Association between inflammatory potential of diet and risk of depression in middle-aged women: the Australian Longitudinal Study on Women’s Health

Nitin Shivappa1,2*, Danielle A. J. M. Schoenaker3, James R. Hebert1,2 and Gita D. Mishra3

1Cancer Prevention and Control Program, University of South Carolina, Columbia, SC 29208, USA
2Department of Epidemiology and Biostatistics, Arnold School of Public Health, University of South Carolina, Columbia, SC 29208, USA
3School of Public Health, University of Queensland, Brisbane, QLD 4006, Australia

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Abstract
Dietary factors and inflammation markers have been shown to play a role in the development of depression. However, there are very few studies that have explored the association between inflammatory potential of diet and risk of depression. In this study, we examined the association between the dietary inflammatory index (DII), which was developed specifically to measure the inflammatory potential of diet, and risk of depression in the middle-aged cohort of the Australian Longitudinal Study on Women’s Health. A total of 6438 women with a mean age of 52.0 (SD 1.4) years at baseline were followed-up at five surveys over 12 years (2001–2013). Depression was defined as a score of ≥10 on the Center for Epidemiologic Studies Depression-10 scale. The DII score, a literature-derived, population-based dietary index that has been validated against several inflammatory markers, was computed on the basis of dietary intake assessed using a validated FFQ. Generalised estimating equations were used to estimate relative risk (RR) of depression according to DII score. Models were adjusted for energy intake, highest education completed, marital status, menopause status and symptoms, personal illness or injury, smoking status, physical activity, BMI and depression diagnosis or treatment. In total, 1156 women (18%) had scores ≥10 on the CESD scale over the course of 9 years. Women with the most anti-inflammatory diet had an approximately 20% lower risk of developing depression compared with women with the most pro-inflammatory diet (RRDII quartile 1 vs. 4: 0.81; 95% CI 0.69, 0.96; \( P \) trend = 0.03). These results suggest that an anti-inflammatory diet is associated with lower risk of depression in middle-aged Australian women.

Key words: Dietary inflammatory index; Diet; Inflammation; Depression; Australia

Unipolar depression affects more than 151 million people worldwide, and it is expected to become the world’s second most common disease, after CVD, by 2020(1). Results from the Global Burden of Disease Study indicate that mental health and substance use disorders are the leading global cause for all non-fatal conditions contributing to the burden of disease, as measured by years lived with disability, with depressive and substance use disorders being the leading global cause for all the prevalence and the incidence of depressive symptoms, associated with lower concentrations of intermediary inflammatory markers(12).

Various studies have been conducted to evaluate dietary exposures in relation to depression(15–17). In general, prospective cohort studies have shown that dietary patterns rich in components such as fruits, vegetables, olive oil and legumes may be protective against depression(18,19). Previous results from the Australian Longitudinal Study on Women’s Health (ALSWH) study showed that increased consumption of fruits reduced both the prevalence and the incidence of depressive symptoms, although no association was observed with vegetable intake(17). In contrast, increased risk has been observed with diets rich in SFA, n-6 fatty acids and refined carbohydrates(20,21). However, a few studies have explored the association between inflammatory potential of diet and depression.

A literature-derived, population-based dietary inflammatory index (DII) was developed to assess the inflammatory potential of an individual’s diet(22). Previously, higher DII scores have

Abbreviations: ALSWH, Australian Longitudinal Study on Women’s Health; CESD-10, Center for Epidemiologic Studies Depression.

* Corresponding author: Dr N. Shivappa, email shivappa@mailbox.sc.edu
been shown to be associated with increased risk of depression in a Spanish study(22). Our aim in this study was to assess the association between the inflammatory potential of dietary patterns of middle-aged Australian women measured through DII scores and the risk of depression.

**Methods**

**Study design and participants**

ALSWH is an ongoing, population-based cohort study examining factors associated with health and well-being of Australian women. Full details on study design, recruitment methods and response rates have been published previously(24,25). Women were randomly selected from the National Medicare Health Insurance database, including all Australian citizens and permanent residents. Women from rural and remote areas were intentionally oversampled. Informed consent was obtained from all participants at each survey, with ethics clearance obtained from the Human Research Ethics Committees of the University of Newcastle and the University of Queensland, Australia.

For this study, we used data from the middle-aged cohort of women born between 1946 and 1951. A total of 13 715 women aged 45–50 years responded to the first survey in 1996. This sample was found to be broadly representative of the general Australian population of women of similar age based on the national census(26). Self-administered questionnaires were sent to participants every 2–3 years, until 2013 (survey 7). Response rates for surveys 2, 3, 4, 5, 6 and 7 were 90.0% (n 12 338), 81.8% (n 11 221), 79.5% (n 10 905), 77.6% (n 10 638), 73.0% (n 10 011) and 66.7% (n 9 151) of baseline, respectively. Information on dietary intake was first collected at survey 3 in 2001, when women were aged 50–55 years; this survey was therefore used as baseline for this study. Women were excluded if they reported a history of depressive symptoms (Center for Epidemiologic Studies Depression (CESD-10) score ≥ 10) from survey 1 (1996) until survey 3 (2001) (n 3459), had missing data on baseline dietary intake or depressive symptoms during follow-up (n 695) or had missing data on relevant confounders (n 634). Our analyses included 6438 women (online Supplementary Fig. S1).

**Dietary assessment**

Diet was assessed using the Dietary Questionnaire for Epidemiologic Studies (DQES) version 2. The development of this 101-item FFQ has been described elsewhere(20). The FFQ was validated against 7-d weighed food records in sixty-three women with a mean age of 33-3 (so 9-5) years. Mean intakes from the weighed food records and FFQ were within ±20% for twenty-one of twenty-seven nutrients(27). Energy-adjusted correlation coefficients for nutrient intake ranged between 0.28 for vitamin A and 0.78 for carbohydrates, indicating that the FFQ performed reasonably well for assessing habitual intake. Participants were asked to report their usual frequency of consumption of seventy-four food items and six alcoholic beverages over the previous 12 months using a ten-point scale ranging from ‘never’ to ‘three or more times per day’. Portion size photographs were used to adjust the serving sizes for vegetables, meat and casseroles. The FFQ included twenty-one additional items on the number of servings of milk, bread, sugar and eggs and the type of milk, bread, fat spreads and cheese consumed. Consumption of 101 food items (g/d) and nutrient intakes were computed from the national government food composition database of Australian foods – the NUTTAB95(28).

**The dietary inflammatory index**

The development and validation of the DII are described in detail elsewhere(22,29). Developing the DII involved reviewing and scoring nearly 2000 scientific articles representing cell culture and laboratory animal experiments, as well as a variety of human studies on diet and six inflammatory markers (i.e., C-reactive protein, IL-1β, IL-4, IL-6, IL-10, TNF-α). One of three possible values was assigned to each article on the basis of the effect of the particular food or constituent on inflammation: +1 if pro-inflammatory, −1 if anti-inflammatory and 0 if it produced no change in the inflammatory marker. In addition, these articles were weighed on the basis of the study design, with experimental cell studies receiving the lowest weight of 3 and human trials the highest weight of 10(22,29). These procedures also entailed the creation of a world standard database that involved obtaining eleven data sets from around the world (Japan, Korea, USA, Canada, Mexico, Taiwan, Australia, New Zealand, Bahrain, Denmark and India) to which individuals’ intakes of forty-five food parameters (consisting of nutrients, spices and whole foods) on which the DII is based could then be compared.

The steps involved in computing DII scores are described in Fig. 1. To summarise, FFQ-derived dietary data were used to calculate DII scores for all participants. First, they were linked to the world database described previously(22). These then became the multipliers to express an individual’s exposure relative to the ‘standard global mean’ as a z-score, which was computed by subtracting the ‘standard global mean’ from the amount reported and dividing this value by the ‘global standard deviation’. To minimise the effect of ‘right skewing’, this value was then converted to a centred percentile score(22) obtained by doubling and subtracting 1. Food parameter-specific DII scores were obtained by multiplying this score(22) for all of the food parameter-specific DII scores were then summed to create the overall DII score for each participant in the study.

For the current study, data on twenty-six of the forty-five DII food parameters could be derived from the FFQ, and were thus used for DII calculation. These included pro-inflammatory components (energy, carbohydrate, protein, fat, SFA, Fe, cholesterol) and anti-inflammatory components (alcohol, fibre, MUFA, PUFA, n-3, n-6, niacin, thiamin, riboflavin, Mg, Zn, vitamin A, vitamin C, vitamin E, folic acid, β-carotene, garlic and onions). The DII score can range from the maximum pro-inflammatory score of +7-98 to the minimum anti-inflammatory score of −8.87(22). Positive DII scores indicate a pro-inflammatory diet and negative scores an anti-inflammatory diet.
The score for each food parameter was calculated giving:
+1 to each article if the effects were pro-inflammatory (significantly increased IL-1\beta, IL-6, TNF-\alpha or CRP, or decreased IL-4 or IL-10),
-1 if the effects were anti-inflammatory (significantly decreased IL-1\beta, IL-6, TNF-\alpha or CRP, or increased IL-4 or IL-10),
0 if the food parameter did not produce any significant change in the inflammatory marker.

The score for each food parameter was weighted according to the study design. The weights were 10 (experimental design), 8 (observational), 7 (case–control), 6 (cross–sectional), 5 (experimental with animal), 3 (cell culture).

A food parameter-specific overall inflammatory effect score was calculated by subtracting the anti-inflammatory fraction from the pro-inflammatory fraction. This score was corrected if the total weighted number of article was < 236. In these cases the raw overall inflammatory score was multiplied by the total weighted number of articles divided by 236.

Z-score and centred-percentiles for each of the 26 available food parameters for each participant were calculated on the basis of the average and standard deviation for each food parameter obtained from the global database which was created from the consumption of the original 45 food parameters from 11 countries.

The centred percentile for each food parameter was multiplied by the respective ‘overall food parameter-specific inflammatory effect score’ to obtain the ‘food parameter-specific DII score’.

All of the ‘food parameter-specific DII scores’ were summed to create the ‘overall DII score’ for each individual.

**Fig. 1.** Sequence of steps in creating the dietary inflammatory index (DII) in the Australian Longitudinal Study on Women’s Health. *The missing food items were anthocyanin, eugenol, flavan-3-ols, flavones, flavonols, flavonones, isoflavones, caffeine, ginger, saffron, turmeric, vitamin D, pepper, thyme and oregano, rosemary and Se. CRP, C-reactive protein.
The DII method has been validated by comparing the DII scores with various inflammatory markers\(^{20-22}\). In addition, the DII score has been shown to be associated with glucose intolerance and dyslipidaemia components of the metabolic syndrome\(^{33,34}\), anthropometric measurements in Spain\(^{35}\), asthma in Australia\(^{32}\), bone mineral density among post-menopausal women in Iran\(^{36}\), two colorectal cancer case-control studies in Spain and Italy\(^{37,38}\) and in a cohort study of women in the USA\(^{39,39}\) and pancreatic and prostate cancers in Italy\(^{40,41}\).

**Depressive symptoms**

Depressive symptoms were measured using the CESD-10 scale. This scale is constructed on the basis of self-reported responses to ten questions on whether, over the past week, women experienced symptoms associated with depression, such as worrying, sleeping difficulty and difficulty relaxing. The CESD-10 has a good test–re-test reliability and predictive validity when compared with the original twenty-item version\(^{42}\), and has high sensitivity (97–100%) and specificity (84–93%) for the screening of major depression in middle-aged and older adults\(^{43}\).

Responses to the ten questions were reported using a four-point Likert scale ranging from ‘rarely’ (scored 0) to ‘most of the time’ (scored 3). Possible scores range from 0 to 30, with higher scores indicating more depressive symptoms. A CESD-10 score of ≥10 was used to define the presence of depressive symptoms. Validation of the CESD-10 against the twenty-item version using a cut-off of ≥10 has been shown to minimise false-positive results with little loss of sensitivity\(^{42}\). Incident depression was defined as new onset of depressive symptoms between survey 4 (2004) and survey 7 (2013).

**Assessment of covariates**

Self-reported information was collected at each survey on a range of socio-demographic variables, symptoms, lifestyle factors and health behaviours. Responses were categorised for the highest qualification obtained as ‘no formal’, (‘higher’ school certificate’, ‘trade/diploma’ or ‘university/higher degree’); for marital status as ‘single/never married’, ‘married or de facto’ or ‘separated/divorced/widowed’; for night sweats as ‘no’ (never) or ‘yes’ (rarely/sometimes/often); for smoking as ‘non-smoker’, ‘ex-smoker’ or ‘current smoker’; and for experience of stressful life events or diagnosis or treatment of depression since the previous survey as ‘no’ or ‘yes’.

BMI was computed as self-reported weight (kg) divided by the square of self-reported height (m) and categorised as ‘normal weight’ (BMI < 25 kg/m²), ‘overweight’ (25 ≤ BMI < 30 kg/m²) or ‘obese’ (BMI ≥ 30 kg/m²). As only a limited proportion of women (1-5%) were classified as ‘underweight’ (BMI < 18.5 kg/m²), they were combined with the ‘normal weight’ group. Self-reported weight and height were validated in a subsample of 159 women (age 54–59 years). Both self-reported weight and height tended to be underestimated compared with measured values, with a mean difference of 0.95 (95% CI 0.44, 1.47) kg and 0.67 (95% CI 0.26, 1.08) cm, respectively. Agreement between BMI categories derived from self-reported and measured data was 84%\(^{44}\).

Physical activity scores were derived from validated questions on frequency and duration of walking (for recreation or transport) and from questions on moderate- and vigorous-intensity activity in the last week and categorised as ‘sedentary/low’ (<600 metabolic equivalent (MET)-min/week), ‘moderate’ (600 ≤ MET < 1200 min/week) or ‘high’ (≥1200 MET-min/week)\(^{45}\). Menopause status was determined using questions on hysterectomy, oophorectomy, hormone therapy and menstrual pattern and categorised as ‘hysterectomy and/or oophorectomy’, ‘hormone therapy use’, ‘premenopausal’, ‘perimenopausal’ or ‘postmenopausal’\(^{46}\).

**Statistical analysis**

Baseline characteristics were described for women according to reporting of depressive symptoms at least once during follow-up and according to quartiles of the DII, and were compared using t tests, ANOVA or χ² tests. All descriptive statistics were weighted by area of residence to account for oversampling of women from rural and remote areas.

Generalised estimating equations were used to examine the prospective associations between the DII score (computed based on survey 3 data) and the incidence of depressive symptoms (survey 4–7). Log-binomial models or log-Poisson models (when log-binomial models did not converge) were used to estimate relative risks (RR) and 95% CI. Models were adjusted for the time-varying confounders described above (survey 3–6). Model 1 adjusted for total energy intake, highest qualification completed, marital status, menopause status, hot flushes and experience of major personal illness or injury; model 2 additionally adjusted for lifestyle factors: smoking status, physical activity and BMI; and model 3 adjusted for depression diagnosis or treatment.

Factors including new onset of hot flushes, occupation, time spent socialising, performing active (walking or exercise) and passive (watching TV, reading, relaxing) leisure activities, and stressful life events including break-up of relationships or divorce and death of a close family member or friend were considered as potential confounders but not included in the final models as adjustments did not substantially change the effect estimates.

To test the robustness of our findings, additional analyses were performed. First, characteristics were compared with women included in the analysis and excluded because of missing data on confounders. Second, a multiple imputation analysis was performed to assess the influence of participant exclusions that resulted from missing data\(^{47}\). Third, in addition to examining the association between DII score and risk of depression in women with no history of depression before survey 3 based on a CESD-10 score of ≥10, we examined the association in women with no history of doctor-diagnosed or treated depression. All statistical tests were conducted using a two-sided 5% level of significance.

**Results**

The mean DII score in this study was 1.11 (sd 1.54), which indicated a slightly pro-inflammatory diet. Table 1 describes
the distribution of several characteristics according to depression status. Women who developed depression during follow-up were more likely to have lower education, be obese, be sedentary/have low physical activity, be a current smoker, to have hot flushes, to have experienced major personal illness or injury and be less likely to be postmenopausal. Table 2 describes the distribution of characteristics across quartiles of DII scores. Women in the most pro-inflammatory group (quartile 4) were more likely to have lower education, be separated/divorced/widowed, be sedentary/have low physical activity and be current smokers.

The distribution of consumption of various nutrients and food items per day across quartiles of DII is shown in the online Supplementary Table S1. Most of the anti-inflammatory dietary components such as PUFA, dietary fibre, Zn, fruits, vegetables, fish and high-fibre grain foods decreased linearly across quartiles of DII. Similar trends were observed for intakes of total energy, total meat, and red and processed meat. Pro-inflammatory components whose intakes were higher among women in the fourth quartile were SFA, cholesterol, Ca, high-fat dairy products and added sugar.

Compared with the most pro-inflammatory group (4th quartile), the risk of developing depression was 26% lower in the most anti-inflammatory group (1st quartile) (RRquartile 1 vs 4 = 0.74; 95% CI 0.61, 0.89, P trend = 0.001) for model 1, 21% (RRquartile 1 vs 4 = 0.79; 95% CI 0.66, 0.95, P trend = 0.03) for model 2 and 19% (RRquartile 1 vs 4 = 0.81; 95% CI 0.69, 0.96, P trend = 0.03) for model 3 (Table 3). The difference in the incidence of depression over 12 years between quartile 4 and quartile 1 of DII was 16%.

Compared with women who were included in the analysis (n 6438), the 10% of women who were excluded because of missing confounding data (n 634) were more likely to have a pro-inflammatory diet, lower education, be separated/divorced/widowed, be postmenopausal (surgical or natural), be obese, smoke tobacco, be sedentary or have low physical activity levels and have a history of major illness or injury. Incidence of depressive symptoms during follow-up did not differ for women who were included in the analysis compared with women who were excluded because of missing data (online Supplementary Table S2). Conclusions on the

| Table 1. Baseline characteristics of middle-aged women according to reporting of depressive symptoms at least once during 12 years of follow-up, Australian Longitudinal Study on Women’s Health, 2001 to 2013* (Mean value and standard deviation; numbers and percentages, weighted by area of residence; n 6438) |
| Baseline characteristics (survey 3) | Depressive symptoms (CESD-10 score ≥10) (n 1156) | No depressive symptoms (CESD-10 score <10) (n 5282) |
| Age (years) | n | % | n | % | P† |
| Mean | 52 | 14 | 52 | 14 | 0.37 |
| Highest qualification completed | | | | | <0.0001 |
| No formal | 178 | 15.4 | 624 | 11.7 | |
| (Higher) school certificate | 574 | 49.6 | 2430 | 45.7 | |
| Trade/diploma | 205 | 17.8 | 1197 | 22.5 | |
| University/higher degree | 200 | 17.3 | 1064 | 20.0 | |
| Marital status | | | | | 0.21 |
| Single/never married | 48 | 4.2 | 178 | 3.4 | |
| Married or de facto | 947 | 81.9 | 4436 | 83.7 | |
| Separated/divorced/widowed | 161 | 14.0 | 687 | 13.0 | |
| Menopause status | | | | | 0.04 |
| Hysterectomy and/or oophorectomy | 310 | 27.0 | 1261 | 23.8 | |
| Hormone therapy use | 225 | 19.6 | 1065 | 20.1 | |
| Premenopausal | 111 | 9.7 | 550 | 10.4 | |
| Perimenopausal | 234 | 20.4 | 1036 | 19.6 | |
| Postmenopausal | 267 | 23.3 | 1376 | 26.0 | |
| BMI | | | | | 0.0001 |
| Normal weight (<25 kg/m²) | 508 | 45.3 | 2441 | 48.2 | |
| Overweight (25 to <30 kg/m²) | 345 | 30.8 | 1698 | 33.6 | |
| Obese (≥30 kg/m²) | 268 | 23.9 | 923 | 18.2 | |
| Physical activity | | | | | 0.0003 |
| Sedentary/low (<600 MET-min/week) | 630 | 55.8 | 2604 | 50.6 | |
| Moderate (600 to <1200 MET-min/week) | 249 | 22.0 | 1132 | 22.0 | |
| High (≥1200 MET-min/week) | 251 | 22.2 | 1408 | 27.4 | |
| Smoking | | | | | <0.0001 |
| Non-smoker | 675 | 58.4 | 3422 | 64.6 | |
| Ex-smoker | 286 | 24.8 | 1296 | 24.5 | |
| Current smoker | 195 | 16.9 | 578 | 10.9 | |
| Night sweats | 406 | 45.1 | 1548 | 39.1 | 0.001 |
| Major personal illness or injury | 115 | 10.0 | 356 | 6.7 | 0.0001 |

† Numbers may vary due to missing baseline values.
* P-values from t test or χ² test. All statistical tests were conducted using a two-sided 5 % level of significance.
The association between DII and depression remained similar in analyses where missing confounding data were imputed (online Supplementary Table S3). Excluding women with a history of doctor-diagnosed or treated depression instead of a history of depressive symptoms based on a CESD-10 score of $\geq 10$ did not alter our findings (data not shown).
Discussion

The present study, the largest prospective investigation on inflammatory potential of diet and depression to date in an Australian population, shows an inverse association between an anti-inflammatory dietary pattern – characterised by consumption of nutrients such as n-3 and n-6 PUFA, dietary fibre and food groups such as fruits, vegetables, meat, red meat, fish and high-fibre grain foods – and a lower risk of depression in a population-based study of middle-aged women.

In our study, most of the nutrient and food group distributions across quartiles of DII were as expected, but the distributions of total meat, processed meat and red meat were contrary to our expectation (levels increased across quartiles of DII). However, it should be noted that red meat is one among several other food items that influence inflammation. There are other food items such as vegetables, fruits and fish that exert strong anti-inflammatory effects per unit exposure and their distributions are as expected. In addition, red meat is rich in nutrients including Zn, vitamin B12 and n-6 fatty acids, all of which have an anti-inflammatory score with DII (48). As the DII takes into account diet as a whole, and red meat eaters also consume vegetable components, spices and other anti-inflammatory foods, overall, their diet was more anti-inflammatory.

These results are consistent with a report from the Nurse’s Health Study, where researchers observed a strong association between an inflammatory dietary pattern and depression (10). Unlike the DII, which is empirically linked to findings in the literature, the inflammatory dietary pattern in that study was identified using reduced-rank regression, which in turn was related to plasma levels of inflammatory markers. To date, there has been only one other study, conducted in Spain, which has examined the association between the inflammatory potential of diet as determined by DII scores and depression. In that study, the hazards ratio for participants in the highest quintile of the DII (strongly pro-inflammatory) was 1.47 (95% CI: 1.17, 1.85) compared with those in the lowest quintile, with a significant dose-response relationship ($P_{\text{for trend}} = 0.01$) (25).

A majority of our participants were well educated, with a high percentage of women in all DII quartiles having at least high school certificate education. It is also evident from the distribution of demographic data across quartiles of DII that women in quartile 4 were more likely to be separated/divorced/widowed, report low levels of physical activity and currently smoking. We also observed decreased energy intake among women in quartile 4 compared with women in quartile 1. This is consistent with the fact that observations across all of the many studies we have conducted to date using the DII reveal two countervailing effects (39, 49, 50). The first is a tendency to eat more of everything as one increases energy intake; this results in a positive correlation between energy intake and nutrient intake. The other is the ‘healthy eater’ effect (e.g., due to the intention of careful, health-conscious people to choose nutrient-dense, energy-sparse foods). This results in negative correlations between energy density and nutrient density. Thus far, we have observed that there is considerable variability in how these eating behaviours are distributed across different populations. In this population, it appears that the former effect of eating more of everything as one increases energy intake is stronger; hence, the mean energy intake is low in higher quartiles, which, in turn, means that the intakes of nutrients that are considered anti-inflammatory, such as vitamins and minerals, are also low.

Previously, in the ALSWH, we showed that increasing fruit consumption and adherence to the Mediterranean dietary pattern (garlic, peppers, mushrooms, salad greens, pasta and red wine) are protective against depression (77, 58). In a recent publication from the ALSWH, no association was observed between Australian Recommended Food Score and Mediterranean dietary pattern and depression, after adjusting for time-varying covariates (51). In addition, pooled results from two Australian cohorts showed people with lower intakes of Zn to be at higher risk of developing depression (52).

Furthermore, a review of the literature on the role of n-3 fatty acids and depression supports an inverse association between the two (45). In a large cohort study conducted in the USA, frequent consumption of sweetened beverages, especially diet drinks, was found to increase the risk of depression among older adults, whereas coffee consumption lowered the risk (153). Various studies have been conducted examining the association between dietary patterns and depression. A systematic review conducted on twenty-one studies indicated that high intakes of fruit, vegetables, fish and wholegrains may be associated with a reduced depression risk (53). In analyses of a cross-sectional study conducted in the USA, which explored the association between various dietary patterns and depression, a healthy dietary pattern characterised by increased fruit and vegetable intakes was inversely associated with depression; increased risk was observed with the sweet dietary pattern characterised by increased intake of cookies, ice cream and chocolates. However, in the same study, no increased risk was observed with an unhealthy pattern characterised by increased intake of fried potatoes, cheese, processed meat and red meat (54). In a US cohort study conducted exclusively among women, no clear association was observed between any dietary patterns (prudent and Western) derived from factor analysis and depression (55). Several studies have shown dietary factors and dietary patterns not to be associated with depression – for example, in a cross-sectional study in Japan, no statistically significant linear association was observed for the intake of riboflavin, pyridoxine, cobalamin, total n-3 PUFA, α-linolenic acid, EPA or DHA (56). Nutrients such as n-3 and Zn and food items such as fruits and vegetables, which are rich in components such as flavonoids and vitamins, and garlic, which are anti-inflammatory, produce more negative DII scores (22).

Several previous studies have established the link between inflammation and depression (5, 6, 57, 58). Possible mechanisms for this observed association could be through the process of hyperleptinaemia or insulin resistance, which are observed in conditions such as obesity, the metabolic syndrome and type 2 diabetes mellitus and have been linked to inflammation processes, which are also common in depressive disorders (59). Similarly, higher levels of cortisol have been observed among obese people, which in turn has resulted in higher hypothalamic–pituitary–adrenal axis reactivity to psychological stress as well as physiological and pharmacological stimulation (60). Indeed, the
pro-inflammatory dietary components could induce sensitisation with increased activation of the inflammatory response system. It has been shown that in people with depression, inflammation is increased and brain-derived neurotrophic factor is reduced.

The strengths of our study include its relatively large sample size and its use of a validated and tested CESD scale to measure depressive symptoms. The prospective nature of the study minimises potential problems resulting from reverse causation.

Limitations include the loss to follow-up (approximately 10%), which may have introduced selection bias. Results from multiple imputations to assess the impact of missing data from confounding variables did not alter the conclusions on the association between DII and depression. Another limitation of our study is that detailed dietary information was collected just once, and thus if changes in diet occurred from survey 4 to 6 over the 6-year period they could not be accounted for in this analysis. However, previous studies have reported that dietary pattern classification is moderately stable over long periods of time during adulthood. For the current study, data were available on only twenty-six of the forty-five food parameters, and therefore missing information on the remaining nineteen food parameters could be a limitation. The missing food items were anthocyanin, eugenol, flavan-3-ols, flavones, flavonols, flavonones, isoflavones, caffeine, ginger, saffron, turmeric, vitamin D, pepper, thyme and oregano, rosemary and Se. Several of these items are presumably not consumed frequently in the Australian population. Other food items such as flavonoids are more frequently consumed, and therefore might have influenced the DII scores computed in our study. Although we could not test and confirm this result in our cohort, other studies have observed relatively strong correlations between inflammatory biomarkers and DII when it was computed with fewer than thirty food parameters. Another limitation could be the absence of evidence of DII scores being associated with inflammatory markers in this study, as no data on inflammatory markers were available. However, the DII has been validated previously with a variety of inflammatory markers in various studies. Although the validity of the FFQ has been evaluated, some degree of misclassification may exist in the dietary assessment. However, owing to the prospective nature of the study, the misclassification is unlikely to be differential. As such, the results probably would be biased towards the null. Another limitation of the FFQ is that it was validated on a small group of women of younger age than those included in the present study. Moreover, some of the results obtained could be biased because of residual confounding from measurement errors with factors such as education, marital status, physical activity and smoking status or unmeasured factors including psychosocial factors that might also be associated with depression (e.g., acquiescence). We think that despite the small difference between quartile 4 and quartile 1, our findings may be relevant at a population level.

In conclusion, anti-inflammatory diet as shown by lower DII scores was associated with a lower risk of developing depression in women from a population-based cohort study of middle-aged Australian women. Therefore, promoting diets with increased consumption of anti-inflammatory foods such as vegetables and fruits could protect against developing depression. However, further studies including both cohort and intervention studies, analysing the link between diet, inflammation and depression, are warranted to elucidate the causal role of diet in the development of depression and other mental disorders.

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G. D. M. designed and conducted the study; N. S. calculated the DII scores and wrote the first draft of the manuscript; D. A. J. M. S. performed analyses and worked on the methods and sections of the manuscript. N. S., D. A. J. M. S., J. R. H. and G. D. M. provided suggestions and revised the manuscript. All the authors approved the final version of the manuscript.

J. R. H. owns controlling interest in Connecting Health Innovations LLC (CHI), a company planning to licence the right to his invention of the DII from the University of South Carolina in order to develop computer and smart phone applications for patient counselling and dietary intervention in clinical settings. N. S. is an employee of CHI.

Supplementary material

For supplementary material/s referred to in this article, please visit http://dx.doi.org/10.1017/S0007114516002853

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