Association between naturally occurring lithium in drinking water and suicide rates: a systematic review and meta-analysis of ecological studies

Anjum Memon, Imogen Rogers, Sophie M. D. D. Fitzsimmons, Ben Carter, Rebecca Strawbridge, Diego Hidalgo-Mazzei and Allan H. Young

Background
The prevalence of mental health conditions and national suicide rates are increasing in many countries. Lithium is widely and effectively used in pharmacological doses for the treatment and prevention of manic/depressive episodes, stabilising mood and reducing the risk of suicide. Since the 1990s, several ecological studies have tested the hypothesis that trace doses of naturally occurring lithium in drinking water may have a protective effect against suicide in the general population.

Aims
To synthesise the global evidence on the association between lithium levels in drinking water and suicide mortality rates.

Method
The MEDLINE, Embase, Web of Science and PsycINFO databases were searched to identify eligible ecological studies published between 1 January 1946 and 10 September 2018. Standardised regression coefficients for total (i.e. both genders combined), male and female suicide mortality rates were extracted and pooled using random-effects meta-analysis. The study was registered with PROSPERO (CRD42016041375).

Results
The literature search identified 415 articles; of these, 15 ecological studies were included in the synthesis. The random-effects meta-analysis showed a consistent protective (or inverse) association between lithium levels/concentration in publicly available drinking water and total (pooled $\beta = -0.27$, 95% CI $-0.47$ to $-0.08$; $P = 0.006$, $I^2 = 83.3\%$), male (pooled $\beta = -0.26$, 95% CI $-0.56$ to $0.03$; $P = 0.08$, $I^2 = 91.9\%$) and female (pooled $\beta = -0.13$, 95% CI $-0.24$ to $-0.02$; $P = 0.03$, $I^2 = 28.5\%$) suicide mortality rates. A similar protective association was observed in the six studies included in the narrative synthesis, and subgroup meta-analyses based on the higher/lower suicide mortality rates and lithium levels/concentration.

Conclusions
This synthesis of ecological studies, which are subject to the ecological fallacy/bias, supports the hypothesis that there is a protective (or inverse) association between lithium intakes from public drinking water and suicide mortality in the population level. Naturally occurring lithium in drinking water may have the potential to reduce the risk of suicide and may possibly help in mood stabilisation, particularly in populations with relatively high suicide rates and geographical areas with a greater range of lithium concentration in the drinking water. All the available evidence suggests that randomised community trials of lithium supplementation of the water supply might be a means of testing the hypothesis, particularly in communities (or settings) with demonstrated high prevalence of mental health conditions, violent criminal behaviour, chronic substance misuse and risk of suicide.

Keywords
Lithium in drinking water; suicide; ecological studies; epidemiology of suicide; suicide prevention.

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improvement in mood when compared with placebo, suggesting that doses of lithium considerably lower than those generally used in psychiatry may have the potential to influence mood and possibly reduce suicide risk.

**Naturally occurring lithium**

Lithium, sometimes referred to as the ‘magic ion’, is a naturally occurring element and is found in variable amounts in vegetables, grains, spices and drinking water. It is present in trace amounts in virtually all rocks and is mobilised by weathering into soils, ground and standing water, and thus into the public water supply in varying concentrations. In certain areas, particularly those close to sources of briny water such as northern Argentina, concentrations of up to 1000 μg/L have been reported. In the first ecological study on the subject, Schrauzer & Shrestha (1990) reported that the average incidence rates of suicide and violence (i.e. homicide and rape) in 27 counties of Texas, USA, over a 10-year period were consistently lower in counties with relatively high natural lithium levels in the drinking water compared with those with medium or low levels. On the basis of these findings, the authors hypothesised that lithium may exert a moderating effect on suicidal and violent criminal behaviour at levels that may be found in public water supplies. Since the publication of this report in 1990, a number of ecological studies from the USA, Japan and Europe have tested the hypothesis that trace doses of naturally occurring lithium in drinking water may have a protective effect against suicide in the general population.

Natural lithium is a mixture of two stable isotopes, lithium-6 and lithium-7. Lithium-7 accounts for over 92% of the natural abundance of the element. The health benefits and curative powers of naturally occurring lithium in water have been known for centuries. For example, the Lithia Springs (in Georgia, USA), an ancient Native American sacred medicinal spring, with its natural lithium-enriched water, is reputed for its health-giving properties and is the source of the brand called Lithia Spring Water. Lithium drinks were so much in demand in the early 20th century that, when the 7-Up commercial drink was created in 1929, it contained the element and was called Bib-Label Lithiated Lemon-Lime Soda. The US Food and Drug Administration banned the use of lithium in soft drinks and beer in 1948 and 7-Up was reformulated. It has been suggested that the 7 in 7-Up referred to the atomic mass of lithium and the Up referred to ‘mood or lithium lift’.

**Aims and objective**

The objective of the present study was to determine the association between lithium levels/concentration in public drinking water and local/regional suicide rates in the general population. To our knowledge, this is the most comprehensive synthesis of the epidemiological evidence on the subject.

**Method**

We conducted a systematic review and meta-analysis of the association between naturally occurring lithium content in publicly accessible sources of drinking water and local/regional suicide rates. This report is in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines. The study protocol was registered with PROSPERO (CRD42016041375, https://www.crd.york.ac.uk/prospero/display_record.php?RecordID=41375).

**Search strategy and selection criteria**

We searched four electronic databases (MEDLINE, Embase, Web of Science and PsyCINFO), without any language restrictions, for articles published between 1 January 1946 and 10 September 2018. The search terms included: lithium AND drinking water OR public water OR water OR tap water AND suicide OR mortality OR violent OR violence. The full search strategy used for Embase in Ovid is given in supplementary Fig. 1, available at https://doi.org/10.1192/bjp.2020.128.

Articles were eligible for inclusion if they were based on original ecological, population-based studies (i.e. analysing aggregate group data defined by geopolitical boundaries rather than individuals) that evaluated the association between lithium levels/concentrations in publicly accessible sources of drinking water and local/regional suicide mortality rates. Articles were excluded if the water samples for lithium measurement were obtained more than 10 years before or after the time period for which suicide rates were measured.

Titles and abstracts of the identified articles were screened by S.F. or I.R. for potentially relevant studies, and the full text was retrieved for articles identified at this screening stage. Two independent reviewers (either S.F. and A.M. or I.R. and A.M.) carried out assessment of the full-text articles for inclusion in the synthesis, and any disagreements were resolved by discussion. The references of all full-text articles and relevant review articles were also searched for additional studies. Authors of ongoing studies that fit the selection criteria were contacted to request unpublished data.

**Data extraction and analysis**

The relevant data from individual articles were extracted using a preconceived and standardised data extraction form. Information extracted included: first author’s name, year of publication, country and region, size of the population studied, methodology of lithium sampling and laboratory analysis, average lithium levels/concentrations in drinking water, mean total suicide mortality rate and/or standardised mortality ratio (SMR), time period of the suicide data, list of covariates that were adjusted for in the analysis, and outcome measure (statistical methods and effect size and s.e.). Corresponding authors were contacted for any data not presented in the published article. For two of the included studies, standardised regression coefficients were not included in the published articles, but were calculated from data supplied by the authors. Where only the standardised regression coefficient and P-value or t-value were reported, the s.e. of the standardised regression coefficient was estimated using the method of Altman and Bland (s.e. calculated by this method was adjusted for one study to account for the small sample size).

**Quality assessment of included studies**

To assess the quality of epidemiological ecological studies, we adapted the checklists produced by Tu & Ko and Betran et al to give five evaluation criteria, which were used to assess each study (supplementary Table 1). The evaluation criteria assessed whether the study participants were representative of the conclusions being drawn, whether statistical methods were used appropriately, whether confounders were adjusted for appropriately, whether key elements of the study design were presented and justified, and discussion of limitations such as the ecological fallacy. Quality assessments were performed by two independent reviewers (either S.F. and A.M. or I.R. and A.M.) and any discrepancies resolved by consensus. The results of the quality assessment were not used to decide on inclusion or exclusion of studies.
Data synthesis

Although we acknowledge the concerns with pooling non-randomised study designs, where the comparisons were deemed reasonably homogeneous the standardised regression coefficients and their corresponding s.e. were pooled using random-effects meta-analyses. Heterogeneity was assessed with $I^2$ and Cochran’s Q. Separate meta-analyses were conducted for total, male and female suicide rates. Where both adjusted and unadjusted regression coefficients were presented, the unadjusted regression coefficient was used in preference. The exposure to lithium was determined according to the lithium levels/concentration in drinking water, or the log lithium level; and the outcome variable was either a standardised regression coefficient of SMR or suicide mortality rate. All analyses were carried out in R 3.5.0 (run on the Windows platform) using the metafor package. Subgroup analyses were conducted to assess heterogeneity according to whether the study took place in a country where the mean age-adjusted total suicide mortality rate in the World Health Organization (WHO) 2015 data was either $\geq 10$ per 100 000 (high suicide rate countries, i.e. Japan, Austria, Lithuania and the USA), or $<10$ per 100 000 (low suicide rate countries, i.e. Greece, Italy and the UK), and according to the highest observed lithium levels in the sampled drinking water ($\geq 80$ μg/L or $<80$ μg/L). Quality assessment was also used to explain heterogeneity between studies. In the primary meta-analyses (all eligible studies of total, male and female suicide rates), we also performed a sensitivity analysis to evaluate robustness and stability by sequentially omitting one study at a time.

Results

A flowchart describing the study selection process is given in Fig. 1. The literature search identified 415 articles – after exclusion of duplicate titles and abstracts, 260 articles were screened and 41 were selected for full-text evaluation. Of these, 15 articles fulfilled the selection criteria and were included in the synthesis, 4 studies were conducted in Japan, 4 in Austria, 3 in USA, and 1 each in England, Greece, Italy and Lithuania (including lithium levels in drinking water samples and corresponding suicide mortality rates for 1286 regions/counties/cities).

In the included ecological studies, drinking water samples were collected from publicly available tap-water sources in a variety of locations or from public wells. Most water samples were analysed by mass spectrometry or inductively coupled plasma emission spectroscopy. Three studies used water supply company’s data on water composition. Eleven studies controlled for one or more confounding variables in the analyses; and six studies controlled for a measure of socioeconomic status (e.g. proportion of population with a college degree, unemployment rate, average income). A variety of other possible confounders were adjusted for in different studies, including measures of local climate, altitude and density of medical professionals. Suicide mortality data were obtained from government statistics in all the studies and covered time periods of between 1 year and 11 years. All studies included the overall male suicide mortality rates (pooled $\beta = -0.27$, 95% CI $-0.47$ to $-0.08$; $P = 0.006$, $I^2 = 83.3$%). To assess the impact of relatively high heterogeneity between the studies, we conducted a sensitivity analysis. The heterogeneity was reduced most on excluding the study by Liaugaudaite et al., and the meta-analysis yielded a pooled $\beta = -0.17$ (95% CI $-0.25$ to $-0.09$; $P = 0.001$, $I^2 = 85.0$%). We obtained a similar result ($\beta = -0.25$, 95% CI $-0.50$ to $-0.01$; $P = 0.04$, $I^2 = 86.0$%) when we repeated the meta-analysis including only the studies considered to be of high quality.

Study quality

Of the nine studies included in the meta-analysis, three fulfilled all five of the evaluation criteria, four fulfilled four of the criteria, one fulfilled three and one fulfilled two. Failure to adjust for covariates was the most common methodological omission. Studies fulfilling at least four of the criteria were considered to be of high quality. Details of the quality assessment are given in supplementary Table 1.

Meta-analysis of association between lithium levels in drinking water and suicide mortality rates

The studies included in the meta-analysis are summarised in Table 1. Eight studies that reported the association between lithium levels in drinking water and suicide mortality as a standardised regression coefficient of either SMR or suicide rates on log lithium levels were eligible for inclusion in the meta-analysis. One further study reported the standardised regression coefficient of age-standardised suicide mortality rate on untransformed lithium levels – the meta-analysis was repeated with and without the inclusion of this study. Four of the articles identified in the literature search were multiple analyses carried out on the same study data – only the original study was included in the meta-analysis. Three studies conducted in Japan had overlapping drinking-water sample areas – only the study with the largest sample size was included in the meta-analysis. One study examined suicide rates over three 10-year periods – only the results for the last period were included in the meta-analysis, as this overlapped with the time when the drinking-water samples were collected. Two of the included studies were conducted in Japan, two in the USA and one each in England, Austria, Greece, Italy and Lithuania. A range of different population sizes and municipalities were evaluated, from entire countries to individual provinces and prefectures. Two studies excluded suicides that were not registered in cities, otherwise all suicides in the regions studied were included.

Total (i.e. both genders combined) suicide mortality rate

Eight eligible studies examined the association between lithium levels in drinking water and total suicide mortality rates – seven of these studies reported a protective (i.e. inverse or negative) association between lithium levels and total suicide rates, which was statistically significant in five of the seven studies (Fig. 2). The random-effects meta-analysis showed a statistically significant protective (or inverse) association between lithium levels and total suicide rates (pooled $\beta = -0.27$, 95% CI $-0.47$ to $-0.08$; $P = 0.006$, $I^2 = 83.3$%). To assess the impact of relatively high heterogeneity between the studies, we conducted a sensitivity analysis. The heterogeneity was reduced most on excluding the study by Liaugaudaite et al., and the meta-analysis yielded a pooled $\beta = -0.17$ (95% CI $-0.25$ to $-0.09$; $P = 0.001$, $I^2 = 85.0$%). We obtained a similar result ($\beta = -0.25$, 95% CI $-0.50$ to $-0.01$; $P = 0.04$, $I^2 = 86.0$%) when we repeated the meta-analysis including only the studies considered to be of high quality.

Male suicide mortality rate

Seven eligible studies examined the association between lithium levels in drinking water and male suicide mortality rates – five of these studies reported a protective (or inverse) association between lithium levels and male suicide rates, which was statistically
significant in three of the five studies (Fig. 3). On the other hand, two studies reported a non-significant positive association. The random-effects meta-analysis showed a protective (or inverse), but statistically non-significant, association between lithium levels and male suicide rates (pooled \( \beta = -0.26, 95\% \text{ CI} -0.56 \text{ to } 0.03; P = 0.08, I^2 = 91.9\% \)). In the sensitivity analyses, the heterogeneity was reduced most on excluding the study by Liaugaudaite et al., and the meta-analysis yielded a pooled \( \beta \) of \(-0.12 (95\% \text{ CI} 0.28\text{ to } -0.03; P = 0.13, I^2 = 64.0\% \)). We obtained a similar result (pooled \( \beta = -0.21, 95\% \text{ CI} -0.53 \text{ to } 0.10; P = 0.19, I^2 = 93.0\% \)) when we repeated the meta-analysis including only the studies considered to be of high quality.

Female suicide mortality rate

Seven eligible studies examined the association between lithium levels in drinking water and female suicide mortality rates – five of these studies reported a protective (or inverse) association between lithium levels and female suicide rates, which was statistically significant in two of the five studies (Fig. 4). On the other hand, two studies reported a non-significant positive association. The random-effects meta-analysis showed a statistically significant protective (or inverse) association between lithium levels and female suicide rates (pooled \( \beta = -0.13, 95\% \text{ CI} -0.24 \text{ to } -0.02; P = 0.03, I^2 = 28.5\% \)). In the sensitivity analyses, the heterogeneity was reduced most on excluding the study by Shiotsuki et al., and the meta-analysis yielded a pooled \( \beta \) of \(-0.17 (95\% \text{ CI} -0.28 \text{ to } -0.07; P = 0.001, I^2 = 0.02\% \)). We obtained a similar result (pooled \( \beta = -0.11, 95\% \text{ CI} -0.22 \text{ to } -0.01; P = 0.05, I^2 = 26.0\% \)) when we repeated the meta-analysis including only the studies considered to be of high quality.

Meta-analysis by higher/lower total suicide mortality rates, according to the WHO data

The random-effects meta-analysis, including five studies from countries with higher suicide rates, showed a statistically significant protective (or inverse) association between lithium levels and total suicide rates (pooled \( \beta = -0.40, 95\% \text{ CI} -0.68 \text{ to } -0.12; P = 0.005, I^2 = 86.0\% \)); including only the three studies from countries with lower suicide rates yielded a pooled \( \beta \) of \(-0.11 (95\% \text{ CI} -0.23 \text{ to } -0.01; P = 0.08, I^2 = 0.0\% \)), with no heterogeneity (supplementary Fig. 2).

Meta-analysis by higher/lower observed lithium levels in drinking water

There were three studies of total suicide mortality rate where the highest observed lithium level in drinking water was \( \geq 80 \mu g/L \).
<table>
<thead>
<tr>
<th>Study</th>
<th>Region and number of locations</th>
<th>Population data</th>
<th>Lithium samples, n; dates and methods of collection, analysis method</th>
<th>Lithium levels (μg/L), mean (range)</th>
<th>Mean suicide rate (per 100 000/year)</th>
<th>SMR, mean (range)</th>
<th>Covariates</th>
<th>Statistical methods and results</th>
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</thead>
<tbody>
<tr>
<td>Kapusta et al, 2011</td>
<td>Austria, 99 districts</td>
<td>All suicides 2005–2009 (total population 8 297 964)</td>
<td>n = 6460; 2005–2010, local drinking water; ICP-OES</td>
<td>11.3 (3.3–82.3)</td>
<td>T: 16.5&lt;sup&gt;b&lt;/sup&gt;</td>
<td>M: 26.4&lt;sup&gt;b&lt;/sup&gt;</td>
<td>F: 7.00&lt;sup&gt;b&lt;/sup&gt;</td>
<td>PWLS regression of SMR on log Li (μg/L)</td>
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<td>Blümle et al, 2013</td>
<td>USA, Texas, 226 counties</td>
<td>All suicides 1999–2007 (total population n.r.)</td>
<td>n = 3123; 1999–2007, from public wells; n.r.</td>
<td>46.3&lt;sup&gt;c&lt;/sup&gt; (2.8–219.0)</td>
<td>13.16&lt;sup&gt;c&lt;/sup&gt;</td>
<td>n.r.</td>
<td>Population density, age, proportion of females, African Americans, Hispanics and Latino Americans, median income per household, poverty, unemployment</td>
<td>PW linear regression of age-standardised suicide rate on log Li</td>
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<td>Giotakis et al, 2013</td>
<td>Greece, 34 prefectures</td>
<td>All suicides 1999–2010 (total population n.r.)</td>
<td>n = 149; 2012; mass spectrometry</td>
<td>11.10 (0.1–121)</td>
<td>n.r.</td>
<td>n.r.</td>
<td>None</td>
<td>Linear regression of age-standardised suicide rate on Li (μg/L)</td>
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<td>Sugawara et al, 2013</td>
<td>Japan, Aomori prefecture, 40 municipalities</td>
<td>All suicides 2008–2010 (total population 1 373 339)</td>
<td>n.r.; n.r.; mass spectrometry</td>
<td>n.r. (0.0 – 12.9)</td>
<td>M: 123 (96–186); F: 105 (72–152)</td>
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<td>Density of medical institutions, unemployment rate</td>
<td>PWLS regression of SMR on log Li</td>
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<th>Study</th>
<th>Region and number of locations</th>
<th>Population data</th>
<th>Lithium samples, n; dates and methods of collection; analysis method</th>
<th>Lithium levels (μg/L), mean (range)</th>
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<td>Unadjusted analyses: T: ( \beta &lt; 0.001, P = 0.997; ) M: ( \beta = 0.046, P = 0.581; ) F: ( \beta = -0.134, P = 0.109 )</td>
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<td>Adjusted analyses: T: ( \beta = 0.079, P = 0.308; ) M: ( \beta = 0.107, P = 0.159; ) F: ( \beta = -0.032, P = 0.703 )</td>
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<td>Unadjusted analyses: T: ( \beta = -0.047, P = 0.578; ) M: ( \beta = -0.099, P = 0.915; ) F: ( \beta = -0.165, P = 0.047 )</td>
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<td>Adjusted analyses: T: ( \beta = 0.079, P = 0.323; ) M: ( \beta = 0.087, P = 0.280; ) F: ( \beta = 0.001, P = 0.998 )</td>
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<td>Unadjusted analyses: T: ( \beta = -0.234, P = 0.005; ) M: ( \beta = -0.161, P = 0.053; ) F: ( \beta = -0.339, P = 0.001 )</td>
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<td>Adjusted analyses: T: ( \beta = -0.044, P = 0.560; ) M: ( \beta = 0.013, P = 0.859; ) F: ( \beta = -0.154, P = 0.043 )</td>
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<td>Unadjusted analyses: T: ( \beta = -0.129, P = 0.070; ) M: ( \beta = -0.164, P = 0.037; ) F: ( \beta = 0.014, P = 0.870 )</td>
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<sup>25</sup> Memon et al, 2015. 672
<table>
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<tr>
<th>Study</th>
<th>Region and number of locations</th>
<th>Population data</th>
<th>Lithium samples, n; dates and methods of collection; analysis method</th>
<th>Lithium levels (μg/L), mean (range)</th>
<th>Mean suicide rate (per 100 000/year)</th>
<th>SMR, mean (range)</th>
<th>Covariates</th>
<th>Statistical methods and results</th>
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<tr>
<td>Liaugaudaite et al, 2017[27]</td>
<td>Lithuania, 9 cities</td>
<td>All suicides 2009–2013 (total population 1 109,261)</td>
<td>n = 22; Nov 2013 to Jan 2014; mass spectrometry</td>
<td>10.9 (0.48–35.53)</td>
<td>T: 27 (range 16–50)</td>
<td>n.r.</td>
<td>Female:male ratio of city population</td>
<td>PWLS regression of log Li on age-standardised suicide rate</td>
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<td>T: $\beta$ (se) = −0.911 (0.156); $\text{P} = 0.034$; ($\text{P} &lt; 0.05$)</td>
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<td>M: $\beta$ (se) = −0.965 (0.100); $\text{P} = 0.013$; ($\text{P} &lt; 0.05$)</td>
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<td>F: $\beta$ (se) = 0.150 (0.374); $\text{P} = 0.523$</td>
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<td>T: $\beta$ = −0.283, $\text{P} = 0.034$; ($\text{P} &lt; 0.05$)</td>
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<td>M: $\beta$ = −0.702, $\text{P} = 0.013$; ($\text{P} &lt; 0.05$)</td>
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<td>F: $\beta$ = 0.253, $\text{P} = 0.523$</td>
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<tr>
<td>Palmer et al, 2019[13]</td>
<td>USA, Alabama, 15 counties</td>
<td>Average suicide rate 1999–2013 (total population n.r.)</td>
<td>n = 75; May 2016 from public locations; plasma emission spectrophotometry</td>
<td>n.r. (0.4–32.9)</td>
<td>n.r. (range 3.3–22.0)</td>
<td>n.r.</td>
<td>None</td>
<td>Spearman’s correlation of age-standardised suicide rate against Li levels</td>
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<td>T: $r = −0.6286$, $\text{P} = 0.0141$</td>
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<td>M: $r = −0.625$, $\text{P} = 0.0148$</td>
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<td>F: $r = −0.4393$, $\text{P} = 0.1032$</td>
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<td>T: $\beta$ (se) = −0.6188 (0.2179); $\text{P} = 0.1032$</td>
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<td>M: $\beta$ (se) = −0.6236 (0.2168); $\text{P} = 0.1032$</td>
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<td>F: $\beta$ (se) = −0.4387 (0.2427); $\text{P} = 0.1032$</td>
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n.r., not reported; SMR, standardised mortality ratio; GP, general practitioner; ICP-OES, inductively coupled plasma optical emission spectrometry; PWLS, population-weighted least squares; T, total (i.e. both genders combined); M, male; F, female; Li, lithium.

* a. Personal communication from Nikolett Kabacs.
* b. Calculated from data supplied by Nikolett Kabacs.
* c. Personal communication from Nestor Kapusta.
* d. SMR has reference 1.00 not 100.
* e. i.e. from personal communication from Vilma Liaugaudaite.
* f. Calculated from data supplied by Greg Gorman.
supplementary Fig. 3). The random-effects meta-analysis showed a statistically significant protective (or inverse) association with a pooled $\beta$ of $-0.20$ (95% CI $-0.31$ to $-0.10$; $P = 0.0002$, $I^2 = 0.0\%$), with no heterogeneity. For the five studies conducted in regions with lithium levels $<80\ \mu g/L$, the pooled $\beta$ was $-0.33$ (95% CI $-0.68$ to 0.01; $P = 0.06$, $I^2 = 87.0\%$).

**Articles included in the narrative synthesis only**

Six articles provided an estimate of the association between lithium levels in drinking water and suicide mortality rates but were not eligible for inclusion in the meta-analysis (supplementary Table 2). Three were additional analyses/reports on the Austrian data-set, which controlled for a variety of additional confounding variables, including altitude and rates of lithium prescriptions. Two studies from Japan$^{23,34}$ were conducted in areas that overlapped with the study area for Shiotsuki et al.$^{29}$ There was also a study that compared suicide rates in Texas, USA, in counties with relatively high, medium or low lithium concentrations in drinking water.$^{26}$ All of these studies found statistically significant protective (or inverse) associations between lithium levels and total suicide rates; where data on gender-specific rates were presented, protective

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**Fig. 2** Random-effects meta-analysis of the association between lithium concentration in drinking water and total suicide rates.

Standardised regression coefficients ($\beta$) for each study are presented as squares, with the position of the square corresponding to the $\beta$ and the 95% CI shown by horizontal lines. 95% CIs for each study shown in the forest plot are obtained by back-transformation using the calculated s.e. used in the analysis and do not always conform exactly to the stated confidence intervals in the paper. The area of the square is inversely proportional to the variance in $\beta$. The diamond represents the pooled $\beta$ and corresponding 95% CI.

<table>
<thead>
<tr>
<th>Study</th>
<th>$\beta$ (95% CI)</th>
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<tbody>
<tr>
<td>Kabacs et al 2011, UK</td>
<td>$-0.06$ ($-0.35$ to $0.22$)</td>
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<tr>
<td>Kapusta et al 2011, Austria</td>
<td>$-0.22$ ($-0.42$ to $-0.02$)</td>
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<tr>
<td>Blüml et al 2013, USA</td>
<td>$-0.24$ ($-0.45$ to $-0.03$)</td>
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<tr>
<td>Giotakos et al 2013, Greece</td>
<td>$-0.17$ ($-0.33$ to $-0.01$)</td>
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<tr>
<td>Pompili et al 2015, Italy</td>
<td>$0.00$ ($-0.24$ to $0.24$)</td>
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<tr>
<td>Shiotsuki et al 2016, Japan</td>
<td>$-0.15$ ($-0.31$ to $0.01$)</td>
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<tr>
<td>Liaugaudaitė et al 2017, Lithuania</td>
<td>$-0.91$ ($-1.22$ to $-0.61$)</td>
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<tr>
<td>Palmer et al 2019, USA</td>
<td>$-0.62$ ($-1.05$ to $-0.19$)</td>
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| RE Model ($I^2 = 83.3\%$)               | $-0.27$ ($-0.47$ to $-0.08$) |

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**Fig. 3** Random-effects meta-analysis of the association between lithium concentration in drinking water and male suicide rates.

Standardised regression coefficients ($\beta$) for each study are presented as squares, with the position of the square corresponding to the $\beta$ and the 95% CI shown by horizontal lines. 95% CIs for each study shown in the forest plot are obtained by back-transformation using the calculated s.e. used in the analysis and do not always conform exactly to the stated confidence intervals in the paper. The area of the square is inversely proportional to the variance in $\beta$. The diamond represents the pooled $\beta$ and corresponding 95% CI.

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<th>Study</th>
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<tr>
<td>Kabacs et al 2011, UK</td>
<td>$-0.06$ ($-0.34$ to $0.22$)</td>
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<tr>
<td>Kapusta et al 2011, Austria</td>
<td>$-0.18$ ($-0.38$ to $0.02$)</td>
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<td>Sugawara et al 2013, Japan</td>
<td>$0.14$ ($-0.19$ to $0.46$)</td>
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<td>Pompili et al 2015, Italy</td>
<td>$0.05$ ($-0.11$ to $0.20$)</td>
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<td>Shiotsuki et al 2016, Japan</td>
<td>$-0.22$ ($-0.38$ to $-0.07$)</td>
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<td>Liaugaudaitė et al 2017, Lithuania</td>
<td>$-0.96$ ($-1.16$ to $-0.77$)</td>
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<tr>
<td>Palmer et al 2019, USA</td>
<td>$-0.62$ ($-1.05$ to $-0.20$)</td>
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| RE Model ($I^2 = 91.9\%$)               | $-0.26$ ($-0.56$ to $0.03$) |
associations were also found with male suicide rates but not with female suicide rates.

**Discussion**

This is the first meta-analysis of the ecological association between the lithium levels/concentration in publicly available drinking water and the incidence of suicide. We found a consistent protective (or inverse) association between lithium levels and total, male and female suicide mortality rates, which was statistically significant for total and female suicide rates. Similar protective association was observed in the six studies included in the narrative synthesis, and in the subgroup meta-analyses based on the higher/lower suicide mortality rates (≥10 per 100,000, <10 per 100,000) and lithium levels (≥280 μg/L, <80 μg/L). These findings, which are consistent with the finding in clinical trials that lithium reduces suicide and related behaviours in people with a mood disorder, suggest that naturally occurring lithium in drinking water may have the potential to reduce the risk of suicide and may possibly help in mood stabilization, particularly in populations with relatively high suicide rates and geographical areas with a greater range of lithium concentration in the drinking water.

**Strengths and limitations**

To our knowledge, this is the most comprehensive synthesis of the ecological association between lithium levels/concentrations in publicly available drinking water and suicide mortality rates. The main limitations of epidemiological ecological studies are that they are observational and are subject to the ecological fallacy (or ecological bias). Ecological studies are essentially conducted to generate hypotheses. They compare aggregate exposure (e.g. population exposed to lithium in drinking water) and disease/health outcome (e.g. suicide mortality rate in the exposed population) across different populations over the same time period or within the same population over time. They are subject to confounding as information on potential confounder(s) may not be available and associations at the population level do not necessarily represent associations at the individual level (ecological fallacy). Populations may also differ in terms of ethnic, religious and social class distribution, prevalence and management of mental disorders, and mobility patterns. As with all systematic reviews and meta-analyses, our study may also be subject to reporting/publication bias. These biases arise when the dissemination of research findings is influenced by the nature and direction of results – statistically significant ‘positive’ results are more likely to be published and cited, whereas non-statistically significant results may be filtered, manipulated or presented in such a way that they become/seem positive. We were unable to conduct a formal assessment of possible publication bias, as our meta-analyses did not meet a key assumption (i.e. a minimum of 10 studies) for the tests of funnel plot asymmetry.

Although most of the studies were assessed to be of good quality using our adapted criteria, and our results were unaltered on excluding the two studies judged to be of lower quality, no standardised criteria for quality assessment of ecological studies were available. The studies varied widely in the range of years for which the suicide rates were measured, and in the number and timing of drinking-water samples taken. Some of the studies utilised drinking-water samples and suicide data from different time periods, thus relying on the assumption that lithium levels do not fluctuate substantially over time. In a study from Japan, Ohgami et al found negligible change in lithium levels in drinking water when they repeated the measurements after 1 year. A study from Denmark also found little variability over time in lithium levels collected from ground water between 1947 and 2012. However, additional evidence on stability over time of lithium levels in drinking water is needed from other geographical areas.

In a nationwide closed (or non-dynamic) historical (or retrospective) cohort study published in 2017, Knudsen et al linked individual-level register-based data on the entire Danish adult population (3.7 million individuals) from 1991 to 2012 with a moving 5-year time-weighted average lithium exposure level from drinking water. Cases of suicides were identified through the Danish Register of Causes of Death and lithium levels in drinking water were ascertained from 158 water samples obtained from 151 public waterworks supplying approximately 42% of the Danish population. Of these, 139 samples were collected via a drinking-water sampling campaign during April–June 2013 (spatially covering the entire country) and 19 samples were collected via a separate campaign at the Greater Copenhagen Utility between October and June 2013.
Possible mechanism(s) behind the association between lithium in drinking water and suicide

A possible mechanism by which lithium in drinking water might prevent suicide is by ameliorating the symptoms of mood disorders. In a randomised controlled trial involving former drug users a 400 μg daily dose of lithium was found to improve mood, with peak effect after 4 weeks of administration. There may also be a cumulative effect of sustained low-dose exposure over longer time periods, as lithium levels in drinking water have been shown to correlate with serum levels of lithium in areas of high drinking-water lithium concentrations. There is little information on the association between lithium in drinking water and the prevalence and severity of mood disorders. In a population-based nested case–control study from Denmark, higher long-term lithium exposure from drinking water was not associated with a lower incidence of mania/bipolar disorder – suggesting that long-term exposure to micro-doses of lithium does not modulate the risk of these conditions. On the other hand, in another population-based nested case–control study by the same research group, higher long-term lithium exposure from drinking water was associated with a lower incidence of dementia in a non-linear way.

It is also possible that trace doses of lithium might reduce suicide rates via its anti-aggressive effects. Therapeutic/pharmacological doses of lithium have been found to reduce aggressive/violent behaviour in a variety of populations, and inverse associations have been observed between lithium levels in drinking water and rates of violent crimes. Recently, a similar cross-sectional study of adolescents in Kochi prefecture in Japan found an inverse association between the lithium content of drinking water available to schools and interpersonal violence and depressive symptoms among adolescents. As violent methods of suicide are more likely to be lethal, lithium ingestion might be expected to reduce the lethality of suicide attempts. This was observed in a recent meta-analysis of the effects of long-term lithium treatment of people with major affective disorder, where the incidence ratio of attempted/completed suicides increased 2.5-fold with lithium treatment, indicating a considerably reduced lethality. The meta-analysis also showed that the risks of attempted and completed suicide were consistently lower (by about 80%) in people with bipolar and other major affective disorders treated with pharmacological doses of lithium for an average of 18 months.

Recommendations for future research

Given that our results suggest a possible protective (or inverse) association between lithium levels in drinking water and suicide mortality at the population level, randomised community trials of lithium supplementation of the water supply might be a possible means of testing the hypothesis, particularly in communities (or settings) with demonstrated high prevalence of mental health conditions, violent criminal behaviour, drug dependency and chronic substance misuse and risk of suicide. This may provide further evidence to support the hypothesis that lithium could be used at the community/population level to reduce or combat the risk of these conditions. It may also be possible to measure lithium levels in people who report to emergency departments with self-harm/suicidal ideation and at inquest for suicide. The synthesised evidence suggests that the protective (or inverse) association between lithium levels in drinking water and suicide mortality rates is likely to be stronger in populations with relatively higher suicide mortality rates. Future studies might benefit by concentrating on areas with a wide range of lithium levels/concentration in drinking water and relatively high suicide mortality rates and would ideally also consider the effect and levels of dietary lithium from food sources.

2009 and June 2010. The lithium levels were measured at a single point in time (i.e. cross-sectional) and it was assumed that they had remained constant over the 22-year study period (from 1991 to 2012). The mean lithium level in the drinking-water samples was 11.6 μg/L (range 0.6–30.7 μg/L). The overall national suicide rate decreased by 38% during the study period – from 29.7 per 100 000 in 1991 to 18.4 per 100 000 in 2012. The spatial regression analysis was adjusted for confounding factors (including gender, age, employment/civil status and calendar year). The study found no statistically significant association between increasing 5-year time-weighted average lithium levels and decreasing suicide rate – although all the incidence rate ratios were ≤1. The authors concluded that there did not seem to be a protective effect of exposure to lithium on the incidence of suicide with levels below 31 μg/L in drinking water. They also noted that in the previous (i.e. ecological) studies that found a significant protective association, the lithium exposure levels were relatively much higher than those found in their study, and the lack of variation in lithium levels in their study may have ‘challenged’ their analyses. The authors acknowledged that the study had several limitations, as it was based on a single cross-sectional measurement of lithium levels from 151 waterworks supplying approximately half of the Danish population; and suggested that future studies could prospectively determine lithium levels from more waterworks on regular basis for a number of years. It is, nevertheless, challenging that the only published study linking exposure to lithium in drinking water with suicide on an individual level was negative.

It has been suggested that the association between lithium in drinking water and suicide might be modified by the rate of lithium prescriptions in the population, as excreted lithium might find its way into the water supply. However, Helbich et al found no evidence that the rate of lithium prescriptions in an area affected the association between lithium levels in drinking water and suicide rates (either directly or via an effect on the mean lithium level in the water). It is also noteworthy that some vegetables, grains/nuts, meat and spices are also rich in lithium, but it is difficult to measure dietary lithium as the content is variable and relates to lithium content in ground water/soil. It is therefore likely that lithium intake from food will be relatively higher in areas that have high lithium content in both ground water and the public water supply. It has been estimated that the average daily intake of lithium from food by adults in the USA ranges from 650 to 3100 μg, whereas in areas of Texas, which have unusually high lithium levels in the public water supply, the contribution from drinking water might be around 340 μg. The association between exposure to dietary lithium and incidence of suicide has not been investigated. Furthermore, bottled drinking water (processed/treated or natural mineral water from springs) often has a much higher lithium content than tap water – the association between exposure to lithium via bottled water and suicide has not been studied.
Supplementary material

Supplementary material is available online at http://doi.org/10.1192/bjp.2020.128.

Data availability

The data that support the findings of this study are available from the corresponding author (A.M.) on reasonable request. The corresponding author had full access to all the data in the study.

Author contributions

A.M. conceived and designed the study, I.R., S.F. and A.M. did the literature search and quality assessment of included studies. I.R. and S.F. selected the studies and extracted the relevant information. I.R. and A.M. synthesised the data. A.M. and I.R. wrote the manuscript. B.C., R.S., D.H.-M. and A.H.Y. critically reviewed the manuscript for intellectual content. All authors approved the final version of the manuscript for publication. A.M. guided and supervised the overall work, and had final responsibility for the decision to submit for publication.

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Declaration of interest

A.H.Y. has received payment for lectures and advisory board membership from AstraZeneca, Eli Lilly, Lundbeck, Sunovion, Servier, Livovara and Janssen; he is the lead investigator for the Embeded Study AUG, BCG Neuroplasticity Study and Aspirina Mania Study, and has led investigator-initiated studies from AstraZeneca, Eli Lilly, Lundbeck, Wyeth and Janssen. R.S. has received an honorarium for speaking at a Lundbeck-sponsored event. investigator-initiated studies from AstraZeneca, Eli Lilly, Lundbeck, Wyeth and Janssen. R.S., D.H.-M. and A.H.Y. critically reviewed the manuscript for intellectual content. All authors approved the final version of the manuscript for publication. A.M. guided and supervised the overall work, and had final responsibility for the decision to submit for publication.

References

15 Altman DG, Bland JM. How to obtain the confidence interval from a P value. BMJ 2011; 343.
Psychiatry and ‘pop’ culture: millennials for mental health

Joel Philip and Vinu Cherian

Pop is generally considered to be the more ‘vanilla’ of musical genres – light-hearted, superficial and created for mass consumption. However, a new generation of millennial pop artistes is now using the platform accorded to them to open up honest conversations about mental health among young people, by writing songs that chronicle their own mental health struggles. Julia Michaels is one such talented singer, who burst onto the global pop music scene in 2017. Her extended plays (EPs), aptly named Nervous System and Inner Monologues, as well as hit singles such as ‘Issues’, are creative works based on her own experiences with anxiety. In her songs, Julia is vocal and brutally honest about her struggles with panic attacks, anxiety and depression. In several interviews, she has spoken about how she spent her initial years in the industry writing chart-topping hits for other artistes, held back by fears of performing on stage and insecurities about how her songs would be received. Today, she gives hope to millions of her fans who suffer the same symptoms of anxiety, by being raw and unapologetic in her lyrical work and encouraging them to seek help and support.

Julia recently collaborated with Selena Gomez to release a single titled ‘Anxiety’. Selena is another young entertainer who has spoken out about the anxiety and depressive episodes that she suffers, partly due to her diagnosis of systemic lupus erythematosus. She has been a long-standing advocate of mental health, speaking openly about her time as an in-patient in a mental health facility, where she voluntarily checked in to seek help for her mental health problems, and how medication and dialectical behaviour therapy have helped her cope with her symptoms. On their collaboration, Julia details how anxiety and depression affect on her social life and relationships:

‘My friends, they wanna take me to the movies,
But I’m holding hands with my depression,
And right when I think I’ve overcome it,
Anxiety starts kicking in.’

In the bridge, Selena joins in as they put into words how millions of people who struggle with anxiety feel every day:

‘I get all these thoughts, running through my mind,
All the damn time, and I can’t seem to shut it off.’

These words hit home for fans who suffer their mental health problems in silence and help to dispel the stigma surrounding a psychiatric diagnosis. The role played by pop artistes and celebrities in de-stigmatising mental illness is worthy of recognition by the psychiatric fraternity. The singer Kehlani, who was admitted to hospital following a suicide attempt and now encourages young people to speak up and seek help for their mental health problems, summed up the sentiment nicely in her lyrics:

‘It’s okay to not be okay,
And it’s alright to not be alright.’

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