# LEAST-SQUARES FITTING SMOOTH CURVES TO DECADAL RADIOCARBON CALIBRATION DATA FROM AD 1145 TO AD 1945 

FB Knox ${ }^{1} \cdot$ B G McFadgen $^{2}$<br>ABSTRACT. Smoothed curves are least-squares fitted to three sets of decadal radiocarbon calibration data from New Zealand and British Isles (AD 1725-1935) and western North America (AD 1145-1945). The curves are compared with each other and with a curve previously calculated from New Zealand data (AD 1335-1745). The smoothing procedure results in reduced standard deviations of the curves, but at the expense of time resolution. The comparison shows a variable ${ }^{14} \mathrm{C}$ offset between the northern and southern hemispheres of $0-70$ years (Southern Hemisphere older), and a Northern Hemisphere longitudinal variation of -20 to +60 years (British Isles generally older than western North America).

## INTRODUCTION

In a previous paper (McFadgen et al. 1994) it was shown that a simple conversion of radiocarbon dates to calendar dates, using ${ }^{14} \mathrm{C}$ calibration data, results in artificial spreading and clumping of the calendar dates. The effect arises from an interaction between the spread of ${ }^{14} \mathrm{C}$ data due to measurement statistics and changes of slope in the calibration curve, whose effect can be mitigated by reducing the slope changes. While some of the change in calibration curve slope is irreducibly set by geophysical factors, an appreciable fraction on the short time-scale is produced by statistical fluctuations from one calibration point to the next, and can be reduced by the application of appropriate filters.

Subsequently (Knox and McFadgen 1997), a method was presented for achieving a reduction in short time-scale slope changes in the curve by least-squares fitting a Wiener filtered Fourier sum to the data: illustrated on a short ${ }^{14} \mathrm{C}$-dated Matai tree-ring decadal calibration sequence from New Zealand (Sparks et al. 1995). The present paper applies the method, slightly modified, to the above calibration sequence, and to three other independently measured sets of decadal calibration data. These are the Southern Hemisphere Cedar and Northern Hemisphere Oak data (AD 1725-1935) of McCormac et al. (1998), and a segment (AD 1145-1945) of the Northern Hemisphere Douglas Fir data of Stuiver et al. (1998).

The Douglas Fir data are from a decadal set that is a component of the INTCAL98 ${ }^{14} \mathrm{C}$ age curve (Stuiver et al. 1998). Our method of fitting a smooth curve in no way replaces the careful assessment of errors that has gone into producing the decadal points of the INTCAL98 curve, but is an improvement over simply joining successive decadal points by straight lines to form that curve. Such joining by straight lines, a traditional aid to the eye, is clearly artificial, and does nothing to smooth away the known statistical fluctuations of measurement between nearby decadal points. By contrast, fitting a Wiener filtered Fourier series to the data optimally removes some of this statistical fluctuation at the expense of time resolution. It produces a smooth curve that is optimal in that it is as close as possible (in the least-squares sense) to the true calibration curve (Press et al. 1994: §13.3).

The Douglas Fir curve calculation is here limited to using data later than AD 1140 as the segment just prior to this includes some irregularly spaced data, the correct processing of which requires modification of our method.

[^0]In addition to their direct use in calibration, the smoothed curves should facilitate comparison of the Northern and Southern Hemisphere data, to find how closely ${ }^{14} \mathrm{C}$ variation matches in the two hemispheres. Such comparison can shed light on the relevant geophysical processes that produce the wiggles in the calibration curves.

## METHOD

## Detrending

The procedure for least-squares fitting Fourier sums to the data points follows that given by Knox and McFadgen (1997) for Sparks et al.'s (1995) Matai data, except for detrending. Detrending in the above paper was achieved simply by subtracting points on the ideal straight line representing ${ }^{14} \mathrm{C}$ versus tree-ring age (AD 1950 equivalent to BP 0 ) from corresponding data points. The procedure works well in the earlier case as the endpoint data values differ from the ideal by no more than one or two years, but this happy coincidence does not hold for data segments in general. Detrending in this paper is achieved by the slightly more complicated procedure of subtracting from the data corresponding points on a straight line joining the endpoints of the data segment. As before, the trend is eventually restored to give the final answer.

Any straight line detrending procedure on a finite set of discrete data will introduce spurious highfrequency components to the data set, not all of which are cancelled when the trend is restored. This comes about through the phenomenon called aliasing (Press et al. 1994:§12.1). The non-cancellable frequencies lie above the Nyquist frequency (the inverse of twice the sampling period), whose frequency, for decadal sampling is $0.05 \mathrm{yr}^{-1}$ (Figure 1). It can be shown that detrending by subtracting the straight line joining the endpoints of the data segment introduces less power at frequencies that can be aliased than the alternative of detrending by subtracting the best fit straight line, with its generally discontinuous jumps to zero at the ends of the data segment. Detrending by subtracting the straight line joining the endpoints only introduces a corresponding discontinuity of slope.


Figure 1 Spectrum of noise power compared with signal+noise power for Matai calibration curve data.

Aliasing errors in a curve detrended by subtracting the straight line joining data end-points is then less (in our case by one to two orders of magnitude) than for a curve detrended by subtracting the best-fit straight line. This explains the small differences, concentrated near the ends, between the differently detrended Matai curves illustrated in Figure 2.


Figure 2 Comparison of Matai calibration curves detrended by the straight line joining data end-points, and detrended by the straight line least-squares fitted to the data points. $\chi^{2}$ statistic for curve detrended by straight line joining data end-points $=9.98$, and for curve detrended by least-squares fitted straight line to the data points $=8.27 .\left(\chi_{0.95}^{2}=11.07\right.$, d.f. $\left.=5\right)$. Note the slight divergence of the two curves at each end of the plot. Standard deviation of the smoothed curves shown as sd.

Since the aliasing errors in the curve detrended by subtracting the straight line joining the data endpoints are one to two orders of magnitude smaller than in the curve detrended by subtracting the least-squares fitted line, most of the difference between the ends of the two curves in Figure 2, a difference of up to 30 years, is expected to be due to aliasing in the least-squares detrended curve. This implies error at the ends of the curve detrended by subtracting the straight line joining data endpoints is one to two orders of magnitude smaller than 30 years, say not more than 3 years at the outside and small compared with the 9 years estimated for the standard deviation of the Matai calibration curve. This 3-year upper limit to the error generated by aliasing is also small compared with the standard deviations of the other calibration curves, except the AD 1515-1945 Douglas Fir curve. In the last case the 3 -year upper limit equals the standard deviation but would be covered by increasing to 4 years the standard deviation at the curve ends.

In situations where there are many thousands or more of data points, the power in the aliased frequencies is proportionately small enough not to matter in either method of detrending, and it is usual to subtract the best-fit straight line. However, here the small number of data points (22-44) allows a detectable error to creep in when detrending is by the least-squares straight line. Although both Matai curves pass the statistical Chi-squared test (see later), in light of the above discussion on aliasing, we have opted for detrending by subtracting the straight line joining data end-points.

## Filtering

The procedure for least-squares fitting Fourier sums to the data involves obtaining the Fourier transform of the chosen section of detrended data, filtering out as much as possible of the statistical fluctuations in the data set, and then inverse Fourier transforming the result to get a least-squares fitted smooth curve to the detrended data. Finally, the trend is added back to give the corresponding curve through the original calibration data.

Along with filtering the statistical fluctuations, opportunity is taken to apply an extra filter to compensate somewhat for the decadal averaging inherent in the data points: $R_{n}$ in equation (8).

The filter used for removing the statistical fluctuations is the optimal Wiener filter, which, in the notation of Press et al. (1994: §13.3), weights a Fourier component of frequency $f$ by multiplying by $\Phi(f)=|S(f)|^{2} /\left(|S(f)|^{2}+|N(f)|^{2}\right)$, where $|S(f)|^{2}$ is the power spectrum of the desired signal, and $|N(f)|^{2}$ is the power spectrum of the statistical fluctuations or noise (Press et al. 1994).
Since the power spectrum of the signal and noise together, say $|C(f)|^{2}$, is what is actually measured, and is approximately equal to $|S(f)|^{2}+|N(f)|^{2}$ (Press et al. 1994), we can put $\Phi(f)=1-|N(f)|^{2} /|C(f)|^{2}$, which, in the notation of our earlier paper (Knox and McFadgen 1997: equation [9]), is

$$
\begin{equation*}
\left.\left.\left\langle\Phi_{n}^{\prime}\right\rangle=1-\left.\langle | Y_{n}\right|^{2}\right\rangle /\left.\langle | T_{n}^{\prime}\right|^{2}\right\rangle \tag{1}
\end{equation*}
$$

In order to calculate $\left.\left.\langle | Y_{n}\right|^{2}\right\rangle$ we take the statistical fluctuations in the decadal data points as being distributed normally about zero with the standard deviations given, and calculate a discrete Fourier transform

$$
\begin{equation*}
Y_{n}=\sum_{k=0}^{N-1} v_{k} e^{2 \pi i k n / N} \tag{2}
\end{equation*}
$$

(Knox and McFadgen 1997, equation [5]), where a value $v_{k}$ is obtained as a number of years selected randomly from a normal distribution (mean zero) with standard deviation equal to that given for the kth data point. All values of $v_{k}$ in the range $0 \leq k \leq N-1$ other than those corresponding to the data points are put equal to zero, the same procedure as for obtaining the Fourier transform of the calibration data set itself (Knox and McFadgen 1997, equation [1]). The power spectrum of $Y_{n}$ is now $\left|Y_{n}\right|^{2}$ the square of the modulus of $Y_{n}$, a discrete function of the frequency $\mathrm{f}=\mathrm{n} / \mathrm{N}$ in units of reciprocal years $\left(\mathrm{yr}^{-1}\right)$.

In any independent run of the $Y_{n}$ power spectrum only one randomly chosen value of $v_{k}$ is used at each value of $k$. Different runs of randomly chosen values were generally found to give very irregular spectra, varying appreciably from one run to the next. However, an average of 500 independent runs produces an acceptably constant and smooth spectrum, denoted here by $\left.\left.\langle | Y_{n}\right|^{2}\right\rangle$.
The same problem occurs in calculating $\left.\left.\langle | T_{n}\right|^{2}\right\rangle$, because each $\tau_{k}$ (see below) is measured only once, and only one set of data is available; but overcoming the problem requires a more elaborate procedure than that given for calculating $\left.\left.\langle | Y_{n}\right|^{2}\right\rangle$ : we first calculate the discrete Fourier transform of the noisy calibration data

$$
T_{n}=\sum_{k=0}^{N-1} \tau_{k} e^{-2 \pi i k n / N}
$$

(Knox and McFadgen 1997, equation [1]), where $\tau_{k}$ are the detrended ${ }^{14} \mathrm{C}$ calibration data values in years, listed in Table 1 of the present paper (see Appendix). As in the case of $Y_{n}$, the power spectrum of the noisy data is now $\left|T_{n}\right|^{2}$ (the square of the modulus of $T_{n}$ ), and a preliminary filter

$$
\begin{equation*}
\left.\Phi_{n}^{\prime}=1-\left.\langle | Y_{n}\right|^{2}\right\rangle /\left|T_{n}\right|^{2} \tag{4}
\end{equation*}
$$

Knox and McFadgen 1997, equation [6]) is used to make a first estimate of the detrended calibration curve, by inverse Fourier transforming the product $\Phi^{\prime}{ }_{n} T_{n}$ (Knox and McFadgen 1997, equation [7]):

$$
\begin{equation*}
\theta_{k}^{\prime}=\frac{D}{N} \sum_{n=0}^{N-1} \Phi_{n}^{\prime} T_{n} e^{-2 \pi i k n / N} \tag{5}
\end{equation*}
$$

where D (the spacing in years between decadal points) is a normalizing factor (Knox and McFadgen 1997:197). Now, as in the case of $Y_{n}$, we calculate a Fourier transform

$$
\begin{equation*}
T_{n}^{\prime}=\sum_{k=0}^{N-1} \tau_{k}^{\prime} e^{2 \pi i k n / N} \tag{6}
\end{equation*}
$$

where a set of $\tau^{\prime}{ }_{k}$ is obtained by taking the set of k values of the above first estimate of the detrended calibration curve, $\theta^{\prime}$, and adding to them a particular set of the $v_{k}$ calculated previously, i.e.

$$
\begin{equation*}
\tau_{k}^{\prime}=\theta^{\prime}{ }_{k}+v_{k} \tag{7}
\end{equation*}
$$

This is repeated for all 500 independent sets of $v_{k}$, and the corresponding 500 independent estimates of $\left|T^{\prime}{ }_{n}\right|^{2}$ averaged to give the $\left.\left.\langle | T^{\prime}{ }_{n}\right|^{2}\right\rangle$ used to calculate $\left\langle\Phi^{\prime}{ }_{n}\right\rangle$ in equation (1).

Since the $v_{k}$ are distributed normally about zero, the $\tau^{\prime}{ }_{k}$ are distributed normally about $\theta^{\prime}{ }_{k}$, making a spectrum which is the sum of both signal and noise power, distinct from $\left|Y_{n}\right|^{2}$ the power spectrum of noise only.

Sufficiently large averages of $\left|T^{\prime}{ }_{n}\right|^{2}$ and $\left|Y_{n}\right|^{2}$ converge to the two distinct power spectra $\left.\left.\langle | T^{\prime}{ }_{n}\right|^{2}\right\rangle$ and $\left.\left.\langle | Y_{n}\right|^{2}\right\rangle$ plotted against frequency $(\mathrm{n} / \mathrm{N})$ in Figure 1. Even averaged over 500 simulations there is still appreciable statistical fluctuation in the spectra, obvious in the high frequency region where at some frequencies the difference between signal plus noise and noise alone is negative. Since this difference must be positive, where it is negative the signal must be negligible compared with the noise. Accordingly we take the signal to be negligible compared with noise for all cases where the difference between signal plus noise and noise alone (regardless of sign) is less than or equal to the largest magnitude of the observed negative difference. In Figure 1 this occurs for all frequencies above $0.01367 \mathrm{yr}^{-1}$, and the corresponding part of the Wiener filter is set equal to zero. Also, since the statistical fluctuations in the long run average to zero, ideally the zero frequency component of the noise is zero, i.e. $|N(0)|^{2}=0$ and $\Phi(0)=1$. The complete Wiener filters, together with signal to noise values, are given in Table 2.

Now, points on the detrended calibration curve are given by

$$
\begin{equation*}
\theta_{k}=\frac{D}{N} \sum_{n=0}^{N-1} \frac{\left\langle\Phi_{n}^{\prime}\right\rangle}{R_{n}} T_{n} e^{-2 \pi i k n / N} \tag{8}
\end{equation*}
$$

Table 2 Wiener filters $\left(=\left\langle\Phi_{n}^{\prime}\right\rangle\right)$ and ratios of signal/noise $\left(=\left(|\mathrm{S}|^{2} /|\mathrm{N}|^{2}\right)\right.$ for smoothed calibration curves. Note: weighting of the terms in the Fourier sum $=\left\langle\Phi_{n}^{\prime}\right\rangle \mathrm{T}_{\mathrm{n}} / \mathrm{R}_{\mathrm{n}}$ (see Equation 8).

| Frequency ( $\mathrm{y}^{-1}$ ) | Matai | $\begin{aligned} & \text { tai } \\ & 5-1745 \end{aligned}$ | Douglas Fir AD 1145-505 |  | Douglas Fir AD 1515-1945 |  | $\begin{gathered} \text { Oak } \\ \text { AD } 1725-1935 \end{gathered}$ |  | Cedar AD 1725-1935 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\left\langle\Phi^{\prime}{ }_{n}\right\rangle$ | $\|\mathrm{S}\|^{2} /\|\mathrm{N}\|^{2}$ | $\left\langle\Phi^{\prime}{ }_{n}\right\rangle$ | $\|S\|^{2} /\|N\|^{2}$ | $\left\langle\Phi^{\prime}{ }_{n}\right\rangle$ | $\|\mathrm{S}\|^{2} / \mid \mathrm{N}^{2}$ | $\left\langle\Phi^{\prime}{ }_{n}\right\rangle$ | $\|\mathrm{S}\|^{2} / \mid \mathrm{N}^{2}$ | $\left\langle\Phi^{\prime}{ }_{n}\right\rangle$ | $\|\mathrm{S}\|^{2} /\|\mathrm{N}\|^{2}$ |
| 0 | 1.000 | $\infty$ | 1.000 | $\infty$ | 1.000 | $\infty$ | 1.000 | $\infty$ | 1.000 | $\infty$ |
| 0.001953125 | 0.969 | 31 | 0.895 | 8.5 | 1.000 | $\infty$ | 0.982 | 55 | 0.983 | 58 |
| 0.003906250 | 0.983 | 58 | 0.886 | 7.8 | 0.997 | 330 | 0.988 | 82 | 0.986 | 70 |
| 0.005859375 | 0.948 | 18 | 0.985 | 66 | 1.000 | $\infty$ | 0.976 | 41 | 0.968 | 30 |
| 0.007812500 | 0.966 | 28 | 0.985 | 66 | 0.995 | 200 | 0.926 | 13 | 0.913 | 10 |
| 0.009765625 | 0.758 | 3.1 | 0.926 | 13 | 0.978 | 44 | 0.938 | 15 | 0.942 | 16 |
| 0.011718750 | 0.369 | 0.58 | 0.750 | 3.0 | 0.991 | 110 | 0.885 | 7.7 | 0.901 | 9.1 |
| 0.013671875 | 0.194 | 0.24 | 0.575 | 1.4 | 0.990 | 99 | 0.438 | 0.78 | 0.574 | 1.3 |
| 0.015625000 | 0 | 0 | 0.732 | 2.7 | 0.983 | 58 | 0.379 | 0.61 | 0.672 | 2.0 |
| 0.017578125 | 0 | 0 | 0.871 | 6.8 | 0.978 | 44 | 0.146 | 0.17 | 0.477 | 0.91 |
| 0.019531250 | 0 | 0 | 0.256 | 0.34 | 0.894 | 8.4 | 0.750 | 3.0 | 0.756 | 3.1 |
| 0.021484375 | 0 | 0 | 0.141 | 0.16 | 0.980 | 49 | 0.794 | 3.9 | 0.852 | 5.8 |
| 0.023437500 | 0 | 0 | 0 | 0 | 0.873 | 6.9 | 0.510 | 1.0 | 0.737 | 2.8 |
| 0.025390625 | 0 | 0 | 0 | 0 | 0.879 | 7.3 | 0 | 0 | 0 | 0 |
| 0.027343750 | 0 | 0 | 0 | 0 | 0.420 | 0.72 | 0 | 0 | 0 | 0 |
| 0.029296875 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 0.031250000 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 0.033203125 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 0.035156250 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 0.037109375 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 0.039062500 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 0.041015625 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 0.042968750 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 0.044921875 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 0.046875000 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 0.048828125 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |

(Knox and McFadgen 1997, equation [10]), where $R_{n}$ is a function correcting decadal averaging in the data (Knox and McFadgen 1997, equation [3]). Finally the optimally least-squares fitted curve is produced by adding back the initially removed trend.

Optimally least-squares fitted curves to the Matai data, detrended in the two ways discussed earlier, are plotted for comparison in Figure 2. The Matai curve, detrended by joining its end-points, is given in the Appendix (Table 6), and is not significantly different from the curve in our earlier paper (Knox and McFadgen 1997, appendix).
Strictly speaking the filter $\left\langle\Phi^{\prime}{ }_{n}\right\rangle$ calculated in equation (1), while optimal for the smooth curve passing through the set of values $\theta^{\prime}{ }_{k}$, is not quite optimal for estimating the true curve giving rise to the original set of calibration data. Nevertheless, for this data it will be near optimal to the extent that the curve passing through the $\theta^{\prime}{ }_{k}$ approximates the true calibration curve; the more so in that, as pointed out by Press et al. (1994), results obtained by optimal filtering differ from the true curve by an amount that is second order in the precision to which the optimal filter is determined.

In fact, Press et al. further point out that a plot of $|C(f)|^{2}$ versus $f$ will often show the spectral signature of a signal rising above a continuous spectrum which can be inferred to be noise (as can be seen in our Figure 1), and suggest that in this case it is sufficient to just extrapolate by eye the noise spectrum through the signal to obtain the full range of $|N(f)|^{2}$. However, in our case we are better off, in that we know the structure and magnitude of the noise, and can calculate $|N(f)|^{2}$ independently of
$|C(f)|^{2}$. The above indicates optimal Wiener filtering is a robust method for separating signal from noise in the situation considered here.

Next we check how well the true calibration curve has been estimated by applying the standard Chisquared statistical test for normality (Snedecor and Cochran 1967:84) to the distribution of the original data points about corresponding decadal averages derived from the estimated curve. The Chisquared check is completely independent of the methods used here to derive the estimated curve. Results of the Chi-squared test are given in Table 3 and Figure 2.

Table 3 Chi-square test statistic, mean standard deviation and standard error of standard deviation of the estimated true calibration curves. $\chi^{2}{ }_{0.95}=11.07$, d.f. $=5$. The chi-square statistic is calculated by taking the observed difference between each data point and the local decadal average of the smoothed curve divided by the standard deviation of the data point, and testing the resulting frequency distribution against that predicted by the normal distribution. Cell expectations at the extremes of the distribution are combined so that their sum is more than 1 (Snedecor and Cochran 1967).

|  | Matai <br> (AD 1335-1745) | Douglas Fir <br> (AD 1145-1505) | Douglas Fir <br> (AD 1515-1945) | Oak <br> (AD 1725-1935) | Cedar <br> (AD 1725-1935) |
| :--- | :--- | :--- | :--- | :--- | :--- |
| Chi-square <br> test statistic | 9.98 | 9.64 | 4.05 | 6.56 |  |
| Mean sd | 9 | 8 | 3 | 8 | 2.73 |
| Se of sd | 3 | 3 | 1 | 4 | 8 |

The likely error in the estimation of a true calibration curve was determined by taking a decadal average of the estimated curve about each calendar date corresponding to an original data point, and adding to it a number of years selected randomly from a normal distribution with standard deviation equal to that of the corresponding data point. These values along with their corresponding calendar dates constitute a simulated data set from which a simulated calibration curve can be recovered by the procedure described in this paper.

For each calibration curve estimated from the real data, 500 simulated curves were constructed, and at any given calendar date of a calibration curve the 500 simulated calibration points were found to be distributed normally with standard deviations not significantly different at the different calendar dates. An average of these standard deviations over the range of calendar dates is given as a measure of error in the corresponding estimation of the true calibration curve.

In our previous paper (Knox and McFadgen 1997) we had not determined that the estimated points of the true calibration curve were distributed normally, and so could only claim the estimated standard deviation to be nominal. Here, and now also in the earlier paper, we can take quoted standard deviations to have their usual strict interpretation.

## RESULTS AND COMPARISON OF CURVES

The same processing as for the Matai data was applied to the Douglas Fir, Oak, and Cedar data. The Douglas Fir data falls into two separate parts: a part AD 1145-1505 having consistently higher standard deviations (root-mean-square value 13 yr), and a part AD 1515-1945 having consistently lower standard deviations (root-mean-square value 4 yr ), see Table 1 . These parts were processed separately.

The standard deviations for the Matai and other estimated curves are given in Table 4, and the Matai calibration curve itself in the Appendix (Table 6) and Figures 2 and 4a. Results of processing the

Douglas Fir, Oak, and Cedar data are given in the Appendix (Tables 7 to 10). The Douglas Fir curve for AD 1140-1510 is plotted in Figure 3, overlaid, for comparison, with the corresponding section of the INTCAL 98 curve. The Matai, Oak and Cedar curves, and both sections of the Douglas Fir curve are plotted and compared in Figures 4a,b.

........ INTCAL 98
-_ Smoothed curve
Figure 3 Comparison of the observed Douglas Fir ${ }^{14} \mathrm{C}$ (AD 1145-1505) INTCAL 98 curve and the smoothed curve. Note the reduced standard deviation of the smoothed curve ( 8 yr ) compared with the root-mean-square standard deviation of the observed data (13yr).

The standard deviation of a point on the estimated true calibration curve, row 4 of Table 4 , is in all cases less than the root-mean-square standard deviation of the raw data, row 2 of Table 4. This is because each point on the estimated calibration curve is effectively a running mean over several data points. The time interval over which the data is effectively averaged is approximately the interval corresponding to the filter frequency at which the ratio of signal to noise power falls below one, and these time intervals are listed in row 1 of Table 4.

Taking the Matai example in Table 4, since the raw data is taken at 10 yr intervals the 45 yr averaging by the Wiener filter reduces the root-mean-square standard deviation of 20 yr to $\sqrt{10 / 45}$ times 20 yr , i.e. 9 yr. This approximate estimate of the standard deviation of a point on the Matai curve, along with corresponding estimates for the other curves, is given in row 3 of Table 4.

Comparison of rows 3 and 4 of Table 4 shows good agreement between the above approximate estimates of the standard deviations and the standard deviations measured over the sets of 500 simulated calibration curves. Thus the reduction in standard deviation for the calibration curves is obtained at the expense of time resolution, but the lost short term variations in the curve estimates are of amplitude so small as to be obscured in any case by the statistical fluctuations in the raw data.
$\Psi$


$\boldsymbol{\omega}$
$\mathrm{dq} \mathrm{I} \kappa \mathrm{\rho}_{\text {๖I }}$
Figure 4 Comparison of smoothed calibration curves for the northern and southern hemispheres. Figure 4 a shows Northern Hemisphere curve from AD 1140
to AD 1950 using North American Douglas Fir (Pseudotsuga menziesii). Southern Hemisphere curve from AD 1330 to AD 1750 using New Zealand Matai (Prumnopitys taxifolia) and from AD 1720 to AD 1940 using New Zealand Cedar (Libocedrus bidwillii). Figure 4b shows smoothed calibration curves for Cedar, Oak (Quercus petroca), and Douglas Fir from AD 1720 to AD 1950.

Time resolution for the calibration curves can be taken as approximately the values given in row 1 of Table 4.

The presence of shorter time variations in the Douglas Fir, Oak and Cedar curves results in larger corrections for the ten year averaging in the data points of these curves than for the ten year averaging in the Matai data points. By evaluating $\theta_{\mathrm{k}}$ with and without the factor $\mathrm{R}_{\mathrm{n}}$ (see Equation 8), and comparing the results, it is found that the corrections can be up to three ${ }^{14} \mathrm{C}$ years for the first three curves compared with up to one year for the Matai curve.

Table 4 Comparison of standard deviation of 500 samples of the estimated true calibration curve with approximate standard deviation obtained by averaging the data over an interval long enough to suppress time variation for which $\left(|S|^{2} /|\mathrm{N}|^{2}\right)<1$.

|  | Matai <br> (AD 1335-1745) | Douglas Fir <br> (AD 1145-1505) | Douglas Fir <br> (AD 1515-1945) | Oak <br> (AD 1725-1935) | Cedar <br> (AD 1725-1935) |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Approx. interval over <br> which data effec- <br> tively averaged (yr) | 45 | 27 | 18 | 21 | 21 |
| Root-mean-square <br> standard deviation of <br> data (yr) | 20 | 13 | 4 | 14 | 13 |
| Expected approx. <br> standard deviation of <br> point on curve (yr) | 9 | 8 | 3 | 10 | 9 |
| Actual standard devi- <br> ation of point on <br> curve (yr) | 9 | 8 | 3 | 8 | 8 |

The reason for processing the Douglas Fir data in two separate parts is as follows. As discussed above, the time interval over which data is effectively averaged is the period corresponding to the frequency at which the ratio of signal to noise power is one, and the noise power is determined by the root-mean-square average of the standard deviations of the data set. In those parts of the data set where the standard deviations are systematically and significantly smaller than elsewhere the time interval over which the local data are averaged, set by the average standard deviation over the whole data set, is longer than is locally optimal. In such parts of the data set the estimated calibration curve in general will not follow the data points as closely as it should. Similarly, in those parts of the data set where the standard deviations are larger than elsewhere, requiring but not having an extra long averaging time to smooth away the noise, the calibration curve will follow the data points more closely than it should. Where the above effect is significant the Chi-squared test for a normal distribution of data points about local decadal averages of the curve will not be passed.

The above defect in our method of estimating a calibration curve is overcome by separating the data into sequences with approximately constant root-mean-square running mean averages of the standard deviations.

In the data sets used in this paper the standard deviations averaged over the appropriate number of data points (determined a posteriori using the time resolution intervals in Table 4 above) vary on average by not more than $13 \%$, with extremes not more than $34 \%$, and the corresponding calibration curves pass the appropriate Chi-squared tests. By contrast the two parts of the Douglas Fir data, with their root-mean-square standard deviations differing by a factor of $\sim 3$, when processed together as one set of data do not pass the Chi-squared test.

Examination of Figure 4 a and the ${ }^{14} \mathrm{C}$ dates at AD 1510 in the Appendix (Tables 7 and 8) shows acceptable agreement between estimates of the true ${ }^{14} \mathrm{C}$ age at AD 1510 in the two abutting Douglas Fir curves. The difference between the two estimates is $16 \pm 9 \mathrm{yr}$, where the standard deviation of 9 yr is taken as the square root of the sum of the variances of the two curves (Table 4).

Figure 4 a also shows the Matai curve continuing approximately into the Cedar curve, as would be expected since both trees used grew in the same region of the globe (Table 5). However, closer examination of the overlap of these curves shows the difference between them, though small, to be statistically significant. The difference (Cedar minus Matai) varies from $-21 \pm 12$ to $+41 \pm 12 \mathrm{yr}$, where the standard deviation of the differences is the square root of the sum of the variances of the Cedar and Matai curves (Table 4).

Table 5 Global locations of trees used: taken from Stuiver and Becker (1986), McCormac et al. (1998), and R J Sparks (personal communication)

| Tree | Latitude | Longitude | Location |
| :--- | :--- | :--- | :--- |
| Douglas Fir | $47^{\circ} 46^{\prime} \mathrm{N}$ | $124^{\circ} 06^{\prime} \mathrm{W}$ | West coast USA (two trees) |
|  | $46^{\circ} 45^{\prime} \mathrm{N}$ | $121^{\circ} 45^{\prime} \mathrm{W}$ | West coast USA |
|  | $43^{\circ} 07^{\prime} \mathrm{N}$ | $123^{\circ} 40^{\prime} \mathrm{W}$ | West coast USA |
|  | $47^{\circ} \mathrm{N}$ | $122^{\circ} \quad \mathrm{W}$ | West coast USA |
|  | $48^{\circ} 40^{\prime} \mathrm{N}$ | $123^{\circ} 40^{\prime} \mathrm{W}$ | West coast Canada |
| English Oak | $53^{\circ} 12^{\prime} \mathrm{N}$ | $01^{\circ} 04^{\prime} \mathrm{W}$ | Middle England |
|  | $54^{\circ} 44^{\prime} \mathrm{N}$ | $06^{\circ} 16^{\prime} \mathrm{W}$ | North coast, Ulster |
| New Zealand Cedar | $39^{\circ} 32^{\prime} \mathrm{S}$ | $175^{\circ} 44^{\prime} \mathrm{E}$ | Middle North Island, |
| New Zealand Matai | $44^{\circ}$ | S | $171^{\circ}$ |

Since the Cedar and Matai curves are detrended by subtracting lines joining their data end points, aliasing errors, which concentrate at the ends of an estimated curve, should be one to two orders of magnitude less than the maximum of 30 yr shown in Figure 2 for a curve detrended by subtracting a best fit straight line. Part of the difference between points on the estimated Cedar and Matai curves may be due to aliased end errors, but by not more than three years.

On the other hand, a more detailed examination of where the trees grew suggests the difference between the above curves could also be real. The Cedar wood is from a tree that grew in the center of North Island, New Zealand, while the Matai wood is from a tree which grew halfway along the east coast of South Island, New Zealand. North Island is surrounded by subtropical water (Carter et al. 1998), while the east coast of South Island faces the subtropical front, a relatively sharp, fluctuating boundary between subtropical and subantarctic water. If the ${ }^{14} \mathrm{C} /{ }^{12} \mathrm{C}$ ratio is different in the ventilated carbon dioxide from the subtropical and subantarctic waters we can expect corresponding differences in this ratio in the Cedar and Matai wood, depending on differences in prevailing wind directions at the two sites.

Furthermore, at times the Cedar tree would have been exposed to carbon dioxide vented from North Island's central volcanic plateau, including Mount Ruapehu, an intermittently active volcano some 25 km away, which carbon dioxide may have a reduced ${ }^{14} \mathrm{C} /{ }^{12} \mathrm{C}$ ratio. However, we cannot say to what extent this effect would be significant. Geochemistry of the tree rings might give some indication.

Coming now to a general comparison of all the calibration curves, Figures 4a,b and 5 show the Matai and Cedar curves to generally give older ${ }^{14} \mathrm{C}$ ages than the Douglas Fir. The average difference in age, up to AD 1900, is seen to support in order of magnitude the recommendation by Stuiver et al.
(1998) to reduce Southern Hemisphere ${ }^{14} \mathrm{C}$ ages by 24 yr prior to calibration. The difference, however, is highly variable: from a maximum of more than 60 yr , which is more than seven standard deviations and highly significant ( $\mathrm{P} \ll 0.0001$ ), to virtually zero (within one standard deviation). Sparks et al. (1995), Stuiver and Braziunas (1998), and McCormac et al. (1998) also report temporal variation in the interhemispheric offset.


Figure 5 Difference in ${ }^{14} \mathrm{C}$ years between curves for northern and southern hemispheres (Douglas Fir-Matai, Douglas Fir-Cedar), and between western North America and Britain (Douglas Fir-Oak). Combined standard errors for differences between pairs of curves shown by error bars: $a=$ Douglas Fir (AD 1330 to 1510)-Matai, b=Douglas Fir (AD 1510 to 1750)-Matai, c=Douglas Fir (AD 1720 to 1940)-Cedar, d=Douglas Fir (AD 1720 to 1940)-Oak.

An interesting feature of the variability is that, with one exception, the minimum difference occurs at times ( $\sim$ AD 1380-1430, $\sim$ AD 1550, $\sim$ AD 1610-1660, $\sim$ AD 1800: see Figure 5) when the calibration curves are most rapidly descending (Figures 4a,b). The exception (~AD 1930-1940) occurs where the curves are rapidly rising, but, as pointed out by McCormac et al. (1998) and Stuiver et al. (1998), by this time there is appreciable injection of fossil fuel carbon into the atmosphere. Consistent with this view Figure 4 b shows all three curves rising to older ${ }^{14} \mathrm{C}$ ages at this time. Thus the exception may not be due entirely to natural causes.

A further interesting feature is that the Northern Hemisphere Oak and Douglas Fir curves do not always agree. There is good agreement between them from AD 1780-1870, but earlier, ~AD 1760, the Oak curve is significantly older by more than five standard deviations ( $\mathrm{P} \ll 0.0001$ ), and actually closer to the Southern Hemisphere Cedar curve. Later, ~AD 1880-1910, there is no significant difference between the Oak and Cedar curves while both give ${ }^{14} \mathrm{C}$ ages some 60 yr older than the Douglas Fir: almost as large a difference as occurs anywhere in Figures 4 and 5. This difference between Oak and Douglas Fir is more than six standard deviations and highly significant $(\mathrm{P} \ll 0.0001)$. Also the average difference between Oak and Douglas Fir ${ }^{14} \mathrm{C}$ ages up to AD 1900, at $19+/-3$ years, falls little short of the average difference between Southern and Northern Hemisphere ages giving rise to Stuiver et al.'s (1998) recommended 24-year correction.

The significant (though small) differences between the Oak and Douglas Fir curves might be expected: the Oaks grew on the prevailing leeward side of the North Atlantic Ocean whereas the Douglas Firs grew on the prevailing leeward side of the North Pacific. Ventilated carbon dioxide from these different oceans may have different ${ }^{14} \mathrm{C} /{ }^{12} \mathrm{C}$ ratios.

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## APPENDIX

Table $1{ }^{14} \mathrm{C}$ age of tree-ring dated wood of New Zealand Matai (Prumnopitys taxifolia) from Sparks et al. (1995: Table 2), Douglas Fir (Pseudotsuga menziesii) from Stuiver et al. (1998), and English Oak (Quercus petroca), and New Zealand Cedar (Libocedrus bidwillii) from McCormac et al. (1998)

| Calendar age (AD) | Conventional ${ }^{14} \mathrm{C}$ age (BP) | Standard error (yr) | $\mathrm{k}^{\text {a }}$ | Detrended data $\tau_{\mathrm{k}}(\mathrm{yr})$ |
| :---: | :---: | :---: | :---: | :---: |
| Matai AD 1335-1745 |  |  |  |  |
| 1335 | 617 | 22 | 50 | 0 |
| 1345 | 635 | 19 | 60 | -28 |
| 1355 | 639 | 20 | 70 | -42 |
| 1365 | 683 | 20 | 80 | -96 |
| 1375 | 637 | 22 | 90 | -60 |
| 1385 | 618 | 17 | 100 | -51 |
| 1395 | 593 | 19 | 110 | -36 |
| 1405 | 599 | 19 | 120 | -52 |
| 1415 | 530 | 20 | 130 | 7 |
| 1425 | 471 | 21 | 140 | 56 |
| 1435 | 484 | 21 | 150 | 33 |
| 1445 | 422 | 21 | 160 | 85 |
| 1455 | 453 | 17 | 170 | 44 |
| 1465 | 450 | 19 | 180 | 37 |
| 1475 | 420 | 22 | 190 | 57 |
| 1485 | 417 | 17 | 200 | 50 |
| 1495 | 380 | 23 | 210 | 77 |
| 1505 | 380 | 21 | 220 | 67 |
| 1515 | 372 | 15 | 230 | 65 |
| 1525 | 334 | 21 | 240 | 93 |
| 1535 | 323 | 15 | 250 | 94 |
| 1545 | 324 | 14 | 260 | 82 |
| 1555 | 322 | 18 | 270 | 74 |
| 1565 | 307 | 22 | 280 | 79 |
| 1575 | 377 | 25 | 290 | -1 |
| 1585 | 385 | 26 | 300 | -19 |
| 1595 | 396 | 21 | 310 | -40 |
| 1605 | 361 | 12 | 320 | -15 |
| 1615 | 367 | 17 | 330 | -31 |
| 1625 | 360 | 21 | 340 | -34 |
| 1635 | 286 | 18 | 350 | 30 |
| 1645 | 288 | 20 | 360 | 18 |
| 1655 | 290 | 16 | 370 | 6 |
| 1665 | 220 | 20 | 380 | 66 |
| 1675 | 163 | 22 | 390 | 113 |
| 1685 | 163 | 20 | 400 | 103 |
| 1695 | 182 | 23 | 410 | 74 |
| 1705 | 167 | 21 | 420 | 79 |

Table $1{ }^{14} \mathrm{C}$ age of tree-ring dated wood of New Zealand Matai (Prumnopitys taxifolia) from Sparks et al. (1995: Table 2), Douglas Fir (Pseudotsuga menziesii) from Stuiver et al. (1998), and English Oak (Quercus petroca), and New Zealand Cedar (Libocedrus bidwillii) from McCormac et al. (1998) (Continued)

| Calendar <br> age (AD) | Conventional <br> ${ }^{14} \mathrm{C}$ age $(\mathrm{BP})$ | Standard <br> error $(\mathrm{yr})$ | $\mathrm{k}^{\mathrm{a}}$ | Detrended <br> data $\tau_{\mathrm{k}}(\mathrm{yr})$ |
| :---: | :---: | :---: | :---: | :---: |
| 1715 | 157 | 17 | 430 | 79 |
| 1725 | 167 | 20 | 440 | 59 |
| 1735 | 176 | 21 | 450 | 40 |
| 1745 | 206 | 17 | 460 | 0 |

Douglas Fir AD 1145-1505

| 1145 | 967 | 10 | 50 | 0 |
| :---: | :---: | :---: | :---: | :---: |
| 1155 | 907 | 15 | 60 | 43 |
| 1165 | 857 | 10 | 70 | 76 |
| 1175 | 881 | 14 | 80 | 35 |
| 1185 | 887 | 13 | 90 | 11 |
| 1195 | 857 | 14 | 100 | 24 |
| 1205 | 867 | 14 | 110 | -3 |
| 1215 | 832 | 14 | 120 | 15 |
| 1225 | 795 | 10 | 130 | 35 |
| 1235 | 804 | 14 | 140 | 9 |
| 1245 | 803 | 14 | 150 | -8 |
| 1255 | 811 | 14 | 160 | -33 |
| 1265 | 755 | 15 | 170 | 6 |
| 1275 | 767 | 15 | 180 | -23 |
| 1285 | 713 | 14 | 190 | 14 |
| 1295 | 661 | 15 | 200 | 49 |
| 1305 | 627 | 8 | 210 | 65 |
| 1315 | 625 | 10 | 220 | 50 |
| 1325 | 613 | 11 | 230 | 45 |
| 1335 | 560 | 10 | 240 | 81 |
| 1345 | 582 | 12 | 250 | 42 |
| 1355 | 617 | 12 | 260 | -11 |
| 1365 | 604 | 11 | 270 | -15 |
| 1375 | 653 | 11 | 280 | -81 |
| 1385 | 633 | 11 | 290 | -78 |
| 1395 | 575 | 14 | 300 | -37 |
| 1405 | 565 | 13 | 310 | -44 |
| 1415 | 520 | 14 | 320 | -17 |
| 1425 | 500 | 13 | 330 | -14 |
| 1435 | 468 | 14 | 340 | 1 |
| 1445 | 418 | 12 | 350 | 34 |
| 1455 | 390 | 10 | 360 | 45 |
| 1465 | 395 | 14 | 370 | 23 |
| 1475 | 390 | 14 | 380 | 11 |
| 1485 | 380 | 13 | 390 | 3 |
| 1495 | 353 | 10 | 400 | 13 |

Table $1{ }^{14} \mathrm{C}$ age of tree-ring dated wood of New Zealand Matai (Prumnopitys taxifolia) from Sparks et al. (1995: Table 2), Douglas Fir (Pseudotsuga menziesii) from Stuiver et al. (1998), and English Oak (Quercus petroca), and New Zealand Cedar (Libocedrus bidwillii) from McCormac et al. (1998) (Continued)

| Calendar age (AD) | Conventional ${ }^{14} \mathrm{C}$ age (BP) | Standard error (yr) | $\mathrm{k}^{\text {a }}$ | Detrended data $\tau_{\mathrm{k}}$ (yr) |
| :---: | :---: | :---: | :---: | :---: |
| 1505 | 349 | 8 | 410 | 0 |
| Douglas Fir AD 1515-1945 |  |  |  |  |
| 1515 | 349 | 5 | 40 | 0 |
| 1525 | 322 | 4 | 50 | 23 |
| 1535 | 301 | 3 | 60 | 41 |
| 1545 | 309 | 4 | 70 | 29 |
| 1555 | 319 | 5 | 80 | 15 |
| 1565 | 326 | 4 | 90 | 5 |
| 1575 | 329 | 5 | 100 | -2 |
| 1585 | 333 | 4 | 110 | -10 |
| 1595 | 340 | 4 | 120 | -21 |
| 1605 | 365 | 4 | 130 | -49 |
| 1615 | 351 | 4 | 140 | -39 |
| 1625 | 333 | 4 | 150 | -25 |
| 1635 | 308 | 4 | 160 | -3 |
| 1645 | 268 | 4 | 170 | 33 |
| 1655 | 241 | 4 | 180 | 56 |
| 1665 | 209 | 4 | 190 | 85 |
| 1675 | 172 | 4 | 200 | 118 |
| 1685 | 139 | 3 | 210 | 147 |
| 1695 | 115 | 2 | 220 | 167 |
| 1705 | 105 | 3 | 230 | 174 |
| 1715 | 95 | 3 | 240 | 180 |
| 1725 | 114 | 3 | 250 | 157 |
| 1735 | 153 | 4 | 260 | 115 |
| 1745 | 163 | 4 | 270 | 101 |
| 1755 | 156 | 3 | 280 | 104 |
| 1765 | 169 | 4 | 290 | 88 |
| 1775 | 167 | 4 | 300 | 86 |
| 1785 | 216 | 5 | 310 | 33 |
| 1795 | 201 | 4 | 320 | 44 |
| 1805 | 159 | 4 | 330 | 83 |
| 1815 | 106 | 4 | 340 | 132 |
| 1825 | 99 | 3 | 350 | 135 |
| 1835 | 116 | 3 | 360 | 115 |
| 1845 | 115 | 4 | 370 | 112 |
| 1855 | 120 | 4 | 380 | 103 |
| 1865 | 117 | 4 | 390 | 103 |
| 1875 | 115 | 4 | 400 | 101 |
| 1885 | 100 | 3 | 410 | 112 |

Table $1{ }^{14} \mathrm{C}$ age of tree-ring dated wood of New Zealand Matai (Prumnopitys taxifolia) from Sparks et al. (1995: Table 2), Douglas Fir (Pseudotsuga menziesii) from Stuiver et al. (1998), and English Oak (Quercus petroca), and New Zealand Cedar (Libocedrus bidwillii) from McCormac et al. (1998) (Continued)

| Calendar age (AD) | Conventional ${ }^{14} \mathrm{C}$ age (BP) | Standard error (yr) | $\mathrm{k}^{\text {a }}$ | Detrended data $\tau_{\mathrm{k}}$ (yr) |
| :---: | :---: | :---: | :---: | :---: |
| 1895 | 76 | 3 | 420 | 132 |
| 1905 | 78 | 3 | 430 | 127 |
| 1915 | 108 | 3 | 440 | 93 |
| 1925 | 138 | 3 | 450 | 59 |
| 1935 | 156 | 4 | 460 | 38 |
| 1945 | 190 | 4 | 470 | 0 |
| Oak AD 1725-1935 |  |  |  |  |
| 1725 | 133 | 14 | 140 | 0 |
| 1735 | 176 | 14 | 150 | -40 |
| 1745 | 205 | 14 | 160 | -67 |
| 1755 | 208 | 18 | 170 | -67 |
| 1765 | 195 | 13 | 180 | -52 |
| 1775 | 184 | 13 | 190 | -38 |
| 1785 | 217 | 13 | 200 | -69 |
| 1795 | 210 | 14 | 210 | -59 |
| 1805 | 122 | 14 | 220 | 32 |
| 1815 | 128 | 14 | 230 | 28 |
| 1825 | 107 | 14 | 240 | 52 |
| 1835 | 103 | 14 | 250 | 58 |
| 1845 | 137 | 14 | 260 | 27 |
| 1855 | 126 | 14 | 270 | 40 |
| 1865 | 126 | 11 | 280 | 43 |
| 1875 | 146 | 13 | 290 | 26 |
| 1885 | 129 | 14 | 300 | 45 |
| 1895 | 148 | 13 | 310 | 29 |
| 1905 | 109 | 13 | 320 | 70 |
| 1915 | 146 | 14 | 330 | 36 |
| 1925 | 145 | 14 | 340 | 39 |
| 1935 | 187 | 14 | 350 | 0 |
| Cedar AD 1745-1935 |  |  |  |  |
| 1725 | 148 | 14 | 140 | 0 |
| 1735 | 204 | 14 | 150 | -55 |
| 1745 | 228 | 13 | 160 | -79 |
| 1755 | 236 | 13 | 170 | -86 |
| 1765 | 223 | 13 | 180 | -73 |
| 1775 | 204 | 13 | 190 | -53 |
| 1785 | 240 | 13 | 200 | -89 |
| 1795 | 235 | 13 | 210 | -83 |
| 1805 | 170 | 13 | 220 | -17 |
| 1815 | 142 | 14 | 230 | 11 |

Table $1{ }^{14} \mathrm{C}$ age of tree-ring dated wood of New Zealand Matai (Prumnopitys taxifolia) from Sparks et al. (1995: Table 2), Douglas Fir (Pseudotsuga menziesii) from Stuiver et al. (1998), and English Oak (Quercus petroca), and New Zealand Cedar (Libocedrus bidwillii) from McCormac et al. (1998) (Continued)

| Calendar <br> age (AD) | Conventional <br> ${ }^{14} \mathrm{C}$ age (BP) | Standard <br> error $(\mathrm{yr})$ | $\mathrm{k}^{\mathrm{a}}$ | Detrended <br> data $\tau_{\mathrm{k}}(\mathrm{yr})$ |
| :---: | :---: | :---: | :---: | :---: |
| 1825 | 146 | 14 | 240 | 8 |
| 1835 | 149 | 13 | 250 | 5 |
| 1845 | 156 | 14 | 260 | -1 |
| 1855 | 167 | 14 | 270 | -12 |
| 1865 | 161 | 11 | 280 | -5 |
| 1875 | 151 | 10 | 290 | 6 |
| 1885 | 155 | 13 | 300 | 2 |
| 1895 | 129 | 12 | 310 | 29 |
| 1905 | 125 | 13 | 320 | 33 |
| 1915 | 118 | 14 | 330 | 41 |
| 1925 | 172 | 13 | 340 | -13 |
| 1935 | 160 | 14 | 350 | 0 |

${ }^{\mathrm{a}} \mathrm{k}=$ the index number of the data point in the extended data set after zero padding.

Table 6 Least-squares smoothed calibration curve at yearly intervals for New Zealand Matai (Prumnopitys taxifolia) (Sparks et al. 1995) between AD 1330 and AD 1750 corrected for a running mean over 10 tree rings. Mean standard error of the curve is $9 \pm 3 \mathrm{yr}$.

| Tree-ring date AD | $\begin{gathered} { }^{14} \mathrm{C} \text { age } \\ (\mathrm{BP}) \end{gathered}$ | Tree-ring date AD | $\begin{gathered} { }^{14} \mathrm{C} \text { age } \\ (\mathrm{BP}) \end{gathered}$ | Tree-ring date AD | ${ }^{14} \mathrm{C}$ age (BP) | Tree-ring date AD | $\begin{gathered} { }^{14} \mathrm{C} \text { age } \\ (\mathrm{BP}) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1330 | 630 | 1363 | 649 | 1396 | 596 | 1429 | 485 |
| 1331 | 630 | 1364 | 649 | 1397 | 593 | 1430 | 482 |
| 1332 | 631 | 1365 | 649 | 1398 | 590 | 1431 | 479 |
| 1333 | 631 | 1366 | 649 | 1399 | 586 | 1432 | 477 |
| 1334 | 632 | 1367 | 649 | 1400 | 583 | 1433 | 474 |
| 1335 | 632 | 1368 | 649 | 1401 | 579 | 1434 | 472 |
| 1336 | 633 | 1369 | 649 | 1402 | 576 | 1435 | 470 |
| 1337 | 634 | 1370 | 648 | 1403 | 572 | 1436 | 468 |
| 1338 | 634 | 1371 | 647 | 1404 | 569 | 1437 | 466 |
| 1339 | 635 | 1372 | 646 | 1405 | 565 | 1438 | 463 |
| 1340 | 636 | 1373 | 645 | 1406 | 562 | 1439 | 462 |
| 1341 | 636 | 1374 | 644 | 1407 | 558 | 1440 | 460 |
| 1342 | 637 | 1375 | 643 | 1408 | 554 | 1441 | 458 |
| 1343 | 638 | 1376 | 642 | 1409 | 551 | 1442 | 456 |
| 1344 | 639 | 1377 | 641 | 1410 | 547 | 1443 | 454 |
| 1345 | 640 | 1378 | 639 | 1411 | 544 | 1444 | 453 |
| 1346 | 640 | 1379 | 638 | 1412 | 540 | 1445 | 451 |
| 1347 | 641 | 1380 | 636 | 1413 | 536 | 1446 | 450 |
| 1348 | 642 | 1381 | 635 | 1414 | 533 | 1447 | 449 |
| 1349 | 643 | 1382 | 633 | 1415 | 529 | 1448 | 447 |
| 1350 | 643 | 1383 | 631 | 1416 | 526 | 1449 | 446 |
| 1351 | 644 | 1384 | 629 | 1417 | 522 | 1450 | 445 |
| 1352 | 645 | 1385 | 626 | 1418 | 519 | 1451 | 444 |
| 1353 | 646 | 1386 | 624 | 1419 | 515 | 1452 | 443 |
| 1354 | 646 | 1387 | 622 | 1420 | 512 | 1453 | 442 |
| 1355 | 647 | 1388 | 619 | 1421 | 509 | 1454 | 441 |
| 1356 | 647 | 1389 | 617 | 1422 | 506 | 1455 | 440 |
| 1357 | 648 | 1390 | 614 | 1423 | 502 | 1456 | 439 |
| 1358 | 648 | 1391 | 611 | 1424 | 499 | 1457 | 438 |
| 1359 | 649 | 1392 | 608 | 1425 | 496 | 1458 | 437 |
| 1360 | 649 | 1393 | 605 | 1426 | 493 | 1459 | 436 |
| 1361 | 649 | 1394 | 602 | 1427 | 490 | 1460 | 435 |
| 1362 | 649 | 1395 | 599 | 1428 | 488 | 1461 | 435 |

Table 6 Least-squares smoothed calibration curve at yearly intervals for New Zealand Matai (Prumnopitys taxifolia) (Sparks et al. 1995) between AD 1330 and AD 1750 corrected for a running mean over 10 tree rings. Mean standard error of the curve is $9 \pm 3$ yr. (Continued)

| Tree-ring date AD | $\begin{gathered} { }^{14} \mathrm{C} \text { age } \\ (\mathrm{BP}) \end{gathered}$ | Tree-ring date AD | $\begin{gathered} { }^{14} \mathrm{C} \text { age } \\ (\mathrm{BP}) \end{gathered}$ | Tree-ring date AD | $\begin{gathered} { }^{14} \mathrm{C} \text { age } \\ (\mathrm{BP}) \end{gathered}$ | Tree-ring date AD | $\begin{gathered} { }^{14} \mathrm{C} \text { age } \\ (\mathrm{BP}) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1462 | 434 | 1496 | 397 | 1530 | 332 | 1564 | 338 |
| 1463 | 433 | 1497 | 395 | 1531 | 331 | 1565 | 339 |
| 1464 | 433 | 1498 | 394 | 1532 | 330 | 1566 | 341 |
| 1465 | 432 | 1499 | 392 | 1533 | 329 | 1567 | 342 |
| 1466 | 431 | 1500 | 390 | 1534 | 328 | 1568 | 344 |
| 1467 | 430 | 1501 | 388 | 1535 | 327 | 1569 | 346 |
| 1468 | 430 | 1502 | 386 | 1536 | 326 | 1570 | 347 |
| 1469 | 429 | 1503 | 384 | 1537 | 325 | 1571 | 349 |
| 1470 | 428 | 1504 | 382 | 1538 | 324 | 1572 | 351 |
| 1471 | 427 | 1505 | 380 | 1539 | 324 | 1573 | 352 |
| 1472 | 427 | 1506 | 378 | 1540 | 323 | 1574 | 354 |
| 1473 | 426 | 1507 | 376 | 1541 | 323 | 1575 | 356 |
| 1474 | 425 | 1508 | 374 | 1542 | 322 | 1576 | 357 |
| 1475 | 424 | 1509 | 372 | 1543 | 322 | 1577 | 359 |
| 1476 | 423 | 1510 | 370 | 1544 | 322 | 1578 | 360 |
| 1477 | 422 | 1511 | 367 | 1545 | 322 | 1579 | 362 |
| 1478 | 422 | 1512 | 365 | 1546 | 322 | 1580 | 364 |
| 1479 | 421 | 1513 | 363 | 1547 | 322 | 1581 | 365 |
| 1480 | 420 | 1514 | 361 | 1548 | 322 | 1582 | 367 |
| 1481 | 418 | 1515 | 359 | 1549 | 323 | 1583 | 368 |
| 1482 | 417 | 1516 | 357 | 1550 | 323 | 1584 | 369 |
| 1483 | 416 | 1517 | 355 | 1551 | 324 | 1585 | 371 |
| 1484 | 415 | 1518 | 353 | 1552 | 324 | 1586 | 372 |
| 1485 | 414 | 1519 | 351 | 1553 | 325 | 1587 | 373 |
| 1486 | 413 | 1520 | 349 | 1554 | 326 | 1588 | 374 |
| 1487 | 411 | 1521 | 347 | 1555 | 327 | 1589 | 375 |
| 1488 | 410 | 1522 | 345 | 1556 | 328 | 1590 | 376 |
| 1489 | 408 | 1523 | 343 | 1557 | 329 | 1591 | 377 |
| 1490 | 407 | 1524 | 342 | 1558 | 330 | 1592 | 378 |
| 1491 | 405 | 1525 | 340 | 1559 | 331 | 1593 | 379 |
| 1492 | 404 | 1526 | 338 | 1560 | 332 | 1594 | 379 |
| 1493 | 402 | 1527 | 337 | 1561 | 333 | 1595 | 380 |
| 1494 | 401 | 1528 | 335 | 1562 | 335 | 1596 | 380 |
| 1495 | 399 | 1529 | 334 | 1563 | 336 | 1597 | 381 |

Table 6 Least-squares smoothed calibration curve at yearly intervals for New Zealand Matai (Prumnopitys taxifolia) (Sparks et al. 1995) between AD 1330 and AD 1750 corrected for a running mean over 10 tree rings. Mean standard error of the curve is $9 \pm 3$ yr. (Continued)

| Tree-ring date AD | $\begin{gathered} { }^{14} \mathrm{C} \text { age } \\ (\mathrm{BP}) \end{gathered}$ | Tree-ring date AD | $\begin{gathered} { }^{14} \mathrm{C} \text { age } \\ (\mathrm{BP}) \end{gathered}$ | Tree-ring date AD | $\begin{gathered} { }^{14} \mathrm{C} \text { age } \\ (\mathrm{BP}) \end{gathered}$ | Tree-ring date AD | ${ }^{14} \mathrm{C}$ age <br> (BP) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1598 | 381 | 1631 | 331 | 1665 | 221 | 1698 | 162 |
| 1599 | 381 | 1632 | 328 | 1666 | 218 | 1699 | 161 |
| 1600 | 381 | 1633 | 325 | 1667 | 215 | 1700 | 161 |
| 1601 | 381 | 1634 | 322 | 1668 | 212 | 1701 | 161 |
| 1602 | 381 | 1635 | 319 | 1669 | 209 | 1702 | 161 |
| 1603 | 380 | 1636 | 316 | 1670 | 206 | 1703 | 161 |
| 1604 | 380 | 1637 | 313 | 1671 | 204 | 1704 | 161 |
| 1605 | 379 | 1638 | 310 | 1672 | 201 | 1705 | 161 |
| 1606 | 379 | 1639 | 306 | 1673 | 199 | 1706 | 161 |
| 1607 | 378 | 1640 | 303 | 1674 | 196 | 1707 | 161 |
| 1608 | 377 | 1641 | 300 | 1675 | 194 | 1708 | 162 |
| 1609 | 376 | 1642 | 296 | 1676 | 191 | 1709 | 162 |
| 1610 | 375 | 1643 | 293 | 1677 | 189 | 1710 | 162 |
| 1611 | 374 | 1644 | 290 | 1678 | 187 | 1711 | 163 |
| 1612 | 373 | 1645 | 286 | 1679 | 185 | 1712 | 163 |
| 1613 | 371 | 1646 | 283 | 1680 | 183 | 1713 | 164 |
| 1614 | 370 | 1647 | 279 | 1681 | 181 | 1714 | 165 |
| 1615 | 368 | 1648 | 276 | 1682 | 179 | 1715 | 165 |
| 1616 | 367 | 1649 | 273 | 1683 | 177 | 1716 | 166 |
| 1617 | 365 | 1650 | 269 | 1684 | 176 | 1717 | 167 |
| 1618 | 363 | 1651 | 266 | 1685 | 174 | 1718 | 168 |
| 1619 | 361 | 1652 | 262 | 1686 | 173 | 1719 | 168 |
| 1620 | 359 | 1653 | 259 | 1687 | 171 | 1720 | 169 |
| 1621 | 357 | 1654 | 256 | 1688 | 170 | 1721 | 170 |
| 1622 | 355 | 1655 | 252 | 1689 | 169 | 1722 | 171 |
| 1623 | 352 | 1656 | 249 | 1690 | 168 | 1723 | 172 |
| 1624 | 350 | 1657 | 246 | 1691 | 167 | 1724 | 173 |
| 1625 | 348 | 1658 | 242 | 1692 | 166 | 1725 | 174 |
| 1626 | 345 | 1659 | 239 | 1693 | 165 | 1726 | 175 |
| 1627 | 342 | 1660 | 236 | 1694 | 164 | 1727 | 176 |
| 1628 | 340 | 1662 | 230 | 1695 | 163 | 1728 | 177 |
| 1629 | 337 | 1663 | 227 | 1696 | 163 | 1729 | 178 |
| 1630 | 334 | 1664 | 224 | 1697 | 162 | 1730 | 179 |

Table 6 Least-squares smoothed calibration curve at yearly intervals for New Zealand Matai (Prumnopitys taxifolia) (Sparks et al. 1995) between AD 1330 and AD 1750 corrected for a running mean over 10 tree rings. Mean standard error of the curve is $9 \pm 3$ yr. (Continued)

| Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{BP})$ | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{BP})$ | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{BP})$ | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{BP})$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1731 | 180 | 1739 | 186 | 1747 | 191 |  |  |
| 1732 | 180 | 1740 | 187 | 1748 | 191 |  |  |
| 1733 | 181 | 1741 | 188 | 1749 | 191 |  |  |
| 1734 | 182 | 1742 | 188 | 1750 | 192 |  |  |
| 1735 | 183 | 1743 | 189 |  |  |  |  |
| 1736 | 184 | 1744 | 189 |  |  |  |  |
| 1737 | 185 | 1745 | 190 |  |  |  |  |
| 1738 | 185 | 1746 | 190 |  |  |  |  |

Table 7 Least-squares smoothed calibration curve at yearly intervals for North American Douglas Fir (Pseudotsuga menziesii) (Stuiver et al. 1998) between AD 1140 and AD 1510 corrected for a running mean over 10 tree rings. Mean standard error of the curve is $8 \pm 3 \mathrm{yr}$.

| Tree-ring date AD | ${ }^{14} \mathrm{C}$ age <br> (BP) | Tree-ring date AD | ${ }^{14} \mathrm{C}$ age <br> (BP) | Tree-ring date AD | ${ }^{14} \mathrm{C}$ age <br> (BP) | Tree-ring date AD | ${ }^{14} \mathrm{C}$ age <br> (BP) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1140 | 970 | 1178 | 871 | 1216 | 828 | 1254 | 804 |
| 1141 | 967 | 1179 | 871 | 1217 | 825 | 1255 | 803 |
| 1142 | 963 | 1180 | 871 | 1218 | 823 | 1256 | 802 |
| 1143 | 960 | 1181 | 872 | 1219 | 821 | 1257 | 800 |
| 1144 | 956 | 1182 | 872 | 1220 | 819 | 1258 | 799 |
| 1145 | 952 | 1183 | 872 | 1221 | 817 | 1259 | 797 |
| 1146 | 949 | 1184 | 873 | 1222 | 815 | 1260 | 795 |
| 1147 | 945 | 1185 | 873 | 1223 | 813 | 1261 | 793 |
| 1148 | 941 | 1186 | 873 | 1224 | 811 | 1262 | 790 |
| 1149 | 937 | 1187 | 874 | 1225 | 810 | 1263 | 788 |
| 1150 | 933 | 1188 | 874 | 1226 | 809 | 1264 | 785 |
| 1151 | 929 | 1189 | 874 | 1227 | 808 | 1265 | 782 |
| 1152 | 925 | 1190 | 874 | 1228 | 807 | 1266 | 779 |
| 1153 | 921 | 1191 | 874 | 1229 | 806 | 1267 | 776 |
| 1154 | 917 | 1192 | 873 | 1230 | 805 | 1268 | 773 |
| 1155 | 913 | 1193 | 873 | 1231 | 805 | 1269 | 769 |
| 1156 | 910 | 1194 | 872 | 1232 | 804 | 1270 | 766 |
| 1157 | 906 | 1195 | 872 | 1233 | 804 | 1271 | 762 |
| 1158 | 903 | 1196 | 871 | 1234 | 804 | 1272 | 758 |
| 1159 | 899 | 1197 | 870 | 1235 | 804 | 1273 | 754 |
| 1160 | 896 | 1198 | 868 | 1236 | 804 | 1274 | 750 |
| 1161 | 893 | 1199 | 867 | 1237 | 804 | 1275 | 746 |
| 1162 | 890 | 1200 | 865 | 1238 | 804 | 1276 | 742 |
| 1163 | 888 | 1201 | 864 | 1239 | 805 | 1277 | 738 |
| 1164 | 885 | 1202 | 862 | 1240 | 805 | 1278 | 734 |
| 1165 | 883 | 1203 | 860 | 1241 | 805 | 1279 | 730 |
| 1166 | 881 | 1204 | 858 | 1242 | 805 | 1280 | 725 |
| 1167 | 879 | 1205 | 856 | 1243 | 806 | 1281 | 721 |
| 1168 | 877 | 1206 | 853 | 1244 | 806 | 1282 | 717 |
| 1169 | 876 | 1207 | 851 | 1245 | 806 | 1283 | 713 |
| 1170 | 875 | 1208 | 848 | 1246 | 806 | 1284 | 709 |
| 1171 | 874 | 1209 | 846 | 1247 | 807 | 1285 | 705 |
| 1172 | 873 | 1210 | 843 | 1248 | 807 | 1286 | 701 |
| 1173 | 872 | 1211 | 841 | 1249 | 806 | 1287 | 697 |
| 1174 | 872 | 1212 | 838 | 1250 | 806 | 1288 | 694 |
| 1175 | 871 | 1213 | 835 | 1251 | 806 | 1289 | 690 |
| 1176 | 871 | 1214 | 833 | 1252 | 805 | 1290 | 686 |
| 1177 | 871 | 1215 | 830 | 1253 | 805 | 1291 | 683 |

Table 7 Least-squares smoothed calibration curve at yearly intervals for North American Douglas Fir (Pseudotsuga menziesii) (Stuiver et al. 1998) between AD 1140 and AD 1510 corrected for a running mean over 10 tree rings. Mean standard error of the curve is $8 \pm 3 \mathrm{yr}$. (Continued)

| Tree-ring date AD | ${ }^{14} \mathrm{C}$ age <br> (BP) | Tree-ring date AD | ${ }^{14} \mathrm{C}$ age <br> (BP) | Tree-ring date AD | ${ }^{14} \mathrm{C}$ age <br> (BP) | Tree-ring date AD | ${ }^{14} \mathrm{C}$ age <br> (BP) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1292 | 680 | 1330 | 630 | 1368 | 649 | 1406 | 553 |
| 1293 | 676 | 1331 | 630 | 1369 | 649 | 1407 | 550 |
| 1294 | 673 | 1332 | 631 | 1370 | 648 | 1408 | 547 |
| 1295 | 670 | 1333 | 631 | 1371 | 647 | 1409 | 543 |
| 1296 | 667 | 1334 | 632 | 1372 | 646 | 1410 | 540 |
| 1297 | 664 | 1335 | 632 | 1373 | 645 | 1411 | 537 |
| 1298 | 661 | 1336 | 633 | 1374 | 644 | 1412 | 534 |
| 1299 | 658 | 1337 | 634 | 1375 | 643 | 1413 | 532 |
| 1300 | 656 | 1338 | 634 | 1376 | 642 | 1414 | 529 |
| 1301 | 653 | 1339 | 635 | 1377 | 641 | 1415 | 526 |
| 1302 | 650 | 1340 | 636 | 1378 | 639 | 1416 | 523 |
| 1303 | 648 | 1341 | 636 | 1379 | 638 | 1417 | 520 |
| 1304 | 645 | 1342 | 637 | 1380 | 636 | 1418 | 517 |
| 1305 | 643 | 1343 | 638 | 1381 | 635 | 1419 | 514 |
| 1306 | 640 | 1344 | 639 | 1382 | 633 | 1420 | 511 |
| 1307 | 638 | 1345 | 640 | 1383 | 631 | 1421 | 508 |
| 1308 | 636 | 1346 | 640 | 1384 | 629 | 1422 | 504 |
| 1309 | 633 | 1347 | 641 | 1385 | 626 | 1423 | 501 |
| 1310 | 631 | 1348 | 642 | 1386 | 624 | 1424 | 498 |
| 1311 | 629 | 1349 | 643 | 1387 | 622 | 1425 | 495 |
| 1312 | 626 | 1350 | 643 | 1388 | 619 | 1426 | 491 |
| 1313 | 624 | 1351 | 644 | 1389 | 617 | 1427 | 488 |
| 1314 | 622 | 1352 | 645 | 1390 | 614 | 1428 | 484 |
| 1315 | 619 | 1353 | 646 | 1391 | 611 | 1429 | 481 |
| 1316 | 617 | 1354 | 598 | 1392 | 608 | 1430 | 477 |
| 1317 | 615 | 1355 | 601 | 1393 | 605 | 1431 | 474 |
| 1318 | 612 | 1356 | 603 | 1394 | 602 | 1432 | 470 |
| 1319 | 610 | 1357 | 606 | 1395 | 599 | 1433 | 466 |
| 1320 | 608 | 1358 | 609 | 1396 | 596 | 1434 | 462 |
| 1321 | 605 | 1359 | 611 | 1397 | 593 | 1435 | 459 |
| 1322 | 603 | 1360 | 614 | 1398 | 590 | 1436 | 455 |
| 1323 | 601 | 1361 | 616 | 1399 | 586 | 1437 | 451 |
| 1324 | 599 | 1362 | 619 | 1400 | 583 | 1438 | 448 |
| 1325 | 597 | 1363 | 621 | 1401 | 579 | 1439 | 444 |
| 1326 | 595 | 1364 | 623 | 1402 | 576 | 1440 | 440 |
| 1327 | 592 | 1365 | 625 | 1403 | 572 | 1441 | 437 |
| 1328 | 591 | 1366 | 627 | 1404 | 569 | 1442 | 433 |
| 1329 | 589 | 1367 | 629 | 1405 | 565 | 1443 | 430 |

Table 7 Least-squares smoothed calibration curve at yearly intervals for North American Douglas Fir (Pseudotsuga menziesii) (Stuiver et al. 1998) between AD 1140 and AD 1510 corrected for a running mean over 10 tree rings. Mean standard error of the curve is $8 \pm 3 \mathrm{yr}$. (Continued)

| Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{BP})$ | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{BP})$ | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{BP})$ | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{BP})$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1444 | 426 | 1461 | 390 | 1478 | 383 | 1495 | 363 |
| 1445 | 423 | 1462 | 389 | 1479 | 382 | 1496 | 361 |
| 1446 | 420 | 1463 | 388 | 1480 | 382 | 1497 | 360 |
| 1447 | 417 | 1464 | 388 | 1481 | 418 | 1498 | 358 |
| 1448 | 414 | 1465 | 387 | 1482 | 417 | 1499 | 356 |
| 1449 | 411 | 1466 | 387 | 1483 | 416 | 1500 | 354 |
| 1450 | 409 | 1467 | 386 | 1484 | 378 | 1501 | 352 |
| 1451 | 406 | 1468 | 386 | 1485 | 377 | 1502 | 350 |
| 1452 | 404 | 1469 | 386 | 1486 | 376 | 1503 | 349 |
| 1453 | 402 | 1470 | 385 | 1487 | 375 | 1504 | 347 |
| 1454 | 400 | 1471 | 385 | 1488 | 374 | 1505 | 345 |
| 1455 | 398 | 1472 | 385 | 1489 | 372 | 1506 | 343 |
| 1456 | 396 | 1473 | 385 | 1490 | 371 | 1507 | 342 |
| 1457 | 395 | 1474 | 384 | 1491 | 370 | 1508 | 340 |
| 1458 | 393 | 1475 | 384 | 1492 | 368 | 1509 | 338 |
| 1459 | 392 | 1476 | 384 | 1493 | 366 | 1510 | 337 |
| 1460 | 391 | 1477 | 383 | 1494 | 365 |  |  |

Table 8 Least-squares smoothed calibration curve at yearly intervals for North American Douglas Fir (Pseudotsuga menziesii) (Stuiver et al. 1998) between AD 1510 and AD 1950 corrected for a running mean over 10 tree rings. Mean standard error of the curve is $3 \pm 1 \mathrm{yr}$.

| Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> (yr BP) | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{yr} \mathrm{BP})$ | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{yr} \mathrm{BP})$ | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> (yr BP) $)$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1510 | 353 | 1553 | 318 | 1596 | 348 | 1639 | 290 |
| 1511 | 352 | 1554 | 319 | 1597 | 349 | 1640 | 287 |
| 1512 | 351 | 1555 | 321 | 1598 | 351 | 1641 | 284 |
| 1513 | 350 | 1556 | 322 | 1599 | 353 | 1642 | 281 |
| 1514 | 348 | 1557 | 323 | 1600 | 354 | 1643 | 278 |
| 1515 | 347 | 1558 | 325 | 1601 | 355 | 1644 | 275 |
| 1516 | 345 | 1559 | 326 | 1602 | 356 | 1645 | 272 |
| 1517 | 343 | 1560 | 326 | 1603 | 357 | 1646 | 269 |
| 1518 | 341 | 1561 | 327 | 1604 | 358 | 1647 | 266 |
| 1519 | 338 | 1562 | 328 | 1605 | 359 | 1648 | 262 |
| 1520 | 336 | 1563 | 328 | 1606 | 359 | 1649 | 259 |
| 1521 | 333 | 1564 | 328 | 1607 | 360 | 1650 | 256 |
| 1522 | 331 | 1565 | 328 | 1608 | 360 | 1651 | 253 |
| 1523 | 328 | 1566 | 328 | 1609 | 360 | 1652 | 250 |
| 1524 | 325 | 1567 | 328 | 1610 | 360 | 1653 | 247 |
| 1525 | 323 | 1568 | 328 | 1611 | 359 | 1654 | 244 |
| 1526 | 320 | 1569 | 328 | 1612 | 359 | 1655 | 241 |
| 1527 | 317 | 1570 | 328 | 1613 | 358 | 1656 | 238 |
| 1528 | 315 | 1571 | 327 | 1614 | 357 | 1657 | 235 |
| 1529 | 312 | 1572 | 327 | 1615 | 356 | 1658 | 231 |
| 1530 | 310 | 1573 | 327 | 1616 | 354 | 1659 | 228 |
| 1531 | 308 | 1574 | 327 | 1617 | 353 | 1660 | 225 |
| 1532 | 306 | 1575 | 326 | 1618 | 351 | 1661 | 221 |
| 1533 | 305 | 1576 | 326 | 1619 | 349 | 1662 | 218 |
| 1534 | 303 | 1577 | 326 | 1620 | 347 | 1663 | 215 |
| 1535 | 302 | 1578 | 326 | 1621 | 345 | 1664 | 211 |
| 1536 | 301 | 1579 | 327 | 1622 | 342 | 1665 | 208 |
| 1537 | 301 | 1580 | 327 | 1623 | 340 | 1666 | 204 |
| 1538 | 300 | 1581 | 328 | 1624 | 337 | 1667 | 200 |
| 1539 | 300 | 1582 | 328 | 1625 | 334 | 1668 | 197 |
| 1540 | 300 | 1583 | 329 | 1626 | 331 | 1669 | 193 |
| 1541 | 301 | 1584 | 330 | 1627 | 329 | 1670 | 190 |
| 1542 | 302 | 1585 | 331 | 1628 | 326 | 1671 | 186 |
| 1543 | 302 | 1586 | 332 | 1629 | 322 | 1672 | 182 |
| 1544 | 304 | 1587 | 333 | 1630 | 319 | 1673 | 179 |
| 1545 | 305 | 1588 | 335 | 1631 | 316 | 1674 | 175 |
| 1546 | 306 | 1589 | 336 | 1632 | 313 | 1675 | 172 |
| 1547 | 308 | 1590 | 338 | 1633 | 310 | 1676 | 168 |
| 1548 | 309 | 1591 | 339 | 1634 | 307 | 1677 | 165 |
| 1549 | 311 | 1592 | 341 | 1635 | 303 | 1678 | 161 |
| 1550 | 313 | 1593 | 343 | 1636 | 300 | 1679 | 158 |
| 1551 | 314 | 1594 | 344 | 1637 | 297 | 1680 | 155 |
| 1552 | 316 | 1595 | 346 | 1638 | 294 | 1681 | 152 |
|  |  |  |  |  |  | 2 |  |

Table 8 Least-squares smoothed calibration curve at yearly intervals for North American Douglas Fir (Pseudotsuga menziesii) (Stuiver et al. 1998) between AD 1510 and AD 1950 corrected for a running mean over 10 tree rings. Mean standard error of the curve is $3 \pm 1 \mathrm{yr}$. (Continued)

| Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> (yr BP) | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> (yr BP) | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> (yr BP) | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> (yr BP) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1682 | 149 | 1725 | 116 | 1768 | 160 | 1811 | 124 |
| 1683 | 146 | 1726 | 119 | 1769 | 162 | 1812 | 119 |
| 1684 | 143 | 1727 | 122 | 1770 | 164 | 1813 | 115 |
| 1685 | 140 | 1728 | 125 | 1771 | 166 | 1814 | 111 |
| 1686 | 137 | 1729 | 128 | 1772 | 169 | 1815 | 108 |
| 1687 | 134 | 1730 | 131 | 1773 | 172 | 1816 | 105 |
| 1688 | 132 | 1731 | 135 | 1774 | 175 | 1817 | 102 |
| 1689 | 129 | 1732 | 138 | 1775 | 178 | 1818 | 101 |
| 1690 | 127 | 1733 | 141 | 1776 | 182 | 1819 | 99 |
| 1691 | 125 | 1734 | 144 | 1777 | 185 | 1820 | 98 |
| 1692 | 122 | 1735 | 147 | 1778 | 189 | 1821 | 97 |
| 1693 | 120 | 1736 | 150 | 1779 | 192 | 1822 | 97 |
| 1694 | 118 | 1737 | 153 | 1780 | 196 | 1823 | 97 |
| 1695 | 116 | 1738 | 155 | 1781 | 199 | 1824 | 98 |
| 1696 | 114 | 1739 | 157 | 1782 | 202 | 1825 | 98 |
| 1697 | 112 | 1740 | 160 | 1783 | 205 | 1827 | 101 |
| 1698 | 110 | 1741 | 161 | 1784 | 208 | 1828 | 102 |
| 1699 | 109 | 1742 | 163 | 1785 | 210 | 1829 | 103 |
| 1700 | 107 | 1743 | 164 | 1786 | 212 | 1830 | 105 |
| 1701 | 106 | 1744 | 165 | 1787 | 213 | 1831 | 107 |
| 1702 | 104 | 1745 | 166 | 1788 | 214 | 1832 | 108 |
| 1703 | 103 | 1746 | 167 | 1789 | 214 | 1833 | 110 |
| 1704 | 102 | 1747 | 167 | 1790 | 214 | 1834 | 111 |
| 1705 | 100 | 1748 | 167 | 1791 | 214 | 1835 | 113 |
| 1706 | 99 | 1749 | 167 | 1792 | 213 | 1836 | 114 |
| 1707 | 99 | 1750 | 166 | 1793 | 211 | 1837 | 115 |
| 1708 | 98 | 1751 | 165 | 1794 | 208 | 1838 | 116 |
| 1709 | 97 | 1752 | 165 | 1795 | 206 | 1839 | 117 |
| 1710 | 97 | 1753 | 164 | 1796 | 202 | 1840 | 118 |
| 1711 | 97 | 1754 | 163 | 1797 | 199 | 1841 | 119 |
| 1712 | 97 | 1755 | 162 | 1798 | 194 | 1842 | 119 |
| 1713 | 97 | 1756 | 160 | 1799 | 190 | 1843 | 119 |
| 1714 | 97 | 1757 | 159 | 1800 | 185 | 1844 | 119 |
| 1715 | 98 | 1758 | 158 | 1801 | 179 | 1845 | 119 |
| 1716 | 99 | 1759 | 158 | 1802 | 174 | 1846 | 119 |
| 1717 | 100 | 1760 | 157 | 1803 | 168 | 1847 | 119 |
| 1718 | 101 | 1761 | 156 | 1804 | 162 | 1848 | 119 |
| 1719 | 103 | 1762 | 156 | 1805 | 156 | 1849 | 118 |
| 1720 | 104 | 1763 | 156 | 1806 | 151 | 1850 | 118 |
| 1721 | 106 | 1764 | 156 | 1807 | 145 | 1851 | 118 |
| 1722 | 108 | 1765 | 157 | 1808 | 139 | 1852 | 117 |
| 1723 | 111 | 1766 | 157 | 1809 | 134 | 1853 | 117 |
| 173 | 113 | 1767 | 159 | 1810 | 129 | 1854 | 117 |
| 173 |  |  |  |  |  |  |  |

Table 8 Least-squares smoothed calibration curve at yearly intervals for North American Douglas Fir (Pseudotsuga menziesii) (Stuiver et al. 1998) between AD 1510 and AD 1950 corrected for a running mean over 10 tree rings. Mean standard error of the curve is $3 \pm 1 \mathrm{yr}$. (Continued)

| Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{yr} \mathrm{BP})$ | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{yr} \mathrm{BP})$ | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{yr} \mathrm{BP})$ | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{yr} \mathrm{BP})$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1855 | 117 | 1879 | 110 | 1903 | 77 | 1927 | 142 |
| 1856 | 116 | 1880 | 109 | 1904 | 78 | 1928 | 145 |
| 1857 | 116 | 1881 | 107 | 1905 | 80 | 1929 | 148 |
| 1858 | 117 | 1882 | 104 | 1906 | 81 | 1930 | 151 |
| 1859 | 117 | 1883 | 102 | 1907 | 83 | 1931 | 153 |
| 1860 | 117 | 1884 | 100 | 1908 | 86 | 1932 | 156 |
| 1861 | 117 | 1885 | 97 | 1909 | 88 | 1933 | 159 |
| 1862 | 117 | 1886 | 95 | 1910 | 91 | 1934 | 161 |
| 1863 | 118 | 1887 | 92 | 1911 | 93 | 1935 | 164 |
| 1864 | 118 | 1888 | 90 | 1912 | 96 | 1936 | 166 |
| 1865 | 118 | 1889 | 87 | 1913 | 99 | 1937 | 168 |
| 1866 | 119 | 1890 | 85 | 1914 | 102 | 1938 | 171 |
| 1867 | 119 | 1891 | 83 | 1915 | 105 | 1939 | 173 |
| 1868 | 119 | 1892 | 81 | 1916 | 108 | 1940 | 175 |
| 1869 | 119 | 1893 | 79 | 1917 | 112 | 1941 | 177 |
| 1870 | 119 | 1894 | 78 | 1918 | 115 | 1942 | 178 |
| 1871 | 119 | 1895 | 77 | 1919 | 118 | 1943 | 180 |
| 1872 | 119 | 1896 | 76 | 1920 | 121 | 1944 | 182 |
| 1873 | 118 | 1897 | 75 | 1921 | 124 | 1945 | 183 |
| 1874 | 117 | 1898 | 74 | 1922 | 127 | 1946 | 184 |
| 1875 | 116 | 1899 | 74 | 1923 | 130 | 1947 | 186 |
| 1876 | 115 | 1900 | 74 | 1924 | 133 | 1948 | 187 |
| 1877 | 114 | 1901 | 75 | 1925 | 136 | 1949 | 187 |
| 1878 | 112 | 1902 | 76 | 1926 | 139 | 1950 | 188 |

Table 9 Least-squares smoothed calibration curve at yearly intervals for English Oak (Quercus petroca) (McCormac et al. 1998) between AD 1720 and AD 1940 corrected for a running mean over 10 tree rings. Mean standard error of the curve is $8 \pm 4 \mathrm{yr}$.

| Tree-ring date AD | ${ }^{14} \mathrm{C}$ age (BP) | Tree-ring date AD | ${ }^{14} \mathrm{C}$ age (BP) | Tree-ring date AD | ${ }^{14} \mathrm{C}$ age (BP) | Tree-ring date AD | ${ }^{14} \mathrm{C}$ age <br> (BP) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1720 | 131 | 1762 | 201 | 1804 | 153 | 1846 | 128 |
| 1721 | 132 | 1763 | 200 | 1805 | 149 | 1847 | 128 |
| 1722 | 134 | 1764 | 200 | 1806 | 145 | 1848 | 127 |
| 1723 | 136 | 1765 | 199 | 1807 | 141 | 1849 | 127 |
| 1724 | 138 | 1766 | 199 | 1808 | 137 | 1850 | 127 |
| 1725 | 140 | 1767 | 199 | 1809 | 133 | 1851 | 126 |
| 1726 | 143 | 1768 | 199 | 1810 | 130 | 1852 | 126 |
| 1727 | 146 | 1769 | 199 | 1811 | 126 | 1853 | 126 |
| 1728 | 148 | 1770 | 199 | 1812 | 123 | 1854 | 125 |
| 1729 | 151 | 1771 | 199 | 1813 | 121 | 1855 | 125 |
| 1730 | 154 | 1772 | 199 | 1814 | 118 | 1856 | 125 |
| 1731 | 158 | 1773 | 199 | 1815 | 116 | 1857 | 124 |
| 1732 | 161 | 1774 | 200 | 1816 | 114 | 1858 | 124 |
| 1733 | 164 | 1775 | 200 | 1817 | 112 | 1859 | 124 |
| 1734 | 168 | 1776 | 201 | 1818 | 111 | 1860 | 124 |
| 1735 | 171 | 1777 | 201 | 1819 | 110 | 1861 | 124 |
| 1736 | 174 | 1778 | 202 | 1820 | 109 | 1862 | 125 |
| 1737 | 178 | 1779 | 202 | 1821 | 109 | 1863 | 125 |
| 1738 | 181 | 1780 | 203 | 1822 | 108 | 1864 | 126 |
| 1739 | 184 | 1781 | 203 | 1823 | 108 | 1865 | 126 |
| 1740 | 187 | 1782 | 203 | 1824 | 109 | 1866 | 127 |
| 1741 | 189 | 1783 | 203 | 1825 | 109 | 1867 | 128 |
| 1742 | 192 | 1784 | 203 | 1826 | 110 | 1868 | 129 |
| 1743 | 194 | 1785 | 203 | 1827 | 111 | 1869 | 130 |
| 1744 | 196 | 1786 | 203 | 1828 | 112 | 1870 | 131 |
| 1745 | 198 | 1787 | 202 | 1829 | 113 | 1871 | 132 |
| 1746 | 200 | 1788 | 201 | 1830 | 114 | 1872 | 133 |
| 1747 | 201 | 1789 | 200 | 1831 | 116 | 1873 | 134 |
| 1748 | 202 | 1790 | 198 | 1832 | 117 | 1874 | 135 |
| 1749 | 203 | 1791 | 196 | 1833 | 118 | 1875 | 136 |
| 1750 | 204 | 1792 | 194 | 1834 | 120 | 1876 | 137 |
| 1751 | 205 | 1793 | 192 | 1835 | 121 | 1877 | 138 |
| 1752 | 205 | 1794 | 190 | 1836 | 122 | 1878 | 139 |
| 1753 | 205 | 1795 | 187 | 1837 | 123 | 1879 | 140 |
| 1754 | 205 | 1796 | 184 | 1838 | 124 | 1880 | 141 |
| 1755 | 205 | 1797 | 180 | 1839 | 125 | 1881 | 141 |
| 1756 | 204 | 1798 | 177 | 1840 | 126 | 1882 | 142 |
| 1757 | 204 | 1799 | 173 | 1841 | 126 | 1883 | 142 |
| 1758 | 203 | 1800 | 169 | 1842 | 127 | 1884 | 142 |
| 1759 | 203 | 1801 | 165 | 1843 | 127 | 1885 | 142 |
| 1760 | 202 | 1802 | 161 | 1844 | 128 | 1886 | 142 |
| 1761 | 201 | 1803 | 157 | 1845 | 128 | 1887 | 142 |

Table 9 Least-squares smoothed calibration curve at yearly intervals for English Oak (Quercus petroca) (McCormac et al. 1998) between AD 1720 and AD 1940 corrected for a running mean over 10 tree rings. Mean standard error of the curve is $8 \pm 4$ yr. (Continued)

| Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{BP})$ | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{BP})$ | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{BP})$ | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{BP})$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1888 | 141 | 1904 | 128 | 1920 | 144 | 1936 | 179 |
| 1889 | 141 | 1905 | 127 | 1921 | 146 | 1937 | 181 |
| 1890 | 140 | 1906 | 127 | 1922 | 149 | 1938 | 182 |
| 1891 | 139 | 1907 | 127 | 1923 | 151 | 1939 | 184 |
| 1892 | 138 | 1908 | 127 | 1924 | 154 | 1940 | 185 |
| 1893 | 137 | 1909 | 128 | 1925 | 156 |  |  |
| 1894 | 136 | 1910 | 128 | 1926 | 159 |  |  |
| 1895 | 135 | 1911 | 129 | 1927 | 161 |  |  |
| 1896 | 134 | 1912 | 130 | 1928 | 163 |  |  |
| 1897 | 133 | 1913 | 131 | 1929 | 166 |  |  |
| 1898 | 132 | 1914 | 133 | 1930 | 168 |  |  |
| 1899 | 131 | 1915 | 134 | 1931 | 170 |  |  |
| 1900 | 130 | 1916 | 136 | 1932 | 172 |  |  |
| 1901 | 129 | 1917 | 138 | 1933 | 174 |  |  |
| 1902 | 129 | 1918 | 140 | 1934 | 176 |  |  |
| 1903 | 128 | 1919 | 142 | 1935 | 178 |  |  |

Table 10 Least-squares smoothed calibration curve at yearly intervals for New Zealand Cedar (Libocedrus bidwillii) (McCormac et al. 1998) between AD 1720 and AD 1940 corrected for a running mean over 10 tree rings. Mean standard error of the curve is $8 \pm 4 \mathrm{yr}$.

| Tree-ring date AD | ${ }^{14} \mathrm{C}$ age <br> (BP) | Tree-ring date AD | ${ }^{14} \mathrm{C}$ age <br> (BP) | Tree-ring date AD | ${ }^{14} \mathrm{C}$ age <br> (BP) | Tree-ring date AD | ${ }^{14} \mathrm{C}$ age <br> (BP) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1720 | 148 | 1762 | 223 | 1804 | 185 | 1846 | 163 |
| 1721 | 149 | 1763 | 222 | 1805 | 181 | 1847 | 162 |
| 1722 | 151 | 1764 | 221 | 1806 | 176 | 1848 | 162 |
| 1723 | 153 | 1765 | 220 | 1807 | 172 | 1849 | 161 |
| 1724 | 156 | 1766 | 219 | 1808 | 168 | 1850 | 160 |
| 1725 | 158 | 1767 | 218 | 1809 | 164 | 1851 | 160 |
| 1726 | 161 | 1768 | 218 | 1810 | 160 | 1852 | 159 |
| 1727 | 164 | 1769 | 218 | 1811 | 157 | 1853 | 158 |
| 1728 | 168 | 1770 | 218 | 1812 | 153 | 1854 | 157 |
| 1729 | 171 | 1771 | 218 | 1813 | 150 | 1855 | 156 |
| 1730 | 175 | 1772 | 218 | 1814 | 148 | 1856 | 156 |
| 1731 | 179 | 1773 | 218 | 1815 | 146 | 1857 | 155 |
| 1732 | 183 | 1774 | 219 | 1816 | 144 | 1858 | 155 |
| 1733 | 187 | 1775 | 220 | 1817 | 142 | 1859 | 154 |
| 1734 | 191 | 1776 | 221 | 1818 | 141 | 1860 | 154 |
| 1735 | 195 | 1777 | 222 | 1819 | 140 | 1861 | 154 |
| 1736 | 199 | 1778 | 223 | 1820 | 139 | 1862 | 154 |
| 1737 | 203 | 1779 | 224 | 1821 | 139 | 1863 | 154 |
| 1738 | 207 | 1780 | 225 | 1822 | 139 | 1864 | 154 |
| 1739 | 210 | 1781 | 226 | 1823 | 140 | 1865 | 154 |
| 1740 | 214 | 1782 | 227 | 1824 | 141 | 1866 | 154 |
| 1741 | 217 | 1783 | 228 | 1825 | 142 | 1867 | 155 |
| 1742 | 220 | 1784 | 229 | 1826 | 143 | 1868 | 155 |
| 1743 | 223 | 1785 | 229 | 1827 | 144 | 1869 | 156 |
| 1744 | 225 | 1786 | 230 | 1828 | 146 | 1870 | 156 |
| 1745 | 227 | 1787 | 230 | 1829 | 147 | 1871 | 157 |
| 1746 | 229 | 1788 | 229 | 1830 | 149 | 1872 | 157 |
| 1747 | 231 | 1789 | 229 | 1831 | 150 | 1873 | 157 |
| 1748 | 232 | 1790 | 228 | 1832 | 152 | 1874 | 158 |
| 1749 | 233 | 1791 | 227 | 1833 | 154 | 1875 | 158 |
| 1750 | 233 | 1792 | 225 | 1834 | 155 | 1876 | 158 |
| 1751 | 233 | 1793 | 224 | 1835 | 157 | 1877 | 158 |
| 1752 | 233 | 1794 | 221 | 1836 | 158 | 1878 | 158 |
| 1753 | 233 | 1795 | 219 | 1837 | 160 | 1879 | 158 |
| 1754 | 232 | 1796 | 216 | 1838 | 161 | 1880 | 157 |
| 1755 | 232 | 1797 | 213 | 1839 | 162 | 1881 | 157 |
| 1756 | 231 | 1798 | 210 | 1840 | 162 | 1882 | 156 |
| 1757 | 230 | 1799 | 206 | 1841 | 163 | 1883 | 155 |
| 1758 | 228 | 1800 | 202 | 1842 | 163 | 1884 | 153 |
| 1759 | 227 | 1801 | 198 | 1843 | 163 | 1885 | 152 |
| 1760 | 226 | 1802 | 194 | 1844 | 163 | 1886 | 151 |
| 1761 | 225 | 1803 | 189 | 1845 | 163 | 1887 | 149 |

Table 10 Least-squares smoothed calibration curve at yearly intervals for New Zealand Cedar (Libocedrus bidwillii) (McCormac et al. 1998) between AD 1720 and AD 1940 corrected for a running mean over 10 tree rings. Mean standard error of the curve is $8 \pm 4 \mathrm{yr}$. (Continued)

| Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{BP})$ | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{BP})$ | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{BP})$ | Tree-ring <br> date AD | ${ }^{14} \mathrm{C}$ age <br> $(\mathrm{BP})$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1888 | 147 | 1907 | 122 | 1926 | 156 |  |  |
| 1889 | 145 | 1908 | 123 | 1927 | 158 |  |  |
| 1890 | 143 | 1909 | 124 | 1928 | 159 |  |  |
| 1891 | 141 | 1910 | 125 | 1929 | 160 |  |  |
| 1892 | 139 | 1911 | 126 | 1930 | 161 |  |  |
| 1893 | 137 | 1912 | 128 | 1931 | 162 |  |  |
| 1894 | 135 | 1913 | 128 | 1932 | 163 |  |  |
| 1895 | 133 | 1914 | 131 | 1933 | 164 |  |  |
| 1896 | 131 | 1915 | 133 | 1934 | 164 |  |  |
| 1897 | 129 | 1916 | 135 | 1935 | 164 |  |  |
| 1898 | 128 | 1917 | 137 | 1936 | 165 |  |  |
| 1899 | 126 | 1918 | 139 | 1937 | 165 |  |  |
| 1900 | 125 | 1919 | 142 | 1938 | 165 |  |  |
| 1901 | 124 | 1920 | 144 | 1939 | 164 |  |  |
| 1902 | 123 | 1921 | 146 | 1940 | 164 |  |  |
| 1903 | 122 | 1922 | 148 |  |  |  |  |
| 1904 | 122 | 1923 | 150 |  |  |  |  |
| 1905 | 122 | 1924 | 152 |  |  |  |  |
| 1906 | 122 | 1925 | 154 |  |  |  |  |
|  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |
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