ON A CLASS OF MARKOV PROCESSES TAKING VALUES ON LINES AND THE CENTRAL LIMIT THEOREM

MASATOSHI FUKUSHIMA and MASUYUKI HITSUDA

To Professor Kiyoshi Noshiro on the occasion of his 60-th birthday

§1. Introduction

We shall consider a class of Markov processes (n(t), x(t)) with the continuous time parameter $t \in [0, \infty)$, whose state space is $\{1, 2, ..., N\} \times R^{1}$. We shall assume that the processes are spacially homogeneous with respect to $x \in R^{1}$. In detail, our assumption is that the transition function

$$F_{ij}(x,t) = P(n(t) = j, x(t) \le x \mid n(0) = i, x(0) = 0), \quad t > 0, \ 1 \le i, j \le N, x \in \mathbb{R}^{1}, x \in \mathbb{R}$$

satisfies following conditions $(1, 1) \sim (1, 4)$.

(1, 1) $F_{ij}(x, t)$ is non-negative, and, for fixed i, j and t, it is monotone nondecreasing and right continuous in $x \in \mathbb{R}^1$.

$$\begin{array}{ll} (1,\,2) & F_{ij}(+\infty,\,t) = \lim_{x \to \infty} F_{ij}(x,\,t) \leq 1, \\ & F_{ij}(-\infty,\,t) = \lim_{x \to \infty} F_{ij}(x,\,t) = 0, \quad 1 \leq i,\,j \leq N, \quad t > 0, \\ & \sum_{i=1}^{N} F_{ij}(+\infty,\,t) = 1, \quad i = 1,\,2,\ldots,\,N, \quad t > 0, \\ (1,\,3) & F_{ij}(x,\,t) = \sum_{k=1}^{N} \int_{\mathbb{R}^{1}} F_{ik}(x-y,\,t) dF_{kj}(y,\,s) \quad t,\,s > 0, \\ & 1 \leq i,\,k \leq N, \quad x \in \mathbb{R}^{1}, \\ (1,\,4) & \lim_{t \downarrow 0} F_{ij}(x,\,t) = \begin{cases} \delta_{ij}, \quad x \in [0,\,+\infty) \\ 0, \quad x \in (-\infty,\,0). \end{cases} \end{array}$$

The central limit theorem for processes of this type, in case of the discrete time parameter and in a special case of the continous time parameter, has been obtained by Keilson and Wishart [3]. In this paper, through introducing a system of generators of the semi-groups related to the processes, we show that the central limit theorem is valid for our cases of the continous time parameter.

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With the transition function $F_{ij}(x, t)$ satisfying $(1, 1) \sim (1, 4)$ we associate its Fourier transform:

$$f_{ij}(z,t) = \int_{R^1} e^{ixz} dF_{ij}(x,t), \ t > 0, \ 1 \leq i, \ j \leq N, \ z \in R^1.$$

Then the matrix $f(z, t) = (f_{ij}(z, t))$ satisfies the followings:

- (1, 5) $\sum_{j=1}^{N} f_{ij}(0, t) = 1, \quad 1 \leq i \leq N, \quad t > 0,$
- (1, 6) f(z, t+s) = f(z, t)f(z, s) t, s > 0,
- (1, 7) f(z, t) converges, as t tends to zero, to the identity matrix E uniformly in $z \in \mathbb{R}^1$ in the wide sense, i.e. each element $f_{ij}(z, t)$ of f(z, t) converges to δ_{ij} uniformly on any compact z-set.

In §2, we shall determine the generator A(z), $z \in \mathbb{R}^{1}$, of the semigroups $f(z, t), z \in \mathbb{R}^{1}$, (Theorem 1). In particular, if N=1, our expression of A(z) is no other than the so-called Lévy-Khintchine formula. $\{A(z), z \in \mathbb{R}^{1}\}$ in Theorem 1 characterizes all the processes whose transition functions satisfy $(1, 1) \sim (1, 4)$.

§3 will be devoted to the proof of the central limit theorem for our processes under some assumptions placed upon A(z) (Theorem 2). Our procedure in this section is essentially due to Keilson and Wishart [3].

In §4, *m* and *v* which are defined by the first and second derivatives respectively of an eigen value $\lambda(z)$ of A(z) will be expressed explicitly by A(z) and its eigenvectors. We shall further solve affirmatively the Harris' conjecture related to the expectation processes of the electron-photon cascade (Harris [2], page 194).

§2. The generators of semigroups f(z, t).

Throughout §2, §3 and the first half of §4, we shall assume that we are given a transition function $F_{ij}(x, t)$ on $\{1, 2, ..., N\} \times R^1$ satisfying $(1, 1) \sim (1, 4)$.

In this section, we shall show the existence of the generator

$$A(z) = \lim_{t \downarrow 0} \frac{f(z, t) - E}{t}$$
 (Lemma 1),

and determine all the possible types of the generator A(z) (Theorem 1).

LEMMA 1. The limit

$$\lim_{t\downarrow 0} \frac{f(z,t)-E}{t} = A(z)$$

exists and the convergence is uniformly in $z \in \mathbb{R}^1$ in the wide sence.

Proof. Let us fix a > 0 arbitrarily. From the property (1, 7), we can choose v > 0 such that, for $z \in [-a, a]$, the matrix

$$\boldsymbol{A}_{v}(\boldsymbol{z}) = \int_{0}^{v} \boldsymbol{f}(\boldsymbol{z}, s) d\boldsymbol{s} = \left(\int_{0}^{v} f_{ij}(\boldsymbol{z}, s) d\boldsymbol{s} \right)$$

has its inverse $\Lambda_v^{-1}(z)$. Since

$$\boldsymbol{\Lambda}_{\boldsymbol{v}}(z)\boldsymbol{f}(z,t) = \int_{0}^{\boldsymbol{v}} \boldsymbol{f}(z,s) \boldsymbol{f}(z,t) ds = \int_{0}^{\boldsymbol{v}} \boldsymbol{f}(z,s+t) ds = \int_{t}^{\boldsymbol{v}+t} \boldsymbol{f}(z,s) ds,$$

we have

$$\boldsymbol{f}(z,t) = \boldsymbol{\Lambda}_{v}^{-1}(z) \int_{t}^{v+t} \boldsymbol{f}(z,s) ds.$$

Therefore we can obtain that

$$\lim_{h \downarrow 0} \frac{f(z, h) - E}{h} = \lim_{h \downarrow 0} \Lambda_v^{-1}(z) \frac{1}{h} \left[\int_h^{v+h} f(z, s) ds - \int_0^v f(z, s) ds \right]$$

= $\Lambda_v^{-1}(z) \lim_{h \downarrow 0} \frac{1}{h} \left[\int_v^{v+h} f(z, s) ds - \int_0^h f(z, s) ds \right]$
= $\Lambda_v^{-1}(z) [f(z, v) - E], \quad z \in [-a, a].$

The convergence in question is uniform since f(z, s) converges to f(z, v) uniformly in $z \in [-a, a]$ as s tends to v. The right side of the above equation should be independent of v because the left does not depend on v.

Thus we have proved Lamma 1.

Remark. We can express f(z, t) in the form $\exp\{A(z)t\}$. In fact, f(z, t) is the unique solution of the equation

$$\frac{\partial}{\partial t} \mathbf{f}(z, t) = \mathbf{A}(z)\mathbf{f}(z, t)$$

with the initial condition

$$\lim_{t\downarrow 0} f(z,t) = E.$$

The following theorem gives us the possible types of A(z).

THEOREM 1. The elements $a_{ij}(z)$, i, j=1, 2, ..., N, of A(z) can be expressed in the form,

$$\begin{aligned} a_{ij}(z) &= \int_{-\infty}^{\infty} e^{izx} \Gamma_{ij}(dx) & i \neq j, \quad 1 \leq i, \ j \leq N, \\ a_{ii}(z) &= -\sum_{j \neq i} \Gamma_{ij}(R^1) + i\nu_i z - \frac{\sigma_i^2}{2} z^2 + \int_{|u| > 1} (e^{izv} - 1) \Pi_i(du) \\ &+ \int_{|u| \leq 1} (e^{izu} - 1 - izu) \Pi_i(du), \qquad i = 1, 2, ..., N \end{aligned}$$

where Γ_{ij} 's, $1 \leq i \neq j \leq N$, are finite measures on R^1 , ν_i 's, i = 1, ..., N are real numbers, σ_i 's, i = 1, ..., N are real and positive, and Π_i 's, i = 1, ..., N are measures on R^1 such that $\int_{|u| \leq 1} x^2 \Pi_i(dx) < \infty$ and $\Pi_i(u; |u| > \varepsilon) < \infty$ for any $\varepsilon > 0$.

Proof. In case that $i \neq j$, we can see by Lemma 1 that the convergence

$$a_{ij}(z) = \lim_{h \downarrow 0} \frac{f_{ij}(z, h)}{h}$$

is uniform in the wide sence in $z \in \mathbb{R}^1$, and therefore $a_{ij}(z)$ is a positive definite function of z because $\frac{f_{ij}(z, h)}{h}$ has the same property. Then there exists a

finite measure Γ_{ij} such that

$$a_{ij}(z) = \int_{-\infty}^{\infty} e^{izx} \Gamma_{ij}(dx).$$

In case that i=j, we note that

$$\frac{f_{ii}(z,h)-1}{h} = \frac{e^{a_{ii}(z)h}-e^{a_{ii}(0)h}}{h} + \frac{e^{a_{ii}(0)h}-1}{h} + \frac{f_{ii}(z,h)-e^{a_{ii}(z)h}}{h}$$

Then, $\frac{f_{ii}(z,h) - e^{a_{ii}(z)h}}{h}$ tends to zero as $h \downarrow 0$ by virtue of the remark after Lemma 1. If we remark that $a_{ii}(0) = -\sum_{j=i}^{N} a_{ij}(0) = -\sum_{j=i}^{N} \Gamma_{ij}(R^1)$, $\frac{e^{a_{ii}(0)} - 1}{h}$ tends to $-\sum_{j=i}^{N} \Gamma_{ij}(R^1)$ as $h \downarrow 0$. The limit of $\frac{e^{a_{ii}(z)h} - e^{a_{ii}(0)h}}{h} = e^{a_{ii}(0)h} \cdot \frac{e^{(a_{ii}(z) - a_{ii}(0))h} - 1}{h}$ as $h \downarrow 0$ is the Lévy-Khintchine formula:

$$i\nu_i z - \frac{\sigma_i^2}{2} z^2 + \int_{|u|>1} (e^{izu} - 1) \Pi_i(du) + \int_{|u|\leq 1} (e^{izu} - 1 - izu) \Pi_i(du)$$

(see for example Gnedenko and Kolmogorov [1]).

Remark: From the above discussion, we can see that Γ_{ij} , ν_i , σ_i ; and Π_i are uniquely determined by the transition function $F_{ij}(x, t)$ satisfying $(1, 1) \sim (1, 4)$.

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§3. The central limit theorem.

Let A(z) be the matrix defined in Lemma 1. We shall assume

Assumption 1.

$$\begin{split} & \int_{-\infty}^{\infty} x^2 \Gamma_{ij}(dx) < +\infty, \quad i \neq j, \ 1 \leq i, \ j \leq N, \\ & \int_{-\infty}^{\infty} x^2 \Pi_i(dx) < +\infty, \quad 1 \leq i \leq N. \end{split}$$

Assumption 2. A(0) is irreducible.

The Assumption 1 is equivalent to the fact that

(*)
$$\int_{-\infty}^{\infty} x^2 dF_{ij}(x, 1) < \infty$$
 for any $i, j = 1, ..., N$.

In fact, Assumption 1 is equivalent to that

(**) A(z) is twice differentiable with respect to z.

And (**) is equivalent to the same property of $e^{A(z)} = f(z, 1)$. This is no other than (*) (see Feller [4] page 485).

By virtue of Assumption 2, $f(0, 1) = e^{A(0)}$ is a positive stochastic matrix and therefore, by Perron-Frobenius' theorem, it has the simple eigenvalue 1 and the absolute values of the other N-1 eigenvalues are less than 1. Therefore, the equation

$$\det(\boldsymbol{A}(0) - \boldsymbol{\lambda}\boldsymbol{E}) = \pi(0, \boldsymbol{\lambda}) = 0$$

has a simple root $\lambda = 0$, and all the other roots have negative real parts. From these facts we can derive the next Lemma.

LEMMA 2. There is a function $\lambda(z)$ defined on some neighbourhood (-a, a) of z=0 which has the continuous first and second derivatives in (-a, a) and satisfies that

$$\pi(z, \lambda(z)) = 0, \qquad z \in (-a, a)$$

$$\lambda(0) = 0$$

where $\pi(z, \lambda) = det (A(z) - \lambda E).$

Further, $\lambda'(0)$ is purely imaginary and $\lambda''(0)$ is real and non-positive.

Proof. As $\lambda = 0$ is a simple root of the equation $\Pi(0, \lambda) = 0$, we know that

$$\pi_{\lambda}(0, 0) = \frac{\partial \pi}{\partial \lambda}(0, 0) \neq 0. \text{ If we put}$$
$$\pi(z, \lambda) = \pi_1(z, \lambda_1, \lambda_2) + i\pi_2(z, \lambda_1, \lambda_2)$$

where $\pi_i(z, \lambda_1, \lambda_2)$, i = 1, 2, are real valued functions and $\lambda_1 = \Re \epsilon \lambda$, $\lambda_2 = \Im \mathfrak{m} \lambda$, then we have

$$\pi_1(0, 0, 0) = 0, \qquad \pi_2(0, 0, 0) = 0.$$

On the other hand, $\pi(z, \lambda)$ is analytic in λ . Therefore we have by Cauchy-Riemann's equation that

$$\begin{vmatrix} \frac{\partial \pi_1}{\partial \lambda_1}(0, 0, 0) & \frac{\partial \pi_2}{\partial \lambda_1}(0, 0, 0) \\ \frac{\partial \pi_1}{\partial \lambda_2}(0, 0, 0) & \frac{\partial \pi_2}{\partial \lambda_2}(0, 0, 0) \end{vmatrix} = \begin{vmatrix} \frac{\partial \pi_1}{\partial \lambda_1}(0, 0, 0) & \frac{\partial \pi_2}{\partial \lambda_1}(0, 0, 0) \\ -\frac{\partial \pi_2}{\partial \lambda_1}(0, 0, 0) & \frac{\partial \pi_1}{\partial \lambda_1}(0, 0, 0) \end{vmatrix}$$
$$= \left(\frac{\partial \pi_1}{\partial \lambda_1}(0, 0, 0)\right)^2 + \left(\frac{\partial \pi_2}{\partial \lambda_1}(0, 0, 0)\right)^2 \neq 0.$$

Now the first part of Lemma 2 is the direct consequence of the implicit function theorem.

Now, it is easy to see by the form of A(z) that $\pi_z(0, 0) = \frac{\partial}{\partial z}\pi(0, 0)$ is pure imaginary and $\pi_\lambda(0, 0)$ is real. Therefore $\lambda'(0) = \frac{-\pi_z(0, 0)}{\pi_\lambda(0, 0)}$ is pure imaginary. By the same method we can see that $\lambda''(0)$ is real. Let us note that $\sum_{i=1}^N |f_{ij}(z, 1)| \leq 1$ where $f_{ij}(z, 1)$ is (i, j)-element of the matrix $e^{A(z)} = f(z, 1)$. This implies, applying Frobenius' theorem again, all the eigenvalues of A(z)have non-positive real parts. Especially, $\Re e \lambda(z) \leq 0$ and $\Re e \lambda(z)$ attains its maximum 0 at z=0. Thus

$$\frac{d^2}{dz} \Re \epsilon \lambda(z)]_{z=0} = \Re \epsilon \lambda^{\prime\prime}(0) = \lambda^{\prime\prime}(0) \leq 0.$$

Remark. We may assume without loss of generality that for $z \in (-a, a)$ $\lambda(z)$ is the eigenvalue of A(z) the real part of which is greater than the real parts of any other eigenvalues, because the eigenvalues of A(z) are continuous in z, and the eigenvalue $\lambda(0)=0$ of A(0) has the maximum real part. In the following we assume the above property of $\lambda(z)$ in $z \in (-a, a)$.

Consider a probability measure $\mu = (\mu_1, \mu_2, \dots, \mu_N)$ on $\{1, 2, \dots, N\} \times R^1$, where μ_i 's are measures on R^1 such that $\sum_{i=1}^N \mu_i(R^1) = 1$. Let us put

$$F_{j}^{\mu}(x,t) = \sum_{i=1}^{N} \int_{-\infty}^{\infty} \mu_{i}(dy) F_{ij}(x-y,t), \quad 1 \leq j \leq N, \quad and$$

$$F^{\mu}(x,t) = \{F^{\mu}_{1}(x,t), F^{\mu}_{2}(x,t), \dots, F^{\mu}_{N}(x,t)\}.$$

Then $F^{\mu}(x, t)$ determines the distribution of a Markov process on $\{1, 2,, N\}$ $\times R^1$ at time t with the initial distribution μ . We further put

$$\boldsymbol{f}^{\mu}(\boldsymbol{z},t) = \Big\{ \int_{-\infty}^{\infty} e^{ixz} dF_{1}^{\mu}(\boldsymbol{x},t), \int_{-\infty}^{\infty} e^{ixz} dF_{2}^{\mu}(\boldsymbol{x},t), \dots, \int_{-\infty}^{\infty} e^{ixz} dF_{N}^{\mu}(\boldsymbol{x},t) \Big\}.$$

Then,

$$f^{\mu}(z, t) = f^{\mu}(z, 0+)f(z, t)$$

where $f^{\mu}(z, 0+)$ is the Fourier transform of μ . Define *m* and *v* by the formula

(3, 1)
$$m = -i\lambda'(0), \quad v = -\lambda''(0).$$

THEOREM 2. For any initial distribution μ on $\{1, 2, ..., N\} \times R^1$,

(3, 2)
$$\lim_{t\to\infty} e^{-it^{\frac{1}{2}}mz} f^{\mu}(t^{-\frac{1}{2}}z, t) = e^{\frac{-1}{2}z^{2}v^{2}} e, \ z \in (-a, a)$$

holds, where $e = (e_1, e_2, \dots, e_N)$ is the left eigenvector belonging to A(0) for the eigenvalue 0, and $\sum_{i=1}^{N} e_i = 1$.

Proof. There is a regulaer matrix T(z) for $z \in (-a, a)$ such that

$$\boldsymbol{A}(\boldsymbol{z}) = \boldsymbol{T}^{-1}(\boldsymbol{z})\boldsymbol{J}(\boldsymbol{z})\boldsymbol{T}(\boldsymbol{z})$$

where J(z) is a matrix of the Jordan's normal form whose (1, 1)-element is $\lambda(z)$. Here we can choose T(z) to be continuous in $z \in (-a, a)$. Now we have

(3, 3)
$$\boldsymbol{f}^{\mu}(t^{-\frac{1}{2}}z, t)e^{-izmt^{\frac{1}{2}}} = \boldsymbol{f}^{\mu}(t^{-\frac{1}{2}}z, 0)\boldsymbol{T}^{-1}(t^{-\frac{1}{2}}z)e^{-izmt^{-\frac{1}{2}}}\exp(t\boldsymbol{J}(t^{-\frac{1}{2}}))\boldsymbol{T}(t^{-\frac{1}{2}}z).$$

In order to investigate the limit of the above expression, we first note the following

(3, 4)
$$\lim_{t\to\infty} f^{\mu}(t^{-\frac{1}{2}}z, 0) = \lim_{z\to 0} f^{\mu}(z, 0) = f^{\mu}(0, 0).$$

Since the first column e(z) of T(z) is the left eigenvector and the first row i(z) of T(z) is the right eigenvector of A(z) belonging to the eigenvalue $\lambda(z)$, we may assume that e(0) and i(0) are of the form

$$\boldsymbol{e}(0) = \boldsymbol{e} = (\boldsymbol{e}_1, \boldsymbol{e}_2, \dots, \boldsymbol{e}_N)$$
$$\boldsymbol{i}(0) = \begin{pmatrix} 1\\ 1\\ \vdots\\ 1 \end{pmatrix}.$$

By the continuity of each element of T(z), we have

(3, 5)
$$\lim_{t \to \infty} T(t^{\frac{-1}{2}}z) = \lim_{z \to 0} T(z) = \begin{pmatrix} e_1, e_2, \dots, e_N \\ * \end{pmatrix}$$

(3, 6)
$$\lim_{t \to \infty} T^{-1}(t^{-\frac{1}{2}}z) = \lim_{z \to 0} T^{-1}(z) = \begin{pmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{pmatrix}$$

The absolute values of the (i, j)-elements of $\exp\{tJ(z)\}$ for $i, j \ge 2$ and $j-i=l \ge 0$ are less than $\frac{1}{l!}te^{-\alpha(z)t}$, where $-\alpha(z)$ denotes the maximum real part of the eigenvalues of A(z) except $\lambda(z)$.

On the other hand the (i, j)-elements of exp $\{tJ(z)\}$ for j-i < 0 or for i=1, $j \neq i$ are equal to 0. Since $\sup_{z \in (-a,a)} e^{-a(z)} < 1$, the (i, j)-elements of exp $\{tJ(z)\}$ for $(i, j) \neq (1, 1)$ converge to 0 as $t \to \infty$. For the (1, 1)-element, we get

$$\lim_{t\to\infty} e^{-izmt^{\frac{1}{2}}} e^{t\lambda(t^{\frac{-1}{2}}z)} = e^{\frac{-1}{2}z^2v^2},$$

with m and v defined by (3, 1). Therefore

(3, 7)
$$\lim_{t \to \infty} e^{-izmt^{\frac{1}{2}}} \exp\{tJ(t^{-\frac{1}{2}}z)\} = e^{-\frac{1}{2}z^{2}t^{2}} \begin{pmatrix} 1 \mid 0, \dots, 0 \\ 0 \mid \\ \vdots \mid 0 \end{pmatrix}$$

holds. Consequently we get (3, 1) by $(3, 3) \sim (3, 7)$ and the proof is complete.

COROLLARY. In case $v = -\lambda''(0) > 0$, we have

$$\lim_{t \to \infty} F^{\mu}(xt^{\frac{1}{2}} + tm, t) = \int_{0}^{\frac{x}{v}} \frac{e^{-\frac{y^{2}}{2}}}{\sqrt{2\pi}} dy \cdot e$$

for any initial distribution μ .

§4. The expression of m and ν , and the expectation process of the eletron-photon cascade.

Let e(z) and i(z) be the left and right eigenvector of A(z) belonging to $\lambda(z)$ respectively, as in the proof of Theorem 2. We can assume that they are twice differentiable in the neighbourhood of the origin and e(0) = e, i(0) = 1. Since $\lambda(0) = 0$, taking dreivatives of the equations on a class of markov processes taking values on lines and the central limit theorem 55

(4, 1)
$$e(z)A(z) = \lambda(z)e(z)$$
 and

(4, 2)
$$A(z)i(z) = \lambda(z)i(z)$$

we have

(4, 3)
$$e'(0)A(0) + eA'(0) = \lambda'(0)e$$
 and

(4, 4)
$$A(0)i'(0) + A'(0)\mathbf{1} = \lambda'(0)\mathbf{1}.$$

By (4, 3), we have

(4,5)
$$m = -i\lambda'(0) = -eA'(0)\mathbf{1}.$$

Differentiating (4, 1) twice and noting (4, 4), we have

(4, 6)
$$v = -\lambda''(0) = -eA''(0)\mathbf{1} + 2e'(0)A(0)i'(0).$$

The following Lemma 3 will illustrate the probabilistic measning of m and v. Let $\{n(t), x(t)\}$ be the Markov process on $\{1, 2, \ldots, N\} \times R^1$ whose transition function are governed by $F_{ij}(x, t)$ and whose initial distrivution is $\delta_0(dx)e$, where δ_0 is the δ -measure concentrated at the origin.

LEMMA 3.
$$m = -ieA'(0)\mathbf{1} = \frac{E(x(t))}{t}, \text{ for any } t > 0$$
$$v = -eA''(0)\mathbf{1} + 2e'(0)A(0)i'(0) = \lim_{t \to +\infty} \frac{E[x(t) - E(x(t))]^2}{t}$$

Proof. We can see that

$$\boldsymbol{E}(\boldsymbol{x}(t)) = -i\boldsymbol{e}\left(\frac{\partial}{\partial z} \boldsymbol{f}(\boldsymbol{z}, t)\right) \mathbf{1} = -i\boldsymbol{e}\left(\frac{\partial}{\partial z} \boldsymbol{e}^{t\mathbf{A}(\boldsymbol{z})}\right) \mathbf{1} = -it\boldsymbol{e}\boldsymbol{A}(\boldsymbol{\theta})\mathbf{1}\boldsymbol{m}\boldsymbol{t},$$

and

$$E[x(t) - E(x(t))]^{2} = E(x(t)^{2}) - m^{2}t^{2} = -e\left(\frac{\partial^{2}}{\partial z^{2}}e^{tA(z)}\right)\mathbf{1} + \lambda'(0)^{2}t^{2}$$
$$= -teA''(0)\mathbf{1} - 2\sum_{k=2}^{\infty} \frac{t^{k}}{k!}eA'(0)A^{k-2}(0)A'(0)\mathbf{1} + \lambda'(0)^{2}t^{2}.$$

While, the second term of the last expression is, by virtue of (4, 3) and (4, 4), equal to $-\lambda'(0)^2 t^2 - 2e'(0)(\exp(tA(0)) - E - tA(0))t'(0)$. Thus, we see that the second equation of Lemma 3 is valid.

In the remainder of this section, let us discuss the case where N=2. In this case, A(0) can be written in the form $A(0) = \begin{pmatrix} -a & a \\ b-b \end{pmatrix}$ and by the assumption

that A(0) is irreducible, a > 0 and b > 0 hold. Therefore

$$A(0)^{n} = (-1)^{n-1}(a+b)^{n-1}A(0), n > 1,$$

and, by (4, 3), (4, 4) and (4, 6), we have

(4, 7)
$$v = -eA''(0)\mathbf{1} + \frac{2}{(a+b)^2}eA'(0)A(0)A'(0)\mathbf{1},$$

The expectation process introduced by Harris ([2] Chap. VII) related to the electron-photon cascade can be dealt with as a special case of our discussion with N=2.

Harris has conjectured that the central limit theorem holds for the expectation process, with m and v given by the right hand sides of the equations (4, 5) and (4, 7) respectively ([2] page 198). Theorem 2 and the above discussions show the validity of his conjecture.

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Tokyo University of Education and Nagoya University