understanding the reasons for these delays. Of these, stigma associated with mental illness is one such potential influence.

This makes help-seeking one important outcome in stigma research. All the more, as early help-seeking might prevent the development of severe social disability and the need for intensive treatment that were both reported to abet stigmatization and discrimination. Our principal finding was that the connection between stigma and active help-seeking varies according to type of stigma and that, in particular, a person’s own attitudes towards mental health help-seeking and towards persons with mental illness are associated with their help-seeking behaviour. Thus, we suggested that these finding might be considered in planning, evaluating and implementing campaigns with the aim of promoting help-seeking.

We certainly admire the work of Evans-Lacko, Thornicroft and colleagues in stigma research. We especially acknowledge that population-level anti-stigma programmes are helpful in reducing negative and stigmatising attitudes and enhancing the public’s knowledge about mental illness. We also acknowledge that different stigma types (such as perceived public attitudes and self-stigma) are unrelated and can influence each other (see the Discussion section of our study). It is important to point out that we understand personal attitudes towards persons with a mental illness as a part of the broader term of public stigma (see the introductory section). With this, we hope to have clarified some initial misperceptions of Evans-Lacko and colleagues.

Second, the authors correctly point out that, on the basis of ROC curves, risk assessment tools performed no better than clinician ratings. The other way of looking at this, however, is that clinician ratings performed no better than risk scales. In particular, the Manchester Self-Harm Rule, a 4-item tool, performed just as well. Importantly, the authors found no evidence of between-hospital heterogeneity for this tool’s performance. Clinician ratings, on the other hand, showed substantial heterogeneity between hospitals, with specificity ranging from 58% to 82%. The lack of variability in the actuarial tools could be argued to be an advantage when performance between clinician rating and assessment tool is no different. Furthermore, tools like this will be considerably quicker, leaving more clinician time for risk management (as opposed to assessment).

Third, the clinicians were based in teaching hospitals (Brighton, Bristol, Derby, Manchester and Oxford) with long-standing research interests in self-harm. Whether the reported predictive accuracy of clinician ratings is generalisable to non-specialist centres is an empirical question.

Fourth, the patient rating may also have been influenced by the questions asked by the tools (which tend to be categorical and therefore it is easy to work out what constitutes a risk factor). In a sense, then, the patient rating is a form of structured judgement.

Comparing risk tools with clinicians may not be informative, or even feasible, as clinical interviews already include many of the items used in risk tools. Instead, future research should compare actuarial scores with or without additional clinician input. In other words, if clinicians disagree with the risk level provided by actuarial tools, does this reclassification lead to an improvement in predictive performance? As the AUCs for the tools in this study ranged from 0.55 to 0.72, there may be considerable room for improvement by incorporating novel and modifiable risk factors, as has been shown in violence risk assessment in patients with severe mental illness. Ultimately, randomised studies will be required to establish the effects of different approaches to risk assessment on patient and service outcomes.

Suicide risk assessment tools do not perform worse than clinical judgement

The study by Quinlivan and colleagues could be interpreted to suggest that clinician and patient ratings are better than actuarial tools in predicting self-harm after an emergency hospital presentation with self-harm. However, we would argue that this is an incorrect interpretation.

First, the clinical evaluation appears to have occurred after these tools were completed by the same clinician, and, although the clinicians were masked to the overall score, the clinical impression will therefore have been strongly informed by the items in these suicide risk assessment tools. In fact, the study does not appear to be a comparison between actuarial tools and a distinct, unstructured clinical judgement, but a comparison between actuarial tools and what is called a structured clinical judgement approach (where structured questions about relevant risk factors are asked, and then a clinical judgement is made about an individual’s overall risk level). Clarification of the exact procedure used is important for interpreting the findings.

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Authors’ reply. We thank Professor Fazel and Dr Wolf for their thoughtful letter and interest in our article. Well, at least we agree on one thing: randomised controlled trials could help to clarify the role of risk assessment scales in the management of people who have harmed themselves. However, we do not think that future research should investigate scores ‘with or without additional clinician input’. Risk should not simply be a score or a colour on a traffic light system. In isolation, such ratings may be worse than useless, especially if they distract clinicians from engaging with their patients. An alternative design might be to investigate what risk scales add to ‘assessment as usual’.

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Just to be clear, we are not suggesting that global clinician or patient judgement should be used in preference to rating scales to predict future self-harm. None of the measures is fit for this purpose in our view. Global risk assessments by clinicians have been found to perform poorly in previous studies. With colleagues in Australia some of us are currently engaged in a systematic review of this very issue. We think that Fazel & Wolf’s observation that there was less variability in the scores on the scales than in the clinician ratings is an interesting one. The problem was that in our study the scales actually performed, on average, a bit worse than the global measures. There may be circumstances when scales are useful, of course. For example, as an aide memoire for new staff or as measures of change.

How might we explain the fact that risk scales performed even less well than unvalidated single-item clinician and patient measures? The obvious explanation is that the risk scales themselves were not very good. But there are other possibilities too. For the clinician ratings, we acknowledged in our discussion that the centres all had a special interest in self-harm and that predictive performance might be different in other hospitals. We also mentioned that clinicians might subliminally have used items on the scales to derive their global assessments. Fazel & Wolf’s idea that patients were also taking such factors into account as the assessment progressed is intriguing and certainly worth exploring in future studies. It could be that we gain useful new insights by asking our patients how they make their own judgements of risk.

There have been a number of large systematic reviews recently all pointing in the same direction. In their editorial on our paper, Owens & Kelley highlighted the poor to worthless performance of all measures. They also asked whether risk assessment scales were on their way out. Over to readers of the BJPsych . . .


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Correction

Association between mental health-related stigma and active help-seeking: systematic review and meta-analysis. BJPsych, 210, 251-268. The second sentence of Conclusions in the summary (p. 261) should read:

‘Campaigns promoting help-seeking by means of fighting mental illness-related stigma should target these personal attitudes rather than broad public opinions.’

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