with folate or homocysteine levels. Neither polymorphism was related to levels of various biomarkers of DNA stability measured in lymphocytes, including strand breaks, misincorporated uracil or global methylation. The evidence from other similar studies on MTHFR and DNA stability is inconsistent. One reason for this may be that these studies, and those of intake, are not assessing folate status in the relevant tissue. What may be important is localised (rather than systemic) folate depletion acting in concert with C677T. Current knowledge on folate status in colonic tissue is limited. For example, the extent to which colonic folate concentrations vary between individuals, or along the colon within a single individual, or correlate with levels in lymphocytes or dietary intakes, is not clear.

The significant interactions between vitamin B6 and MTHFR genotype followed a similar pattern and had risk estimates similar to those for folate. This may reflect the fact that dietary sources for the two nutrients overlap substantially; after adjusting for total energy intake the correlation between folate and vitamin B6 exceeded 0.70.

Some investigators have reported an alcohol–C677T interaction, such that higher alcohol intake abolished the reduced disease risk associated with the TT genotype. Previous studies have suggested this modification, but were not significant. The median alcohol intake among consumers was lower than the ‘high’ intake groups in previous studies (present study, 3.9 g/d; about 0.5 units/d; Chen et al., five drinks/week; Ma et al., one drink/d).

Most previous studies of folate and folate–MTHFR interactions have been conducted in populations where intake is relatively high, and a substantial proportion comes in the form of folic acid. The results of the present study, undertaken in a population with relatively low folate intake, suggest that the role of folate metabolism in colorectal cancer aetiology may be even more complex than previously thought. The challenge now is to unravel this complexity. 5,10-MTHFR irreversibly converts 5,10-methylene-tetrahydrofolate to 5-methyl-tetrahydrofolate, the former vital for DNA synthesis and the latter key in DNA methylation. Thus, the C677T polymorphism alters the distribution of erythrocyte folates, suggesting that particular folate species might be important in colorectal cancer. The direction of the MTHFR–colorectal cancer association – which is the opposite of that expected a priori – suggests that the availability of tetrahydrofolate and 5,10-methylene-tetrahydrofolate may be particularly pertinent. This is supported by a study in colorectal tumour tissue, which found that 677TT patients had significantly lower concentrations of tetrahydrofolate and 5,10-methylene-tetrahydrofolate. Expansion of the evidence on individual folate species with regard to: (1) distribution in the diet, blood and other bodily tissues in different populations; (2) factors that influence intracellular distribution and availability; and (3) disease risk, might help advance understanding of folate metabolism in colorectal neoplasia and, ultimately, shed light on carcinogenesis pathways.

Acknowledgements

The present study was funded by a grant from the National Hospitals Trust. We thank Janet Kyle, Margaret Smith, Martin Barnes and Jennifer Grzybowski for assistance with
data entry and David Grubb for processing the FFQ. Graeme McHardy and Diane Thom kindly provided downloads of cases and controls. Rachel Melvin, Joanne Tomany and Mhairi Gilmour assisted with compiling data on the content of dietary supplements.

References

2. Ames BN (2001) DNA damage from micronutrient deficiencies is likely to be a major cause of cancer. Mutat Res 475, 7–20.


