

Invited Article

Cite this article: Mulvey, L. P. A., M. C. Nikolic, B. J. Allen, T. A. Heath, and R. C. M. Warnock (2025). From fossils to phylogenies: exploring the integration of paleontological data into Bayesian phylogenetic inference. *Paleobiology* **51**, 214–236.

<https://doi.org/10.1017/pab.2024.47>

Received: 28 March 2024

Revised: 03 September 2024

Accepted: 10 September 2024

Handling Editor:

Mark Patzkowsky

Corresponding author:

Rachel C. M. Warnock;

Email: rachel.warnock@fau.de

From fossils to phylogenies: exploring the integration of paleontological data into Bayesian phylogenetic inference

Laura P. A. Mulvey¹ , Mark C. Nikolic², Bethany J. Allen^{3,4} , Tracy A. Heath⁵ and Rachel C. M. Warnock¹ 

¹GeoZentrum Nordbayern, Department of Geography and Geosciences, Friedrich-Alexander Universität Erlangen-Nürnberg, Erlangen, Germany

²Division of Paleontology (Invertebrates), American Museum of Natural History, New York, New York, U.S.A

³Department of Biosystems Science and Engineering, ETH Zurich, Basel, Switzerland

⁴Computational Evolution Group, Swiss Institute of Bioinformatics, Quartier Sorge, Bâtiment Amphipôle, Lausanne, Switzerland

⁵Department of Ecology, Evolution, & Organismal Biology, Iowa State University, Ames, Iowa, U.S.A

Abstract

Incorporating paleontological data into phylogenetic inference can greatly enrich our understanding of evolutionary relationships by providing insights into the diversity and morphological evolution of a clade over geological timescales. Phylogenetic analysis of fossil data has been significantly aided by the introduction of the fossilized birth–death (FBD) process, a model that accounts for fossil sampling through time. A decade on from the first implementation of the FBD model, we explore its use in more than 170 empirical studies, summarizing insights gained through its application. We identify a number of challenges in applying the model in practice: it requires a working knowledge of paleontological data and their complex properties, Bayesian phylogenetics, and the mechanics of evolutionary models. To address some of these difficulties, we provide an introduction to the Bayesian phylogenetic framework, discuss important aspects of paleontological data, and finally describe the assumptions of the models used in paleobiology. We also present a number of exemplar empirical studies that have used the FBD model in different ways. Through this review, we aim to provide clarity on how paleontological data can best be used in phylogenetic inference. We hope to encourage communication between model developers and empirical researchers, with the ultimate goal of developing models that better reflect the data we have and the processes that generated them.

Non-technical Summary

Reconstructing evolutionary relationships among organisms provides important insight into the history of life on Earth. Evolutionary (phylogenetic) trees can be used to show the relationships between extinct and extant organisms. By incorporating fossils, we can then estimate the timing of significant events such as a speciation event. The models used to generate trees in paleobiology combine different data sources, including molecular sequences from living organisms, the ages of fossils, and morphological information from both. Using this framework, we can learn about the rate of evolutionary dynamics of organisms, for example, the rate of speciation or diversification, or geographic movements through time. In this review article, we describe details of the statistical modeling framework used to integrate observations from living and fossil taxa. In particular, we focus on the use of the fossilized birth–death process, which is the only available model that allows us to include knowledge about the structure of the paleontological record into phylogenetic analyses. Because not all organisms and environments are equally well preserved, a flexible framework for working with fossils is essential for obtaining reliably dated phylogenies. In the decade since the model first became available in phylogenetic software, it has been used in more than 170 studies. We celebrate different applications of the model and highlight practical challenges. An important point that emerges from our discussion is that both the complexity of fossil data and details of the assumptions made by different models are crucial to consider. We hope to stimulate the exchange of ideas between researchers collecting and curating paleontological data and those developing models and software, with a view to further improving approaches to studying evolution in deep time.

© The Author(s), 2025. Published by Cambridge University Press on behalf of Paleontological Society. This is an Open Access article, distributed under the terms of the Creative Commons Attribution licence (<http://creativecommons.org/licenses/by/4.0/>), which permits unrestricted re-use, distribution and reproduction, provided the original article is properly cited.

PALEOBIOLOGY
A PUBLICATION OF THE
 PALEONTOLOGICAL SOCIETY

 **CAMBRIDGE**
UNIVERSITY PRESS

Introduction

Phylogenetic analysis has been widely utilized in paleontological research, as it provides an intuitive method for understanding the relationships among species (López-Antónanzas et al. 2022).



Not only can a phylogenetic framework guide taxonomic classification, but it can also be used to estimate divergence times (Wolfe et al. 2023); rates of diversification (Herrera 2017; Tougaard 2017); and various other evolutionary dynamics, such as biogeographic patterns (Tavares et al. 2018; Landis et al. 2021; de Faria Santos et al. 2022; Coiro et al. 2023) or trait evolution (Slater et al. 2017; Sterli et al. 2018; Farina et al. 2023). Despite their ubiquity, phylogenetic analyses can be conceptually challenging and difficult to apply in practice. This comes down to properties inherent in both the data and the methods themselves. The typical data used in a phylogenetic analysis are a combination of molecular and/or morphological characters,

Table 1. Application of fossilized birth–death (FBD) models to different types of phylogenetic character data. To date, studies have included molecular data for extant samples or morphological data for extant and extinct (†) samples. (Molecular data could theoretically also be included for extinct samples, if ancient DNA is available.) In “total-evidence” analyses, character data are included for both extant (molecular and morphology) and extinct (morphology only) samples. “Extant only” refers to analyses in which molecular data are included for extant samples only and fossils are placed using constraints. “Morphology” refers to analyses in which morphology is included for both extant and extinct samples. “Extinct only” refers to analyses of fully extinct trees in which morphological data are available for extinct samples. “No phylogenetic data” refers to analyses in which no phylogenetic character data are included (see Boxes 2 and 5). All analyses using the FBD model must also include temporal data. See Section The Data for more information

Analysis type	Phylogenetic data			No. of analyses
	Molecular	Morphology	Morphology [†]	
Total evidence	✓	✓	✓	53
Extant only	✓			78
Morphology		✓	✓	26
Extinct only			✓	35
No phylogenetic data				16

taxonomic identity, and fossil age information (see Table 1). These data are labor-intensive to collect and curate, non-uniformly incomplete, and associated with complex uncertainties. Integrating these different data types into a single analysis is not straightforward and requires a statistically advanced approach, often applied within a Bayesian framework.

A major breakthrough in our ability to integrate these data sources came with the introduction of the fossilized birth–death (FBD) family of models, which allow for the joint estimation of a phylogeny and divergence times using extinct and extant taxa (Stadler 2010; Didier et al. 2012; Gavryushkina et al. 2014, 2017; Heath et al. 2014; Zhang et al. 2016; Stadler et al. 2018). FBD models offer a number of advantages over other approaches. They are typically implemented in a Bayesian framework, which provides a robust approach to incorporating uncertainty in our data and estimates. The most important property of these models is that they explicitly account for both fossil and extant sampling probabilities. In this way, it is possible to model both extinct and extant observations under the same generating process (Stadler 2010). There have been a number of extensions to the original model, and we note that we consider any model that includes the fossil sampling process explicitly (i.e., has a parameter for the rate of fossil sampling) to be a member of the FBD model family (see Table 2). Phylogenetic analysis using FBD models offers great potential for furthering our understanding of evolution in deep time. However, as these methods become more complex, it can be challenging to apply them to empirical data in practice.

In this review, we aim to provide a guide on how FBD models work, what they assume, and how all of this fits with the nature of empirical data from the fossil record. To this end, we first quantified the use of the FBD model to date through a literature survey. This allowed us to understand the types of questions asked by (paleo) biologists using the models, as well as determine areas of confusion regarding the application of the models. Next, we provide an introduction to Bayesian inference, along with some relevant terminology and concepts. We then discuss aspects of paleontological data that allow for their inclusion in analyses using FBD models,

Table 2. Available fossilized birth–death (FBD) models and extensions

Model name	Description	Reference
FBD specimen process	Model assumes constant rates of diversification (λ , μ) and sampling through time (ψ) and assumes uniform sampling of extant taxa.	Stadler (2010); Heath et al. (2014)
FBD skyline	Model allows for diversification (λ_t , μ_t) and sampling rates (ψ_t) to vary through time across discrete intervals (i.e., in a piecewise-constant fashion).	Stadler et al. (2012); Gavryushkina et al. (2014)
Diversified sampling	Model accounts for a non-uniform sampling strategy that aims to maximize the phylogenetic diversity represented in the extant tree. This is achieved by introducing the parameter x_{cut} , which represents the youngest node age, after which we have no more sampled nodes.	Höhna et al. (2011); Zhang et al. (2016)
Multi-type birth–death process	Model allows for variation in diversification and fossil sampling rates across different parts of the tree associated with different “types”. Types may correspond to any trait that effects differences in rates. Lineages can switch between types with rate m .	Kühnert et al. (2016); Barido-Sottani et al. (2020c)
Occurrence birth–death process	Model allows for the incorporation of data from fossil observations outside the tree, i.e., those for which we have no taxonomic information or morphological data, such as trace fossils. These observations are modeled using a separate sampling process, with rate parameter ω .	Manceau et al. (2021); Gupta et al. (2020); Andréoletti et al. (2022)
Episodic FBD process	Model allows for instantaneous speciation, extinction, or sampling events to affect the entire tree, corresponding to events such as mass extinctions or species radiations.	Magee and Höhna (2021)
FBD multispecies coalescent	Model allows separate genes to evolve independently under a coalescent model.	Ogilvie et al. (2022)
FBD range process	Model provides the framework to explicitly take stratigraphic range data into account, thereby associating multiple occurrences within a range of the same taxon.	Stadler et al. (2018)

including more complicated properties that cannot yet be accounted for by existing models. Following this, we discuss the models used in Bayesian phylogenetic inference, with a focus on the FBD model, which is grounded in discussions about the realities of paleontological data. We discuss different software and resources available for Bayesian phylogenetic analysis of fossil data (Box 1) and how to include information on tree topology (Box 2). Further, we present three exemplar case studies that showcase different applications of the FBD model in paleobiology, emphasizing the connection between the data and the models in each case, (see Box 3–5). Finally, we end with a perspective on outstanding challenges and future directions for integrating paleontological data into FBD models and modeling fossil sampling.

Through this review, we hope to celebrate the FBD model family, inform readers of its nuances and potential uses, and stimulate discussion around best practices for answering empirical questions. Within this, we encourage researchers to critically assess the models they use and how those models might apply to their data. We also hope to improve collective understanding of the properties of paleontological data, in order to facilitate the development of models that make more realistic assumptions.

Application of FBD Models

A decade on from the first implementations of the FBD model (Heath et al. 2014; Gavryushkina et al. 2014), it continues to receive significant citation and has been applied to a wide range of clades across the tree of life. To quantify its use to date, we

carried out a literature survey (see [Supplementary Material](#) for details) and collated information from 176 studies, with 208 empirical analyses ([Supplementary Table S2](#)). The results are summarized in [Figures 1 and 2](#). The term “samples” is used here to encompass both extant and extinct data. This distinction is important, because some taxa may be represented by multiple individuals, and some fossils may represent sampled ancestors rather than just terminal taxa. In this way, increasing the number of samples may be different from increasing the number of tips/taxa in an analysis.

We found that FBD models have been most frequently used on datasets of vertebrates, with 135 empirical analyses carried out, while invertebrates and plants were the focus of 45 and 27 studies, respectively. Datasets range in size, with a mean of 74 extant and 60 extinct samples. FBD models can be applied to fossils with or without morphological characters; across all studies, 117 used morphological data, with an average of 246 characters. Our survey shows that a wide range of dataset sizes have successfully been analyzed using FBD models and that these models are extremely flexible in terms of the number of taxa and/or characters used. We split the analysis type into five categories, based on the type of phylogenetic character data used (for more details, see [Section The Data, Table 1](#)).

[Figure 2](#) shows the temporal range of the empirical analyses applying the FBD model. Only 20% of the studies used data from the Paleozoic, suggesting that the challenges involved in applying the FBD may increase with clade age and/or may be due to the limited availability of phylogenetic datasets further back in time.

Box 1. Software and resources for Bayesian phylogenetic analysis of fossil data.

The fossilized birth–death (FBD) model has been implemented in a number of phylogenetic software programs. Choosing which might be most appropriate for running a specific analysis should include considering the component models available in the different software programs and the accessibility of relevant documentation and example configuration files. Here we describe a selection of software programs that implement the FBD model.

BEAST2 (Bouckaert et al. 2019) is a common choice for divergence time estimation using a ready-built phylogeny and for joint estimation of tree topology and divergence times. The program is modularized, with a core supported by packages containing specific models and model components. BEAST2 configuration files, which are .xml, can be created using BEAUti, its accompanying GUI. This makes building models more user-friendly, particularly features such as the live visualization of prior distributions. Skyline and stratigraphic range implementations are available in BEAST2, but at present, the software offers limited options for morphological substitution models. While some tutorials are available on the core BEAST2 website (<https://www.beast2.org/tutorials>), additional tutorials and training courses are available through Taming the BEAST (Barido-Sottani et al. 2018; <https://taming-the-beast.org>).

MrBayes (Ronquist et al. 2012b) is the most popular choice for FBD analyses using extinct clades, based on our literature review. MrBayes is operated through the command line, meaning that it can be run easily on all operating systems but requires some coding expertise to use. MrBayes configuration files are NEXUS, or .nex, files. MrBayes includes a wide range of evolutionary models. The R package PASTIS (Thomas et al. 2013) can be used to add taxonomic constraints to tips without genetic data in a MrBayes analysis. A user manual (<http://nbisweden.github.io/MrBayes/manual.html>) and example configuration files are supplied with the program. A comprehensive tutorial is also available for the FBD model (Zhang 2016).

RevBayes (Höhna et al. 2016) allows users to build phylogenetic models using an internal R-like language, Rev. This enables the easy translation of models into graphical objects, allowing visualization to improve the interpretability of models. It also aims to balance the convenience of using ready-built modules with the customizability afforded by a programming language. Configuration files are provided in the unique .Rev format. There are also RevBayes-specific companion R packages, RevGadgets (Tribble et al. 2022) and Revticulate (Charpentier and Wright 2022). Extensive documentation, tutorials, and videos are available on the RevBayes website (<https://revbayes.github.io>).

DateFBD (Didier and Laurin 2020), available as both a stand-alone program and an R package, can also be used to conduct divergence time estimation on preconstructed phylogenies using the FBD.

Box 2. Previously inferred trees.

Previously inferred trees can be used to constrain the entire tree topology or just parts of the tree, for example, constraining specific key clades or even using a backbone constraint for extant taxa. When taking this approach, the user needs to decide whether or not to use the entire posterior distribution of trees (most plausible set of trees estimated) or a summary tree, and if so, which type of summary tree (see [Section Assessing and Interpreting the Output](#) for information on summary trees). Constraining the topology using a previously inferred tree is reasonable when working with whole-genome data or large alignments (e.g., Fleming et al. 2018; Coiro et al. 2023; see [Box 4](#)) or, more commonly in paleobiology, using supertrees (e.g., Allen et al. 2024; see [Box 5](#)). Supertrees can be either informal or formal, depending on how they are constructed. Informal supertrees are often manually constructed, using software such as Mesquite (Maddison and Maddison 2023), although new methods are available to automate this step (Castiglione et al. 2022). This involves combining trees from different studies without explicitly addressing conflict between topologies. In cases where there are conflicting hypotheses about the phylogenetic history of a group, it is common to construct more than one supertree to reflect these controversies. Any downstream analysis would thus be carried out using all supertrees (e.g., Godoy et al. 2019; Dunne et al. 2023). Formal supertrees are also constructed using trees from different studies but rely on algorithms that explicitly account for conflicting topologies to construct them (e.g., Lloyd and Slater 2021; Stockdale and Benton 2021).

Box 3. Case study 1: Inference of a fully extinct phylogeny.

Study Objectives. One of the first applications of the fossilized birth–death (FBD) model to a completely extinct group was carried out by Wright (2017). Here, the main objective was to explore phylogenetic relationships among Ordovician non-camerate crinoids and to revisit the systematics of this group based on trees inferred using the FBD model. While the phylogeny and taxonomy of this group had been studied previously using a phylogenetic approach, the FBD model allowed both character and temporal evidence to inform the topology. This study also included a comprehensive description of the FBD model from a paleontological perspective and a detailed explanation of how it could be applied to fully extinct groups.

Data. The data used in this study consisted of a discrete morphological character matrix and taxon age information. The morphological data consisted of 87 characters coded for 42 species of non-camerate crinoids. Wright (2017) gathered age information from the literature for each of the 42 species in the matrix comprising 14 singletons and 28 taxa whose stratigraphic range spanned more than one geological interval (see Section Temporal Data, Fig. 4).

Bayesian Inference. The analysis was carried out using MrBayes 3.2 (Ronquist et al. 2012b). The FBD specimen model was used as the tree model (Stadler 2010; see Section The FBD Process). The FBD process was parameterized in terms of diversification, turnover, and sampling proportion; see Box 6 for details. Phylogenetic software usually scales the tree (irrespective of the input ages) such that the youngest sample is always at age zero. The probability of sampling at the “present,” that is, at the end of the sampling interval, p , was set to 1. An alternative would be to set $p = 0$ and co-estimate the age of the youngest (ψ sampled) fossil, which later became possible in BEAST2, with the introduction of an offset parameter (Barido-Sottani et al. 2020b). A single age range was specified for each taxon, representing the upper limit of the oldest possible fossil age and the lower limit of the youngest possible fossil age.

The author carried out model selection to identify the most suitable clock model. A relaxed clock model was found to provide the best fit to the data, which matched the author’s expectation that morphological rates of evolution vary across the tree (see Section Clock Models). An MkV model (Lewis 2001) with a lognormal distribution describing rate variation was used to describe the evolution of the character data (see Section Substitution models). As all taxa included in the analysis had character data, the position of each fossil was inferred during the inference. Several well-established clades (Disparida, Flexibilia) were constrained to improve efficiency. Support for the monophyly of taxonomic groups was evaluated using the resulting summary tree and the posterior probability of internal nodes.

Sensitivity analyses were performed to assess the impact of the priors on the FBD and clock model parameters.

Study Outcomes. Wright (2017) found support for several of the clades identified in previous studies. For example, strong support was found for the split between disparids (and disparid-like taxa) and all other non-camerate crinoids and for the monophyly of several other previously classified groups. However, there were several cases in which the monophyly of previously described taxonomic groups was not supported, and the position of some groups relative to others shifted compared with previous phylogenetic analyses. We note that because the majority of taxa included in this study span more than one geological stage, it would be interesting to reanalyze this dataset using the FBD range model and observe whether this results in any further changes to the topology.

Box 4. Case study 2: Total-evidence analysis.

Study Objectives. Cycads (Cycadales) are seed plants that once exhibited high diversity and had a global distribution, but are now restricted to low latitudes. Coiro et al. (2023) aimed to infer the origins of the group and the evolution of their geographic range over the clade’s history. They used a total-evidence approach, which allowed them to combine molecular data, morphological data, and temporal information from the fossil record into a single phylogenetic inference (Ronquist et al. 2012a). The resulting time-calibrated tree, including extant and fossil taxa, was used in downstream biogeographic analyses.

Data. The morphological matrix used in this analysis consisted of 60 fossil and 321 extant cycad species, with 31 morphological characters covering leaf morphology and cuticular anatomy. The molecular matrix contained 18 loci for each of the 321 extant species. Temporal occurrence data associated with fossils in the morphological alignment were collected from the literature (see Section The Data).

Bayesian Inference. The analysis was carried out in MrBayes 3.2 (Ronquist et al. 2012b) using the fossilized birth–death (FBD) model with diversified sampling (Zhang et al. 2016), which accounts for non-uniform sampling of extant species (see Section The FBD Process, Table 2). The model was parameterized using speciation (λ), extinction (μ), and fossil sampling rates (ψ). The prior on λ was an exponential distribution with a mean of 0.1. The priors on μ and ψ rates were beta distributions with a mean of 0.5. The probability of extant species sampling, p , was set to 0.87, as the analysis contained 321 of the 370 recognized extant species. The prior on the origin time (t_{origin}) was set to a uniform distribution between 359 Ma (the early Famennian, corresponding to the earliest evidence of stem seeds) and 259 Ma (during the late Lopingian). A single age range was specified for each fossil, based on the upper and lower limits of the fossil sample localities. For the character data, the authors used six separate partitions for the molecular data and one for morphological data. A relaxed clock model was used for the molecular data, with the average rate and variance linked across partitions, while a strict clock model was used for the morphological partition (see Section Clock Models). Model selection was used to identify the best-suited substitution models for the molecular partitions, and an MkV+I model (Yang 1994; Lewis 2001) was used for the morphological partition (see Section Substitution Models). As all fossils were associated with character data, their positions in the tree could be inferred during the inference. To improve computational efficiency, the relationships among extant taxa were fixed, based on a tree inferred using a maximum-likelihood approach (see Section Application of FBD Models). The resulting dated phylogeny was used for further analysis. Patterns of historical biogeography were inferred using the dispersal–extinction–cladogenesis (DEC) model (Ree and Smith 2008), using the software DECX (Beeravolu and Condamine 2016) and the R package BioGeoBEARS (Matzke 2013). Ancestral latitudes were also estimated using a Brownian motion model in the software fossilBM (Zhang et al. 2022).

Study Outcomes. The inferred phylogeny obtained using the FBD model, combined with the biogeographic analyses, provide a rich picture of the temporal and spatial history of the group. Cycads were estimated to have originated in the Carboniferous on the Laurasian continent, with two major lineages (Cycadaceae, Zamiaceae) later originating in the Triassic and Jurassic, as the group expanded its range across Gondwana. The authors were able to identify key vicariance events, as well as intervals during which latitude ranges expanded and contracted. This study also served to demonstrate the importance of incorporating fossils into phylogenies for testing biogeographic hypotheses. Our literature survey revealed that the FBD model is often used to reconstruct trees for this purpose (Supplementary Table S2).

The time-calibrated trees inferred in FBD analyses are often used in secondary analyses such as biogeographic inference (Areces-Berazain and Ackerman 2016; Xiang et al. 2019; Thomas et al. 2020), ancestral-state estimation (Fleming et al. 2018; Farina et al. 2023; Jiang et al. 2023; Wolfe et al. 2023), or other phylogenetic comparative methods (Slater et al. 2017; Song et al. 2020; Dunne et al. 2023). BEAST2 (Drummond et al. 2006) is the most popular software for these analyses (109 studies), followed by MrBayes (Ronquist et al. 2012b) (93), and RevBayes (Höhna et al. 2016) (10).

Collating past applications of the FBD model allowed us not only to explore its use but also to pinpoint potential sources of difficulty

regarding how the models are applied in practice. A major challenge in its application appears to stem from how temporal information, that is, fossil ages, relate to the assumptions of the FBD model. This was determined by the way in which these data are described in papers, often providing a lack of information or justification regarding the use of specific fossils, or whether temporal intervals referred to stratigraphic ranges or fossil age uncertainty. Presenting well-justified fossil age information is important for any fossil-based approach to divergence time estimation (Parham et al. 2012).

Reporting of the temporal data used and other information has important consequences for the interpretation of results and

Box 5. Case study 3: Inferring diversification dynamics.

Study objectives. The diversity of non-avian dinosaurs through time, between their origin during the Triassic and their ultimate extinction at the end-Cretaceous mass extinction, has been extensively studied. However, these studies have reached a wide range of conclusions about dinosaur diversity in the latest Cretaceous, with some suggesting that the clade was still diversifying right up until the mass extinction event, while others indicate that dinosaurs may have been declining in diversity for tens of millions of years prior. Allen *et al.* (2024) applied two skyline models, a fossilized birth–death (FBD) skyline and a coalescent skyline, to dinosaur supertrees to investigate the role of model sampling assumptions in determining the outcomes of analyses of dinosaur diversification.

Data. The data used consisted of four supertrees (three informal and one formal; see Section Application of FBD Models) and fossil age information in the form of stratigraphic ranges obtained for each species from the PBDB. Following taxon alignment between these two datasets, the supertrees contained 391 (Lloyd *et al.* 2008), 542 (two topologies; Benson *et al.* 2014), and 750 (Lloyd *et al.* 2016) dinosaur species.

Bayesian Inference. The FBD analysis was carried out using a skyline model (Gavryushkina *et al.* 2014; see Section The FBD Process, Table 2) in BEAST2 (Bouckaert *et al.* 2019). The model was parameterized using speciation (λ) and extinction (μ) rates, along with a fossil sampling rate (ψ), with diversification rates (speciation – extinction) calculated post hoc. The probability of sampling extant species, ρ , was set to 0. Seven change times, placed at geological stage boundaries, allowed piecewise constant rates to be calculated for eight approximately equal intervals, ranging from the Triassic to the Cretaceous. The priors on λ and μ were exponential distributions with a mean of 1.0. The prior on ψ was an exponential distribution with a mean of 0.2 to reflect the prior expectation that sampling is low relative to diversification. The prior on the origin time (t_{origin}) of the clade was set to a uniform distribution with an upper bound of 266 Ma (during the mid-Permian) and a lower bound of 66 Ma (the end of the Cretaceous). The topologies of the supertrees remained fixed during the analyses, but fossil age constraints were used to provide temporal information. In each iteration, the age of each species was sampled from a uniform distribution ranging from the upper limit of the oldest possible fossil age to the lower limit of the youngest possible fossil age. Node ages (and hence branch lengths) and the phylodynamic (λ , μ , ψ) parameters were also sampled. This allowed dinosaur diversification rates to be inferred while taking into account uncertainty in fossil ages.

Study Outcomes. Allen *et al.* (2024) found that the FBD skyline models suggested that dinosaurs were still experiencing positive diversification rates during the latest Cretaceous, but the accompanying coalescent skyline models inferred negative diversification rates during this same interval. The dinosaur fossil record is highly uneven, and the different results produced are likely linked to the different sampling assumptions underlying these two models. This is supported by the fact that the FBD skyline model's highest posterior density (HPD) intervals reached unrealistically high sampling rates in the latest Cretaceous. However, this was not the case for the largest phylogeny, indicating that tree size likely also has a large impact on our ability to infer diversification rates using skyline models. Overall, these results demonstrate the importance of understanding model assumptions when applying phylogenetic models.

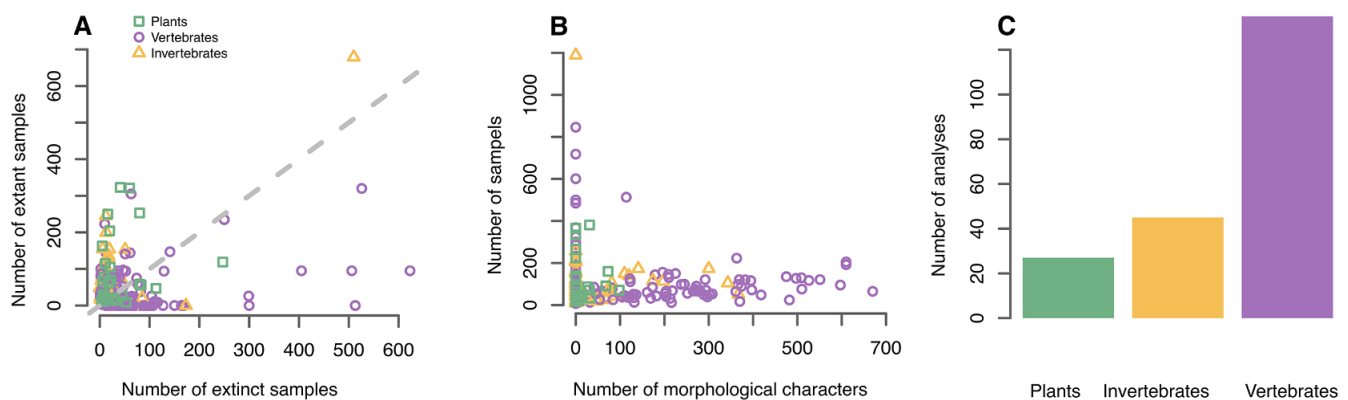


Figure 1. **A**, The number of extant vs. extinct samples. Each point represents an individual analysis. Points above the dashed line have more extant samples relative to extinct, whereas those under the line have more extinct samples. **B**, The number of samples (extant plus extinct samples) vs. the number of morphological characters used in an analysis. **C**, The number of analyses using plants, invertebrates, and vertebrates.

reproducibility of phylogenetic analyses (DeBiasse and Ryan 2019; Drummond *et al.* 2023; Jenkins *et al.* 2023). There are many decisions to be made in any phylogenetic analysis, and each is important; these decisions, therefore, should be reported in a paper to ensure reproducibility. In our literature survey, it was often not possible to fully interpret which data were included and how they were used in an analysis. We appreciate that in many cases the estimation of a time-calibrated tree was a step in a larger analytical workflow that includes further analyses such as those mentioned earlier. Yet ensuring the analysis is fully and correctly described is essential for the reader to interpret and assess the reliability of the results. Many phylogenetic software are run using configuration files, and providing these files in data supplements is a simple and significant way of improving transparency and reproducibility.

Bayesian Framework

Bayesian statistical inference seeks to estimate the posterior probability of a model conditioned on observed data (for more background on Bayesian phylogenetics, see Huelsenbeck *et al.* 2001; Alfaro and Holder 2006; Yang 2014; Nascimento *et al.* 2017; Warnock and Wright 2020). The posterior probability is an expression of the researcher's degree of belief in the model, and incorporates both prior knowledge and the information contributed by the data. The foundation of Bayesian statistics is Bayes theorem, which defines the posterior probability of a model conditioned on observed data as

$$P(\text{Model}|\text{Data}) = \frac{P(\text{Data}|\text{Model})P(\text{Model})}{P(\text{Data})}, \quad (1)$$

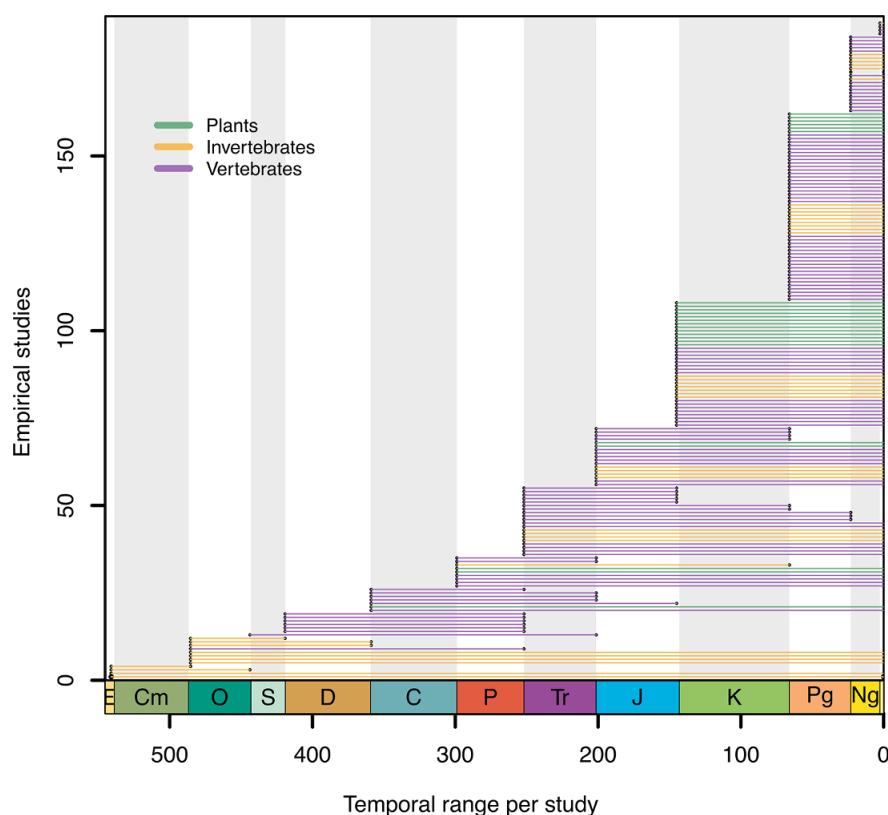


Figure 2. The temporal range, based on the oldest fossil age to the tips of the tree from each of the empirical studies in the literature survey. Timeline plotted using R package *divDyn* Kocsis 2019. CM, Cambrian; O, Ordovician; S, Silurian; D, Devonian; C, Carboniferous; P, Permian; Tr, Triassic; J, Jurassic; K, Cretaceous; Pg, Paleogene; Ng, Neogene.

where $P(\text{Model} \mid \text{Data})$ is the posterior probability of the model given the data, the information in the data appears in the likelihood $P(\text{Data} \mid \text{Model})$, a priori information is defined in the prior probability of the model $P(\text{Model})$, and the normalizing constant $P(\text{Data})$ is the marginal probability of the data. Figure 3 illustrates a phylogenetic application of Bayes theorem using character data and fossil occurrence times. Here, we are interested in estimating the joint posterior probability of the phylogeny and divergence times for a set of samples (fossil and extant), the diversification and sampling rates of the FBD process, the parameters of the character substitution model, and the lineage-specific substitution rates (the clock model parameters). Each component of our hierarchical Bayesian model includes parameters for which the values are not known (also called free or stochastic parameters) and will be estimated (e.g., extinction rate). The outcome of an analysis will yield posterior distributions for each of the free parameters and thus will quantify the statistical uncertainty conditional on all other model components and the data. We will use Bayes theorem (eq. 1, Fig. 3) to organize our exploration of the components of a Bayesian statistical analysis, and then we will elaborate on the models specific to phylogenetic inference using data from the fossil record in Section The Models.

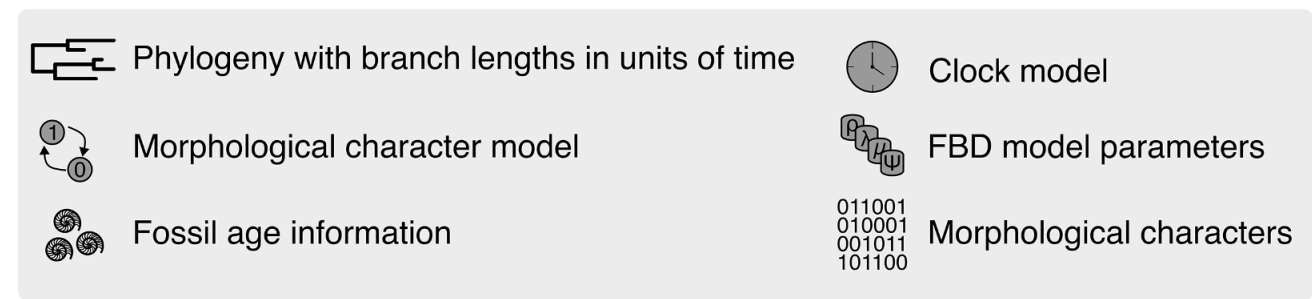
The Likelihood

The likelihood represents the probability of the observed data given the model parameters. That is, it is the probability that a particular evolutionary hypotheses (in Fig. 3, a given tree topology and divergence times, rates of the FBD process, substitution model

parameters, and clock rates) could have given rise to the observed data (i.e., morphological character states and fossil ages in Fig. 3). This type of hierarchical Bayesian modeling allows us to consider multiple sources of information in a unified framework, accounting for the dependencies and interactions of different data sources under a holistic model. The likelihood is then a product of the individual likelihoods associated with the different data types. For every set of observations in our model, the information coming from these data will be combined with the information from our prior knowledge in the calculation of the posterior.

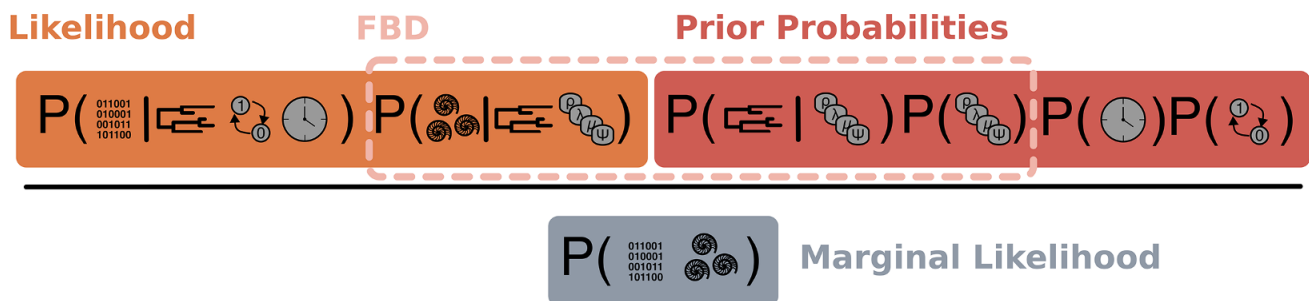
The Prior

A prior probability is a mathematical expression that represents our knowledge about the parameters of interest before considering the information in the data. For any unknown variable in our model (e.g., speciation or extinction rate), we define a prior probability distribution to quantify our statistical uncertainty in the value of that parameter, thus allowing us to integrate our existing knowledge (or lack of knowledge) about the model into the inference. While the ability to add prior information to an analysis is a major strength of Bayesian phylogenetics, transforming biological knowledge into a mathematical expression can be daunting in practice. Priors can be chosen by taking into account information from previous studies or relying on theoretical knowledge regarding the evolution of the system and properties of model parameters. There are a number of standard or default probability distributions used as priors in Bayesian phylogenetics. For example, exponential distributions are frequently applied as priors to speciation or extinction rates. By assuming that a parameter is drawn from an



Posterior distribution

$$P(\text{Phylogeny, FBD model parameters, Morphological character model, Clock model} \mid \text{Morphological characters, Fossil age information}) =$$



The **posterior distribution** is the probability of the model given the data

The **prior probabilities** associated with the free parameters

The **likelihood** is probability of the data given the model parameters

FBD model components

The **marginal likelihood** is the probability of observing the data under all possible parameter values

Figure 3. Symbolic representation of Bayes theorem (eq. 1) for a phylogenetic analysis of fossil ages and morphological characters. The data components include a morphological character matrix and fossil ages. The model parameters are a phylogeny with branch times, the diversification and sampling parameters of the fossilized birth–death (FBD) model, the lineage-specific branch rates of the clock model, and the parameters of the morphological substitution model (Mk model; Lewis 2001). The probabilities are delineated to highlight the joint posterior distribution, likelihood, and prior probability distributions. The FBD probability density includes some components for which we calculate prior probabilities (the tree topology, branch times, and diversification and sampling parameters) and some that are observations in the likelihood (fossil ages). Thus, these are separated to clarify the contributions to the posterior density coming from the prior and those coming from the data.

exponential distribution, we assign higher probability density to smaller values and lower density to larger values, such that we expect a priori that speciation and extinction rates will generally be small. Note that for the sake of clarity in Figure 3, we show the prior probabilities for the FBD rate parameters as a single joint prior, $P(\lambda, \mu, \psi, \rho)$, but typically, we assume independent generating models for each parameter. For example, we may have different exponential priors on each of the three rates and fix the extant sampling probability to an approximate value based on our

knowledge of the current extant diversity (e.g., Barido-Sottani et al. 2020a).

There are numerous papers offering discussions on the importance of judiciousness in prior specification (e.g., Holder and Lewis 2003; Lemmon and Moriarty 2004; Alfaro and Holder 2006), strategies for constructing hierarchical Bayesian models for phylogenetics and selecting priors (e.g., Wang and Yang 2014; Nascimento et al. 2017; Bromham et al. 2018), and assessments of the impact of model violations and misspecified priors on posterior

estimates (e.g., Zwickl and Holder 2004; Ritchie et al. 2017; Möller et al. 2018; Yang and Zhu 2018; Sarver et al. 2019). While a detailed exploration of these topics is beyond the scope of this paper, we hope to emphasize the important role that priors play in Bayesian phylogenetic inference. Thus, it is critical that researchers applying these methods understand the models they select as priors and their assumptions and behaviors. Many of the software tools for Bayesian phylogenetic inference provide detailed documentation and tutorials on the available priors (e.g., Barido-Sottani et al. 2018, 2020a), and while there is no “one-size-fits-all” set of priors for a Bayesian phylogenetic analysis, these resources provide intuition for the various choices that must be made and can guide users in translating their own prior knowledge into probabilistic models. With this knowledge, researchers can then provide sufficient justification for their priors and assumptions when reporting their methods and results (Gelman 2012).

The Marginal Likelihood

The marginal likelihood is the probability of observing the data given the model, averaged over all possible values of the model parameters (Oaks et al. 2019). This quantity acts as a normalizing constant and is very difficult (and often intractable) to calculate, as it involves integrating over every possible value in a complex and hyperdimensional parameter space. Sampling algorithms offer a way to circumvent directly calculating marginal likelihoods. In Bayesian phylogenetics, Markov chain Monte Carlo (MCMC) algorithms, especially the Metropolis-Hastings algorithm (Metropolis et al. 1953; Hastings 1970), are the most common methods for approximating the posterior distribution of a hierarchical model. MCMC uses simulation to sample parameter values in proportion to their posterior probability, and in the evaluation of simulated parameter draws, the marginal likelihoods cancel out (for an explanation, see Nascimento et al. [2017] or Section Inference using MCMC). While the marginal likelihoods are not computed during inference, there are other cases when we would want to use these quantities.

Because the marginal likelihood is a measure of the fit of a model to a given set of observations (averaged over all model parameter values), they are important in Bayesian methods for model comparison (Oaks et al. 2019). Bayes factors are a way to evaluate the relative fit between two models for a given dataset. We can compute the Bayes factor for two models, M_0 and M_1 , as the ratio of the marginal likelihoods of the data for each model:

$$BF(M_0, M_1) = \frac{P(\text{data} | M_0)}{P(\text{data} | M_1)}, \quad (2)$$

where $BF(M_0, M_1) < 1$ corresponds to support for M_1 and $BF(M_0, M_1) > 1$ indicates support for M_0 (Jeffreys 1961; Kass and Raftery 1995). There are several approaches for estimating marginal likelihoods that range in their degree of computational intensity and accuracy (Oaks et al. 2019; Fourment et al. 2020). In particular, path sampling (Lartillot and Philippe 2006) and stepping stone sampling (Xie et al. 2011) have been shown to be accurate when comparing clock models (Baele et al. 2012) or substitution models for molecular data (Fourment et al. 2020). However, using Bayes factors to choose between morphological models has been shown to be unreliable when comparing different partitioning schemes (Mulvey et al. 2024).

It is also very difficult to estimate marginal likelihoods to compare tree models (e.g., two different FBD models). Because

the FBD model probability contains both likelihood elements and prior probability elements (Fig. 3), the marginal likelihoods are not easily estimated (May and Rothfels 2023). Methods of estimating marginal likelihoods also take abundant time and computational resources to run. Therefore, quantitative model comparison is not always viable; however, in this case, it is critically important that models are chosen carefully with justifiable rationale.

The Posterior

For any given set of values for the parameters in the model, the posterior is the joint probability density of those values conditioned on the observed data. A typical Bayesian analysis will use algorithmic approaches for approximating the posterior (such as MCMC, as explored in Section Inference using MCMC). These methods return a series of samples for every free parameter. With an MCMC sample, one can then examine the marginal posterior for each free parameter. For instance, a summarization of the marginal posterior density for the speciation rate shows the posterior probability density for speciation rate values while averaging over all other parameter values (e.g., the extinction rate, tree topology, node ages). This provides an intuitive way to quantify the uncertainty associated with each of our parameters. If the variance of the posterior density for the speciation rate is extremely large or matches the prior, there may not have been a strong signal in the data supporting any particular range of values. See Section Assessing and Interpreting the Output for more information.

Inference Using MCMC

When we use Bayesian inference and MCMC to estimate the posterior probability of our phylogenetic model given observed data, our analysis generates a series of sampled values for every free parameter defined in our model. The algorithm essentially performs a random walk over parameter space by proposing new values for each stochastic parameter (i.e., every parameter that is assigned a prior distribution) and accepting or rejecting those values in proportion to their posterior probability (Metropolis et al. 1953; Hastings 1970). We will not go into all the details about the workings of MCMC in a Bayesian phylogenetic analysis, as this information is readily available in the literature (for overviews, see Huelsenbeck et al. 2002; Holder and Lewis 2003; Ronquist et al. 2009; Yang 2014; Nascimento et al. 2017; Barido-Sottani et al. 2024). Instead, we will touch on some aspects of MCMC that may be helpful to researchers setting up Bayesian phylogenetic analyses.

Every stochastic variable in the model that is assigned a prior distribution is also associated with one or more MCMC proposal mechanisms that will sample the value of that parameter over the course of the Markov chain. Note that these mechanisms are called “proposals” in MrBayes (Ronquist et al. 2012b), “moves” in RevBayes (Höhna et al. 2016), and “operators” in BEAST1 (Suchard et al. 2018) and BEAST2 (Bouckaert et al. 2019). Every MCMC move is assigned a “weight” that determines the frequency with which that particular move will be made over the course of the run. For example, if the model has four stochastic parameters that are each operated on by a single move with a weight of 1.0, and the total number of moves is set to 1000, then the MCMC will perform each move 250 times, on average. However, if three of the moves are assigned a weight of 1.0 and one has a weight of 2.0, then after 1000 moves, the higher-weight parameter will be sampled approximately 400 times and each of the others will be sampled 200 times, on

average. The move weights are important for ensuring that your analysis will sample your parameters of interest frequently enough to obtain a sufficient sample. Moreover, there may be parameters in our model that we simply wish to integrate over to account for uncertainty, but we do not intend to study or report. These are often referred to as “nuisance” parameters. For example, in most phylogenetic analyses, researchers may not report the estimates for the state frequencies or substitution rates of their character-evolution model. Additionally, character data are often very informative for these parameters. Thus, when assigning weights to the MCMC, we may not need to sample these parameters as frequently as others.

Some parts of a phylogenetic model are very complex and require multiple types of moves to effectively sample the posterior distribution. This is particularly true for the tree topology and branch times. To explore this complex parameter space, one may assign multiple different moves that may have higher or lower variance in the values they propose. For example, to sample the tree topology (Lakner et al. 2008), we may choose a move that proposes a new state by moving one branch to a neighboring branch (i.e., a nearest-neighbor interchange move) and combine this with a move that makes larger changes like moving a subtree to a more distant branch (i.e., the subtree pruning and regrafting move). By combining local and global tree topology moves, we can allow the Markov chain to reach different parts of parameter space while also sampling the high-probability regions more exhaustively.

Any Bayesian software will require the user to specify a stopping criterion for the Markov chain. Typically, this is done by setting the number of iterations the MCMC will perform. It is important to note, however, that different Bayesian phylogenetics programs have different default definitions of an MCMC “iteration” (these may also be called “generations,” “replicates,” or “steps”). In BEAST2 (Bouckaert et al. 2019) and MrBayes (Ronquist et al. 2012b), a single MCMC iteration performs one move, thus requiring the user to set the chain length accordingly. For very complex models with a large number of parameters and moves, the chain length should be longer than for a simpler model. For instance, analyzing a tree with 20 taxa corresponds to 19 internal node ages (including the root) that must be sampled. Often each internal node age is associated with a proposal that slides the node age up or down. If considering a single move per iteration, the total number of moves must be increased if you consider a larger tree (a tree with 50 taxa has 49 internal node ages) and wish for the same average number of proposals per node. RevBayes (Höhna et al. 2016), on the other hand, takes all of the moves and adds up their weights to determine the number of moves performed within a single MCMC iteration. If you specify 10 moves with weight 1.0 and 10 moves with weight 2.0, a single iteration will perform 30 moves, and each new proposal that is executed is drawn from the list of moves in proportion to its weight. With this approach, to achieve a desired number of moves per parameter, the user will specify a chain length with fewer “iterations” than if they perform a single move at each step. Ultimately, no matter how this is handled, one should carefully consider the total number of moves and recognize that thousands of proposals may be required to obtain a sufficient sample from the posterior distribution for a given parameter.

Understanding the behavior of MCMC for different parameters and how proposals are implemented in phylogenetic software is helpful for diagnosing issues and ensuring that your results are sensible. The nuances of MCMC that we have touched on also go hand in hand with protocols for processing and summarizing the output of your Bayesian phylogenetic analysis (see Section Assessing and Interpreting the Output).

If you are new to these methods, we highly recommend reviewing the introductory tutorials and documentation for available software to learn the specifics of the MCMC algorithms used (see Box 1).

The Data

In a Bayesian framework, phylogenetic inference with an FBD model allows us to incorporate different types of data. These can include temporal information (fossil ages), morphological data, and/or molecular data. In Figure 3, we include fossil ages and morphology in our representation of Bayes theorem. The data are the only component that we can directly observe. In this section, we describe the paleobiological data available for phylogenetic analyses, along with important aspects of their implications for analyses.

Character Data

Molecular Data. Recent advances in DNA-sequencing technologies have reduced many barriers to generating high-quality molecular data for taxa spanning the tree of life. Typically, molecular phylogenetic datasets are aligned nucleotide sequences of homologous regions of the genome. Nucleotides across the genome share biochemical and physiochemical properties, which have subsequently informed an array of generalizable substitution models (Liò and Goldman 1998; Arenas 2015). Proteins are also sources of molecular data for phylogenetic analysis, particularly for studies examining distantly related species or groups (e.g., Williams et al. 2020). These molecular datasets constitute alignments comprising linear sequences of amino acids that form homologous proteins. Similar to nucleotide models, several amino acid substitution models have been introduced based on detailed study of protein structure and function (Blanquart and Lartillot 2008; Huelsenbeck et al. 2008; Trivedi and Nagarajaram 2020). It is important to note that there are a number of different approaches for aligning molecular sequences; see Stadler et al. (2024) for an in-depth review of available methods.

Molecular data can be extracted from the tissues and cells of living organisms or preserved specimens housed within natural history museums, and an enormous trove of archived data is available in widely used public databases (e.g., NCBI; Sayers et al. 2022). While the vast majority of molecular data are observations of extant and recently extinct species, scientific and technological advances have enabled integration of fossil taxa in molecular sequence alignments (Dalén et al. 2023). However, the degradation of DNA molecules limits paleogenomic sequencing to short fragments from well-preserved specimens that lived within the last 1 Myr (though the oldest sequenced specimen is ~2 Myr old; Kjær et al. 2022). For molecular data from much older fossils, researchers can turn to paleoproteomics techniques, which allow paleontologists to elucidate amino acid sequences from ancient fossilized proteins as old as 80 Myr (Schweitzer et al. 2009). While proteins have the potential to illuminate the evolutionary history of many of the dark branches on the tree of life, protein analysis of fossil material is labor-intensive, expensive, and destructive (Schroeter et al. 2021). Moreover, the DNA fragments and proteins accessible to paleogenomic and paleoproteomics studies will always be specific to certain tissue types and preservation conditions and influenced by all of the factors that structure the fossil record. Thus, the lion’s share of molecular data used in phylogenetic analyses will be sampled from modern taxa, and the evolutionary history of most

ancient fossil specimens will be analyzed using matrices of morphological characters.

Morphological Data. Morphological data are highly valuable for phylogenetic inference, as, unlike DNA sequences, morphological characters are available for a wider range of long-extinct and extant taxa (Lee and Palci 2015). However, compiling morphological data from fossils and/or extant taxa is an extremely labor-intensive task, which requires a thorough understanding of the morphology and taxonomy of the clade. First, phenotypic traits must be selected that describe an appropriate amount of variation in taxa across the clade (Serenó 2007). These traits must then be manually coded or measured for each taxon, using software such as Mesquite (Maddison and Maddison 2023) or MorphoBank (O’Leary and Kaufman 2011) to construct the matrix. This process requires access to adequately preserved specimens, which could involve visiting museum collections or sourcing morphological images, computed tomographic data, or descriptions from the literature. The coding scheme may require revision as additional taxa are included. Fragmentary fossil preservation as a result of taphonomic processes can present a major barrier to coding morphological characters and can result in characters being coded using multiple individual specimens believed to belong to the same operational taxonomic unit. Higher levels of matrix incompleteness are expected for morphological characters compared with genetic data (O’Reilly and Donoghue 2021).

Morphological data for phylogenetics can be either discrete or continuous. Discrete data may represent, for instance, whether a trait is present or absent, while continuous data are typically used for traits such as body size or the length of specific anatomical features. Different data types require different evolutionary models, so often only one data type is used, and typically these are discrete morphological data. Continuous traits are sometimes discretized in order to be included in such analyses. Here we focus on models of discrete character evolution, although continuous models are growing in popularity and have recently become available for use in FBD analyses (Zhang et al. 2024; see also Parins-Fukuchi 2017; Alvarez-Carretero et al. 2019).

In a discrete morphological matrix, characters can be binary or multistate, depending on the trait described. Binary characters exhibit only two states, such as presence or absence of a trait, while multistate characters can assume multiple states, allowing for greater variation and complexity in trait expression. Multistate characters are advantageous for capturing a wider range of morphological variation, but they also require more detailed trait description and coding and can be more complex to model accurately in phylogenetic analyses (Wright 2019). There are also a number of special characters coded in morphological matrices by taxonomists, which have distinct and specific meanings (Brazeau et al. 2019; Hopkins and St. John 2021; Simões et al. 2023b). For example, throughout morphological datasets, it is common to see “?” and “-”, to represent missing and non-applicable characters, respectively. These two types of characters are typically treated in the same way by available software; however, they represent fundamentally different pieces of information. A missing character, “?”, is used when all available specimens of a taxon are too decayed or fragmentary to determine whether it has a certain character trait or not. Non-applicable characters, “-”, are used when the trait is not associated with a taxon. An example of a non-applicable trait may be traits associated with ornamentation. One character trait may be used to represent the presence or absence of ornamentation. Some taxa

will have ornamentation, and therefore be coded as 1, while others may have none and be coded with a 0. A second character could then be coded to determine the type of ornamentation. In this case, all taxa that have a 0 for the previously described trait would be coded as non-applicable. There are several strategies for coding such morphological characters (see Hopkins and St. John 2021), but it is good practice to be consistent and explicit about the coding strategy (e.g., missing vs. non-applicable characters). Generally, however, contingent coding (where hierarchies of characters are included through the presence and absence of controlling/primary characters) offers the best solution to handling non-applicable characters (Simões et al. 2023b).

Given the intensive nature of collating morphological matrices and the incompleteness of preservation, these datasets tend to be quite small relative to molecular datasets. Additionally, the variation in matrix size will be dependent on the number of characters one can reasonably come up with for a given group. For instance, a group of marine invertebrates, like bivalves, may possess fewer obvious anatomical features compared with a group of vertebrates, such as large mammals, that have multiple distinct anatomical regions that can potentially be well preserved. From our survey, we found the number of morphological characters used in FBD analyses was typically greater for vertebrates than invertebrates (Fig. 1B).

Morphological matrices can be published and made available for download in a standardized format on databases such as MorphoBank (O’Leary and Kaufman 2011) or are often included in the supplementary material associated with a study.

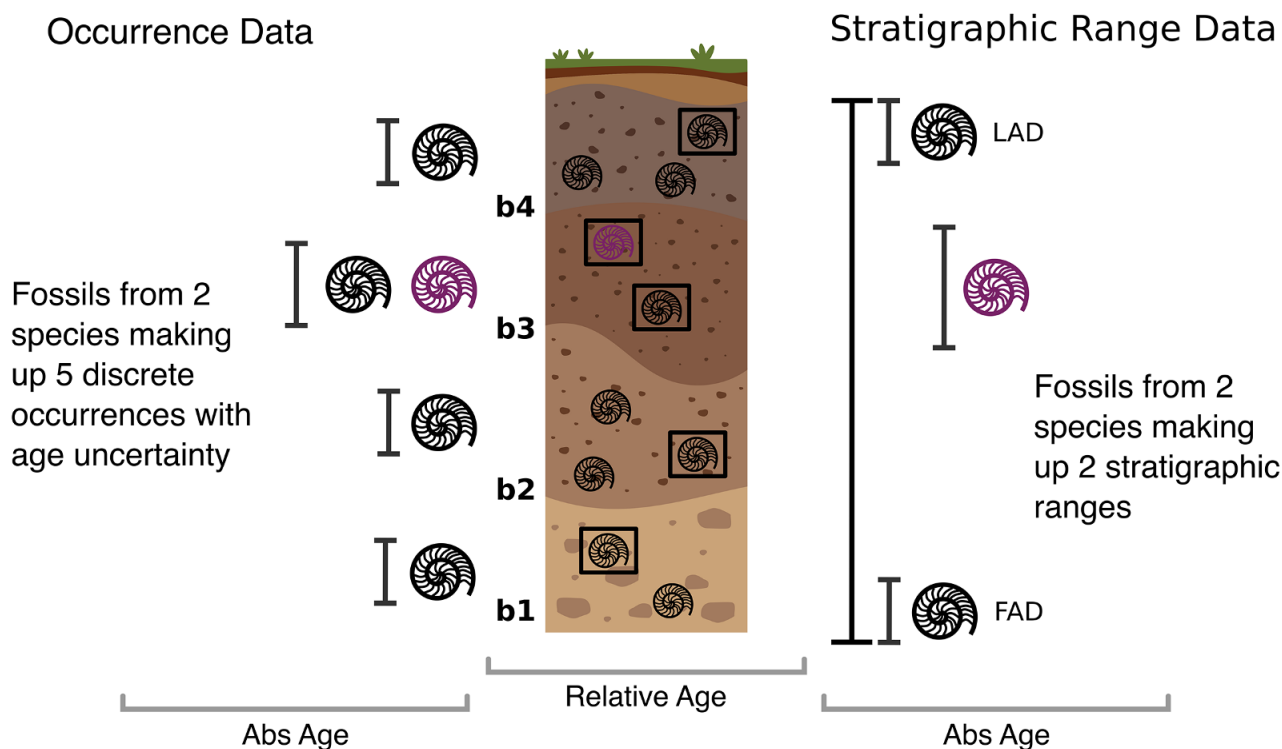
Temporal Data

The temporal information associated with a fossil taxon can fall into one of two categories: (1) occurrence data, where each occurrence is associated with a specific time or interval from which it was sampled; or (2) stratigraphic ranges, which describe the interval between the first and last occurrences of a taxon (see Fig. 4). Understanding the differences between these two data types and how they reflect fossil sampling is extremely important for carrying out an analysis using an FBD model.

Fossil occurrences each represent a single point in time corresponding to a known fossil sampling event and, in an FBD model, constitute evidence that the given lineage was extant at that time. Because fossil ages are often imprecise, occurrences are usually sampled from a discrete interval or range of ages. This interval should be informed by prior knowledge of uncertainty of the age of the fossil and often (but does not have to) represents a named geological interval. In this context, an occurrence may represent multiple specimens of the same taxon sampled from a single geological locality and/or time interval, or just a single fossil: FBD models do not currently include a way of incorporating information on abundance data. Having multiple specimens might be common for groups with high preservation potential, such as skeletonizing marine invertebrates or microfossils, while groups such as terrestrial tetrapods that have a much sparser fossil record may be more commonly associated with a single specimen per occurrence. If the same taxon is found in multiple time intervals, multiple occurrence times can be included in an analysis, each associated with its own upper and lower age bounds.

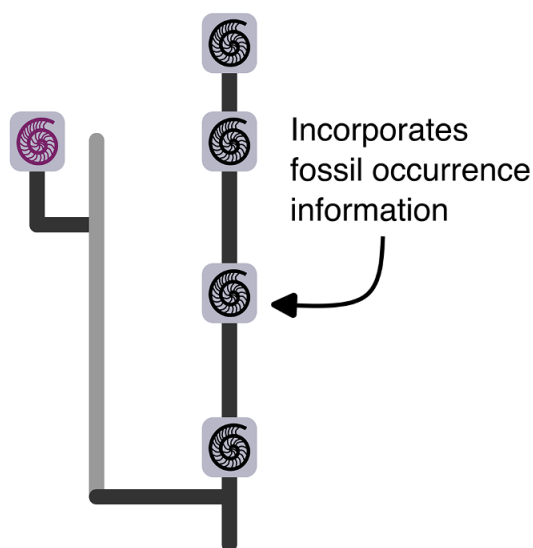
Stratigraphic ranges instead constitute an overview of all available occurrence data for each taxon, describing the age of the first occurrence (often termed “first appearance datum,” or FAD) and the last occurrence (often termed “last appearance datum,” or

A Fossil Ages



B FBD Models

FBD specimen model



FBD range model

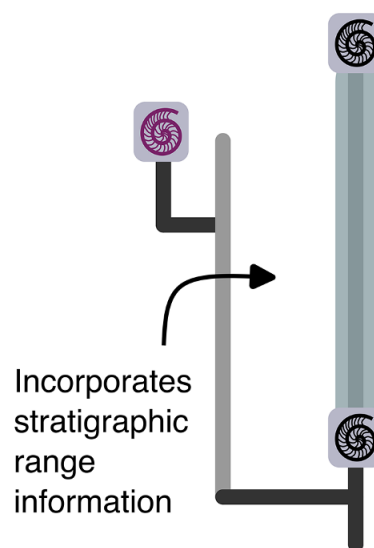


Figure 4 The temporal information available from fossils and how it can be incorporated into fossilized birth–death (FBD) models. **A**, Section with four fossil beds, b1–b4. Within each bed, there are fossils that can be used to provide temporal information for an FBD analysis. In this section, there are two different fossil taxa depicted as purple and black ammonites. Fossil age information can be taken as either occurrence data or stratigraphic range data. Occurrence data describe the age uncertainty associated with an individual sample or a discrete interval (shown to the left of the section). Stratigraphic range data describe the age around multiple fossils of the same taxon. The lower and upper bounds of the range (i.e., the first and last appearances) will also have a degree of age uncertainty around each of them (shown to the right of the section). Different FBD models are available to incorporate these are fundamentally different way of using fossil age information. **B**, How these different models incorporate the temporal information. The FBD specimen model uses occurrence information. Note that multiple fossil specimens from the same bed that are associated with the same age uncertainty should only be incorporated into the analysis once. FBD models do not currently have a way to account for abundance information. The FBD range model uses stratigraphic range information. In this case, it uses the first and last appearance fossil ages. Note, for the taxon in purple, there is only one fossil (i.e., a singleton); therefore, the occurrence and range information are the same. The gray branches on the tree represent unsampled lineages or taxa.

LAD). In some cases, only first and last occurrences are known, while for others, many occurrences may lie within this range. If a taxon only has one occurrence (known as a singleton), the stratigraphic range is analogous to the age information attached to that single occurrence. Stratigraphic ranges better reflect the age information we have available for many groups in paleobiology. This includes classic datasets, such as Sepkoski's 2002 compendium for marine genera (available via the *sepkoski* R package; <https://sepkoski.palaeoverse.org>), as well as datasets for which we have a huge number of occurrences spanning multiple intervals, for example, some groups of Cenozoic mammals.

The original implementation of the FBD model (here referred to as the FBD specimen model) assumes the fossil information constitutes occurrences (Stadler 2010). The FBD specimen model can use all available occurrences (computational limits permitting), but does not incorporate any information about how these occurrences may be associated with each other. The FBD range model instead uses stratigraphic ranges (Stadler et al. 2018) and explicitly takes into account the sampling of taxa through time, better reflecting the data available for some groups. Figure 4 illustrates the age information used by these different models. Note that the choice to use either occurrence data or stratigraphic ranges will have a strong impact on the sampling rate inferred by the model. If all taxa are represented by singletons, the input will be the same for both models.

Collating the necessary age information from the fossil record, regardless of whether occurrences or stratigraphic ranges are used, can be very involved. Determining the age of a fossil often requires synthesis of information from a combination of different sources (Benton et al. 2009; Parham et al. 2012; Hopkins et al. 2018). Under rare circumstances, it may be possible to directly date the bed in which a fossil was found, for example, using isotopic analysis of an ash layer, but this is uncommon. Often a single date will be measured for a stratigraphic section, and the ages of the rest of the stratigraphic column in that section will be inferred based on this age. Alternatively, biostratigraphic correlation can be used. This method relies on identifying index fossils in the same or nearby beds as the fossil of interest. Index fossils are defined as globally abundant, easy to recognize, and associated with a short interval in time, and can therefore be used to determine the *relative* age of a stratigraphic section. To establish the *absolute* age of the fossil, we then rely on chronostratigraphy. The global geological timescale allows us to use the index fossils to correlate their occurrences in a section to absolute time (Gradstein et al. 2020). The absolute ages have been determined in previous studies, using methods such as radiometric dating or Bayesian stratigraphic age models (Deino et al. 2021; Halverson et al. 2022; Lee et al. 2023). Given that a fossil age is inherently reliant on a number of external calibrations that combine information from different sources, the uncertainty around this age is extremely important to consider (Barido Sottani et al. 2019a, 2020b). As such, it is essential to account for the associated uncertainty when defining the minimum and maximum bounds for each of the individual occurrences in the analysis. Similarly, you also need to incorporate the uncertainty around the first and last occurrences when using stratigraphic ranges.

There are different approaches to collecting age information. It may be necessary to compile age information directly from the literature. Descriptions of fossils, particularly type specimens, usually include the stratigraphic context in which the fossil was found. For example, the biozone in which it was found can provide relatively accurate and high-resolution bounds on the fossil's age. However, if one is working with fossils from different biogeographic regions, biozones may need to be correlated and/or translated directly into numerical ages, which requires expertise

and careful consideration to be done appropriately. Alternatively, age information can be obtained from large databases of fossil occurrences, such as the Paleobiology Database (PBDB; Uhen et al. 2023).

Getting Data from the PBDB. Taking age data from the PBDB requires some understanding of the structure of the database. Comprehensive guidance can be found in the database's User Guide (Uhen et al. 2023). Age information can be accessed using the PBDB web portal or API (application programming interface), either as individual occurrences or summarized as stratigraphic ranges (first and last occurrences) for each taxon. We recommend downloading full occurrence data, as this is more amenable to vetting than age ranges. Other important factors to consider are whether to include or exclude trace fossils e.g., Simões and Pierce 2021. The identity of the tracemaker can rarely be established with high taxonomic resolution, but such fossils could still be incorporated as age constraints in an analysis.

Age information is provided in two different sets of columns: "early interval" and "late interval", and "min ma" and "max ma". The "early interval" and "late interval" are the named geological intervals provided by the data enterer corresponding to the age range of the occurrence; these might be internationally recognized interval names or regional biozone/formation names. When the age uncertainty of an occurrence spans a single interval, it is named in the "early interval" column, while the "late interval" column is left blank. The "min ma" and "max ma" columns give the numerical ages, in millions of years, corresponding to the full temporal range of the time intervals according to the PBDB's internal chronostratigraphic scheme. Tools also exist for cleaning and processing these data, such as the R package *palaeoverse* (Jones et al. 2023), which can be used to update interval ages according to more recent chronostratigraphies and to summarize occurrence data into first and last occurrence dates.

All of the fossil occurrences then need to be assigned to a taxonomic lineage. Key information about the identification of each fossil occurrence can be found in the "identified name" and "accepted name" columns. The "identified name" is the identity given to the occurrence in the original reference describing it. This might include some indicator of uncertainty in the identification, such as "cf." or "?". Such occurrences should be evaluated carefully before a judgment is made as to whether they should be included in the age information used in a phylogenetic analysis. The "accepted name" is the identity given to the occurrence following processing by the database's internal taxonomic structure. This reallocates occurrences of e.g., synonyms, and invalid taxa, to provide an updated identification with respect to taxonomic opinions in the recent literature. Usually this column is the one used when placing occurrences onto the phylogeny.

When data from the PBDB are used, it is crucial that the data downloaded are validated by an expert. As with any big database, errors can be introduced to the data in a number of ways. Given that taxonomy and stratigraphy are dynamic fields, the classification of taxa and time intervals can also change over the years, resulting in some taxa being identified or dated incorrectly. Having a taxonomic expert verify the data should be a standard step in any phylogenetic analysis using these data.

Placing Fossils in a Tree

When we analyze paleontological data in a phylogenetic analysis using the FBD model, the placement of fossil observations is either

specified a priori or estimated as part of the inference. If we have character data for the fossil, we can directly infer its position in the phylogeny during the analysis. Alternatively, we can rely on taxonomic placement, which does not necessarily require any character data to be available for the fossil, but still allows its age information to be incorporated into the analysis. Because we specify taxonomic affiliation or phylogenetic placement based on prior information, we consider this part of the prior probabilities rather than data in the construction of our Bayesian model (Fig. 3).

Taxonomic placement, or bracketing, can be used to constrain parts of the tree topology (Soul and Friedman 2015). This can be applied to both extant and extinct samples in the tree, although we will focus on extinct (i.e., fossil) placement here. Fossils (occurrences or ranges) can be fixed within a given clade in the tree, denoted using a node that we are confident represents the common ancestor of the clade to which our fossil belongs. The taxonomic resolution of the fossil will influence its placement in the tree: fossils that can be easily identified might be restricted to a specific genus, however, for other fossils, it may only be possible to assign them to a specific family or order. As this approach does not require fossils to have any associated character data, the placement of a fossil taxon cannot be precisely inferred, and instead we account for the possibility that it could attach anywhere within the clade defined using the taxonomic constraints (Gavryushkina et al. 2014; Heath et al. 2014). Fossils without character data should be pruned from the inferred trees before any downstream interpretation of taxonomic relationships.

Determining phylogenetic relationships among extant species can be difficult, even with molecular data in hand (Bergsten 2005; Huson and Bryant 2006; Degnan and Rosenberg 2009). Establishing the position of extinct taxa is even more challenging, both with other extinct and extant taxa. No fossil is a perfect representation of the living organism: from the time the organism dies, anatomical information is lost through the decay process, in a manner that can bias its placement in phylogenies (Sansom et al. 2010; Pattinson et al. 2014). Similarly, geological processes such as metamorphism and erosion can negatively impact the preservation of a fossil. Highly trained taxonomists who know their study group and understand these processes are therefore crucial for establishing taxonomic affiliation and to ensure the correct incorporation of fossils into an analysis. Simulations have demonstrated the importance of well-described anatomical data, showing that without fundamental taxonomic research, even the current “best” available models will produce erroneous results (Barido-Sottani et al. 2023).

Both character and temporal data are informative about the phylogenetic position of fossils (Barido-Sottani et al. 2020b; Mongiardino Koch et al. 2021). However, when character information is available, the position of a fossil in the tree may be more precisely inferred. In this case, samples may be associated with both morphological data and occurrence times, which can be true for fully extinct trees or trees combining extant and fossil observations. Analyses including molecular data for extant tips, alongside morphological data for at least a subset of those taxa, and fossils with morphological and temporal data, are often referred to as total-evidence analyses (Ronquist et al. 2012a; Zhang et al. 2016; Gavryushkina et al. 2017; see Box 4). When fossils or extant samples are associated with character data, it is still possible to constrain parts of the tree topology in the case of strong prior knowledge. This may be more computationally efficient, as the analysis searches a smaller region of tree space and can therefore help a Bayesian analysis to converge faster and improve overall accuracy (Barido-Sottani et al. 2023). Table 1 summarizes the different types of data used in

applications of the FBD model to date. See Box 2 for a discussion on the use of previously inferred or fixed trees and Box 5 for an example.

The Models

The FBD Process

Stadler (2010) was the first to present a stochastic model of cladogenesis that also accounted for sampling lineages through time. This serially sampled constant-rate birth–death model was called the “fossilized birth–death process” by Heath et al. (2014) to denote applications of the model in macroevolutionary analyses including fossil taxa. The FBD model is an extension of the birth–death processes that are important models for describing the diversification of a set of contemporaneous (extant) taxa (Yule 1924; Kendall 1948; Thompson 1975; Raup 1985; Nee et al. 1994; Patzkowsky 1995; Yang and Rannala 1997; Gernhard 2008; Stadler 2009). Birth–death models also play an important role in quantitative paleobiology. For example, these models underpin the theory behind the boundary-crosser rate estimators (Foote 2000) and the software PyRate (Silvestro et al. 2014, 2019), both used to estimate origination and extinction rates.

The FBD process is a mechanistic model that combines diversification and sampling processes, such that the parameters have both a biological and geological interpretation (Stadler 2010; Heath et al. 2014). The model generates rooted phylogenies with branch lengths proportional to time and a set of sampled fossil and extant lineages. It is modulated by five parameters: origination (or speciation) rate λ , extinction rate μ , fossil sampling rate ψ , extant sampling probability ρ , and the origin time of the process t_{origin} . In this section, we describe the process of simulation under the FBD model to provide a more intuitive explanation for the model’s assumptions and behavior.

When we simulate a tree forward in time, the process starts with a single lineage. This lineage can then speciate or go extinct according to rates λ and μ , respectively. Lineages can also be sampled at any instance according to rate ψ . An origination event assumes budding speciation and results in a single new lineage, which will then be associated with its own FBD process, that is, it can speciate, go extinct, or be sampled, independent of other lineages. The simulation can continue until a given duration of time or a given number of coexisting (extant) lineages is reached (Hartmann et al. 2010). The lineages extant at the end of the process ($t = 0$) can be assumed to represent living species and be sampled with probability ρ . The time between the start and end of the simulation is equivalent to the origin time, $t = t_{\text{origin}}$. The origin represents the beginning of the process, before any other events (speciation, extinction, or sampling) have occurred and is therefore older than the root age of the tree. At the end of this simulation process, the generated complete (true) tree contains all of the samples associated with the simulated lineages. Some internal branches will have fossil sampling events along them, and some extant tips will have associated sampling events (if $0 < \rho < 1$). Not all extinct branches will necessarily have fossil sampling events associated with them. To obtain the reconstructed tree—that is, the tree that can be inferred using the observable set of samples—we prune all unobserved lineages from the complete tree (see Stadler 2010; Fig. 1). The reconstructed tree only includes tips that are associated with fossil or extant sampling events. For any dataset generated by the FBD model where $\psi > 0$, there is a nonzero probability that some fossil

samples will also have sampled descendants (Foote 1996). The proportion of these “sampled ancestors” increases with higher values of ψ and ρ (Pett and Heath 2020). Additionally, as extinction rates decrease relative to speciation rates, sampled ancestors become more prevalent. When we apply the FBD model to empirical data, we will estimate the reconstructed tree, including the placement of fossils as sampled ancestors. Because we cannot observe unsampled lineages, we also cannot estimate the complete tree. Forward simulation has played an important role in the development of FBD models, providing the necessary tools for examining the accuracy of Bayesian phylogenetic methods and helping to characterize model behavior and assumptions (FossilSim [Barido-Sottani et al. 2019b] and paleobuddy [do Rosario Petrucci et al. 2022] are R packages that simulate under a wide range of FBD models).

When simulating under the FBD process, we specify known values for the diversification and sampling parameters to generate a tree. When we use the model for inference, however, our aim is to infer these rates, along with the reconstructed tree topology and divergence times. The FBD model allows us to calculate the probability of a reconstructed tree, given the FBD process parameters, t_{origin} , λ , μ , ψ , ρ , and the probability of the fossil sampling times, given the tree and the model parameters. As the fossil ages are considered to be data, we include them as part of the likelihood calculation in our representation of the FBD model in Bayes theorem (Fig. 3). In any birth–death process, the parameters for speciation, extinction, and sampling are highly correlated, and fixing one of them to a single value can ensure that the model is fully identifiable (Gavryushkina et al. 2014; Truman et al. 2024). Typically, it is recommended to specify a fixed value for ρ , as researchers often have prior knowledge of the known extant diversity of the clade of interest. Thus, ρ is often set to the proportion of sampled taxa as an approximation of the probability an extant

taxon is sampled. In the original model, extant taxon sampling is assumed to be uniform among branches surviving to $t = 0$, but this assumption, along with many others, can be relaxed. Table 2 presents the FBD family of models and their assumptions, and Box 6 presents information on different parameterizations and prior choices.

Clock Models

The clock model describes the rate of character change along a tree with branch lengths in units of time. These models generate a set of branch-specific substitution rates. With these values, it is possible to transform the branch times of the FBD tree into units of evolutionary distance (i.e., the number of substitutions/site). The likelihood of the character data is then calculated using these branch lengths and the parameters of the character substitution model (Fig. 3). A number of different models have been introduced to describe the distribution of substitution rates among branches (see Heath and Moore 2014), all of which can be applied to molecular or morphological data. In an analysis using both data types, that is, a total-evidence analysis, the clock model governing variation in molecular substitution rates is typically assumed to be independent from the process generating variation in morphological branch rates (see Box 4 for an example specifying separate molecular and morphological clock models). This is because we do not expect the rate of morphological and molecular evolution to be exactly the same for any given branch. It is also possible to apply different clock models to different subsets of molecular or morphological data (Simões et al. 2023a).

Branch-rate models can be categorized into two types: strict or relaxed clock models (for comprehensive reviews of clock models, see Heath and Moore 2014; Bromham et al. 2018). Under a strict clock model, the rate of evolution is constant across every branch in the tree (Zuckerkandl and Pauling 1962, 1965). Thus, this model

Box 6. Choosing priors.

Under the standard fossilized birth–death (FBD) parameterization, we place prior distributions directly on the FBD model rate parameters, namely origination (λ), extinction (μ), and fossil sampling (ψ). Often, we use exponential priors for these rates—this distribution places a high probability on values close to zero, which is typical of estimates obtained from fossil data, but does not preclude larger values. The rates λ , μ , and ψ are always used to calculate the probability under the FBD model, but we also have the option to parameterize the model using different combinations of transformed parameters, enabling us to instead set priors on these values. For instance, we can place priors on diversification (d), turnover (r), and the probability of sampling before extinction (s), which can be transformed during inference using the formulas shown below to recover λ , μ , and ψ . One advantage of parameterizing the model using d , r , and s is that r and s can be bounded within the range $[0, 1]$, if we assume $\lambda > \mu$, that is, that net diversification is positive, in contrast to λ , μ , and ψ , which are all in the range $(0, \infty)$ (Heath et al. 2014). It will not always be appropriate to assume $\lambda > \mu$ (see Marshall 2017), but users can and should make the choice based on their specific datasets. Many possibilities for constraining and transforming model parameters exist within a flexible Bayesian framework. This is particularly true of more parameter-rich models, such as skyline (see Table 2, Box 5), which can allow for a more complex set of priors. In an analysis of Cambrian echinoderms, Wright et al. (2021) used the FBD skyline model, which permits variation in evolutionary rates between time intervals. They used an exponential prior for speciation, and constrained extinction such that turnover was within the range $(0.90, 1.05)$, reflecting the observation that λ and μ tend to be correlated. Rates can also be linked across adjacent intervals (for details, see Zhang et al. 2023). Finally, it is possible to constrain or even fix the FBD model parameters based on independent estimates, which is often done for extant sampling probability, ρ (see Section The FBD Process), but can easily be done for other parameters, such as fossil sampling (e.g., O’Reilly and Donoghue 2021). Within paleobiology, we often work with per-interval sampling probabilities (Foote and Sepkoski 1999; Alroy 2008), which can also be transformed to recover ψ (e.g., Warnock et al. 2020), allowing us to take advantage of previous work in paleobiology.

Net diversification	$d = \lambda - \mu$
Turnover	$r = \lambda / \mu$
Probability of sampling before extinction	$s = \psi / (\mu + \psi)$
Origination/speciation	$\lambda = d / (1 - r)$
Extinction	$\mu = r d / (1 - r)$
Fossil sampling	$\psi = (s / (1 - s)) (r d / (1 - r))$
Per-interval sampling probability	$P_t = 1 - e^{-\psi_t \Delta t}$
Per-interval rate	$\psi = -\ln(1 - P_t) / \Delta t$

has a single free parameter—the global substitution rate—that is applied to every branch on the tree. While this model's simplicity is attractive, the assumption of constant rates is unlikely to hold for most molecular or morphological datasets, especially over deep macroevolutionary timescales.

Several relaxed clock models have been introduced that allow rates of character evolution to change over the tree. These models have different assumptions about the pattern of rate variation and the underlying processes driving rate change, including local molecular clocks (Hasegawa et al. 1989; Yang and Yoder 2003; Drummond and Suchard 2010), autocorrelated rate change (Thorne et al. 1998; Thorne and Kishino 2002; Huelsenbeck et al. 2000; Lepage et al. 2007), time-dependent rates (Membrebe et al. 2019), mixture models on branch rates (Heath et al. 2012), and uncorrelated-rate models (Drummond et al. 2006; Lepage et al. 2007; Rannala and Yang 2007). In particular, uncorrelated-rate relaxed clock models assume that the substitution rates associated with each branch in the tree are independently drawn from an underlying probability distribution. Because the branch rates are independent, these models can be sampled efficiently with MCMC and are the most common relaxed clock models used in Bayesian divergence time estimation. Specifically, in paleobiology, one of the more frequently used uncorrelated-rate models is the uncorrelated lognormal (UCLN) model (Drummond et al. 2006). In the UCLN clock model, independent branch rates are drawn from a lognormal distribution (Drummond et al. 2006). The mean of the lognormal distribution corresponds to the average substitution rate, while the variance of the distribution reflects the degree of rate variation across the tree. The mean and variance are specified using separate prior distributions and are informed by the character data in the likelihood (Fig. 3).

Uncorrelated-rate models characterize rate patterns and are not based on any particular biological process. These models are also widely available across different Bayesian phylogenetics software, making them a popular choice for divergence time analyses. However, the wide range of relaxed clock models mentioned earlier may have other properties that are worth considering for analysis of fossil (and extant) taxa. Autocorrelated-rate models, in particular, seek to account for an underlying biological process driving variation in rates of character evolution over the tree. Under these models, closely related lineages are more likely to have similar rates, as these values change gradually along the tree. Moreover, some studies have shown that the use of autocorrelated versus uncorrelated molecular clock models can have a large impact on divergence times (e.g., Lepage et al. 2007; Dohrmann and Wörheide 2017). It is important to note, however, that much of the work investigating the fit and adequacy of clock models has focused on applications to molecular data. Thus, the performance of different morphological clock models for datasets including fossil taxa remains underexamined (however, see, Lee 2016; Simões et al. 2020; Congreve et al. 2021).

Substitution Models

Substitution models describe the evolution of character states along a tree with branch lengths that correspond to evolutionary distance. They are used to calculate the likelihood of observing our phylogenetic character data in Bayes theorem (Fig. 3). As described in The Data section, the character data used in phylogenetic analyses are typically either morphological or molecular. There are a number of substitution models for molecular data; however, as these data are not the core focus of this review we do not go into detail here

(but see Sullivan and Joyce 2005; Yang and Rannala 2012; Arenas 2015). There are comparatively fewer models available for morphological data, as developing models that generalize the underlying processes of morphological evolution is difficult (Wright 2019). In part this is because the individual characters defined in morphological matrices are much more variable in definition and scope than those in molecular data. The name “substitution” model stems from the description of substitutions between nucleotides. Despite its poor translation, this same name is often also applied to analogous models for morphological data where we are not modeling substitutions, but rather transitions between morphological phenotypes. To address this mismatch, these models are sometimes referred to as morphological substitution models, discrete models, or discrete models of character evolution, as well as simply substitution models, often interchangeably. It is important to be aware that they all refer to the same type of model.

The most common model applied to morphological data is the Mk model (Lewis 2001), where M refers to Markov, and k to the number of states. That is, if this model is used with a dataset with four character states, it may be referred to as an M4 model. This model is equivalent to the Jukes Cantor model used for molecular data (Jukes et al. 1969). It assumes equal transition rates between character states, that is, transitioning from state 1 to 0 is equally probable as transitioning from 0 to 1, or from 0 to 2, and so on. This simplistic model has subsequently been extended to better reflect the properties inherent to morphological data. The first extension, described in the same paper as the standard Mk model (Lewis 2001) is the MkV model. This extension accounts for the fact that in a morphological matrix, most (if not all) of the sites will vary. That is, when taxonomists are creating morphological matrices, it is common to code preferentially for phenotypic traits that vary among species. This results in a matrix in which all sites have varying characters. This failure to include characters that are shared by all species in the dataset introduces what is known as an “ascertainment bias”, which can result in overestimation of the substitution rate and therefore branch lengths (evolutionary distance) in the inferred tree (Lewis 2001). The MkV model instead conditions on the fact that there are only variable sites in the alignment, ensuring that branch lengths are not overestimated. This is achieved by calculating the conditional likelihood of a site, given that it is variable. In practice, a dummy character is added to the alignment to calculate the probability of a constant character (given a tree and set of branch lengths), which can then be used to calculate the conditional likelihood.

In a morphological matrix, different character traits can describe many aspects of a phenotype, and so might be expected to transition more or less frequently. To account for variation in transition rates between traits, researchers can allow for rates among characters to vary according to a gamma distribution (Yang 1994), the MkV+ Γ model. Wagner (2012) explored the use of a lognormal rate distribution to describe this variation in rates, although a gamma distribution remains the most commonly applied. Aside from character rate variation, it is also important to consider whether or not it makes sense to partition the data. Partitioning a dataset involves allocating traits to groups that are thought to have evolved under similar conditions or pressures. Each group of traits is then modeled with a separate Mk model, such that the likelihood is then a product of a number of Mk models. Partitioning morphological data by the maximum observed state is the most common approach and is the default in several phylogenetic software programs, for example, MrBayes (Ronquist et al. 2012b) and BEAST2 (Bouckaert et al. 2019). Characters can also be partitioned based on some prior

knowledge or hypotheses based on anatomical region or function, for example, cranial and postcranial or feeding and non-feeding structures (Klopfstein et al. 2015; Simões and Pierce 2021; Casali et al. 2023). Choice of morphological model has been shown to impact both the branch lengths and topology of inferred phylogenies (Khakurel et al. 2024). It is important to provide justification for substitution model choice through approaches such as model adequacy (Mulvey et al. 2024).

As described in the Section on Character Data, there are often some special characters in morphological matrices, such as missing and non-applicable characters, denoted as “?” and “-”, respectively. These characters are often treated the same way by standard Mk models, despite the fact they have distinct meanings. Further, non-applicable characters violate the Mk model, which assumes that each character trait is independent of all others (Hopkins and St. John 2021). In spite of this model violation, Bayesian inference still outperforms parsimony methods specifically designed to handle non-applicable characters (Simões et al. 2023b). New models are also being developed to account for nested characters, such as the embedded dependency model, which takes into account the dependency one trait can have on another (Tarasov 2023), although this model is yet to be applied in an analysis using the FBD model. Further development is required in these areas to enable us to more accurately model morphological evolution.

Assessing and Interpreting the Output

Examining MCMC Efficiency and Performance. When Bayesian statistical inference is performed using sampling algorithms like MCMC, it is very important to evaluate the performance and efficiency of the sampler (Cowles and Carlin 1996; Gelman 1996). This involves a range of different techniques for assessing convergence (often called “convergence diagnostics”). Ideally, we would like for our Markov chain to converge on the true posterior distribution and sample parameter values in proportion to their posterior probability. However, we can only be certain that this type of convergence is reached after an infinite number of MCMC iterations. Because we are restricted to finite timelines to complete our research, we instead use different qualitative and quantitative approaches for identifying signs that our chains are yielding sensible samples and are not stuck in narrow regions of parameter space. Convergence diagnostic methods ultimately tell us if our Markov chain has converged on a stationary distribution and is sampling values from that distribution in proportion to their probability.

Perhaps one of the most useful approaches for evaluating the performance of MCMC is to run multiple, independent chains and compare their samples and probabilities (Gelman 1996). For this to work effectively, each analysis must start with different values for the free parameters. Once each run has collected a desired number of samples, we can then compare the marginal posterior distributions of our numerical parameters (e.g., speciation rate, clock rate). For example, if the distributions of samples for speciation rate match between multiple independent Markov chains, then that is evidence that the runs have sampled from the same stationary distribution. For more complex parts of our model like the tree topology, we can compare bipartitions or posterior probabilities or use other topology diagnostics (Nylander et al. 2008; Lanfear et al. 2016; Warren et al. 2017; Fabreti and Höhna 2022).

Another important attribute to examine is correlation among samples for a given parameter within a single run. Because MCMC simulates dependent samples and we do not want our samples to be highly correlated, we typically specify the frequency at which we

save values to a file. Thus, if this frequency is set to 100, then our MCMC output will only save our parameters at every 100 iterations. However, even after our MCMC sample is thinned, the values for some parameters may still be autocorrelated. An MCMC algorithm’s ability to efficiently sample a target distribution (the posterior) is often referred to as “mixing”. We can visually look for good or bad mixing in a numerical parameter by plotting the sampled value as a function of the timeline of the Markov chain. If the chain appears to sample clusters separated by large valleys or show a very gradual trend, then this is often called “bad mixing”. Conversely, if we plot the values sampled by a run that is mixing well, we will see regular oscillation over the range of probable values (it is often said that this pattern resembles a “fuzzy caterpillar” like the larvae of the garden tiger moth) without any clusters or gradual trends (Ali et al. 2017; Roy 2020). In addition to employing visual inspection, we can also quantify the degree of autocorrelation among samples of a given parameter (for a single run) using a summary statistic called the effective sample size (ESS). This value is a measure of the effective number of independent draws from the stationary distribution for a parameter. In phylogenetics, an ESS of 200 is generally considered to be the minimum threshold indicating good mixing. However, this may not be reasonable as a universal rule of thumb across all parameters and phylogenetic models (Fabreti and Höhna 2022).

One important tool for assessing MCMC convergence is the program Tracer (Rambaut et al. 2018), which allows for visual inspection of MCMC samples, in addition to assessment via metrics like the ESS. Tracer takes the MCMC output files of numerical parameters and renders plots of the sampled parameter traces and marginal densities, as well as many other visualizations. A wide range of convergence diagnostics (for numerical and tree parameters) are also available in the R packages RWTY (Warren et al. 2017), Convenience (Fabreti and Höhna 2022), and RevGadgets (Tribble et al. 2022).

After a range of different MCMC performance assessments are carried out, the resulting diagnosis may be that your runs are mixing well and have converged on the same stationary distribution. Then, the next step is to summarize the MCMC samples and proceed to interpreting your results. When your evaluation of MCMC output indicates poor mixing and non-convergence, however, it is not appropriate to assume that you have sufficient samples of your posterior distribution. There are many strategies for improving MCMC efficiency, including running the analysis much longer, increasing the number of moves for the parameters with poor mixing (by increasing proposal weights), adding different types of moves to parameters that are not converging, removing the initial samples (often called the “burn in”), increasing or decreasing the variance in proposed values for some moves, or introducing more prior information to reduce the complexity of parameter space (for more on diagnosing non-convergence, see Barido-Sottani et al. 2024). Phylogenetic data, models, and analyses are extraordinarily complex, with many parameters and computationally intensive calculations. Accordingly, MCMC analyses of phylogenetic problems require a lot of time and computing power (and patience!) to sufficiently sample a complex posterior distribution.

Summarizing MCMC Samples and Interpreting Results. Depending on your research question, you may be interested in different aspects of the inference output. Estimates of numeric parameter values, such as branch lengths, divergence times, and substitution rates, can be summarized using Bayesian credible intervals. The highest

posterior density (HPD) interval is widely used in Bayesian phylogenetic inference, which is the shortest range within the posterior that captures 95% of the posterior distribution. This eliminates the tails of the posterior distribution, while still reflecting uncertainty in the parameter estimates. In addition, it can be useful to visually explore the posterior output, which can also be done using Tracer or R, to examine the overall shape of the distribution. This will indicate whether your posterior is uni- or multimodal. You can also compare the posterior to your prior distributions to determine the level of signal in your data. It is conventional to run your inference “under the prior” (many Bayesian phylogenetic software packages have an option for doing so) and to compare the prior and posterior distributions. Note that this typically only excludes the substitution model likelihood and not the likelihood associated with the FBD model (Fig. 3). It therefore allows you to assess the signal for the model parameters (including divergence times) coming from the fossil ages versus the character data.

Summarizing across the posterior distribution of tree topologies is more difficult, especially when there is a high degree of uncertainty around the topology (O'Reilly and Donoghue 2018). Two of the more common summary tree approaches used in paleobiology are the maximum clade credibility (MCC) tree and the maximum a posteriori (MAP) tree. The MCC tree consists of clades that have the highest probability across all trees to best represent the evolutionary history of the group, while taking into account the uncertainty. The MAP tree is an actual sampled tree in the posterior distribution that has the highest posterior probability. Summary trees can be generated using inference tools and visualized using R or more specialist tree visualization software, such as FigTree (<http://tree.bio.ed.ac.uk/software/figtree>) or IcyTree (Vaughan 2017). There are also methods for visually exploring the entire posterior tree space (Wright and Lloyd 2020; Smith 2022). This can be a good way to visualize the impact of different models or priors on the posterior tree space (e.g., May et al. 2021; Mulvey et al. 2024).

It is always important to consider the uncertainty associated with phylogenetic parameter estimates, including the topology, regardless of whether the tree topology is the main focus of the study. Although trees inferred using molecular data can be highly uncertain, datasets used for tree inference in paleobiology are often relatively small (e.g., <100 characters), and therefore more likely to result in a high degree of uncertainty around the topology (Barido-Sottani et al. 2023). When time-calibrated trees are used for downstream analyses, such as phylogenetic comparative analysis or ancestral-state estimation, it is important to represent this uncertainty by using the entire posterior, or a subset, in secondary analyses, depending on computational limitations (Soul and Wright 2021).

Challenges and Perspectives

FBD models provide a robust way of incorporating paleontological data into phylogenetic analyses, thereby allowing us to better study evolution through time. While these models have a number of advantages over alternative approaches, it is important to appreciate their limitations, along with the challenges associated with modeling fossil data. There are a number of important questions still to be fully addressed or resolved. For example: Are our models sufficiently complex to adequately represent biological phenomena? How do identifiability issues impact parameter estimates? Do available models fit our empirical datasets?

Developing models that encompass biological, as well as geological and sampling processes, is challenging. While FBD models

offer an advanced framework for inferring phylogenetic parameters in paleobiology, at their core, the assumptions regarding the generating processes are relatively simplistic (see Section The FBD Process). When considering whether or not our models make reasonable assumptions, it is worth reflecting on the complexity of the system we are modeling. The most widely applied FBD model assumes a constant rate of fossil sampling through time (Supplementary Table S2), which does not reflect the true nature of fossil data. The geological and stratigraphic records are non-uniformly incomplete (Benton et al. 2011; Holland 2016; Benson et al. 2021). Physiological properties, such as the differential mineralization of hard parts, as well as the environment in which an organism degrades, also influence the probability of becoming a fossil (Nanglu and Cullen 2023). In addition, a range of factors impact the probability of a fossil being sampled, for example, rock exposure (Dunhill et al. 2012) or the socioeconomic status of the country in which the fossil is now deposited (Raja et al. 2022). As a result of these biases, the fossil record is highly unevenly sampled across time, space, and lineages in the tree of life. Identifying generalizable processes that determine the structure of the fossil record will be crucial to the development of more realistic sampling regimes in FBD models.

There have already been numerous extensions to the FBD model that relax the assumption of uniform extant or fossil sampling (see Table 2). In particular, the FBD skyline model can accommodate variation in rates through time (Gavryushkina et al. 2014), among discrete time intervals, analogous to many approaches for inferring evolutionary rates from fossil occurrence data (e.g., Foote 2000; Alroy 2008, 2010; Liow and Nichols 2010). However, thus far, FBD skyline models have rarely been applied in empirical studies (only nine times according to our literature survey; Supplementary Table S2). Model extensions can also account for diversified sampling of extant taxa (Höhna et al. 2011; Zhang et al. 2016), which has been applied more often in empirical studies (34 times according to our literature survey; Supplementary Table S2). Beyond this, further extensions are currently in development. For example, multi-type FBD processes allow for variation in diversification or sampling rates across lineages (Kühnert et al. 2016; Barido-Sottani et al. 2020c). This framework could, in theory, be used to account for variation in spatial sampling (Antell et al. 2024) or other causes of variation in preservation potential (Nanglu and Cullen 2023) or diversification across lineages. However, as model complexity increases to better account for complex generating processes and the sizes of our datasets increase, the computational expense of running an inference also increases. Analyses can quickly become expensive to the point where using them may be impractical. Approaches for alleviating some of the computational expenses involved in running these analyses could include constraining or fixing a large portion of the tree (Lloyd and Slater 2021; Farina et al. 2023), using parsimony or maximum-likelihood approaches to generate a starting tree, or uniformly subsampling fossil occurrences (O'Reilly and Donoghue 2020).

Increased model complexity can also lead to identifiability issues. A model is identifiable if unique sets of parameter values produce unique observations. Non-identifiability can therefore be a major concern, arising when our data are compatible with multiple sets of parameter values. We can see how this may occur fairly intuitively using the FBD model parameters: an increase in the number of fossils over time could be caused by an increase in the speciation rate or an increase in the sampling rate. Our reliance on small, incomplete datasets exacerbates this issue, as often there is insufficient signal in the data to help recover precise parameter

estimates, especially for complex models with many parameters. In addition, the broad structure of birth–death models results in inherent identifiability issues that cannot necessarily be resolved with the addition of more data (Louca and Pennell 2020; Louca et al. 2021; Andréoletti and Morlon 2023; Kopperud et al. 2023). Prior choice also plays an important role in the identifiability of Bayesian models. For example, in analyses that employ the FBD model, it is common to fix the sampling probability ρ in order to ensure that the model is identifiable. In addition, we could consider the potential role of prior information on other parameters. For example, fossil-based estimates of sampling probability can be used to inform the sampling rate (Wagner and Marcot 2010; O'Reilly and Donoghue 2020; Wright et al. 2021; Thuy et al. 2022). Any implications of prior choice for the reliability of the results are rarely discussed in the literature. The relationship between prior choice and identifiability in FBD models is therefore one of the more challenging aspects of their application and needs to be investigated more thoroughly in future.

Because we can simulate data under FBD models, as well as perform inference, simulations are a valuable tool for exploring the behavior and limitations of these models (Barido-Sottani et al. 2019b; do Rosario Petrucci et al. 2022). Several simulation studies have explored scenarios violating FBD model assumptions to determine whether they lead to inaccurate results (Matschiner 2019; O'Reilly and Donoghue 2020; Barido-Sottani et al. 2023). Simulation studies have also been used to explore how resilient the model is to nuances of fossil sampling (Barido-Sottani et al. 2020b; O'Reilly and Donoghue 2020; Luo et al. 2020, 2023; Warnock et al. 2020; Mongiardino Koch et al. 2021). Together, these studies show that increasing overall fossil or taxon sampling and/or the number of morphological characters available for fossil samples improves the accuracy of parameter estimates under the FBD model. Further, simulations have shown that stratigraphic age or taxonomic information used to inform the inference must be accurate (Barido-Sottani et al. 2019a, 2020b, 2023). Collectively, these simulation studies demonstrate that empirical expertise is key to obtaining reliable results when using FBD models. It is therefore crucial that fundamental stratigraphic and taxonomic research remain well supported and that we, as a scientific community, find ways of incentivizing data collection and supporting the infrastructure around paleontological databases (Smith et al. 2023; Uhen et al. 2023).

Ultimately, the question remains as to whether available FBD models provide a reasonable fit to empirical fossil datasets. Standard model selection approaches cannot be used to compare the relative fit of alternative FBD models (for details, see May and Rothfels 2023). This means we need alternative approaches for assessing the fit of FBD models to empirical fossil data. An alternative to model selection is to use a model adequacy approach, such as posterior predictive simulations, which allow us to compare the absolute (rather than relative) fit of different models to our data (Brown 2014a,b; Duchêne et al. 2018; Höhna et al. 2018). Although model adequacy has mainly been used to assess the fit of substitution models in phylogenetics (for an example in paleobiology, see Mulvey et al. 2024), it has also been applied to tree models in the context of epidemiology (Duchêne et al. 2019) and to explore the fit of trait-dependent models of diversification (Schwery et al. 2023) and, in principle, could be extended to birth–death process models that include the fossil sampling process.

The development of new models and tools for morphological data is also an active area of research (Simões et al. 2023a; Tarasov 2023; Khakurel et al. 2024; Mulvey et al. 2024). Although there is evidence to suggest that discrete morphology can evolve in a clocklike manner and can therefore be modeled

using standard clock models (Drummond and Stadler 2016), some interesting issues emerge with the development of new FBD models. For example, how do we model morphological evolution along a tree with stratigraphic ranges? The FBD range model allows us to incorporate information about taxa through time (Stadler et al. 2018). If stratigraphic ranges are associated with a suite of characters that do not change between the FAD and LAD, this could imply an interval of evolutionary stasis, or it could be that morphological change occurs in bursts, coincident with speciation events, that is, according to a model of punctuated equilibrium (Gingerich 1984). We cannot account for these scenarios using standard clock models. Further, we tend not to link parameters across models of trait evolution and diversification (or sampling), although we could, in theory, within the Bayesian framework described here (Fig. 3). More research is needed to develop models that better reflect the nature of sampled morphological data and to assess the impact of different aspects of model violation.

Moving forward, both modeling and empirical perspectives are needed to resolve outstanding questions regarding the integration of fossil data in phylogenetic models. This will involve further theoretical work, simulation studies grounded in empirical observations, better support for empirical data collection and taxonomic research, and more exchange between paleobiologists and model developers about the assumptions the models make about the data. In writing this review, we hope to provide a common ground for discussing Bayesian phylogenetic inference using the FBD process with paleontological data.

Acknowledgments. We would like to thank the editor M. Patzkowsky as well as R. Dearden and two other anonymous reviewers for their helpful comments that greatly improved this article. The contributions of B.J.A. were supported by ETH Zürich.

Competing Interests. The authors declare no conflicts of interest.

Data Availability Statement. The data gathered from the literature survey, as well as the criteria used for its collection, are available on Dryad: <https://doi.org/10.5061/dryad.gf1vhhmxv> and on github: https://github.com/rachelwarnock/FBD_survey.

Literature Cited

- Alfaro, M. E., and M. T. Holder. 2006. The posterior and the prior in Bayesian phylogenetics. *Annual Review of Ecology, Evolution, and Systematics* 37: 19–42.
- Ali, R. H., M. Bark, J. Miró, S. A. Muhammad, J. Sjöstrand, S. M. Zubair, R. M. Abbas, and L. Arvestad. 2017. VMCMC: a graphical and statistical analysis tool for Markov chain Monte Carlo traces. *BMC Bioinformatics* 18:1–8.
- Allen, B. J., M. V. Volkova Oliveira, T. Stadler, T. G. Vaughan, and R. C. Warnock. 2024. Mechanistic phylodynamic models do not provide conclusive evidence that dinosaurs were in decline before their final extinction. *Cambridge Prisms: Extinction* 1.
- Alroy, J. 2008. Dynamics of origination and extinction in the marine fossil record. *Proceedings of the National Academy of Sciences USA* 105: 11536–11542.
- Alroy, J. 2010. Fair sampling of taxonomic richness and unbiased estimation of origination and extinction rates. *Paleontological Society Papers* 16:55–80.
- Alvarez-Carretero, S., A. Goswami, Z. Yang, and M. dos Reis. 2019. Bayesian estimation of species' divergence times using correlated quantitative characters. *Systematic Biology* 68:967–986.
- Andréoletti, J., and H. Morlon. 2023. Exploring congruent diversification histories with flexibility and parsimony. *Methods in Ecology and Evolution* 14:2931–2941.

- Andréoletti, J., A. Zwaans, R. C. Warnock, G. Aguirre-Fernández, J. Barido-Sottani, A. Gupta, T. Stadler, and M. Manceau. 2022. The occurrence birth-death process for combined-evidence analysis in macroevolution and epidemiology. *Systematic Biology* 71:1440–1452.
- Antell, G. T., R. B. Benson, and E. E. Saupe. 2024. Spatial standardization of taxon occurrence data—a call to action. *Paleobiology* 50:177–193.
- Areces-Berzain, F., and J. Ackerman. 2016. Phylogenetics, delimitation and historical biogeography of the pantropical tree genus *Thespesia* (Malvaceae, Gossypieae). *Botanical Journal of the Linnean Society* 181:171–198.
- Arenas, M. 2015. Trends in substitution models of molecular evolution. *Frontiers in Genetics* 6:319.
- Baele, G., W. L. S. Li, A. J. Drummond, M. A. Suchard, and P. Lemey. 2012. Accurate model selection of relaxed molecular clocks in Bayesian phylogenetics. *Molecular Biology and Evolution* 30:239–243.
- Barido-Sottani, J., V. Bošková, L. du Plessis, D. Kühnert, C. Magnus, V. Mitov, N. F. Müller, et al. 2018. Taming the BEAST—a community teaching material resource for BEAST2. *Systematic Biology* 67:170–174.
- Barido-Sottani, J., G. Aguirre-Fernández, M. J. Hopkins, T. Stadler, and R. Warnock. 2019a. Ignoring stratigraphic age uncertainty leads to erroneous estimates of species divergence times under the fossilized birth–death process. *Proceedings of the Royal Society B* 286:20190685.
- Barido-Sottani, J., W. Pett, J. E. O'Reilly, and R. C. Warnock. 2019b. Fossil-Sim: an R package for simulating fossil occurrence data under mechanistic models of preservation and recovery. *Methods in Ecology and Evolution* 10: 835–840.
- Barido-Sottani, J., J. A. Justison, A. M. Wright, R. C. Warnock, W. Pett, and T. A. Heath. 2020a. Estimating a time-calibrated phylogeny of fossil and extant taxa using RevBayes. Pp. 5.2:1–5.2:23 in C. Scornavacca, F. Delsuc, and N. Galtier, eds. *Phylogenetics in the genomic era*. No commercial publisher/open access book. <https://hal.science/hal-02535070v3>.
- Barido-Sottani, J., N. M. Van Tiel, M. J. Hopkins, D. F. Wright, T. Stadler, and R. C. Warnock. 2020b. Ignoring fossil age uncertainty leads to inaccurate topology and divergence time estimates in time calibrated tree inference. *Frontiers in Ecology and Evolution* 8:183.
- Barido-Sottani, J., T. G. Vaughan, and T. Stadler. 2020c. A multitype birth–death model for Bayesian inference of lineage-specific birth and death rates. *Systematic Biology* 69:973–986.
- Barido-Sottani, J., A. Pohle, K. De Baets, D. Murdock, and R. C. Warnock. 2023. Putting the F into FBD analysis: tree constraints or morphological data? *Palaeontology* 66:e12679.
- Barido-Sottani, J., O. Schwery, R. C. Warnock, C. Zhang, and A. M. Wright. 2024. Practical guidelines for Bayesian phylogenetic inference using Markov chain Monte Carlo (MCMC). *Open Research Europe* 3:204.
- Beeravolu, C. R., and F. L. Condamine. 2016. An extended maximum likelihood inference of geographic range evolution by dispersal, local extinction and cladogenesis. bioRxiv 2016–04. <https://doi.org/10.1101/038695>.
- Benson, R. B., N. E. Campione, M. T. Carrano, P. D. Mannion, C. Sullivan, P. Upchurch, and D. C. Evans. 2014. Rates of dinosaur body mass evolution indicate 170 million years of sustained ecological innovation on the avian stem lineage. *PLoS Biology* 12:e1001853.
- Benson, R. B., R. J. Butler, R. A. Close, E. E. Saupe, and D. L. Rabosky. 2021. Biodiversity across space and time in the fossil record. *Current Biology* 31: R1225–R1236.
- Benton, M. J., P. C. Donoghue, and R. J. Asher. 2009. Calibrating and constraining molecular clocks. Pp. 35–86 in S. Blair Hedges and S. Kumar, eds. *The timetree of life*, online ed. Oxford Academic, Oxford. <https://doi.org/10.1093/oso/9780199535033.003.0004>.
- Benton, M. J., A. M. Dunhill, G. T. Lloyd, and F. G. Marx. 2011. Assessing the quality of the fossil record: insights from vertebrates. *Geological Society of London Special Publication* 358:63–94.
- Bergsten, J. 2005. A review of long-branch attraction. *Cladistics* 21:163–193.
- Blanquart, S., and N. Lartillot. 2008. A site-and time-heterogeneous model of amino acid replacement. *Molecular Biology and Evolution* 25:842–858.
- Bouckaert, R., T. G. Vaughan, J. Barido-Sottani, S. Duchêne, M. Fourment, A. Gavryushkina, J. Heled, et al. 2019. BEAST 2.5: an advanced software platform for Bayesian evolutionary analysis. *PLoS Computational Biology* 15:e1006650.
- Brazeau, M. D., T. Guillerme, and M. R. Smith. 2019. An algorithm for morphological phylogenetic analysis with inapplicable data. *Systematic Biology* 68:619–631.
- Bromham, L., S. Duchêne, X. Hua, A. M. Ritchie, D. A. Duchêne, and S. Y. Ho. 2018. Bayesian molecular dating: opening up the black box. *Biological Reviews* 93:1165–1191.
- Brown, J. M. 2014a. Detection of implausible phylogenetic inferences using posterior predictive assessment of model fit. *Systematic Biology* 63:334–348.
- Brown, J. M. 2014b. Predictive approaches to assessing the fit of evolutionary models. *Systematic Biology* 63:289–292.
- Casali, D. M., F. V. Freitas, and F. A. Perini. 2023. Evaluating the impact of anatomical partitioning on summary topologies obtained with Bayesian phylogenetic analyses of morphological data. *Systematic Biology* 72:62–77.
- Castiglione, S., C. Serio, A. Mondanaro, M. Melchionna, and P. Raia. 2022. Fast production of large, time-calibrated, informal supertrees with tree-merger. Technical report. *Wiley Online Library* 65(1), e12588.
- Charpentier, C. P., and A. M. Wright. 2022. Revtulate: an R framework for interaction with RevBayes. *Methods in Ecology and Evolution* 13: 1177–1184.
- Coiro, M., R. Allio, N. Mazet, L. J. Seyfullah, and F. L. Condamine. 2023. Reconciling fossils with phylogenies reveals the origin and macroevolutionary processes explaining the global cycad biodiversity. *New Phytologist* 240: 1616–1635.
- Congreve, C. R., M. E. Patzkowsky, and P. J. Wagner. 2021. An early burst in brachiopod evolution corresponding with significant climatic shifts during the great Ordovician biodiversification event. *Proceedings of the Royal Society B* 288:20211450.
- Cowles, M. K., and B. P. Carlin. 1996. Markov chain Monte Carlo convergence diagnostics: a comparative review. *Journal of the American Statistical Association* 91:883–904.
- Dalén, L., P. D. Heintzman, J. D. Kapp, and B. Shapiro. 2023. Deep-time paleogenomics and the limits of DNA survival. *Science* 382:48–53.
- DeBiasse, M. B., and J. F. Ryan. 2019. Phylotocol: promoting transparency and overcoming bias in phylogenetics. *Systematic Biology* 68:672–678.
- de Faria Santos, A., E. M. Canello, and A. C. Morales. 2022. Phylogeography of *Nasutitermes ephratae* (Termitidae: Nasutitermitinae) in neotropical region. *Scientific Reports* 12:11656.
- Degnan, J. H., and N. A. Rosenberg. 2009. Gene tree discordance, phylogenetic inference and the multispecies coalescent. *Trends in Ecology and Evolution* 24:332–340.
- Deino, A. L., C. Heil Jr., J. King, L. J. McHenry, I. G. Stanistreet, H. Stollhofen, J. K. Njau, J. Mwankunda, K. D. Schick, and N. Toth. 2021. Chronostratigraphy and age modeling of Pleistocene drill cores from the Olduvai basin, Tanzania (Olduvai Gorge coring project). *Palaeogeography, Palaeoclimatology, Palaeoecology* 571:109990.
- Didier, G., and M. Laurin. 2020. Exact distribution of divergence times from fossil ages and tree topologies. *Systematic Biology* 69:1068–1087.
- Didier, G., M. Royer-Carenzi, and M. Laurin. 2012. The reconstructed evolutionary process with the fossil record. *Journal of Theoretical Biology* 315:26–37.
- Dohrmann, M., and G. Wörheide. 2017. Dating early animal evolution using phylogenomic data. *Scientific Reports* 7:3599.
- do Rosario Petrucci, B., M. Januario, and T. Quental. 2022. paleobuddy: an R package for flexible simulations of diversification and fossil sampling. *Methods in Ecology and Evolution* 13:2692–2698.
- Drummond, A. J., and T. Stadler. 2016. Bayesian phylogenetic estimation of fossil ages. *Philosophical Transactions of the Royal Society B* 371:20150129.
- Drummond, A. J., and M. A. Suchard. 2010. Bayesian random local clocks, or one rate to rule them all. *BMC Biology* 8:114.
- Drummond, A. J., S. Y. W. Ho, M. J. Phillips, and A. Rambaut. 2006. Relaxed phylogenetics and dating with confidence. *PLoS Biology* 4:e88.
- Drummond, A. J., K. Chen, F. K. Mendes, and D. Xie. 2023. LinguaPhylo: a probabilistic model specification language for reproducible phylogenetic analyses. *PLoS Computational Biology* 19:e1011226.
- Duchêne, D. A., S. Duchêne, and S. Y. Ho. 2018. Differences in performance among test statistics for assessing phylogenomic model adequacy. *Genome Biology and Evolution* 10:1375–1388.
- Duchêne, S., R. Bouckaert, D. A. Duchêne, T. Stadler, and A. J. Drummond. 2019. Phylodynamic model adequacy using posterior predictive simulations. *Systematic Biology* 68:358–364.
- Dunhill, A. M., M. J. Benton, R. J. Twitchett, and A. J. Newell. 2012. Completeness of the fossil record and the validity of sampling proxies at outcrop level. *Palaeontology* 55:1155–1175.

- Dunne, E. M., A. Farnsworth, R. B. Benson, P. L. Godoy, S. E. Greene, P. J. Valdes, D. J. Lunt, and R. J. Butler. 2023. Climatic controls on the ecological ascendancy of dinosaurs. *Current Biology* 33:206–214.
- Fabreti, L. G., and S. Höhna. 2022. Convergence assessment for Bayesian phylogenetic analysis using MCMC simulation. *Methods in Ecology and Evolution* 13:77–90.
- Farina, B. M., P. L. Godoy, R. B. Benson, M. C. Langer, and G. S. Ferreira. 2023. Turtle body size evolution is determined by lineage-specific specializations rather than global trends. *Ecology and Evolution* 13:e10201.
- Fleming, J. F., R. M. Kristensen, M. V. Sørensen, T.-Y. S. Park, K. Arakawa, M. Blaxter, L. Rebecchi, et al. 2018. Molecular palaeontology illuminates the evolution of ecdysozoan vision. *Proceedings of the Royal Society B* 285:20182180.
- Foote, M. 1996. On the probability of ancestors in the fossil record. *Paleobiology* 22:141–151.
- Foote, M. 2000. Origination and extinction components of taxonomic diversity: general problems. *Paleobiology* 26:74–102.
- Foote, M., and J. J. Sepkoski. 1999. Absolute measures of the completeness of the fossil record. *Nature* 398:415–417.
- Fourment, M., A. F. Magee, C. Whidden, A. Bilge, F. A. Matsen IV, and V. N. Minin. 2020. 19 dubious ways to compute the marginal likelihood of a phylogenetic tree topology. *Systematic Biology* 69:209–220.
- Gavryushkina, A., D. Welch, T. Stadler, and A. J. Drummond. 2014. Bayesian inference of sampled ancestor trees for epidemiology and fossil calibration. *PLoS Computational Biology* 10:e1003919.
- Gavryushkina, A., T. A. Heath, D. T. Ksepka, T. Stadler, D. Welch, and A. J. Drummond. 2017. Bayesian total-evidence dating reveals the recent crown radiation of penguins. *Systematic Biology* 66:57–73.
- Gelman, A. 1996. Inference and monitoring convergence. Pp. 131–144 in W. R. Gilks, S. Richardson, and David Spiegelhalter, eds. *Markov chain Monte Carlo in practice*. Chapman & Hall, Boca Raton, Fla.
- Gelman, A. 2012. Ethics and statistics: ethics and the statistical use of prior information. *Chance* 25:52–54.
- Gernhard, T. 2008. The conditioned reconstructed process. *Journal of Theoretical Biology* 253:769–778.
- Gingerich, P. D. 1984. Punctuated equilibria—where is the evidence? *Systematic Zoology* 33:335–338.
- Godoy, P. L., R. B. Benson, M. Bronzati, and R. J. Butler. 2019. The multi-peak adaptive landscape of crocodylomorph body size evolution. *BMC Evolutionary Biology* 19:1–29.
- Gradstein, F. M., J. G. Ogg, M. D. Schmitz, and G. M. Ogg. 2020. *Geologic time scale 2020*. Elsevier, Amsterdam.
- Gupta, A., M. Manceau, T. G. Vaughan, M. Khammash, and T. Stadler. 2020. The probability distribution of the reconstructed phylogenetic tree with occurrence data. *Journal of Theoretical Biology* 488:110115.
- Halverson, G. P., C. Shen, J. H. Davies, and L. Wu. 2022. A Bayesian approach to inferring depositional ages applied to a Late Tonian reference section in Svalbard. *Frontiers in Earth Science* 10:798739.
- Hartmann, K., D. Wong, and T. Stadler. 2010. Sampling trees from evolutionary models. *Systematic Biology* 59:465–476.
- Hasegawa, M., H. Kishino, and T. Yano. 1989. Estimation of branching dates among primates by molecular clocks of nuclear DNA which slowed down in Hominoidea. *Journal of Human Evolution* 18:461–476.
- Hastings, W. K. 1970. Monte Carlo sampling methods using Markov chains and their applications. *Biometrika* 57:97–109.
- Heath, T. A., and B. R. Moore. 2014. Bayesian inference of species divergence times. Pp. 277–318 in M.-H. Chen, L. Kuo, and P. O. Lewis, eds. *Bayesian phylogenetics: methods, algorithms, and applications*. CRC Press, Boca Raton, Fla.
- Heath, T. A., M. T. Holder, and J. P. Huelsenbeck. 2012. A Dirichlet process prior for estimating lineage-specific substitution rates. *Molecular Biology and Evolution* 29:939–955.
- Heath, T. A., J. P. Huelsenbeck, and T. Stadler. 2014. The fossilized birth-death process for coherent calibration of divergence-time estimates. *Proceedings of the National Academy of Sciences USA* 111:E2957–E2966.
- Herrera, J. P. 2017. Testing the adaptive radiation hypothesis for the lemurs of Madagascar. *Royal Society Open Science* 4:161014.
- Höhna, S., T. Stadler, F. Ronquist, and T. Britton. 2011. Inferring speciation and extinction rates under different sampling schemes. *Molecular Biology and Evolution* 28:2577–2589.
- Höhna, S., M. J. Landis, T. A. Heath, B. Boussau, N. Lartillot, B. R. Moore, J. P. Huelsenbeck, and F. Ronquist. 2016. RevBayes: Bayesian phylogenetic inference using graphical models and an interactive model-specification language. *Systematic Biology* 65:726–736.
- Höhna, S., L. M. Coghill, G. G. Mount, R. C. Thomson, and J. M. Brown. 2018. P3: phylogenetic posterior prediction in RevBayes. *Molecular Biology and Evolution* 35:1028–1034.
- Holder, M., and P. O. Lewis. 2003. Phylogeny estimation: traditional and Bayesian approaches. *Nature Reviews Genetics* 4:275–284.
- Holland, S. M. 2016. The non-uniformity of fossil preservation. *Philosophical Transactions of the Royal Society B* 371:20150130.
- Hopkins, M. J., and K. St. John. 2021. Incorporating hierarchical characters into phylogenetic analysis. *Systematic Biology* 70:1163–1180.
- Hopkins, M. J., D. W. Bapst, C. Simpson, and R. C. Warnock. 2018. The inseparability of sampling and time and its influence on attempts to unify the molecular and fossil records. *Paleobiology* 44:561–574.
- Huelsenbeck, J. P., B. Larget, and D. L. Swofford. 2000. A compound Poisson process for relaxing the molecular clock. *Genetics* 154:1879–1892.
- Huelsenbeck, J. P., F. Ronquist, R. Nielsen, and J. P. Bollback. 2001. Bayesian inference of phylogeny and its impact on evolutionary biology. *Science* 294:2310–2314.
- Huelsenbeck, J. P., B. Larget, R. E. Miller, and F. Ronquist. 2002. Potential applications and pitfalls of Bayesian inference of phylogeny. *Systematic Biology* 51:673–688.
- Huelsenbeck, J. P., P. Joyce, C. Lakner, and F. Ronquist. 2008. Bayesian analysis of amino acid substitution models. *Philosophical Transactions of the Royal Society B* 363:3941–3953.
- Huson, D. H., and D. Bryant. 2006. Application of phylogenetic networks in evolutionary studies. *Molecular Biology and Evolution* 23:254–267.
- Jeffreys, H. 1961. *Theory of probability*. Oxford University Press, Oxford.
- Jenkins, G. B., A. P. Beckerman, C. Bellard, A. Benítez-López, A. M. Ellison, C. G. Foote, A. L. Hufton, et al. 2023. Reproducibility in ecology and evolution: minimum standards for data and code. *Ecology and Evolution* 13:e9961.
- Jiang, B., Y. He, A. Elsler, S. Wang, J. N. Keating, J. Song, S. L. Kearns, and M. J. Benton. 2023. Extended embryo retention and viviparity in the first amniotes. *Nature Ecology and Evolution* 7:1131–1140.
- Jones, L. A., W. Gearty, B. J. Allen, K. Eichenseer, C. D. Dean, S. Galván, M. Kouvari, et al. 2023. palaeoverse: a community-driven R package to support palaeobiological analysis. *Methods in Ecology and Evolution* 14:2205–2215.
- Jukes, T. H., C. R. Cantor. 1969. Evolution of protein molecules. *Mammalian Protein Metabolism* 3:21–132.
- Kass, R. E., and A. E. Raftery. 1995. Bayes factors. *Journal of the American Statistical Association* 90:773–795.
- Kendall, D. G. 1948. On the generalized “birth-and-death” process. *Annals of Mathematical Statistics* 19:1–15.
- Khakurel, B., C. Grigsby, T. D. Tran, J. Zariwala, S. Höhna, and A. M. Wright. 2024. The fundamental role of character coding in Bayesian morphological phylogenetics. *Systematic Biology* 73:861–871.
- Kjær, K. H., M. Winther Pedersen, B. De Sanctis, B. De Cahsan, T. S. Korneliussen, C. S. Michelsen, K. K. Sand, et al. 2022. A 2-million-year-old ecosystem in Greenland uncovered by environmental DNA. *Nature* 612:283–291.
- Klopfstein, S., L. Vilhelmsen, and F. Ronquist. 2015. A nonstationary Markov model detects directional evolution in hymenopteran morphology. *Systematic Biology* 64:1089–1103.
- Kocsis, A. T., C. J. Reddin, J. Alroy, and W. Kiessling. 2019. The R package divDyn for quantifying diversity dynamics using fossil sampling data. *Methods in Ecology and Evolution* 10:735–743.
- Kopperud, B. T., A. F. Magee, and S. Höhna. 2023. Rapidly changing speciation and extinction rates can be inferred in spite of nonidentifiability. *Proceedings of the National Academy of Sciences USA* 120:e2208851120.
- Kühnert, D., T. Stadler, T. G. Vaughan, and A. J. Drummond. 2016. Phylogenetics with migration: a computational framework to quantify population structure from genomic data. *Molecular Biology and Evolution* 33:2102–2116.
- Lakner, C., P. Van Der Mark, J. P. Huelsenbeck, B. Larget, and F. Ronquist. 2008. Efficiency of Markov chain Monte Carlo tree proposals in Bayesian phylogenetics. *Systematic Biology* 57:86–103.

- Landis, M. J., D. A. Eaton, W. L. Clement, B. Park, E. L. Spriggs, P. W. Sweeney, E. J. Edwards, and M. J. Donoghue. 2021. Joint phylogenetic estimation of geographic movements and biome shifts during the global diversification of *Viburnum*. *Systematic Biology* 70:67–85.
- Lanfear, R., X. Hua, and D. L. Warren. 2016. Estimating the effective sample size of tree topologies from Bayesian phylogenetic analyses. *Genome Biology and Evolution* 8:2319–2332.
- Lartillot, N., and H. Philippe. 2006. Computing Bayes factors using thermodynamic integration. *Systematic Biology* 55:195–207.
- Lee, M. S. 2016. Multiple morphological clocks and total-evidence tip-dating in mammals. *Biology Letters* 12:20160033.
- Lee, M. S., and A. Palci. 2015. Morphological phylogenetics in the genomic age. *Current Biology* 25:R922–R929.
- Lee, T., D. Rand, L. E. Lisiecki, G. Gebbie, and C. Lawrence. 2023. Bayesian age models and stacks: combining age inferences from radiocarbon and benthic $\delta^{18}\text{O}$ stratigraphic alignment. *Climate of the Past* 19:1993–2012.
- Lemmon, A. R., and E. C. Moriarty. 2004. The importance of proper model assumption in Bayesian phylogenetics. *Oxford lists Systematic Biology* 53: 278–298.
- Lepage, T., D. Bryant, H. Philippe, and N. Lartillot. 2007. A general comparison of relaxed molecular clock models. *Molecular Biology and Evolution* 24: 2669–2680.
- Lewis, P. O. 2001. A likelihood approach to estimating phylogeny from discrete morphological character data. *Systematic Biology* 50:913–925.
- Liò, P., and N. Goldman. 1998. Models of molecular evolution and phylogeny. *Genome Research* 8:1233–1244.
- Liow, L. H., and J. D. Nichols. 2010. Estimating rates and probabilities of origination and extinction using taxonomic occurrence data: capture-mark-recapture (CMR) approaches. *Paleontological Society Papers* 16:81–94.
- Lloyd, G., D. Bapst, M. Friedman, and K. Davis. 2016. Probabilistic divergence time estimation without branch lengths: dating the origins of dinosaurs, avian flight and crown birds. *Biology Letters* 12:20160609.
- Lloyd, G. T., and G. J. Slater. 2021. A total-group phylogenetic metatree for Cetacea and the importance of fossil data in diversification analyses. *Systematic Biology* 70:922–939.
- Lloyd, G. T., K. E. Davis, D. Pisani, J. E. Tarver, M. Ruta, M. Sakamoto, D. W. Hone, R. Jennings, and M. J. Benton. 2008. Dinosaurs and the Cretaceous terrestrial revolution. *Proceedings of the Royal Society B* 275:2483–2490.
- López-Antónanzas, R., J. Mitchell, T. R. Simões, F. L. Condamine, R. Aguilée, P. Peláez Campomanes, S. Renaud, J. Rolland, and P. C. Donoghue. 2022. Integrative phylogenetics: tools for palaeontologists to explore the tree of life. *Biology* 11:1185.
- Louca, S., and M. W. Pennell. 2020. Extant timetrees are consistent with a myriad of diversification histories. *Nature* 580:502–505.
- Louca, S., A. McLaughlin, A. MacPherson, J. B. Joy, and M. W. Pennell. 2021. Fundamental identifiability limits in molecular epidemiology. *Molecular Biology and Evolution* 38:4010–4024.
- Luo, A., D. A. Duchêne, C. Zhang, C.-D. Zhu, and S. Y. Ho. 2020. A simulation-based evaluation of tip-dating under the fossilized birth-death process. *Systematic Biology* 69:325–344.
- Luo, A., C. Zhang, Q.-S. Zhou, S. Y. Ho, and C.-D. Zhu. 2023. Impacts of taxon-sampling schemes on Bayesian tip dating under the fossilized birth-death process. *Systematic Biology* 72:781–801.
- Manceau, M., A. Gupta, T. Vaughan, and T. Stadler. 2021. The probability distribution of ancestral population size under birth-death processes. *J. Theor. Biol.* 509:110400.
- Maddison, W. P., and D. Maddison. 2023. Mesquite: a modular system for evolutionary analysis. <http://mesquiteproject.org>.
- Magee, A. F., and S. Höhna. 2021. Impact of K-Pg mass extinction event on Crocodylomorpha inferred from phylogeny of extinct and extant taxa. *bioRxiv* 2021–01. <https://doi.org/10.1101/2021.01.14.426715>.
- Marshall, C. R. 2017. Five palaeobiological laws needed to understand the evolution of the living biota. *Nature Ecology and Evolution* 1:0165.
- Matschiner, M. 2019. Selective sampling of species and fossils influences age estimates under the fossilized birth-death model. *Frontiers in Genetics* 10:1064.
- Matzke, N. J. 2013. BioGeoBEARS: biogeography with Bayesian (and likelihood) evolutionary analysis in R scripts, R package version 0.2.1. <https://rdrr.io/cran/BioGeoBEARS>.
- May, M. R., and C. J. Rothfels. 2023. Diversification models conflate likelihood and prior, and cannot be compared using conventional model-comparison tools. *Systematic Biology* 72:713–722.
- May, M. R., D. L. Contreras, M. A. Sundue, N. S. Nagalingum, C. V. Looy, and C. J. Rothfels. 2021. Inferring the total-evidence timescale of marattialean fern evolution in the face of model sensitivity. *Systematic Biology* 70:1232–1255.
- Membrebe, J. V., M. A. Suchard, A. Rambaut, G. Baele, and P. Lemey. 2019. Bayesian inference of evolutionary histories under time-dependent substitution rates. *Molecular Biology and Evolution* 36:1793–1803.
- Metropolis, N., A. W. Rosenbluth, M. N. Rosenbluth, A. H. Teller, and E. Teller. 1953. Equation of state calculations by fast computing machines. *Journal of Chemical Physics* 21:1087–1092.
- Möller, S., L. du Plessis, and T. Stadler. 2018. Impact of the tree prior on estimating clock rates during epidemic outbreaks. *Proceedings of the National Academy of Sciences USA* 115:4200–4205.
- Mongiardino Koch, N., R. J. Garwood, and L. A. Parry. 2021. Fossils improve phylogenetic analyses of morphological characters. *Proceedings of the Royal Society B* 288:20210044.
- Mulvey, L. P., M. R. May, J. M. Brown, S. Höhna, A. M. Wright, and R. C. Warnock. 2024. Assessing the Adequacy of Morphological Models using posterior predictive simulations. *Systematic Biology* 74:34–52. <https://doi.org/10.1101/2024.01.25.577179>.
- Nanglu, K., and T. M. Cullen. 2023. Across space and time: a review of sampling, preservational, analytical, and anthropogenic biases in fossil data across macroecological scales. *Earth Science Reviews* 244:104537.
- Nascimento, F. F., M. d. Reis, and Z. Yang. 2017. A biologist's guide to Bayesian phylogenetic analysis. *Nature Ecology and Evolution* 1:1446–1454.
- Nee, S., R. M. May, and P. H. Harvey. 1994. The reconstructed evolutionary process. *Philosophical Transactions of the Royal Society B* 344:305–311.
- Nylander, J. A., J. C. Wilgenbusch, D. L. Warren, and D. L. Swofford. 2008. AWTY (are we there yet?): a system for graphical exploration of MCMC convergence in Bayesian phylogenetics. *Bioinformatics* 24:581–583.
- Oaks, J. R., K. A. Cobb, V. N. Minin, and A. D. Leaché. 2019. Marginal likelihoods in phylogenetics: a review of methods and applications. *Systematic Biology* 68:681–697.
- Ogilvie, H. A., F. K. Mendes, T. G. Vaughan, N. J. Matzke, T. Stadler, D. Welch, and A. R. Drummond. 2022. Novel integrative modeling of molecules and morphology across evolutionary timescales. *Systematic Biology* 71:208–220.
- O'Leary, M. A., and S. Kaufman. 2011. MorphoBank: phylophenomics in the “cloud.” *Cladistics* 27:529–537.
- O'Reilly, J. E., and P. C. Donoghue. 2018. The efficacy of consensus tree methods for summarizing phylogenetic relationships from a posterior sample of trees estimated from morphological data. *Systematic Biology* 67: 354–362.
- O'Reilly, J. E., and P. C. Donoghue. 2020. The effect of fossil sampling on the estimation of divergence times with the fossilized birth–death process. *Systematic Biology* 69:124–138.
- O'Reilly, J. E., and P. C. Donoghue. 2021. Fossilization processes have little impact on tip-calibrated divergence time analyses. *Palaeontology* 64: 687–697.
- Parham, J. F., P. C. Donoghue, C. J. Bell, T. D. Calway, J. J. Head, P. A. Holroyd, J. G. Inoue, et al. 2012. Best practices for justifying fossil calibrations. *Systematic Biology* 61:346–359.
- Parins-Fukuchi, C. T. 2017. Use of continuous traits can improve morphological phylogenetics. *Systematic Biology* 67:328–339.
- Pattinson, D. J., R. S. Thompson, A. K. Piotrowski, and R. J. Asher. 2014. Phylogeny, paleontology, and primates: do incomplete fossils bias the tree of life? *Systematic Biology* 64:169–186.
- Patzkowski, M. E. 1995. A hierarchical branching model of evolutionary radiations. *Paleobiology* 21:440–460.
- Pett, W., and T. A. Heath. 2020. Inferring the timescale of phylogenetic trees from fossil data. Pp. 5.1:1–5.1:18 in C. Scornavacca, F. Delsuc, and N. Galtier, eds. *Phylogenetics in the genomic era*. No commercial publisher/open access book. <https://hal.science/hal-02535070v3>.
- Raja, N. B., E. M. Dunne, A. Mtiwane, T. M. Khan, P. S. Natscher, A. M. Ghilardi, and D. Chattopadhyay. 2022. Colonial history and global economics distort our understanding of deep-time biodiversity. *Nature Ecology and Evolution* 6:145–154.

- Rambaut, A., A. J. Drummond, D. Xie, G. Baele, and M. A. Suchard. 2018. Posterior summarization in Bayesian phylogenetics using Tracer 1.7. *Systematic Biology* 67:901–904.
- Rannala, B., and Z. Yang. 2007. Inferring speciation times under an episodic molecular clock. *Systematic Biology* 56:453–466.
- Raup, D. M. 1985. Mathematical models of cladogenesis. *Paleobiology* 11: 42–52.
- Ree, R. H., and S. A. Smith. 2008. Maximum likelihood inference of geographic range evolution by dispersal, local extinction, and cladogenesis. *Systematic Biology* 57:4–14.
- Ritchie, A. M., N. Lo, and S. Y. Ho. 2017. The impact of the tree prior on molecular dating of data sets containing a mixture of inter- and intraspecies sampling. *Systematic Biology* 66:413–425.
- Ronquist, F., P. van der Mark, and J. P. Huelsenbeck. 2009. Bayesian phylogenetic analysis using MrBayes. *The phylogenetic handbook*, 2(210–266), 723. Cambridge University Press Cambridge.
- Ronquist, F., S. Klopfstein, L. Vilhelmsen, S. Schulmeister, D. L. Murray, and A. P. Rasnitsyn. 2012a. A total-evidence approach to dating with fossils, applied to the early radiation of the Hymenoptera. *Systematic Biology* 61: 973–999.
- Ronquist, F., M. Teslenko, P. Van Der Mark, D. L. Ayres, A. Darling, S. Höhna, B. Larget, L. Liu, M. A. Suchard, and J. P. Huelsenbeck. 2012b. MrBayes 3.2: efficient Bayesian phylogenetic inference and model choice across a large model space. *Systematic Biology* 61:539–542.
- Roy, V. 2020. Convergence diagnostics for Markov chain Monte Carlo. *Annual Review of Statistics and Its Application* 7:387–412.
- Sansom, R. S., S. E. Gabbott, and M. A. Purnell. 2010. Non-random decay of chordate characters causes bias in fossil interpretation. *Nature* 463: 797–800.
- Sarver, B. A., M. W. Pennell, J. W. Brown, S. Keeble, K. M. Hardwick, J. Sullivan, and L. J. Harmon. 2019. The choice of tree prior and molecular clock does not substantially affect phylogenetic inferences of diversification rates. *PeerJ* 7:e6334.
- Sayers, E. W., E. E. Bolton, J. R. Brister, K. Canese, J. Chan, D. C. Comeau, R. Connor, et al. 2022. Database resources of the national center for biotechnology information. *Nucleic Acids Research* 50:D20.
- Schroeter, E. R., T. P. Cleland, and M. H. Schweitzer. 2021. Deep time paleoproteomics: looking forward. *Journal of Proteome Research* 21:9–19.
- Schweitzer, M. H., W. Zheng, C. L. Organ, R. Avci, Z. Suo, L. M. Freimark, V. S. Lebleu, et al. 2009. Biomolecular characterization and protein sequences of the Campanian hadrosaur *B. canadensis*. *Science* 324:626–631.
- Schwery, O., W. Freyman, and E. E. Goldberg. 2023. adequaSSE: model adequacy testing for trait-dependent diversification models. bioRxiv 2023–03. <https://doi.org/10.1101/2023.03.06.531416>.
- Sepkoski, J. J., Jr. 2002. A compendium of fossil marine animal genera. *Bulletins of American Paleontology* 363:1–560.
- Sereno, P. C. 2007. Logical basis for morphological characters in phylogenetics. *Cladistics* 23:565–587.
- Silvestro, D., J. Schnitzler, L. H. Liow, A. Antonelli, and N. Salamin. 2014. Bayesian estimation of speciation and extinction from incomplete fossil occurrence data. *Systematic Biology* 63:349–367.
- Silvestro, D., N. Salamin, A. Antonelli, and X. Meyer. 2019. Improved estimation of macroevolutionary rates from fossil data using a Bayesian framework. *Paleobiology* 45:546–570.
- Simões, T. R., and S. E. Pierce. 2021. Sustained high rates of morphological evolution during the rise of tetrapods. *Nature Ecology and Evolution* 5: 1403–1414.
- Simões, T. R., M. W. Caldwell, and S. E. Pierce. 2020. Sphenodontian phylogeny and the impact of model choice in Bayesian morphological clock estimates of divergence times and evolutionary rates. *BMC Biology* 18:1–30.
- Simões, T. R., N. Greifer, J. Barido-Sottani, and S. E. Pierce. 2023a. Evophylo: an R package for pre- and postprocessing of morphological data from relaxed clock Bayesian phylogenetics. *Methods in Ecology and Evolution* 14:1981–1993.
- Simões, T. R., O. V. Vernygora, B. A. de Medeiros, and A. M. Wright. 2023b. Handling logical character dependency in phylogenetic inference: extensive performance testing of assumptions and solutions using simulated and empirical data. *Systematic Biology* 72:662–680.
- Slater, G. J., J. A. Goldbogen, and N. D. Pyenson. 2017. Independent evolution of baleen whale gigantism linked to Plio-Pleistocene ocean dynamics. *Proceedings of the Royal Society B* 284:20170546.
- Smith, J. A., N. B. Raja, T. Clements, D. Dimitrijević, E. M. Dowding, E. M. Dunne, B. M. Gee, et al. 2023. Increasing the equitability of data citation in paleontology: capacity building for the big data future. *Paleobiology* 50:165–176.
- Smith, M. R. 2022. Robust analysis of phylogenetic tree space. *Systematic Biology* 71:1255–1270.
- Song, Y.-G., Y. Fragnière, H.-H. Meng, Y. Li, S. Bétrisey, A. Corrales, S. Manchester, et al. 2020. Global biogeographic synthesis and priority conservation regions of the relict tree family Juglandaceae. *Journal of Biogeography* 47:643–657.
- Soul, L. C., and M. Friedman. 2015. Taxonomy and phylogeny can yield comparable results in comparative paleontological analyses. *Systematic Biology* 64:608–620.
- Soul, L. C., and D. F. Wright. 2021. *Phylogenetic comparative methods: a user's guide for paleontologists*. Cambridge University Press, Cambridge.
- Stadler, T. 2009. On incomplete sampling under birth-death models and connections to the sampling-based coalescent. *Journal of Theoretical Biology* 261:58–66.
- Stadler, T. 2010. Sampling-through-time in birth-death trees. *Journal of Theoretical Biology* 267:396–404.
- Stadler, T., D. Kühnert, S. Bonhoeffer, and A. J. Drummond. 2012. Birth-death skyline plot reveals temporal changes of epidemic spread in HIV and hepatitis C virus (HCV). *Proceedings of the National Academy of Sciences USA* 110:228–3233.
- Stadler, T., A. Gavryushkina, R. C. Warnock, A. J. Drummond, and T. A. Heath. 2018. The fossilized birth-death model for the analysis of stratigraphic range data under different speciation modes. *Journal of Theoretical Biology* 447:41–55.
- Stadler, T., C. Magnus, T. Vaughan, J. Barido-Sottani, V. Bošková, J. Huisman, and J. Pečerska. 2024. *Decoding genomes: from sequences to phylogenomics*. Self-published. [10.3929/ethz-b-000664449](https://doi.org/10.3929/ethz-b-000664449).
- Sterli, J., M. S. De La Fuente, and G. W. Rougier. 2018. New remains of *Condorchelys antiqua* (Testudinata) from the Early-Middle Jurassic of Patagonia: anatomy, phylogeny, and pedomorphosis in the early evolution of turtles. *Journal of Vertebrate Paleontology* 38:1–17.
- Stockdale, M. T., and M. J. Benton. 2021. Environmental drivers of body size evolution in crocodile line archosaurs. *Communications Biology* 4:38.
- Suchard, M. A., P. Lemey, G. Baele, D. L. Ayres, A. J. Drummond, and A. Rambaut. 2018. Bayesian phylogenetic and phylodynamic data integration using BEAST 1.10. *Virus Evolution* 4:vey016.
- Sullivan, J., and P. Joyce. 2005. Model selection in phylogenetics. *Annual Review of Ecology, Evolution, and Systematics* 36:445–466.
- Tarasov, S. 2023. New phylogenetic Markov models for inapplicable morphological characters. *Systematic Biology* 72:681–693.
- Tavares, V. d. C., O. M. Warsi, F. Balseiro, C. A. Mancina, and L. M. Dávalos. 2018. Out of the Antilles: fossil phylogenies support reverse colonization of bats to South America. *Journal of Biogeography* 45:859–873.
- Thomas, D. B., A. J. Tennyson, R. P. Scofield, T. A. Heath, W. Pett, and D. T. Ksepka. 2020. Ancient crested penguin constrains timing of recruitment into seabird hotspot. *Proceedings of the Royal Society B* 287:20201497.
- Thomas, G. H., K. Hartmann, W. Jetz, J. B. Joy, A. Mimoto, and A. O. Mooers. 2013. PASTIS: an R package to facilitate phylogenetic assembly with soft taxonomic inferences. *Methods in Ecology and Evolution* 4: 1011–1017.
- Thompson, E. A. 1975. *Human evolutionary trees*. Cambridge University Press, Cambridge.
- Thorne, J., and H. Kishino. 2002. Divergence time and evolutionary rate estimation with multilocus data. *Systematic Biology* 51:689–702.
- Thorne, J. L., H. Kishino, and I. S. Painter. 1998. Estimating the rate of evolution of the rate of molecular evolution. *Molecular Biology and Evolution* 15:1647–1657.
- Thuy, B., M. E. Eriksson, M. Kutscher, J. Lindgren, L. D. Nummerger-Thuy, and D. F. Wright. 2022. Miniaturization during a Silurian environmental crisis generated the modern brittle star body plan. *Communications Biology* 5:14.

- Tougaard, C. 2017. Did the Quaternary climatic fluctuations really influence the tempo and mode of diversification in European rodents? *Journal of Zoological Systematics and Evolutionary Research* 55:46–56.
- Tribble, C. M., W. A. Freyman, M. J. Landis, J. Ying Lim, J. Barido-Sottani, B. T. Kopperud, S. Höhna, and M. R. May. 2022. RevGadgets: an R package for visualising Bayesian phylogenetic analyses from RevBayes. *Methods in Ecology and Evolution* 13:314–323.
- Trivedi, R., and H. A. Nagarajaram. 2020. Substitution scoring matrices for proteins—an overview. *Protein Science* 29:2150–2163.
- Truman, K., T. G. Vaughan, A. Gavryushkin, and A. Gavryushkina. 2024. The fossilised birth death model is identifiable. *Systematic Biology* 74:112–123. <https://doi.org/10.1101/2024.02.08.579547>.
- Uhen, M. D., B. Allen, N. Behboudi, M. E. Clapham, E. Dunne, A. Hendy, P. A. Holroyd, M. Hopkins, P. Mannion, P. Novack-Gottshall, C. Pimiento, and P. Wagner. 2023. Paleobiology Database User Guide Version 1.0. *PaleBios* 40:1–56.
- Vaughan, T. G. 2017. IcyTree: rapid browser-based visualization for phylogenetic trees and networks. *Bioinformatics* 33:2392–2394.
- Wagner, P. J. 2012. Modelling rate distributions using character compatibility: implications for morphological evolution among fossil invertebrates. *Biology Letters* 8:143–146.
- Wagner, P. J., and J. D. Marcot. 2010. Probabilistic phylogenetic inference in the fossil record: current and future applications. *Paleontological Society Papers* 16:189–211.
- Wang, Y., and Z. Yang. 2014. Priors in Bayesian phylogenetics. Pp. 5–24 in M.-H. Chen, L. Kuo, and P. O. Lewis, eds. *Bayesian phylogenetics: methods, algorithms, and applications*. CRC Press, Boca Raton, Fla.
- Warnock, R. C., and A. M. Wright. 2020. *Understanding the tripartite approach to Bayesian divergence time estimation*. Cambridge University Press, Cambridge.
- Warnock, R. C., T. A. Heath, and T. Stadler. 2020. Assessing the impact of incomplete species sampling on estimates of speciation and extinction rates. *Paleobiology* 46:137–157.
- Warren, D. L., A. J. Geneva, and R. Lanfear. 2017. RWTY (R we there yet): an R package for examining convergence of Bayesian phylogenetic analyses. *Molecular Biology and Evolution* 34:1016–1020.
- Williams, T. A., C. J. Cox, P. G. Foster, G. J. Szöllösi, and T. M. Embley. 2020. Phylogenomics provides robust support for a two-domains tree of life. *Nature Ecology and Evolution* 4:138–147.
- Wolfe, J. M., L. Ballou, J. Luque, V. M. Watson-Zink, S. T. Ah Yong, J. Barido-Sottani, T.-Y. Chan, et al. 2023. Convergent adaptation of true crabs (Decapoda: Brachyura) to a gradient of terrestrial environments. *Systematic Biology* 73:247–262.
- Wright, A., P. J. Wagner, and D. F. Wright. 2021. *Testing character evolution models in phylogenetic paleobiology: a case study with Cambrian echinoderms*. Cambridge University Press, Cambridge.
- Wright, A. M. 2019. A systematist's guide to estimating Bayesian phylogenies from morphological data. *Insect Systematics and Diversity* 3:2.
- Wright, A. M., and G. T. Lloyd. 2020. Bayesian analyses in phylogenetic palaeontology: interpreting the posterior sample. *Palaeontology* 63: 997–1006.
- Wright, D. F. 2017. Bayesian estimation of fossil phylogenies and the evolution of early to middle Paleozoic crinoids (Echinodermata). *Journal of Paleontology* 91:799–814.
- Xiang, X., K. Xiang, R. D. C. Ortiz, F. Jabbour, and W. Wang. 2019. Integrating palaeontological and molecular data uncovers multiple ancient and recent dispersals in the pantropical Hamamelidaceae. *Journal of Biogeography* 46:2622–2631.
- Xie, W., P. O. Lewis, Y. Fan, L. Kuo, and M.-H. Chen. 2011. Improving marginal likelihood estimation for Bayesian phylogenetic model selection. *Systematic Biology* 60:150–160.
- Yang, Z. 1994. Maximum likelihood phylogenetic estimation from DNA sequences with variable rates over sites: approximate methods. *Journal of Molecular Evolution* 39:306–314.
- Yang, Z. 2014. *Molecular evolution: a statistical approach*. Oxford University Press, Oxford.
- Yang, Z., and B. Rannala. 1997. Bayesian phylogenetic inference using DNA sequences: a Markov chain Monte Carlo method. *Molecular Biology and Evolution* 14:717–724.
- Yang, Z., and B. Rannala. 2012. Molecular phylogenetics: principles and practice. *Nature Reviews Genetics* 13:303–314.
- Yang, Z., and A. D. Yoder. 2003. Comparison of likelihood and Bayesian methods for estimating divergence times using multiple gene loci and calibration points, with application to a radiation of cute-looking mouse lemur species. *Systematic Biology* 52:705–716.
- Yang, Z., and T. Zhu. 2018. Bayesian selection of misspecified models is overconfident and may cause spurious posterior probabilities for phylogenetic trees. *Proceedings of the National Academy of Sciences USA* 115: 1854–1859.
- Yule, G. U. 1924. A mathematical theory of evolution, based on the conclusions of Dr. J. C. Wills, F. R. S. *Philosophical Transactions of the Royal Society B* 213:21–87.
- Zhang, C. 2016. Molecular clock dating using MrBayes. arXiv 1603.05707.
- Zhang, C., T. Stadler, S. Klopstein, T. A. Heath, and F. Ronquist. 2016. Total-evidence dating under the fossilized birth–death process. *Systematic Biology* 65:228–249.
- Zhang, C., F. Ronquist, and T. Stadler. 2023. Skyline fossilized birth–death model is robust to violations of sampling assumptions in total-evidence dating. *Systematic Biology* 72:1316–1336.
- Zhang, Q., R. H. Ree, N. Salamin, Y. Xing, and D. Silvestro. 2022. Fossil-informed models reveal a boreotropical origin and divergent evolutionary trajectories in the walnut family (Juglandaceae). *Systematic Biology* 71:242–258.
- Zhang, R., A. J. Drummond, and F. K. Mendes. 2024. Fast Bayesian inference of phylogenies from multiple continuous characters. *Systematic Biology* 73: 102–124.
- Zuckerkandl, E., and L. Pauling. 1962. Molecular disease, evolution, and genetic heterogeneity. Pp. 189–225 in M. Kasha and B. Pullman, eds. *Horizons in biochemistry*. Academic Press, New York.
- Zuckerkandl, E., and L. Pauling. 1965. Evolutionary divergence and convergence in proteins. Pp. 97–166 in *Evolving genes and proteins*. Academic press.
- Zwickl, D. J., and M. T. Holder. 2004. Model parameterization, prior distributions, and the general time-reversible model in Bayesian phylogenetics. *Systematic Biology* 53:877–888.