The impact of salt, fat and sugar levels on toddler food intake

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Abstract
Understanding the early determinants of food intake, in particular the role of food sensory quality, is a necessary step to improve the prevention of unhealthy food habits. However, the extent to which food intake varies according to salt, fat and sugar content is imperfectly known. The present study aimed at evaluating whether toddler food intake varied during lunches or snacks in which salt, fat or sugar contents had been modified in common foods. Seventy-four children (30 (± 0.5) months old) participated in the study in their usual day-care centres. Every other week, they were served lunches composed, among other items, of green beans and pasta with varying salt (0, 0.6 and 1.2 % added salt) or fat (0, 2.5 and 5 % added butter) levels and afternoon snacks composed of fruit purée varying in sugar level (0, 5 and 10 % added sugar). During each meal, children could eat as much as they wanted from the target foods. Each food was weighed before and after the meal. Salt level had a positive impact on the intake of the target foods. On the contrary, no impact of added fat or added sugar levels was observed. This implies that fat and sugar addition could be avoided in foods for children without having an impact on palatability, allowing the energy density of children's diet to be limited. Salt addition should be limited, but its suppression in vegetables, whose intake is to be promoted, should be considered cautiously.

Key words: Children; Food intake; Salt; Fat; Sugar; BMI

Over the past few years, many countries have developed governmental policies aimed at decreasing the intake of salt-rich, fat-rich and sugar-rich foods (1). In France in particular, the National Nutrition and Health Programme aims at limiting the quantity of salt, fat and sugar in foods (2). However, data measuring the impact of these reductions on food consumption, in particular in young children, are not available. Regarding the influence of sensory properties of foods on intake, especially in young children (3), and the importance of early childhood in the development of long-standing food behaviour (4), studying the impact of the recommendations on child food intake seems necessary. Few studies to date have evaluated the link between ingredient levels and intake especially among young children. In the present context of the universal prevalence of paediatric obesity and of the wide availability of palatable processed foods, it is essential to take into account the sensory drivers of eating.

In infants, previous studies using aqueous solutions showed that a preference for salty taste in a solution emerges at about the age of 4 months and persists till the age of 2 years (5–7). Unlike what is observed for salty taste, humans exhibit an innate preference for sweet solutions, already observed a few hours after birth (8), with newborn babies preferring highly sweet solutions (9). Furthermore, during the first year of life, sweet and salty solutions are consumed on average more than water (7). In children under the age of 2 years, sweet solutions generally elicit acceptance (10), and the sweeter the solution the more they consume (11). This preference may be linked to a high consumption of sweet foods: 4-year-old children who prefer the sweetest solution choose the sweetest apple juice at snack time (12).

Beyond studies using solutions, research on foods showed that the preference for salt in foods, which appears at the age of 6 months, is reinforced by repeated exposure to salty foods, at least till 12 months of age (13). Moreover, a context-specific preference for salt appears in early childhood. On the one hand, 2-year-old children tend to ingest more salty soup and carrots than the plain versions. On the other hand, children between 3 and 6 years old reject salty solutions but not salty soup (10,14). This suggests the emergence of a preference for salty solutions between 3 and 6 months, followed by a preference for salty over non-salty versions of foods.

In other respects, fat-rich and sugar-rich foods, frequently associated with the high prevalence of overweight and obesity, are both highly palatable and appealing (15),...
while providing dietary energy at low cost\(^{(16)}\). Sugar and fat are liked food features by children and the sweet and fat taste preference in childhood can be explained by the pleasure or by the energy they provide, considering the energy needs while growing\(^{(17)}\). Furthermore, an age-related decline in sweet preference was documented: it is high in childhood\(^{(18)}\) and early adolescence and decreases in early adulthood\(^{(19)}\). As far as fat is concerned, it has positive post-ingestive consequences, since it conditions the preference for new flavours in preschool children\(^{(20,21)}\).

However, the impact of fat level on child preference was not studied directly as was that of salt or sugar level. In a recent study, decreasing the fat content of a food (to decrease its energy density) was associated with a decrease in energy intake and also with a slight increase in weight intake in relation to energy compensation\(^{(22)}\). Moreover, child preference and consumption of fats appear to be linked to those of their parents and to their own BMI\(^{(23)}\).

The present study aims at evaluating the impact of salt, fat or sugar levels in common foods on children’s intake, during a normal meal in a familiar environment. The general hypothesis is that salt, fat and sugar increase food palatability, which can potentially lead to an increase of food intake\(^{(24)}\). Three ingredient levels were studied: no addition, usual level used while preparing foods for children and twice the usual level. Our hypotheses were that children would (i) consume more of the saltier version and less of the salt-free one than of the usual one and (ii) consume more of the high-fat version and less of the fat-free one compared with the usual one considering that a positive effect of fat on palatability might be observed. Concerning added sugar variations, they were tested in a fruit puree. It could be supposed that children would (iii) consume more of the sweeter versions of this food than of the unsweetened version. Moreover, the existence of a food-specific effect of ingredient level on intake can be hypothesised, as for instance context-specific effects were shown for salt preference. In particular, the ingredient level effect may be different according to the hedonic value of the target food, since sensory variations in a liked food might affect the intake differently compared with one that is rather disliked. Hence, variations of salt and fat were tested in two foods: one vegetable and one starchy food.

### Participants and methods

#### Participants

All children attending one of three nurseries in Dijon (France) were proposed to take part in the present experiment. The children’s body weight (kg) and height (m) were measured by the nursery doctor and z-scores for BMI were calculated using the French and International Obesity Task Force standards for child overweight and obesity\(^{(25)}\). Their parents received an information sheet and signed a consent form. One of the children’s parents did not allow their child participate due to food allergies. The study procedure was conducted according to the guidelines laid down in the Declaration of Helsinki, and approved by a local ethical committee (Comité de Protection des Personnes Est 1 Bourgogne).

A total of seventy-four children (forty-two girls and thirty-two boys) participated in the study. Their average age was 30 (SE 4) months old (range 18–37 months), with average BMI of 15.9 (SE 0.1) kg/m\(^2\) (range 13.5–18.9 kg/m\(^2\)) and average BMI z-score –0.2 (SE 0.1) (range –2.3–2.0).

### General experimental design

The present study aimed at evaluating toddler food intake according to variations of salt, fat and sugar levels in usual foods in a natural setting. It took place in the children’s usual nursery in order to minimise the effects due to context changes. Before the beginning of the study, the experimenters visited the nurseries several times in order for the children and the nursery staff to become familiar to them. The nurses were asked to behave as normal (keep an eye on the children during the meal and help maintain calm conditions). Moreover, they were instructed neither to exhort children to eat nor to urge them to eat. Children were seated by groups of four to six around a table and one nurse was in charge of one table as usual. All the children were able to eat alone with a spoon or a fork, without any help. All the menu items were prepared in each nursery kitchen by the nursery cook and were pre-weighed by experimenters before lunch. The children could eat as much as they wanted from the target foods by asking for additional servings if they wanted to, which was a common practice in each nursery. All the servings were registered during the meal or snack for each child on a record sheet. Intake (g) of each food item of the lunch or snack was measured by weighing each food before and after consumption to the nearest 1 g (Soehnle, Leifheit Group, Germany). The nurses were blind to the modifications brought to the target foods.

### Added salt and fat level variations during lunches: procedure

For salt or fat level variations, the target foods were green beans and pasta served at lunch. The experimental fixed menu was served every other week at 11.00 hours. It was composed of vegetable salad (a unique serving of 50 g), chicken (a unique serving of 50 g), green beans (an initial serving of 50 g and up to a complement of 125 g on the basis of 25 g for each additional serving), pasta (an initial serving of 50 g and up to a complement of 125 g on the basis of 25 g for each additional serving), yoghurt (a unique serving of 125 g), bread (\textit{ad libitum}) and water (\textit{ad libitum}). Portion sizes were adapted from the French guidelines for school canteen usage\(^{(26)}\). Vegetable salad and chicken were prepared as usual. Frozen green beans and pasta (small macaroni) were boiled with tap water (\textit{ad libitum}). Portion sizes were adapted from the French guidelines for school canteen usage\(^{(26)}\).
water (11 and 8 min, respectively) without adding salt, oil or any flavouring. The experimenters further added salt or butter according to the experimental design and stirred the foods homogeneously.

Added salt level variations were studied in a first series of five observations and added fat level variations were studied in a second series of five observations. According to the nursery cooks, the average levels used for preparing foods for children were 0.6 g of salt/100 g of food and 2.5 g of butter/100 g of food. These usual levels were then chosen to be the 'control' values (0.6% salt and 2.5% butter). We varied these levels by adding no salt or no fat (0%) or by doubling them (1.2% salt and 5% butter). Thus, three levels were tested for each target food and each ingredient. For both the series, the first observation was a control for salt and fat. Then the order of the four other observations was different across the nurseries (Table 1). For each observation, salt or fat level varied only in green beans or in pasta. For instance, if salt level varied in green beans, their fat content was the control level and pasta was served with control levels of salt and of fat.

After the serving of the fixed amount of vegetable salad, children were served chicken and green beans; afterwards, they were served pasta and then the yogurt. This order of presentation was chosen in order to allow the children to eat the vegetable first, which was generally less consumed by children compared with pasta (27) – this also being the usual practice in the nurseries.

**Added sugar level variation during snack time: procedure.** For sugar level variations, the target food was a fruit purée served during the afternoon snack every other week at 15.30 hours. In order to evaluate the children's intake depending on added sugar level, an unsweetened apple–peach purée (Hero, 26400 ALLEX, France) was chosen and three levels of added sugar were tested. One level corresponded to the plain fruit purée (unsweetened; variant named 0%); for the second one, 5 g of sugar was added to 100 g of purée (variant named 5%); and for the third one, 10 g of sugar was added to 100 g of purée (variant named 10%). The order of presentation for the three variants was different across the nurseries. A first portion of 100 g was served and additional servings of 50 g were available, up to a complement of 200 g. Afterwards, the children were served a maximum of 20 g of biscuits then *ad libitum* milk and/or water.

**Sensory description of food variants**

A sensory description of the food variants was performed with an adult panel. In fact, regarding the young age of the children participating in the main study, sensory description was not possible (28,29). A trained sensory panel of twelve adults (six women and six men), aged between 27 and 66 years, evaluated all variants of the three target foods. The panel members were screened for their olfactory abilities, their abilities to perceive and identify the basic tastes and for their capability to use a linear scale to score intensities. The intensity scales from the Spectrum™ method for descriptive analysis were used to describe sweet, salty, sour and bitter tastes (30). A Spectrum™-like scale was developed for fattiness perception. After eleven training sessions on the way to use the scale for each taste or perception, the panellists were asked to score the perceived intensities of each variant of each target food presented monadically, from 'not perceived' (0) to 'very intense' (10). Two replications were performed.

**Statistical analyses**

Statistical analyses were carried out using SAS System for Windows version 9.1 (SAS Institute, Inc., Cary, NC, USA). Significance was set at *P* < 0.05. For sensory description, an ANOVA was performed using a linear model (SAS GLM

### Table 1. Experimental design for varying added (g/100 g) salt and fat (butter) in green beans and pasta for each nursery

<table>
<thead>
<tr>
<th>Nursery</th>
<th>Food</th>
<th>Ingredient</th>
<th>Salt level variations</th>
<th>Fat level variations</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td></td>
<td>Observation number</td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td>1  2  3  4  5</td>
<td>6  7  8  9  10</td>
</tr>
<tr>
<td>A</td>
<td>Green beans</td>
<td>Salt</td>
<td>0.6 0.6 0.6 1.2</td>
<td>0.6 0.6 0.6 0.6 0.6</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Butter</td>
<td>2.5 2.5 2.5 2.5</td>
<td>2.5 2.5 2.5 2.5 2.5</td>
</tr>
<tr>
<td></td>
<td>Pasta</td>
<td>Salt</td>
<td>0.6 1.2 0.6 0.6</td>
<td>0.6 0.6 0.6 0.6 0.6</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Butter</td>
<td>2.5 2.5 2.5 2.5</td>
<td>2.5 2.5 2.5 2.5 2.5</td>
</tr>
<tr>
<td>B</td>
<td>Green beans</td>
<td>Salt</td>
<td>0.6 0 0.6 1.2 0.6</td>
<td>0.6 0.6 0.6 0.6 0.6</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Butter</td>
<td>2.5 2.5 2.5 2.5</td>
<td>2.5 2.5 2.5 2.5 2.5</td>
</tr>
<tr>
<td></td>
<td>Pasta</td>
<td>Salt</td>
<td>0.6 0.6 1.2 0.6</td>
<td>0.6 0.6 0.6 0.6 0.6</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Butter</td>
<td>2.5 2.5 2.5 2.5</td>
<td>2.5 2.5 2.5 2.5 2.5</td>
</tr>
<tr>
<td>C</td>
<td>Green beans</td>
<td>Salt</td>
<td>0.6 0.6 1.2 0.6</td>
<td>0.6 0.6 0.6 0.6 0.6</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Butter</td>
<td>2.5 2.5 2.5 2.5</td>
<td>2.5 2.5 2.5 2.5 2.5</td>
</tr>
<tr>
<td></td>
<td>Pasta</td>
<td>Salt</td>
<td>0.6 0.6 1.2 0.6</td>
<td>0.6 0.6 0.6 0.6 0.6</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Butter</td>
<td>2.5 2.5 2.5 2.5</td>
<td>2.5 2.5 2.5 2.5 2.5</td>
</tr>
</tbody>
</table>

* Control observations.
procedure) for each taste quality and for the three variants of each target food. The model was:

\[
\text{Taste quality} = \text{panellist} + \text{ingredient level} + \text{panellist} \\
\times \text{ingredient level} + \text{error},
\]

where ‘panellist’ was considered as a random effect. Means for the three variants were compared with one another using \( t \) tests.

To analyse the children’s consumption, the outcome was intake (g) of green beans, pasta or fruit puree. For each series, paired \( t \) tests (\( t_{66} \)) were used to determine, over the five lunches, whether the average intake of green beans was different from the average intake of pasta. For each ingredient (fat, salt or sugar) and each target food intake (green beans, pasta or fruit puree), data were analysed using a mixed linear model (SAS Mixed procedure), with ‘child effect’ considered as random. The primary factor tested in the model was ingredient level. The children’s BMI \( z \)-score was introduced as a covariate to evaluate to what extent it might affect the target food intake. Interactions between ingredient level and BMI were tested in the model, if significant. For the comparison of means, the Dunnett test was applied: the control variant (0·6 % added salt, 2·5 % added butter or 5 % added sugar) was compared with the other variants. Data are reported as mean values with their standard errors.

**Results**

**Sensory description of food variants**

**Salt level variations in green beans and pasta.** Saltiness intensity differed according to added salt levels in green beans and in pasta (\( F(2,36) = 4·10; P=0·025 \) and \( F(2,36) = 24·30; P<0·0001 \), respectively), as shown in Table 2. However, panellists were not able to distinguish the 2·5 % and the 5 % added fat green beans, whereas the three variants of pasta were distinguished from each other.

**Fat level variations in green beans and pasta.** Fattiness perception differed according to added fat levels in green beans and in pasta (\( F(2,36) = 7·49; P<0·0001 \) as reported in Table 2: the three variants were significantly different, as the more the added fat the higher the perceived fattiness. They also differed with regard to sourness and bitterness intensities, (\( F(2,36) = 8·34; P=0·0011 \) and \( F(2,36) = 9·57; P=0·0005 \), respectively): the 0 % added salt variant was perceived more sour and more bitter than the other two variants.

**Sugar level variations in fruit puree.** Sweetness intensity differed according to added sugar level in fruit puree (\( F(2,36) = 74·89; P<0·0001 \) as reported in Table 2: the three variants were significantly different, as the more the added sugar the higher the perceived sweetness. They also differed with regard to sourness and bitterness intensities, (\( F(2,36) = 71·47; P<0·0001 \) and \( F(2,36) = 4·35; P=0·025 \), respectively). The 2·5 % and the 5 % added fat green beans, whereas the three variants of pasta were distinguished from each other.

**Children’s intake during experimental lunches**

**Intake when added salt level varied.** From the initial sample of seventy-four children, only those who were present for at least two meals out of five were considered for the present analyses, resulting in a final group of sixty-six children. No child reached the maximum serving of 175 g for green beans and only 7 % reached this maximum for pasta. A paired \( t \) test comparing the average intake of green beans and only 7 % reached this maximum for pasta. A paired \( t \) test comparing the average intake of green beans and pasta over the five lunches showed, as expected, that pasta intake, 87 (SE 4) g, was greater than that of green beans, 32 (SE 3) g, where \( t_{65} = 12·5 \) (\( P<0·0001 \)).

### Table 2. Sensory description of food variants

<table>
<thead>
<tr>
<th>Food</th>
<th>Ingredient</th>
<th>Level (%)</th>
<th>Salty*</th>
<th>Sweet</th>
<th>Sour</th>
<th>Bitter</th>
<th>Fatty</th>
</tr>
</thead>
<tbody>
<tr>
<td>Green beans</td>
<td>Salt</td>
<td>0</td>
<td>0·25a</td>
<td>0·38a</td>
<td>0·62b</td>
<td>0·21a</td>
<td>0·50a</td>
</tr>
<tr>
<td></td>
<td>0·6</td>
<td>2·10a</td>
<td>0·59a</td>
<td>0·40a</td>
<td>0·15a</td>
<td>1·15a</td>
<td></td>
</tr>
<tr>
<td></td>
<td>1·2</td>
<td>3·81b</td>
<td>0·55a</td>
<td>0·49a</td>
<td>0·17a</td>
<td>1·11a</td>
<td></td>
</tr>
<tr>
<td></td>
<td>Butter</td>
<td>0</td>
<td>2·22a</td>
<td>0·76a</td>
<td>0·63a</td>
<td>0·18a</td>
<td>0·77a</td>
</tr>
<tr>
<td></td>
<td>2·5</td>
<td>2·24a</td>
<td>0·73a</td>
<td>0·44a</td>
<td>0·17a</td>
<td>1·08a</td>
<td></td>
</tr>
<tr>
<td></td>
<td>5</td>
<td>2·20a</td>
<td>0·79a</td>
<td>0·53a</td>
<td>0·18a</td>
<td>1·16a</td>
<td></td>
</tr>
<tr>
<td>Pasta</td>
<td>Salt</td>
<td>0</td>
<td>0·26a</td>
<td>0·69a</td>
<td>0·38a</td>
<td>0·21a</td>
<td>1·58a</td>
</tr>
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<td></td>
<td>0·6</td>
<td>2·54b</td>
<td>0·79a</td>
<td>0·26a</td>
<td>0·14a</td>
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<td></td>
<td>1·2</td>
<td>3·94c</td>
<td>0·86a</td>
<td>0·31a</td>
<td>0·44a</td>
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<td></td>
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<td>2·45a</td>
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<tr>
<td></td>
<td>5</td>
<td>2·47a</td>
<td>0·59a</td>
<td>0·32a</td>
<td>0·08a</td>
<td>1·99a</td>
<td></td>
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<tr>
<td>Fruit purée</td>
<td>Sugar</td>
<td>0</td>
<td>0·19a</td>
<td>1·45a</td>
<td>3·96b</td>
<td>0·65a</td>
<td>0·21a</td>
</tr>
<tr>
<td></td>
<td>5</td>
<td>0·14a</td>
<td>3·10a</td>
<td>2·84a</td>
<td>0·37a</td>
<td>0·27a</td>
<td></td>
</tr>
<tr>
<td></td>
<td>10</td>
<td>0·12a</td>
<td>4·64a</td>
<td>2·33a</td>
<td>0·25a</td>
<td>0·34a</td>
<td></td>
</tr>
</tbody>
</table>

* For each sensory descriptor, the used scale ranged from 0 to 10.

\( a, b, c \): Mean values within columns with unlike superscript letters were significantly different for each food, each ingredient and each sensory descriptor (\( P<0·05 \)).
The impact of salt level variations on intake was evaluated separately for each target food. The results are shown in Fig. 1(a). Concerning green beans, the salt level had an impact on the intake ($R^2(1,106) = 0.89; P<0.0001; n = 63$): 0% salt green beans were less consumed than 0.6% salt green beans. Concerning pasta, the salt level also had an impact on the intake ($R^2(2,92) = 3.19; P=0.045; n = 60$): 1.2% salt pasta was more consumed than 0.6% salt pasta. Hence, the suppression of salt decreased green bean intake and the addition of salt increased pasta intake, compared with the usual salt level. Despite this difference in salt level effect observed through the analyses conducted separately for each target food, a global analysis performed on data of both target foods did not reveal a significant food $\times$ salt level interaction.

**Intake when added fat level varied.** From the initial sample of seventy-four children, only those who attended at least two meals out of five were considered for the present analyses, resulting in a final group of sixty-nine children. No child reached the maximum serving of 175 g for green beans and only 11% reached this maximum for pasta. A paired t-test comparing the average intake of green beans and pasta over the five lunches showed, as expected, that pasta intake, 99 (SE 4) g was greater than green bean intake, 43 (SE 3) g, where $t_{62} = 12.23 (P<0.0001)$.

The impact of fat level variations on intake was evaluated separately for each target food. Fat level had an impact on neither the intake of green beans ($R^2(2,103) = 0.45; P=0.63; n = 62$) nor on that of pasta ($R^2(2,99) = 0.48; P=0.62; n = 60$). The results are shown in Fig. 1(b).

*Children's intake during experimental snacks when added sugar level varied*

From the initial sample of seventy-four children, only those who attended the three snack sessions in which sugar level varied in fruit purée were considered for the present analysis ($n = 61$). Only 4% of children reached the maximum serving of 300 g of fruit purée.

Sugar level variations did not have an impact on fruit purée intake ($R^2(2,120) = 1.43; P=0.24$), as shown in Fig. 2.

*Effect of BMI z-score on intake*

The children’s BMI z-score was not related to intake when added salt level varied, neither in green beans ($R^2(1,106) = 0.78; P=0.38$) nor in pasta ($R^2(1,92) = 0.61; P=0.44$). When added fat level varied, BMI z-score did not affect the green bean intake ($R^2(1,103) = 1.77; P=0.18$), but it had an impact on pasta intake ($R^2(1,99) = 6.01; P=0.01$): the higher the BMI z-score the higher the pasta intake. Further examination of data revealed that BMI z-score was especially related to intake of the 5% added butter pasta. When added sugar level varied in fruit purée, no impact of BMI z-score on intake was observed ($R^2(1,120) = 0.32; P=0.57$).
intake in children by maintaining a salty perception without adding large amounts of salt.

Concerning the absence of fat impact on intake, two hypotheses can be formulated: either the changes of fat content over different meals were not perceived by children or fat is not an immediate hedonic determinant of food intake in children. Here, the trained panel could differentiate the ‘fattiness perception’ in the three samples of pasta varying in fat content, and concerning green beans, the panel could only differentiate the 0% added butter green beans from the other two samples. Despite these sensory differences, children’s intake was not affected by fat content. Hence, fat sensory properties may not be well detected by human subjects. In fact, studies in adults using dairy products showed that participants could not distinguish between the samples according to their fat content, but liked the fattiest samples. Yet, it is admitted that the preference for high-fat foods is not based on a ‘conscious’ perception of fat content. However, as for salt, fat in foods is also known to enhance their flavour and their acceptability and may lead to an increased preference, and by extension, to an increased intake. However, we did not observe such an effect. That the fat level in food is poorly perceived is also revealed in a study examining the link between energy density associated with varying fat content and intake of macaroni and cheese in 2- to 5-year-old children. Intake for two fat levels (39.8 and 19.7% total fat) was measured in separate meals: when the fat level decreased, a small but significant increase in intake was observed (+7%). In the present study, the addition of 5% butter corresponded to 41% fat; thus when total fat varied from 41 to 0%, a 6.5 % increase in intake was noticed, which did not reach significance. Nevertheless, these fat levels correspond to those generally used in nursery canteens for preparing food for children. With such a range of fat content, the presentation of an ad libitum meal was not associated with children’s adjustment for varying palatability and/or energy density in the short term. Similarly, concerning the green beans, the decrease in fat content in green beans (from 41 to 0% total fat) resulted in a 4.5% increase in intake, which again was NS. Hence, contrarily to our initial hypothesis, no direct effect of fat on intake was observed.

The present data also showed that preschool children with a higher BMI z-score consumed more pasta when fat level was higher, thus confirming the previous result highlighting that fatter children prefer high-fat foods. Moreover, child ‘fat appetite’ was previously linked to their parents’ body mass, and a link between child preference for high-fat snack foods and their consumption of a high percentage of total energy from fat was also shown. A similar link between the preference for high-fat foods and BMI was described in 9- to 12-year-old children. Longitudinal approaches would help us to understand the contribution of fatty food intake in...
the development of obesity. Alternatively, this effect could be related to sensitivity to fatty acids which is lower in heavier adult subjects, as reported recently \cite{43}. To date, such an effect has not been studied in children.

In other respects, two main reasons are generally put forward to explain the low intake of vegetables among young children: the first one is the absence of reinforcement resulting from their low energy density \cite{44,45} and the second one is their disliked taste that might be due to their bitterness \cite{46} (as was discussed earlier). The lower intake of green beans compared with that of pasta confirms the previous data showing that, in a free choice situation, vegetables are among the least selected foods by young children compared with starchy foods and animal products \cite{27}, despite their recognised beneficial effects on health \cite{47}. Moreover, the low variety of consumed vegetables observed in early childhood persists through to adulthood \cite{44,45}. Hence, increasing vegetable intake by children is a major issue and more knowledge is needed about factors influencing their acceptance. Several studies have investigated the methods for improving preschool children's vegetable consumption by evaluating the effect of learning practices (for a review, see Havermans \cite{48}) as well as the effect of varying ways of preparation \cite{49}. The last study showed that children prefer boiled vegetables, which was the preparation method chosen for the present study. Clearly, sensory properties of vegetables are largely modified by preparation and much is to be learned to fully understand their impact on vegetable intake by children.

**Added sugar level variations**

The present findings indicate that 2- to 3-year-old children's intake of a fruit purée was not affected by its added sugar content. Contrarily to our initial hypothesis, no increase in intake with increasing added sugar level was observed. With regard to the recognised effect of sweetening on enhancing the palatability of foods, three hypotheses may explain this result. First, when tasting the fruit purée, children may not have perceived the difference in sweetness intensity between the variants over separate snack days. However, the results obtained by the trained panel showed a clear distinction between the samples in terms of sweetness and to a lesser extent of sourness and bitterness. Secondly, the unsweetened variant may have been sweet enough for children to like and consume it, due to the natural sweet taste of the fruits. Thirdly, the intake of all versions of the fruit purée, 163 g on average, was far above the median portion of commercialised products (100 g), therefore suggesting a ceiling effect.

Individual differences in liking for sweet taste were previously reported \cite{24,50}. In a study comparing children's and adults' liking related to sugar concentration, responses fitted an inverted U-shape for preference with a maximum hedonic response for both groups, but preferred concentrations were higher for children than for adults \cite{17}. In the present study, taken as a group, the intake by children levelled up with increasing added sugar level, with neither increase nor decrease depending on the sugar level, suggesting an indifference to sugar level. Hence, the hypothesis of a tolerance zone to sweet taste, in fact a 'large tolerance plateau', could be considered.

The high consumption of fruit purée is consistent with the fact that liking sweet-tasting foods is strongly linked to the child's basic biology (for a review, see Pepino & Mennella \cite{51}), and the absence of sugar level effect on intake is in line with the observation that most studies in adults failed to find an effect of sugar level on food intake \cite{24}. That is to say, one may not need to add sugar to fruit purée to enhance its intake. Furthermore, an interventional study in children aged 7–11 years showed that the more children consume sweetened drinks, the higher their risk of being overweight or obese \cite{52}. Altogether, these findings suggest that in order to promote a healthy energy balance in children and to favour the consumption of healthy foods such as fruits, limiting or suppressing added sugar may help to prevent excess energy input without altering their intake. Similar data concerning different foods with different textures and different levels of sugar, such as drinks, cakes or dairy products would also be useful and should be investigated.

**Methodological considerations**

Here, an initial portion of 50 g was served of green beans and pasta and of 100 g of fruit purée. This did not lead to a systematic consumption of this portion: at the group level, green bean intake was always below the served portion, revealing that children did not clean up their plates but pasta and fruit purée intake was always above the served portions; so, children did not hesitate to ask for additional servings. The fact that 'additional serving' was a usual practice in nurseries could account for this observation, which is consistent with the results showing that children below 3 years old are not systematically influenced by portion size \cite{53}. Thus, this procedure seems appropriate to observe the effects of subtle sensory manipulations.

Conducting such a study where children did not have the opportunity to compare variants directly, no difference in intake was observed when fat and sugar level varied. Another protocol using a paired preference test might have yielded different results. However, studying intake appears more relevant in order to understand the actual children's behaviour in response to a given level of ingredient, because it is closer to everyday eating experiences where children are not able to compare the products.
Salt level had a positive impact on the intake of target foods, whereas fat (butter) content did not, either for green beans or for pasta. Besides, sugar level did not elicit any difference in the intake of fruit purée. That is to say, lowering the addition of fat or sugar while preparing children’s foods, especially when the food is liked without any extra ingredients being added (such as starchy foods or fruit purée), may appear to be a useful way to reduce energy intake and to avoid the development of a habit of consuming foods rich in fat and sugar. Furthermore, these findings suggest that there is no need to add salt to pasta which is consumed anyway. On the contrary, salt suppression in vegetables, whose intake is to be promoted, should be considered cautiously. Alternatively, seasoning vegetables using salt-associated spices may be an efficient and necessary strategy to promote their consumption in children. Hence, nutritional policies regarding the use of salt, fat and sugar aimed at children could take these results into account and could emphasise that salt addition should be limited and that fat and sugar addition can be avoided, helping to prevent obesity from an early age on and to establish healthier food habits.

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References

Salt, fat, sugar and toddler food intake


