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# APPROXIMATION OF PASSAGE TIMES OF *y*-REFLECTED PROCESSES WITH FBM INPUT

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## Abstract

Define a  $\gamma$ -reflected process  $W_{\gamma}(t) = Y_H(t) - \gamma \inf_{s \in [0,t]} Y_H(s), t \ge 0$ , with input process  $\{Y_H(t), t \ge 0\}$ , which is a fractional Brownian motion with Hurst index  $H \in$ (0, 1) and a negative linear trend. In risk theory  $R_{\gamma}(u) = u - W_{\gamma}(t), t \ge 0$ , is referred to as the risk process with tax payments of a loss-carry-forward type. For various risk processes, numerous results are known for the approximation of the first and last passage times to 0 (ruin times) when the initial reserve u goes to  $\infty$ . In this paper we show that, for the  $\gamma$ -reflected process, the conditional (standardized) first and last passage times are jointly asymptotically Gaussian and completely dependent. An important contribution of this paper is that it links ruin problems with extremes of nonhomogeneous Gaussian random fields defined by  $Y_H$ , which we also investigate.

*Keywords:* Gaussian approximation; passage time;  $\gamma$ -reflected process; workload process; risk process with tax; fractional Brownian motion; Piterbarg constant; Pickands constant

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# 1. Introduction and main result

Let  $\{X_H(t), t \ge 0\}$  be a standard fractional Brownian motion (FBM) with Hurst index  $H \in (0, 1)$ , meaning that  $X_H$  is a centered Gaussian process with almost surely continuous sample paths and covariance function

$$\operatorname{cov}(X_H(t), X_H(s)) = \frac{1}{2}(t^{2H} + s^{2H} - |t - s|^{2H}), \quad t, s \ge 0$$

We will define the  $\gamma$ -reflected process with input process  $Y_H(t) = X_H(t) - ct$  by

$$W_{\gamma}(t) = Y_H(t) - \gamma \inf_{s \in [0,t]} Y_H(s), \qquad t \ge 0,$$
(1.1)

where  $\gamma \in [0, 1]$  and c > 0 are two fixed constants.

Motivations for studying  $W_{\gamma}$  come from both risk and queueing theory. For instance, in queueing theory  $W_1$  is the so-called workload process (or queue length process); see, e.g. Harrison (1985), Asmussen (1987), Zeevi and Glynn (2000), Whitt (2002), and Awad and Glynn (2009). In advanced risk theory the process  $R_{\gamma}(t) = u - W_{\gamma}(t)$ ,  $t \ge 0$ ,  $u \ge 0$ , is referred to as the risk process with tax payments of a loss-carry-forward type; see, e.g. Asmussen and Albrecher (2010). Recently, Hashorva *et al.* (2013) studied the asymptotics of the probability

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 $\mathbb{P}\{\sup_{t \in [0,T]} W_{\gamma}(t) > u\}$  as  $u \to \infty$  for both  $T < \infty$  and  $T = \infty$ . Continuing the study of the aforementioned paper in this contribution we will investigate the approximations of the first and last passage times of  $W_{\gamma}$ . Specifically, define the first and last passage times of  $W_{\gamma}$  to a constant threshold u > 0 by

$$\tau_1(u) = \inf\{t \ge 0, W_{\gamma}(t) > u\}$$
 and  $\tau_2(u) = \sup\{t \ge 0, W_{\gamma}(t) > u\},\$ 

respectively (here we use the convention that  $\inf\{\emptyset\} = \infty$  and  $\sup\{\emptyset\} = 0$ ). Furthermore, define  $\tau_1^*(u)$ ,  $\tau_2^*(u)$ , u > 0, in the same probability space such that

$$(\tau_1^*(u), \tau_2^*(u)) \stackrel{\mathrm{D}}{=} (\tau_1(u), \tau_2(u)) \mid (\tau_1(u) < \infty), \tag{1.2}$$

where  $\stackrel{\text{`}}{=}$  denotes equality in distribution.

The first and last passage times of Gaussian processes conditioned on the event  $\{\tau_1(u) < \infty\}$ are analysed in Hüsler and Piterbarg (2008) and Hüsler and Zhang (2008) when  $\gamma = 0$ . Therein, the Guassian approximations of both  $\tau_1^*(u)$  and  $\tau_2^*(u)$  are derived. First passage times (sometimes called ruin times) are also studied extensively in the framework of insurance risk processes; see the recent articles Griffin and Maller (2012), Griffin (2013), Griffin et al. (2013), and Debicki et al. (2014), and the monographs Embrechts et al. (1997) and Asmussen and Albrecher (2010) for approximations of ruin times of various risk processes. In our framework,  $\tau_1^*(u)$  can be interpreted as the conditional ruin time of the FBM risk process with tax payments of a loss-carry-forward type. With motivation from the aforementioned contributions, this paper is concerned with the Gaussian approximation of the random vector  $(\tau_1^*(u), \tau_2^*(u))$ . For the derivation of the tail asymptotics of  $\sup_{t \in [0,T]} W_{\gamma}(t)$ , Hashorva et al. (2013) showed that the investigation of the supremum of certain nonstationary Gaussian random fields is crucial. One key merit of our problem of approximating the joint distribution function of  $(\tau_1^*(u), \tau_2^*(u))$  is that it leads, as in the case of the analysis of the tail asymptotics of  $\sup_{t \in [0,T]} W_{\gamma}(t)$ , to an interesting unsolved problem of asymptotic theory of Gaussian random fields. Although the latter investigation was not initially in the scope of this paper, the corresponding result derived in Theorem 2.1 is important for various theoretical questions. Next, set

$$A(u) = \frac{H^{H+1/2}}{(1-H)^{H+1/2}c^{H+1}}u^H \text{ and } \tilde{t}_0 = \frac{H}{c(1-H)}$$

and denote by  $\stackrel{\text{D}}{\to}$  and  $\stackrel{\text{P}}{\to}$  the convergences in distribution and in probability, respectively. Furthermore, let  $\mathcal{N}$  be an N(0, 1) random variable. Our principal result is the following theorem.

**Theorem 1.1.** Let the  $\gamma$ -reflected process  $\{W_{\gamma}(t), t \ge 0\}$  be given as in (1.1) with  $\gamma \in (0, 1)$ , and let  $\tau_1^*(u)$  and  $\tau_2^*(u)$  be defined as in (1.2). Then, as  $u \to \infty$ ,

$$\left(\frac{\tau_1^*(u) - \tilde{t}_0 u}{A(u)}, \frac{\tau_2^*(u) - \tilde{t}_0 u}{A(u)}\right) \xrightarrow{\mathrm{D}} (\mathcal{N}, \mathcal{N}).$$
(1.3)

**Remarks.** (a) The joint convergence in (1.3) implies that  $(\tau_2^*(u) - \tau_1^*(u))/A(u) \xrightarrow{\mathbb{P}} 0$  as  $u \to \infty$ .

(b) For any  $u \ge 0$ ,  $\mathbb{P}{\tau_1(u) < \infty} = 1$  when  $\gamma = 1$  (cf. Duncan and Jin (2008)), which is the reason for considering only the case that  $\gamma \in (0, 1)$ . Under the latter assumption on  $\gamma$ , we further

have  $\mathbb{P}\{\tau_2(u) < \infty \mid \tau_1(u) < \infty\} = 1$ , which follows from the fact that  $\lim_{t\to\infty} W_{\gamma}(t) = -\infty$  almost surely since in view of Remark 5 of Kozachenko *et al.* (2014)

$$\lim_{t \to \infty} \frac{\sup_{s \in [0,t]} |X_H(s)|}{t} = 0 \quad \text{for all } H \in (0,1).$$

(c) It is somewhat surprising that the Gaussian approximation of the conditional first and last passage times does not involve the reflection constant  $\gamma$ .

The rest of the paper is organized as follows. In the next section we present a key result on the supremum of some Gaussian random fields defined by  $Y_H$  and then provide the proof of Theorem 1.1. Section 3 is dedicated to the proof of Theorem 2.1. A variant of the Piterbarg lemma suitable for Gaussian random fields is presented in Appendix A.

#### 2. Further results and the proof of Theorem 1.1

Following the idea of Hüsler and Piterbarg (1999), (2008), and as discussed in Hashorva *et al.* (2013), it is convenient to introduce the following family of Gaussian random fields:

$$Y_u(s,t) := \frac{X_H(ut) - \gamma X_H(us)}{(1 + ct - c\gamma s)u^H}, \qquad s, t \ge 0, \ u > 0.$$

The variance function of  $\{Y_u(s, t), s, t \ge 0\}$  is given by

$$V_Y^2(s,t) = \frac{(1-\gamma)t^{2H} + (\gamma^2 - \gamma)s^{2H} + \gamma(t-s)^{2H}}{(1+ct-c\gamma s)^2}, \qquad s,t \ge 0.$$

Moreover, on the set  $\{(s, t): 0 \le s \le t < \infty\}$ , it attains its maximum at the unique point  $(0, \tilde{t}_0)$  with  $\tilde{t}_0 = H/c(1 - H)$ , and, furthermore,

$$V_Y(0, \tilde{t}_0) = \frac{H^H (1 - H)^{1 - H}}{c^H}$$

By the change of variables t = t'u and s = s'u, and noting that the distribution of  $Y_u$  does not depend on u, we obtain

$$\mathbb{P}\{\tau_1(u) < \infty\} = \mathbb{P}\{\text{there exists } t \in [0, \infty) \text{ such that } W_{\gamma}(t) > u\}$$
$$= \mathbb{P}\{\text{there exists } t' \in [0, \infty) \text{ such that } Y_u(s', t') > u^{1-H} \text{ for some } s' \in [0, t']\}$$
$$= \mathbb{P}\{\text{there exists } t \in [0, \infty) \text{ such that } Y(s, t) > u^{1-H} \text{ for some } s \in [0, t]\},$$

where

$$Y(s,t) := \frac{X_H(t) - \gamma X_H(s)}{1 + c(t - \gamma s)}, \qquad s, t \ge 0.$$

In order to complete the proof of Theorem 1.1, we need to know the tail asymptotic behavior of the supremum of the Gaussian random field Y over a region which might depend on u. Therefore, we will first investigate the tail asymptotic behavior of the supremum of certain nonstationary Gaussian random fields (including Y as a special case) over a region depending on u in Theorem 2.1 followed then by the proof of Theorem 1.1.

Hereafter, we assume that all considered Gaussian random fields (or processes) have almost surely continuous sample paths. We need to introduce some more notation, starting with the well-known Pickands constant  $\mathcal{H}_{\alpha}$  given by

$$\mathcal{H}_{\alpha} := \lim_{T \to \infty} \frac{1}{T} \mathcal{H}_{\alpha}[0, T] \in (0, \infty), \qquad \alpha \in (0, 2],$$

where

$$\mathcal{H}_{\alpha}[0,T] = \mathbb{E}\Big(\exp\Big(\sup_{t\in[0,T]}(\sqrt{2}B_{\alpha}(t)-t^{\alpha})\Big)\Big), \qquad T\in(0,\infty),$$

with  $\{B_{\alpha}(t), t \ge 0\}$  an FBM with Hurst index  $\alpha/2 \in (0, 1]$ . It is known that  $\mathcal{H}_1 = 1$  and  $\mathcal{H}_2 = 1/\sqrt{\pi}$ ; see Pickands (1969), Albin (1990), Piterbarg (1996), Dębicki (2002), Dębicki *et al.* (2003), Mandjes (2007), Dębicki and Mandjes (2011), and Dieker and Yakir (2014) for various properties of Pickands' constant and its generalizations. Next, we introduce another constant, usually referred to as the Piterbarg constant, given by

$$\mathcal{P}^{a}_{\alpha} := \lim_{S \to \infty} \mathcal{P}^{a}_{\alpha}[0, S] \in (0, \infty), \qquad \alpha \in (0, 2], \ a > 0,$$

where

$$\mathcal{P}^a_{\alpha}[S,T] = \mathbb{E}\Big(\exp\Big(\sup_{t \in [S,T]} (\sqrt{2}B_{\alpha}(t) - (1+a)|t|^{\alpha})\Big)\Big), \qquad 0 \le S < T < \infty.$$

It is also known that

$$\mathcal{P}_{1}^{a} = 1 + \frac{1}{a}$$
 and  $\mathcal{P}_{2}^{a} = \frac{1}{2} \left( 1 + \sqrt{1 + \frac{1}{a}} \right);$ 

see, e.g. Dębicki and Mandjes (2003) and Dębicki and Tabiś (2011). As will be seen in Theorem 2.1 below, both Pickands' constant and Piterbarg's constant are important for our study. We denote by  $\Phi(\cdot)$  the standard normal distribution (of an N(0, 1) random variable), and, furthermore, set  $\Psi(\cdot) := 1 - \Phi(\cdot)$ .

In the following we investigate the tail asymptotic behavior of the supremum of nonstationary Gaussian random fields over a region which depends on u. Our next result is of interest on its own, and, furthermore, is the key to the proof of Theorem 1.1.

**Theorem 2.1.** Let *S* and *T* be two positive constants, and let  $\{X(s, t), (s, t) \in [0, S] \times [0, T]\}$  be a centered Gaussian random field, with standard deviation function  $\sigma(\cdot, \cdot)$  and correlation function  $r(\cdot, \cdot, \cdot, \cdot)$ . Assume that  $\sigma(\cdot, \cdot)$  attains its maximum on  $[0, S] \times [0, T]$  at the unique point  $(0, t_0)$  with  $t_0 \in (0, T)$ , and, furthermore, that

$$\sigma(s,t) = 1 - b_1 s^{\beta} (1 + o(1)) - b_2 |t - t_0|^2 (1 + o(1)) - b_3 s |t - t_0| (1 + o(1))$$
(2.1)

as  $(s, t) \rightarrow (0, t_0)$  for some constants  $\beta \in (1, 2)$ ,  $b_i > 0$ , i = 1, 2, and  $b_3 \in \mathbb{R}$  satisfying  $b_2 + b_3/2 > 0$ . Suppose, furthermore, that

$$r(s, s', t, t') = 1 - (a_1|s - s'|^{\beta} + a_2|t - t'|^{\beta})(1 + o(1)) \quad as(s, t), (s', t') \to (0, t_0) \quad (2.2)$$

for some constants  $a_i > 0$ , i = 1, 2. Then, for any  $x \in \mathbb{R}$ ,

$$\mathbb{P}\left\{\sup_{(s,t)\in\widetilde{\Delta}_{x}^{1}(u)}X(s,t) > u\right\} = \sqrt{\frac{\pi}{b_{2}}}a_{2}^{1/\beta}\mathcal{P}_{\beta}^{b_{1}/a_{1}}\mathcal{H}_{\beta}u^{2/\beta-1}\Psi(u)\Phi(\sqrt{2b_{2}}x)(1+o(1)), \quad (2.3)$$

$$\mathbb{P}\left\{\sup_{(s,t)\in\widetilde{\Delta}_{x}^{2}(u)}X(s,t)>u\right\} = \sqrt{\frac{\pi}{b_{2}}}a_{2}^{1/\beta}\mathcal{P}_{\beta}^{b_{1}/a_{1}}\mathcal{H}_{\beta}u^{2/\beta-1}\Psi(u)\Psi(\sqrt{2b_{2}}x)(1+o(1)),\quad(2.4)$$

as 
$$u \to \infty$$
, where  $\delta_1(u) = (\ln u/u)^{2/\beta}$ ,  $\delta_2(u) = \ln u/u$ , and  
 $\widetilde{\Delta}_x^1(u) = [0, \delta_1(u)] \times [t_0 - \delta_2(u), t_0 + xu^{-1}], \qquad \widetilde{\Delta}_x^2(u) = [0, \delta_1(u)] \times [t_0 + xu^{-1}, t_0 + \delta_2(u)].$ 

**Remarks.** (a) If  $\beta \in (0, 1)$  then (2.1) becomes

$$\sigma(s,t) = 1 - b_1 s^{\beta} (1 + o(1)) - b_2 |t - t_0|^2 (1 + o(1)) \quad \text{as } (s,t) \to (0,t_0).$$

We mention that in this case both (2.3) and (2.4) are still valid.

(b) It can be shown along the lines of the proof of Theorem 2.1 that if x = x(u) satisfies the conditions

$$\lim_{u \to \infty} x(u) = \infty, \qquad x(u) = o(u^{\varepsilon}), \quad \text{as } u \to \infty \text{ for any } \varepsilon > 0, \tag{2.5}$$

then (2.3) still holds with  $\Phi(\sqrt{2b_2x})$  replaced by 1. Similarly, if x = -x(u) with x(u) satisfying (2.5) then (2.4) holds with  $\Psi(\sqrt{2b_2x})$  replaced by 1.

# 2.1. Proof of Theorem 1.1

Define

$$T_1(u) = \inf\{t \ge 0 \colon Y(s, t) > u^{1-H} \text{ for some } s \in [0, t]\}$$

and

$$T_2(u) = \sup\{t \ge 0 : Y(s, t) > u^{1-H} \text{ for some } s \in [0, t]\}.$$

Clearly,  $\tau_i(u) \stackrel{\text{D}}{=} uT_i(u)$ , i = 1, 2. Consider first the approximation of  $\tau_1(u)$ . For any  $x \in \mathbb{R}$  and u > 0, we have

$$\mathbb{P}\left\{\frac{\tau_{1}(u) - \tilde{t}_{0}u}{A(u)} \le x \mid \tau_{1}(u) < \infty\right\} = \mathbb{P}\{T_{1}(u) \le \tilde{t}_{0} + xA(u)u^{-1} \mid T_{1}(u) < \infty\}$$
$$= \frac{\mathbb{P}\{\sup_{0 \le s \le t \le \tilde{t}_{0} + xA(u)u^{-1}} Y(s, t) > u^{1-H}\}}{\mathbb{P}\{\tau_{1}(u) < \infty\}}.$$

In view of Hashorva *et al.* (2013), for any  $H, \gamma \in (0, 1)$ ,

$$\mathbb{P}\{\tau_1(u) < \infty\} = \mathbb{P}\left\{\sup_{t \ge 0} W_{\gamma}(t) > u\right\} = W_H(u)\Psi\left(\frac{c^H u^{1-H}}{H^H (1-H)^{1-H}}\right)(1+o(1))$$
(2.6)

as  $u \to \infty$ , where

$$\mathcal{W}_{H}(u) = 2^{1/2 - 1/2H} \frac{\sqrt{\pi}}{\sqrt{H(1 - H)}} \mathcal{H}_{2H} \mathcal{P}_{2H}^{(1 - \gamma)/\gamma} \left(\frac{c^{H} u^{1 - H}}{H^{H} (1 - H)^{1 - H}}\right)^{1/H - 1}$$

Next, we focus on the analysis of  $\mathbb{P}\{\sup_{0 \le s \le t \le \tilde{t}_0 + xA(u)u^{-1}} Y(s, t) > u^{1-H}\}$ . By Bonforroni's inequality,

$$p_{3}(u) \leq \mathbb{P}\left\{\sup_{0 \leq s \leq t \leq \tilde{t}_{0} + xA(u)u^{-1}} Y(s, t) > u^{1-H}\right\} \leq p_{1}(u) + p_{2}(u) + p_{3}(u),$$
(2.7)

where  $p_i(u)$ , i = 1, 2, 3, are defined in (2.8), (2.12), and (2.13) below. In the following we will derive the asymptotics of  $p_3(u)$  as  $u \to \infty$ , and give bounds for both  $p_1(u)$  and  $p_2(u)$  for large u, assuring that they are relatively negligible.

We first consider bounds for  $p_1(u)$  and  $p_2(u)$ . Since on the set  $\{(s, t): 0 \le s \le t < \infty\}$  the maximum of the variance function  $V_Y^2(s, t)$  is attained uniquely at  $(0, \tilde{t}_0)$ , we obtain from the Borell-TIS inequality (see, e.g. Adler and Taylor (2007)) that, for any constant  $K \ge 2\tilde{t}_0$ , there exist constants  $\rho > 0$  small enough and  $\theta \in (0, 1)$  such that, for sufficiently large u,

$$p_1(u) := \mathbb{P}\left\{\sup_{\substack{0 \le s \le t \le K\\s \in [\rho, K] \text{ or } t \in [0, \tilde{t}_0 - \rho]}} Y(s, t) > u^{1-H}\right\} \le \exp\left(-\frac{(u^{1-H} - d)^2}{2\theta V_Y^2(0, \tilde{t}_0)}\right),$$
(2.8)

with  $d = \mathbb{E}(\sup_{0 \le s \le t \le K} Y(s, t)) < \infty$ . It follows that

$$1 - \frac{V_Y(s,t)}{V_Y(0,\tilde{t}_0)} = \begin{cases} \frac{c^2(1-H)^3}{2H}(\tilde{t}_0-t)^2(1+o(1)) \\ + \frac{(\gamma-\gamma^2)(1-H)^{2H}c^{2H}}{2H^{2H}}s^{2H}(1+o(1)), & H \le \frac{1}{2}, \\ \frac{c^2(1-H)^3}{2H}(\tilde{t}_0-t+\gamma s)^2(1+o(1)) \\ + \frac{(\gamma-\gamma^2)(1-H)^{2H}c^{2H}}{2H^{2H}}s^{2H}(1+o(1)), & H > \frac{1}{2}, \end{cases}$$
(2.9)

as  $(s, t) \rightarrow (0, \tilde{t}_0)$ , and, furthermore, the correlation function of Y satisfies

$$1 - \operatorname{cov}\left(\frac{Y(s,t)}{V_Y(s,t)}, \frac{Y(s',t')}{V_Y(s',t')}\right) = \frac{1}{2\tilde{t}_0^{2H}} (|t-t'|^{2H} + \gamma^2 |s-s'|^{2H})(1+o(1))$$
(2.10)

as  $(s, t), (s', t') \rightarrow (0, \tilde{t}_0)$ . In addition, for  $\rho > 0$  chosen small enough, there exists some Q > 0 such that, for any  $(s, t), (s', t') \in [0, \rho] \times [\tilde{t}_0 - \rho, \tilde{t}_0 + \rho]$ ,

$$\mathbb{E}(Y(s,t) - Y(s',t'))^2 \le Q(|t-t'|^{2H} + |s-s'|^{2H}).$$
(2.11)

Next, let

$$A = \frac{H^{1/2}}{c(1-H)^{3/2}}, \qquad \tilde{u} = \frac{u^{1-H}}{V_Y(0, \tilde{t}_0)}.$$

In light of (2.9) and (2.11), by the Piterbarg inequality (see Theorem 8.1 of Piterbarg (1996) or Theorem 8.1 of Piterbarg (2001)), for all sufficiently large u,

$$p_{2}(u) := \mathbb{P}\left\{\sup_{\substack{(s,t)\in[0,\rho]\times[\tilde{\iota}_{0}-\rho,\tilde{\iota}_{0}+xA(u)u^{-1}]\\s\in[\tilde{\delta}_{1}(\tilde{u}),\rho] \text{ or } t\in[\tilde{\iota}_{0}-\rho,\tilde{\iota}_{0}-\tilde{\delta}_{2}(\tilde{u})]}{s \in C_{1}u^{2(1-H)/H} \exp\left(-\frac{u^{2(1-H)}}{2V_{Y}^{2}(0,\tilde{\iota}_{0})}-C_{2}(\ln u)^{2}\right)\right\}$$

$$(2.12)$$

for some positive constants  $C_i$ , i = 1, 2, where  $\tilde{\delta}_1(\tilde{u}) = (\ln \tilde{u}/\tilde{u})^{1/H}$  and  $\tilde{\delta}_2(\tilde{u}) = \ln \tilde{u}/\tilde{u}$ . Furthermore, we have

$$p_{3}(u) := \mathbb{P}\left\{\sup_{\substack{(s,t)\in[0,\tilde{\delta}_{1}(\tilde{u})]\times[\tilde{t}_{0}-\tilde{\delta}_{2}(\tilde{u}),\tilde{t}_{0}+xA(u)u^{-1}]}} Y(s,t) > u^{1-H}\right\}$$
$$= \mathbb{P}\left\{\sup_{\substack{(s,t)\in\widehat{\Delta}_{Ax}^{1}(\tilde{u})}} \frac{Y(s,t)}{V_{Y}(0,\tilde{t}_{0})} > \tilde{u}\right\},$$
(2.13)

where  $\widehat{\Delta_{Ax}^1}(\tilde{u}) = [0, \tilde{\delta}_1(\tilde{u})] \times [\tilde{t}_0 - \tilde{\delta}_2(\tilde{u}), \tilde{t}_0 + Ax\tilde{u}^{-1}]$ . Utilizing (2.9) and (2.10), it follows from Theorem 2.1 that

$$\mathbb{P}\left\{\sup_{(s,t)\in\widehat{\Delta_{Ax}^{1}}(\tilde{u})}\frac{Y(s,t)}{V_{Y}(0,\tilde{t}_{0})} > \tilde{u}\right\} = \mathcal{H}_{2H}\mathcal{P}_{2H}^{(1-\gamma)/\gamma}2^{-1/2H}\sqrt{2\pi}A\frac{c(1-H)}{H}\Psi(\tilde{u})\tilde{u}^{1/H-1}\Phi(x)(1+o(1)) \qquad (2.14)$$

as  $u \to \infty$ . Consequently, we conclude from (2.7)–(2.8) and (2.12)–(2.14) that

$$\mathbb{P}\left\{\sup_{\substack{0 \le s \le t \le \tilde{t}_0 + xA(u)u^{-1} \\ H}} Y(s,t) > u^{1-H}\right\}$$
  
=  $\mathcal{H}_{2H}\mathcal{P}_{2H}^{(1-\gamma)/\gamma} 2^{-1/2H} \sqrt{2\pi} A \frac{c(1-H)}{H} \Psi(\tilde{u})\tilde{u}^{1/H-1} \Phi(x)(1+o(1))$ 

as  $u \to \infty$ , and, thus, in light of (2.6),

$$\lim_{u \to \infty} \sup_{x \in \mathbb{R}} \left| \mathbb{P}\left\{ \frac{\tau_1(u) - \tilde{t}_0 u}{A(u)} \le x \mid \tau_1(u) < \infty \right\} - \Phi(x) \right| = 0.$$

Using similar arguments, we conclude from the properties of the random field Y and (2.4) that

$$\mathbb{P} \Big\{ \sup_{t \ge \tilde{t}_0 + xA(u)u^{-1}, s \in [0, t]} Y(s, t) > u^{1-H} \Big\}$$
  
=  $\mathcal{H}_{2H} \mathcal{P}_{2H}^{(1-\gamma)/\gamma} 2^{-1/2H} \sqrt{2\pi} A \frac{c(1-H)}{H} \Psi(\tilde{u}) \tilde{u}^{1/H-1} \Psi(x)(1+o(1))$ 

as  $u \to \infty$ , where we used the fact that, for any large enough integer  $K > \tilde{t}_0$ ,

$$\mathbb{P}\left\{\sup_{0 \le s \le t < \infty} Y(s, t) > u^{1-H}\right\} = \mathbb{P}\left\{\sup_{0 \le s \le t < K} Y(s, t) > u^{1-H}\right\} (1 + o(1)) \quad \text{as } u \to \infty;$$

see Hashorva et al. (2013). Therefore,

$$\mathbb{P}\left\{\frac{\tau_{2}(u) - \tilde{t}_{0}u}{A(u)} \le x \mid \tau_{1}(u) < \infty\right\} = 1 - \mathbb{P}\left\{\frac{\tau_{2}(u) - \tilde{t}_{0}u}{A(u)} \ge x \mid \tau_{1}(u) < \infty\right\}$$
$$= 1 - \mathbb{P}\{T_{2}(u) \ge \tilde{t}_{0} + xA(u)u^{-1} \mid T_{1}(u) < \infty\}$$
$$= 1 - \frac{\mathbb{P}\{\sup_{t \ge \tilde{t}_{0} + xA(u)u^{-1}, s \in [0,t]} Y(s,t) > u^{1-H}\}}{\mathbb{P}\{\tau_{1}(u) < \infty\}}$$
$$\to \Phi(x) \quad \text{as } u \to \infty$$

for any  $x \in \mathbb{R}$ . Hence, the proof follows by a direct application of Lemma 2.1 below.

**Lemma 2.1.** Let  $(Z_{u1}, Z_{u2})$ , u > 0, be a bivariate random sequence such that  $Z_{u2} \ge Z_{u1}$  almost surely for all large u. If the convergence in distribution

$$Z_{ui} \xrightarrow{\mathrm{D}} \mathcal{Z} \quad as \ u \to \infty$$

holds for i = 1, 2 with Z a nondegenerate random variable, then we have the joint convergence in distribution

$$(Z_{u1}, Z_{u2}) \xrightarrow{\mathrm{D}} (\mathbb{Z}, \mathbb{Z}) \quad as \ u \to \infty.$$

*Proof.* Let x and y be any two continuous points of the distribution function  $\mathbb{P}\{\mathbb{Z} \leq t\}$ ,  $t \in \mathbb{R}$ . It is sufficient to show that

$$\lim_{u \to \infty} \mathbb{P}\{Z_{u1} \le x, \ Z_{u2} \le y\} = \mathbb{P}\{Z \le \min(x, y)\}.$$

In fact, if  $x \ge y$  by the assumption that  $Z_{u2} \ge Z_{u1}$  holds for all large u, we have

$$\mathbb{P}\{Z_{u1} \le x, Z_{u2} \le y\} = \mathbb{P}\{Z_{u2} \le y\} \to \mathbb{P}\{Z \le y\} \quad \text{as } u \to \infty$$

Furthermore, if  $x \leq y$ ,

$$\mathbb{P}\{Z_{u1} \le x, \ Z_{u2} \le y\} = \mathbb{P}\{Z_{u1} \le x\} - \mathbb{P}\{Z_{u1} \le x, \ Z_{u2} > y\}$$
$$\geq \mathbb{P}\{Z_{u1} \le x\} - \mathbb{P}\{Z_{u1} \le y, \ Z_{u2} > y\}$$
$$= \mathbb{P}\{Z_{u1} \le x\} - (\mathbb{P}\{Z_{u2} > y\} - \mathbb{P}\{Z_{u1} > y\})$$
$$\rightarrow \mathbb{P}\{Z \le x\} \quad \text{as } u \to \infty$$

and

$$\mathbb{P}\{Z_{u1} \le x, \ Z_{u2} \le y\} \le \mathbb{P}\{Z_{u1} \le x\} \to \mathbb{P}\{Z \le x\} \quad \text{as } u \to \infty$$

hold; hence, the claim follows.

### 3. Proof of Theorem 2.1

We present only the proof of (2.3) with  $x \ge 0$ , since the other cases can be dealt with similarly. For simplicity, we will assume that  $a_1 = a_2 = 1$ ; the general case follows by a time scaling.

Since our approach is asymptotic in nature, and both  $\delta_1(u)$  and  $\delta_2(u)$  converge to 0 as u tends to  $\infty$ , properties (2.1) and (2.2) are the only necessary properties of the Gaussian random field X needed for the asymptotics (which can be seen from the proof below). Therefore, we conclude that

$$\mathbb{P}\left\{\sup_{(s,t)\in\widetilde{\Delta}_{x}^{1}(u)}X(s,t)>u\right\} = \mathbb{P}\left\{\sup_{(s,t)\in\widetilde{\Delta}_{x}^{1}(u)}\widetilde{\xi}(s,t)>u\right\}(1+o(1)) =: \pi(u)(1+o(1))$$

as  $u \to \infty$ , with { $\tilde{\xi}(s, t)$ ,  $s, t \ge 0$ } any Gaussian random field possessing properties (2.1) and (2.2). In particular, we choose

$$\tilde{\xi}(s,t) = \frac{\xi(s,t)}{(1+b_1s^{\beta})(1+b_2|t-t_0|^2+b_3|t-t_0|s)}, \qquad s,t \ge 0$$

with  $\{\xi(s, t), s, t \in \mathbb{R}\}$  a centered homogeneous Gaussian random field with covariance function

$$r_{\xi}(s,t) = \exp(-|s|^{\beta} - |t|^{\beta}), \qquad s, t \in \mathbb{R}.$$

Since  $\beta < 2$ , for any positive constants  $S_1$  and  $S_2$ , we can divide the intervals  $[0, \delta_1(u)]$  and  $[t_0 - \delta_2(u), t_0 + xu^{-1}]$  into several subintervals of length  $S_1u^{-2/\beta}$  and  $S_2u^{-2/\beta}$ , respectively. Specifically, for  $S_1, S_2 > 0$ , let

$$\Delta_0^i = u^{-2/\beta}[0, S_i], \qquad \Delta_k^i = u^{-2/\beta}[kS_i, (k+1)S_i], \quad k \in \mathbb{Z}, \ i = 1, 2.$$

Furthermore, for any u > 0, let

$$h_1(u) = \lfloor S_1^{-1}(\ln u)^{2/\beta} \rfloor + 1, \qquad h_2(u) = \lfloor S_2^{-1}(\ln u)u^{2/\beta - 1} \rfloor + 1,$$
$$h_{2,x}(u) = \lfloor S_2^{-1}xu^{2/\beta - 1} \rfloor + 1.$$

Here [.] denotes the ceiling function. Applying Bonferroni's inequality we obtain

$$\pi(u) \leq \sum_{k_1=0}^{h_1(u)} \sum_{k_2=-h_2(u)}^{h_{2,x}(u)} \mathbb{P}\left\{\sup_{(s,t)\in\Delta_{k_1}^1\times(t_0+\Delta_{k_2}^2)}\tilde{\xi}(s,t) > u\right\}$$
$$= \sum_{k_2=-h_2(u)}^{h_{2,x}(u)} \mathbb{P}\left\{\sup_{(s,t)\in\Delta_0^1\times(t_0+\Delta_{k_2}^2)}\tilde{\xi}(s,t) > u\right\}$$
$$+ \sum_{k_1=1}^{h_1(u)} \sum_{k_2=-h_2(u)}^{h_{2,x}(u)} \mathbb{P}\left\{\sup_{(s,t)\in\Delta_{k_1}^1\times(t_0+\Delta_{k_2}^2)}\tilde{\xi}(s,t) > u\right\}$$
$$=: I_{1,x}(u) + I_{2,x}(u)$$

and

$$\pi(u) \ge J_{1,x}(u) - J_{2,x}(u),$$

where

Next, we derive the required asymptotic bounds of  $I_{1,x}(u)$  and  $J_{1,x}(u)$ , and show that

$$I_{2,x}(u) = J_{2,x}(u)(1+o(1)) = o(I_{1,x}(u)) = o(J_{1,x}(u))$$
(3.1)

as  $u \to \infty$  and  $S_i \to \infty$ , i = 1, 2. Assuming, furthermore, that  $b_3 > 0$ , we have

$$\begin{aligned} J_{1,x}(u) &\geq \sum_{k_2=0}^{h_{2,x}(u)-1} \mathbb{P} \bigg\{ \sup_{(s,t)\in\Delta_0^1\times\Delta_{k_2}^2} \frac{\xi(s,t)}{1+b_1s^{\beta}} > u(1+b_2((k_2+1)S_2u^{-2/\beta})^2 \\ &\quad + b_3((k_2+1)S_2u^{-2/\beta})(S_1u^{-2/\beta})) \bigg\} \\ &\quad + \sum_{k_2=-h_2(u)+1}^{-1} \mathbb{P} \bigg\{ \sup_{(s,t)\in\Delta_0^1\times\Delta_{k_2}^2} \frac{\xi(s,t)}{1+b_1s^{\beta}} > u(1+b_2(-k_2S_2u^{-2/\beta})^2 \\ &\quad + b_3(-k_2S_2u^{-2/\beta})(S_1u^{-2/\beta})) \bigg\} \end{aligned}$$

$$=: J_{1,1,x}(u) + J_{1,2,x}(u).$$

In view of Lemma A.1 in Appendix A,

$$J_{1,1,x}(u) = (1+o(1))\mathcal{P}_{\beta}^{b_{1}}[0, S_{1}]\mathcal{H}_{\beta}[0, S_{2}]\frac{1}{\sqrt{2\pi u}} \\ \times \sum_{k_{2}=0}^{h_{2,x}(u)-1} \frac{1}{1+b_{2}((k_{2}+1)S_{2}u^{-2/\beta})^{2}+b_{3}((k_{2}+1)S_{2}u^{-2/\beta})S_{1}u^{-2/\beta}} \\ \times \exp\left(-\frac{u^{2}(1+b_{2}((k_{2}+1)S_{2}u^{-2/\beta})^{2}+b_{3}((k_{2}+1)S_{2}u^{-2/\beta})S_{1}u^{-2/\beta})^{2}}{2}\right) \\ = \mathcal{P}_{\beta}^{b_{1}}[0, S_{1}]\mathcal{H}_{\beta}[0, S_{2}]\Psi(u)(1+o(1)) \\ \times \sum_{k_{2}=0}^{h_{2,x}(u)-1} \exp(-b_{2}((k_{2}+1)S_{2}u^{1-2/\beta})^{2}-b_{3}u^{2}((k_{2}+1)S_{2}u^{-2/\beta})S_{1}u^{-2/\beta}) \\ = \mathcal{P}_{\beta}^{b_{1}}[0, S_{1}]\frac{\mathcal{H}_{\beta}[0, S_{2}]}{S_{2}}\Psi(u)u^{2/\beta-1}\int_{0}^{x} e^{-b_{2}y^{2}} dy(1+o(1))$$
(3.2)

as  $u \to \infty$ , where in the last equation we utilized the facts that

$$h_{2,x}(u) \to \infty, \qquad h_{2,x}(u)S_2u^{1-2/\beta} \to x, \qquad u^2(h_{2,x}(u)S_2u^{-2/\beta})(S_1u^{-2/\beta}) \to 0,$$

as  $u \to \infty$ . Similarly,

$$J_{1,2,x}(u) = \mathcal{P}_{\beta}^{b_1}[0, S_1] \frac{\mathcal{H}_{\beta}[0, S_2]}{S_2} \Psi(u) u^{2/\beta - 1} \int_{-\infty}^{0} e^{-b_2 y^2} dy (1 + o(1))$$
(3.3)

as  $u \to \infty$ . Therefore, we conclude that

$$J_{1,x}(u) \ge \mathcal{P}_{\beta}^{b_1}[0, S_1] \frac{\mathcal{H}_{\beta}[0, S_2]}{S_2} \Psi(u) u^{2/\beta - 1} \int_{-\infty}^x e^{-b_2 y^2} dy (1 + o(1))$$
(3.4)

as  $u \to \infty$ . Using similar arguments, we further obtain

$$I_{1,x}(u) \leq \sum_{k_2=0}^{h_{2,x}(u)-1} \mathbb{P}\left\{\sup_{(s,t)\in\Delta_0^1\times\Delta_{k_2}^2} \frac{\xi(s,t)}{1+b_1s^{\beta}} > u(1+b_2(k_2S_2u^{-2/\beta})^2)\right\} \\ + \sum_{k_2=-h_2(u)}^{-1} \mathbb{P}\left\{\sup_{(s,t)\in\Delta_0^1\times\Delta_{k_2}^2} \frac{\xi(s,t)}{1+b_1s^{\beta}} > u(1+b_2(-(k_2+1)S_2u^{-2/\beta})^2)\right\} \\ = \mathcal{P}_{\beta}^{b_1}[0,S_1] \frac{\mathcal{H}_{\beta}[0,S_2]}{S_2} \Psi(u)u^{2/\beta-1} \int_{-\infty}^{x} e^{-b_2y^2} dy(1+o(1))$$
(3.5)

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as  $u \to \infty$ . Next we verify (3.1). Specifically,

$$\begin{split} I_{2,x}(u) &\leq \sum_{k_1=1}^{h_1(u)} \sum_{k_2=0}^{h_{2,x}(u)} \mathbb{P}\Big\{ \sup_{(s,t) \in \Delta_{k_1}^1 \times \Delta_{k_2}^2} \xi(s,t) > u(1+b_1(k_1S_1u^{-2/\beta})^\beta + b_2(k_2S_2u^{-2/\beta})^2) \Big\} \\ &+ \sum_{k_1=1}^{h_1(u)} \sum_{k_2=-h_2(u)}^{-1} \mathbb{P}\Big\{ \sup_{(s,t) \in \Delta_{k_1}^1 \times \Delta_{k_2}^2} \xi(s,t) > u(1+b_1(k_1S_1u^{-2/\beta})^\beta \\ &+ b_2(-(k_2+1)S_2u^{-2/\beta})^2) \Big\}. \end{split}$$

Applying similar arguments as in (3.4) yield

$$I_{2,x}(u) \leq \mathcal{H}_{\beta}[0, S_{1}]\mathcal{H}_{\beta}[0, S_{2}]\Psi(u)(S_{2}^{-1}u^{2/\beta-1})(1+o(1))$$
$$\times \int_{-\infty}^{x} e^{-b_{2}y^{2}} dy \sum_{k_{1}=1}^{h_{1}(u)} \exp(-b_{1}(k_{1}S_{1})^{\beta})$$
(3.6)

as  $u \to \infty$ . Furthermore, we write

$$J_{2,x}(u) = \sum_{-h_2(u)+1 \le i < j \le h_{2,x}(u)-1} \mathbb{P}\left\{\sup_{(s,t) \in \Delta_0^1 \times (t_0 + \Delta_i^2)} \tilde{\xi}(s,t) > u, \sup_{(s,t) \in \Delta_0^1 \times (t_0 + \Delta_j^2)} \tilde{\xi}(s,t) > u\right\}$$
  
=:  $\Sigma_{1,x}(u) + \Sigma_{2,x}(u),$ 

where  $\Sigma_{1,x}(u)$  is the sum over indexes j = i + 1, and  $\Sigma_{2,x}(u)$  is the sum over indexes j > i + 1. Let  $B(i, S_2, u) = u(1 + b_2(|i|S_2u^{-2/\beta})^2), i \in \mathbb{Z}, S_2 > 0, u > 0$ . It follows that

$$\sup_{(s,t)\in \triangle_0^1\times \triangle_{i+1}^2} \frac{\xi(s,t)}{1+b_1 s^{\beta}} > B(i+2,S_2,u) \bigg\}$$

and, for any  $i, j \in \mathbb{Z}$ ,

$$\begin{split} & \mathbb{P}\bigg\{\sup_{(s,t)\in\Delta_{0}^{1}\times\Delta_{i}^{2}}\frac{\xi(s,t)}{1+b_{1}s^{\beta}} > B(j,S_{2},u), \sup_{(s,t)\in\Delta_{0}^{1}\times\Delta_{i+1}^{2}}\frac{\xi(s,t)}{1+b_{1}s^{\beta}} > B(j,S_{2},u)\bigg\} \\ & = \mathbb{P}\bigg\{\sup_{(s,t)\in\Delta_{0}^{1}\times\Delta_{0}^{2}}\frac{\xi(s,t)}{1+b_{1}s^{\beta}} > B(j,S_{2},u)\bigg\} + \mathbb{P}\bigg\{\sup_{(s,t)\in\Delta_{0}^{1}\times\Delta_{1}^{2}}\frac{\xi(s,t)}{1+b_{1}s^{\beta}} > B(j,S_{2},u)\bigg\} \\ & - \mathbb{P}\bigg\{\sup_{(s,t)\in\Delta_{0}^{1}\times(\Delta_{0}^{2}\cup\Delta_{1}^{2})}\frac{\xi(s,t)}{1+b_{1}s^{\beta}} > B(j,S_{2},u)\bigg\}. \end{split}$$

Therefore, analogous to the derivation of (3.4), we obtain

$$\limsup_{u \to \infty} \frac{\Sigma_{1,x}(u)}{\Psi(u)u^{2/\beta - 1}} \le \mathcal{P}_{\beta}^{b_1}[0, S_1] \frac{2\mathcal{H}_{\beta}[0, S_2] - \mathcal{H}_{\beta}[0, 2S_2]}{S_2} \left( x + \int_{-\infty}^0 e^{-b_2 y^2} \, \mathrm{d}y \right).$$
(3.7)

Furthermore, for any u > 0,

$$\begin{split} \Sigma_{2,x}(u) &\leq \sum_{i=-1}^{h_{2,x}(u)-1} \sum_{j\geq 2} \mathbb{P}\Big\{ \sup_{(s,t)\in \Delta_0^1 \times \Delta_0^2} \xi(s,t) > u, \sup_{(s,t)\in \Delta_0^1 \times \Delta_j^2} \xi(s,t) > u \Big\} \\ &+ \sum_{i=-h_2(u)+1}^{-2} \sum_{j\geq 2} \mathbb{P}\Big\{ \sup_{(s,t)\in \Delta_0^1 \times \Delta_0^2} \xi(s,t) > B(i+1,S_2,u), \sup_{(s,t)\in \Delta_0^1 \times \Delta_j^2} \xi(s,t) > u \Big\} \\ &\leq \sum_{i=-1}^{h_{2,x}(u)-1} \sum_{j\geq 2} \mathbb{P}\Big\{ \sup_{\substack{(s',t')\in \Delta_0^1 \times \Delta_j^2\\(s,t)\in \Delta_0^1 \times \Delta_0^2}} \zeta(s,t,s',t') > 2u \Big\} \\ &+ \sum_{i=-h_2(u)+1}^{-2} \sum_{j\geq 2} \mathbb{P}\Big\{ \sup_{\substack{(s',t')\in \Delta_0^1 \times \Delta_j^2\\(s,t)\in \Delta_0^1 \times \Delta_0^2}} \zeta(s,t,s',t') > B(i+1,S_2,u) + u \Big\}, \end{split}$$

where

$$\xi(s, t, s', t') = \xi(s, t) + \xi(s', t'), \qquad s, s', t, t' \ge 0.$$

It is easy to check that, for sufficiently large *u*,

$$2 \le \mathbb{E}((\zeta(s,t,s',t'))^2) = 4 - 2(1 - r_{\xi}(s-s',t-t')) \le 4 - ((j-1)S_2)^{\beta}u^{-2}$$

for any  $(s, t) \in \Delta_0^1 \times \Delta_0^2$ ,  $(s', t') \in \Delta_0^1 \times \Delta_j^2$ . Applying the same arguments as in the proof of Lemma 6.3 of Piterbarg (1996) we conclude that

$$\limsup_{u \to \infty} \frac{\Sigma_{2,x}(u)}{\Psi(u)u^{2/\beta - 1}} \le Q_x(\mathcal{H}_{\beta}[0, S_1])^2 S_2 \sum_{j \ge 1} \exp\left(-\frac{1}{8}(jS_2)^{\beta}\right)$$
(3.8)

for some positive constant Q. Hence, the claim follows from (3.1)–(3.8) when  $b_3 > 0$  by letting  $S_2, S_1 \rightarrow \infty$ . When  $b_3 < 0$ , the same results can be obtained using similar arguments as above and the fact that

$$1 - \sigma(s, t) \ge b_1 s^{\beta} (1 + o(1)) + \left(b_2 + \frac{b_3}{2}\right) |t - t_0|^2 (1 + o(1))$$

as  $(s, t) \rightarrow (0, t_0)$ , which is utilized for verifying (3.1). This completes the proof.

## Appendix A. Piterbarg's lemma for Gaussian random fields

In order to find the asymptotics of the supremum of centered nonsmooth Gaussian processes, two crucial results are important, namely the Pickands lemma and the Piterbarg lemma. Although for experts in this field the results are well known, we would like to briefly mention them. Let  $\{X(t), t \ge 0\}$  be a centered stationary Gaussian process with almost surely continuous sample paths and correlation function r(t) which satisfies  $r(t) = 1 - t^{\alpha}(1 + o(1))$ as  $t \to 0$  with  $\alpha \in (0, 2]$  and r(t) < 1 for all t > 0. In the seminal paper Pickands (1969) it was shown that, for any  $T \in (0, \infty)$ ,

$$\mathbb{P}\left\{\sup_{t\in[0,T]}X(t)>u\right\} = \mathcal{H}_{\alpha}Tu^{2/\alpha}\Psi(u)(1+o(1)) \quad \text{as } u\to\infty.$$
(A.1)

The proof of (A.1) strongly relies on Pickands' lemma which states that

$$\mathbb{P}\left\{\sup_{t\in[0,u^{-2/\alpha}T]}X(t)>u\right\} = \mathcal{H}_{\alpha}[0,T]\Psi(u)(1+o(1)) \quad \text{as } u\to\infty.$$
(A.2)

In the seminal contribution Piterbarg (1972), Piterbarg rigorously proved (A.1) and then extended (A.2) to a result which we refer to as the Piterbarg lemma, namely, for any constant b > 0,

$$\mathbb{P}\left\{\sup_{t\in[0,u^{-2/\alpha}T]}\frac{X(t)}{1+bt^{\alpha}}>u\right\}=\mathcal{P}^b_{\alpha}[0,T]\Psi(u)(1+o(1))\quad\text{as }u\to\infty.$$

Our next result is a variant of the Piterbarg lemma for the two-dimensional case. We omit its proof since it follows exactly the same arguments as those used in the proof of Lemma 6.1 of Piterbarg (1996).

**Lemma A.1.** Let  $\{\xi(s, t), s, t \in \mathbb{R}\}$  be a centered homogeneous Gaussian random field with covariance function

$$r_{\xi}(s,t) = \exp(-|s|^{\alpha_1} - |t|^{\alpha_2}), \qquad s, t \in \mathbb{R}, \, \alpha_1, \alpha_2 \in (0,2].$$

Furthermore, let S, T, and  $T_0$  be three constants such that S, T > 0 and  $T_0 \in \mathbb{R}$ . Then, for any constant  $b \ge 0$  and any positive function g(u),  $u \ge 0$ , satisfying  $\lim_{u\to\infty} g(u)/u = 1$ , we have

$$\mathbb{P}\left\{\sup_{(s,t)\in[0,u^{-2/\alpha_{1}}S]\times[u^{-2/\alpha_{2}}T_{0},u^{-2/\alpha_{2}}(T_{0}+T)]}\frac{\xi(s,t)}{1+bs^{\alpha_{1}}}>g(u)\right\}$$
$$=\mathcal{P}_{\alpha_{1}}^{b}[0,S]\mathcal{H}_{\alpha_{2}}[0,T]\Psi(g(u))(1+o(1))$$

as  $u \to \infty$ .

**Remark.** In the last formula we identify  $\mathcal{P}^{b}_{\alpha_{1}}[0, S]$  to be  $\mathcal{H}_{\alpha_{1}}[0, S]$  when b = 0.

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